# **Changing Evidence Accumulation and Sharing Behaviour Through Incentives**

## **Laura Katharina Globig**

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Department of Experimental Psychology University College London

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## <span id="page-1-0"></span>**Declaration**

I, Laura K. Globig, confirm that the work presented in this thesis is my own. Where information has been derived from other sources, I confirm that this has been indicated in the thesis.

## <span id="page-2-0"></span>**Abstract**

Behaviour is driven by external incentives (e.g., money, social feedback) and internal incentives (e.g., emotions). We monitor the stock market for profitable investments, we share posts online to receive positive feedback, and we accumulate evidence for the success of a medical procedure to ease presurgery anxiety. This thesis explores how incentives alter evidence accumulation and sharing behaviour.

People accumulate evidence to form beliefs which elicit positive emotions, even if these beliefs are incorrect. This bias is adaptive, as the benefits for well-being typically outweigh the harm of inaccuracy. Chapter 2 examines if evidence accumulation becomes less biased in threatening environments, where severe harm is probable. Combining a social-threat manipulation with a sequential sampling task and Drift-Diffusion Modelling (DDM), I find that under threat, participants are less biased and require weaker evidence for negative conclusions. This may increase precautionary actions.

Although financial accuracy incentives are thought to reduce biases, the empirical evidence is mixed. Chapter 3 examines why they may fail to reduce biased evidence accumulation. Coupling a perceptual task with DDM, I show that while accuracy incentives increase caution, they modulate a separate element of the accumulation process and thus do not reduce the bias, possibly due to its unconscious nature.

I propose that when accuracy incentives are coupled with feedback, decisions become more accurate. External incentives are pervasive online: people share information to receive 'likes'. I hypothesize that the incentive structure of social media platforms, whereby social rewards ('likes') and punishments ('dislikes') are dissociated from accuracy, contributes to the spread of misinformation. Chapter 4 combines simulated social media environments and DDM to show that when feedback is contingent on accuracy, sharing becomes more discerning, reducing misinformation spread.

This thesis unveils the mechanisms through which incentives influence evidence accumulation and sharing behaviour, offering insights for interventions to mitigate biased decision-making.

## <span id="page-4-0"></span>**Impact Statement**

In today's digital world we are perpetually faced with opportunities to accumulate evidence and share information. This behaviour is motivated by external incentives, such as money or social feedback, and internal incentives, such as emotions. These incentives sometimes lead to biased decisions. This thesis investigates how altering the incentive structure people face can mitigate biased evidence accumulation and make sharing behaviour more discerning.

When accumulating evidence, individuals are biased towards 'desirable' conclusions that make them feel good and thus prioritize 'desirable' evidence over 'undesirable' evidence. This can have detrimental consequences. For example, wanting to avoid negative emotions may cause you to dismiss critical signs of disease, and thus fail to get help. Here I examine whether and how changing the incentive structure people face — by (1) exposing participants to a threatening environment, or (2) rewarding accurate responses — reduces the bias in evidence accumulation.

I find that under threat weaker evidence is required to reach undesirable conclusions. This can be advantageous as it leads to increased precautionary actions in threatening environments. While the negative effects of stress have been repeatedly underscored this study shows that stress, induced by perceived threat, can be adaptive. These findings may also explain overly pessimistic decisions in those with anxiety and depression and highlight possible target mechanisms for therapeutic interventions.

While rewarding accuracy is commonly thought to reduce biased decisions, the empirical evidence is mixed. I show that accuracy incentives do not reduce biased evidence accumulation and provide a mechanistic explanation for why this is the case. Specifically, I find that accuracy incentives and the bias towards desirable conclusions alter orthogonal aspects of the accumulation process. This suggests that participants are unaware of their own bias. As accuracy incentives are commonly employed to improve decision-making, these findings are relevant to academics, policymakers and industry leaders and may help inform the development of novel interventions against biased decisions.

Building on this research, I propose that when accuracy incentives are coupled with direct feedback, decision-making becomes more accurate. I demonstrate that the spread of misinformation online is facilitated by the existing incentive structure of social media platforms in which existing social rewards (e.g., 'likes') and punishments (e.g., 'dislikes') are dissociated from the accuracy of the information shared. We share information others will react positively to and avoid sharing information others will react negatively to. However, because this feedback is not tied to accuracy, we sometimes share information even if we suspect it may be false. I find that an intervention which slightly changes this incentive structure, such that feedback is contingent on accuracy increases the proportion of true relative to false information shared, without reducing user engagement. The results offer a framework for an intervention that could be adopted to reduce misinformation spread, which in turn could reduce violence, vaccine hesitancy and political polarization.

Bridging theoretical insights with practical solutions, this research provides a mechanistic account of how changing incentives alters evidence accumulation and sharing behaviour, thereby contributing to the development of theorybased behavioural interventions to improve decision-making.

## <span id="page-6-0"></span>**Acknowledgments**

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## <span id="page-7-0"></span>**Notes To Examiners**

The findings from Chapter 2 have been published in a peer-reviewed article: Globig, L. K., Witte, K., Feng, G., & Sharot, T. (2021). Under threat, weaker evidence is required to reach undesirable conclusions. *Journal of Neuroscience*, *41*(30), 6502-6510. https://doi.org/10.1523/JNEUROSCI.3194- 20.2021

The findings from Chapter 4 have been published in a peer-reviewed article: Globig, L. K., Holtz, N., & Sharot, T. (2023). Changing the incentive structure of social media platforms to halt the spread of misinformation. *Elife*, *12*, e85767. https://doi.org/10.7554/eLife.85767

All chapters have benefited from the guidance and advice of my supervisor Tali Sharot. The 'Factory Task' used in Chapter 2 was designed by Donal Cahill, Filip Gesiarz, and Tali Sharot (Gesiarz et al., 2019). The data presented in Chapter 2 was partially collected by Gloria Feng and Kristin Witte. Nora Holtz assisted with creation of stimuli and task design for Chapter 4.

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## <span id="page-12-0"></span>**Chapter 1: Introduction**

## <span id="page-12-1"></span>**1.1 Reward-Oriented Evidence Accumulation**

In today's *age of information* (Castells, 1996) data is constantly at our fingertips. Recent reports show that every day, internet users spend nearly seven hours online (Statista, 2023), able to effortlessly accumulate an abundance of noisy information (or 'evidence'). This raises a fundamental question: *Why* do we accumulate evidence?

### <span id="page-12-2"></span>**1.1.1 External Incentives**

Humans accumulate evidence by continuously sampling noisy information, in order to form beliefs and make decisions (Platt & Glimcher, 1999; Ratcliff, 1978; Usher & McClelland, 2001) which can gain them rewards and help them avoid punishments ( = *external outcomes*, **see Figure 1.1**; Gold & Shadlen, 2002, 2007). These outcomes can be tangible, such as monetary gains or losses, and/or intangible, for instance in the form of social feedback, which at times can have tangible consequences. For example, an individual might accumulate evidence to decide whether to invest in emerging technology. To make this decision, they need to arbitrate between two alternatives: 'to invest' or 'to withhold'. In doing so, they may be motivated by financial profit and by others' reactions to their decision. When the evidence in favour of one alternative relative to the other is large enough, they make their decision (see also Ratcliff & McKoon, 2008). For instance, if the individual finds strong evidence that an investment will yield substantial financial and/or social returns, they form the belief that the investment is profitable, decide to invest, and may later reap the rewards.

### <span id="page-12-3"></span>**1.1.2 Internal Incentives**

Evidence accumulation is not solely motivated by external incentives. Rather, information also carries intrinsic utility (Blanchard et al., 2015; Bromberg-Martin et al., 2024; Bromberg-Martin & Hikosaka, 2009, 2011; Charpentier et al., 2018; Eliaz & Schotter, 2007; Grant et al., 1998) even if it provides no instrumental benefit. It can deepen our sense of understanding the world around us (Dörnemann et al., 2022; Johnston & Davey, 1997; Kobayashi & Hsu, 2019; Sharot & Sunstein, 2020). For example, imagine someone gathers information about black holes. While this information is unlikely to motivate decisions which lead to financial or social rewards, it increases the individual's understanding of the world around them. Information can also alter our emotions. For example, learning that we are at high risk of cancer will make us feel anxious, while learning that we received a high score on a test can make us feel happy. Unsurprisingly therefore, people prefer seeking desirable information (Charpentier et al., 2018) - which elicits positive emotions (**see Figure 1.1**).

It has also been shown that individuals selectively accumulate evidence to form positive beliefs from which they derive internal rewards, such as positive emotions (Gesiarz et al., 2019; Leong et al., 2019). For example, in Gesiarz et al. (2019), participants completed a sequential sampling task, in which they had to determine whether they were in a desirable state, associated with greater financial rewards than losses, or an undesirable state, associated with greater financial losses than rewards. Crucially, they had no control over which state they were in and were financially incentivized to accurately identify whether they were in a desirable or undesirable state. In spite of this, participants were biased towards concluding they were in a desirable state and required weaker evidence to reach a desirable compared to an undesirable conclusion. Computational modelling revealed that this desirability bias was due to participants placing more weight on evidence that aligned with their preferred (positive) belief (i.e., believing that they were in a desirable state), compared to evidence that opposed it. As a result, they were faster to reach a desirable, internally rewarding conclusion and continued to accumulate evidence for longer when it pointed to an undesirable conclusion, perhaps hoping to refute it. This tendency is also observed in the real world. For instance, some missed cancer diagnoses are due to doctors deciding not to conduct additional tests when the initial evidence suggests that the patient may not be ill (Lyratzopoulos et al., 2015). Neurocomputational evidence attributes this bias in evidence accumulation to a selective increase in neural activity for the desirable state,

which encompasses both anticipatory activity in the nucleus accumbens (nACC), as well as activity in sensory regions of the brain which track the accumulation of evidence towards the desired conclusion (Calabro et al., 2023; Leong et al., 2019). This suggests that internal incentives not only influence the way in which we accumulate evidence but also alter the neural representation of our current state. Simply put, people accumulate evidence in a way that fosters positive emotions and reduces negative emotions.

### <span id="page-14-0"></span>**1.1.3 A Bias towards Internal Incentives: Adaptation or Flaw?**

Prior work suggests that when accumulating evidence, internal incentives may take precedence over external incentives (Gesiarz et al., 2019; Leong et al., 2019). The tasks used in these studies were designed such that a bias towards desirable conclusions was financially suboptimal. Participants preferred forming positive beliefs from which they derived internal rewards, even if these beliefs caused them to incur financial losses. This can have detrimental consequences. For instance, it is thought that the tendency to prioritize internal incentives, such as positive emotions or increased sense of self-efficacy, has contributed to falsely optimistic investment decisions preceding the financial crisis (Shefrin, 2015), as well as a false sense of security which resulted in failure to take preventative measures against natural disasters (Paton, 2003). So why then do humans prioritize positive beliefs, and thus internal incentives over external incentives when they accumulate evidence?

Despite the potential negative repercussions, there are numerous benefits associated with forming positive beliefs even when they are false. It has been suggested that positive beliefs from which individuals derive internal rewards promote mental (Carver & Scheier, 2014; Taylor et al., 2000; Taylor & Brown, 1994), as well as physical well-being (Hernandez et al., 2015; Tindle et al., 2009). Patients who had positive expectations of their future, were less likely to develop coronary heart disease (Tindle et al., 2009), had a lower chance of rehospitalization after surgery (Scheier et al., 1999), and had higher survival rates after cancer diagnosis (Allison et al., 2003; Novotny et al., 2010). When individuals hold positive beliefs, they feel a greater sense of control and selfefficacy (McFarland & Ross, 1982). This motivates them to adopt measures to protect their well-being, such as engaging in regular physical exercise (Giltay et al., 2007; Steptoe et al., 2006). Positive beliefs have also been shown to reduce stress (Jobin et al., 2014), which, over time, improves wellbeing (Taylor et al., 2000; Taylor & Brown, 1994). By contrast, individuals with major depression tend to be more pessimistic about their future (Cropley & MacLeod, 2003; Strunk et al., 2006; Strunk & Adler, 2009). In light of their positive influence on physical and mental well-being, it has been argued, that positive beliefs are adaptive even when inaccurate (McKay & Dennett, 2009). For instance, if a patient erroneously believes that they have a high probability of fully recovering from a severe injury, this can lead to a more positive attitude and greater effort in the rehabilitation process. As such, accumulating evidence in a way that prioritizes the formation of positive beliefs ('biased evidence accumulation') and thus generates internal rewards, may too be adaptive, even when those beliefs are inaccurate.

When vying for resources, biased individuals may be more persistent and assertive in staking their claims on otherwise unattainable resources, while their more capable, yet less confident rivals might withdraw from the competition (Johnson & Fowler, 2011). Compared to unbiased individuals, those who overestimate their ability to succeed, may also be more likely to apply to challenging jobs, thereby increasing their chances of securing employment. In contrast, when the potential costs of erroneously held positive beliefs clearly exceed the benefits it is advantageous to be unbiased (Johnson & Fowler, 2011). In light of this, it has been suggested that whether individuals prioritize internal incentives over external incentives depends on the environment (Basten et al., 2010; Bromberg-Martin & Sharot, 2020; Garrett et al., 2018; Sharot & Garrett, 2016). The same may be true for evidence accumulation (**see Figure 1.1**). In environments where the benefits of being biased towards positive beliefs outweigh the costs, as is the case in safe environments, individuals place less weight on negative, undesirable evidence, which may be advantageous for their physical and mental wellbeing. By contrast, in environments where the costs of being biased towards positive beliefs clearly

outweigh the benefits, such as threatening environments, they shift their attention towards negative evidence and prioritize external incentives. Depending on the outcomes they obtain, they can then adjust their behaviour to maximize overall utility (Bond et al., 2023).

<span id="page-16-1"></span>

**Figure 1.1. Reward-Oriented Evidence Accumulation.** Humans accumulate evidence to form beliefs which motivate decisions and actions towards desirable external outcomes (e.g., monetary rewards). However, they are also motivated to accumulate evidence to form beliefs which give rise to desirable internal outcomes (e.g., positive emotions), sometimes even at the expense of external outcomes. Informed by the outcomes they obtain; they adaptively adjust their behaviour to maximize overall utility in a given environment.

### <span id="page-16-0"></span>**1.1.4 Changing the Incentives to Accumulate Evidence**

Thus far, it has been shown that internal incentives influence evidence accumulation, such that participants are biased towards desirable ( = internally rewarding) conclusions (Gesiarz et al., 2019; Leong et al., 2019). Prior research has also underscored the influence of contextual features on value-based decision-making (for a review see Engelmann & Hein, 2013). If individuals adaptively prioritize internal and external incentives based on their

environment, evidence accumulation should be less biased when false beliefs resulting in false conclusions are costly. Based on this assumption, this thesis investigates whether incentives for evidence accumulation can be altered to make evidence accumulation less biased. First, I will test the hypothesis that exposing participants to an environment in which the costs of false beliefs are high, they are less biased towards desirable conclusions. Then, I will investigate whether increasing the external incentives to provide correct responses can mitigate the influence of internal incentives on evidence accumulation.

#### Perceived Threat

In threatening environments, the costs of false beliefs can be especially high (Dunning, 2009; Haselton & Nettle, 2005; Johnson & Fowler, 2011). These environments elicit a physiological stress response which signals a high risksituation (Nesse et al., 2016). In these situations, selective sampling of positive, desirable evidence could be detrimental, as the potential for adverse outcomes is high. Rather, it is adaptive to err on the side of caution. Imagine, you are walking through a high-crime area at night, and you see someone walking closely behind you. It might just be a coincidence, but it could also be someone trying to accost you. Assuming the latter, you will likely change your route and as a result avert potential danger. It then follows that under threat individuals prioritize the accumulation of negative, undesirable evidence to pre-emptively mitigate potential losses, thereby reducing the bias in evidence accumulation towards desirable conclusions. This hypothesis aligns with prior research. When individuals are anxious, anticipated adverse consequences override the importance of possible desirable outcomes in risky decision-making (Engelmann et al., 2015). In line with this, Robinson et al., (2011) observed a threat-induced bias towards aversive stimuli in an emotional Stroop task in which participants had to identify appetitive and aversive faces. Stress reportedly increases attention to negative stimuli (Robinson et al., 2012; Robinson, Vytal, et al., 2013), thereby reducing preferential belief updating for desirable information (Garrett et al., 2018), and improving response inhibition (Robinson, Krimsky, et al., 2013). However, until now, the hypothesis that perceived threat mitigates biased evidence accumulation has not been tested.

In Chapter 2, I therefore examine how perceived threat impacts evidence accumulation. I combine a social-threat manipulation with a sequential sampling task in which participants have to determine whether they are in a desirable state, associated with greater financial rewards than losses, or an undesirable state, associated with greater financial losses than rewards. Using computational modelling, I then examine if and how acute stress alters the process by which evidence is accumulated when participants are motivated to reach desirable, internally rewarding conclusions.

#### Accuracy Incentives

If evidence accumulation becomes less biased in environments, in which false beliefs are costly, it is plausible, that directly increasing the external incentives for accurate beliefs might have a similar effect. The rationale for this is that when external incentives are tied to accuracy, the costs of false beliefs increase. One way this may be achieved, is by financially incentivizing people to make more accurate decisions. Historically, many companies have employed financial bonuses not only to improve overall performance but also as a targeted approach to diminish biases. For instance, Goldman Sachs implemented financial incentives in an attempt to reduce bias in hiring practices (AFR, 2015). This strategy is predicated on the understanding that biases and heuristics, can be moderated through deliberate and effortful thinking when performance-related rewards are at stake (Bonner & Sprinkle, 2002; Botvinick & Braver, 2015; Smith & Walker, 1993).

Yet, despite the intuitive appeal of this approach, prior research using financial accuracy incentives presents a conflicting picture (Dale et al., 2007; Engelmann et al., 2019; Epley & Gilovich, 2005; Lefebvre et al., 2011; Meub et al., 2013, 2013; Wright & Anderson, 1989, 1989; Zhang & Rand, 2023). On the one hand, accuracy incentives foster deliberation and (slow) rational thinking thereby reducing the reliance on cognitive shortcuts (Dale et al., 2007; Epley & Gilovich, 2005; Lefebvre et al., 2011; Meub et al., 2013; Wright & Anderson, 1989). On the other hand, there is some evidence to suggest that accuracy incentives are ineffective against motivational biases (Engelmann et al., 2019; Enke et al., 2023; Zhang & Rand, 2023; but see Prior et al., 2015; Rathje et al., 2023), which occur because individuals *want* to hold certain beliefs, such as wanting to believe they will be promoted (Montibeller & Von Winterfeldt, 2015).

In Chapter 3, I test a possible explanation for why financial accuracy incentives may fail to mitigate motivational biases. I hypothesize that financial incentives motivate participants to invest more cognitive effort. While this may aid in reducing the use of heuristics, which are the result of fast, careless processing; increased effort will fail to mitigate biases, when the cause of the error is a bias that is beyond participants' awareness, such as a bias in how evidence itself is processed (for example as when people put greater weight on desirable than undesirable evidence as shown in Gesiarz et al., 2019; Globig et al., 2021). To test this hypothesis, I examine the effect of varying accuracy incentives in a perceptual evidence accumulation task, in which participants have to determine whether they are in a desirable state, associated with financial rewards, or an undesirable state, associated with no rewards. Using computational modelling, I then examine if and how accuracy incentives alter the process by which evidence is accumulated when participants are motivated to reach a desirable, internally rewarding conclusion. The results will be particularly relevant in areas such as finance, politics, and healthcare, where biased evidence accumulation can have detrimental consequences and financial incentives are frequently used to alter behaviour (AFR, 2015; Fainman & Kucukyazici, 2020; Hasnain & Pierskalla Henryk, 2012).

Together, the results from Chapter 2 and 3 will enhance our understanding of internal and external incentives for evidence accumulation. The results can inform the development of interventions to prevent biased accumulation in situations in which it can have detrimental consequences.

### <span id="page-20-0"></span>**1.2 Reward-Oriented Information-Sharing**

As the reach of digital information technologies widens, it has become increasingly easy not only to accumulate evidence, but also to share information with others. Whether it is telling a friend about a new restaurant or sharing information about a disease outbreak, the social transmission of information has become a ubiquitous feature of everyday life. While this mounting ease of sharing has facilitated globalization and social connection, it has also given rise to an unprecedented explosion of misinformation (Lazer et al., 2018). This has had drastic consequences such as increased polarization, racism, and resistance to climate action and vaccines (Barreto et al., 2021; Rapp & Salovich, 2018; Tsfati et al., 2020; Van Bavel et al., 2021).

Unsurprisingly therefore, researchers, policymakers and industry leaders alike are working on mitigating the spread of misinformation (Saltz et al., 2021; Traberg et al., 2022). As yet, however, existing interventions to halt the spread, such as flagging misleading content, have only had limited impact (Chan et al., 2017; Grady et al., 2021; Lees et al., 2022). Here, I propose that to effectively improve sharing behaviour, a deep understanding of the underlying incentives driving sharing decisions is necessary. This type of 'from-theory-to-practice' approach has long been adopted in the medical sector, where researchers devote significant amounts of time and resources to investigating the pathways of health and disease in order to develop and optimize treatments. Such theoryinformed behavioural change interventions have been shown to be more successful in promoting health-related behaviour than non-theory-based interventions (Webb et al., 2010). Yet, despite frequently drawing analogies between infectious diseases and the spread of misinformation (Bonnevie et al., 2021; van der Linden, 2022; Zarocostas, 2020), over 70% of existing interventions are not informed by basic theory (Ziemer & Rothmund, 2024).

#### <span id="page-20-1"></span>**1.2.1 Incentives for Information-Sharing**

Classical reinforcement theory stipulates that humans seek actions that result in the greatest rewards and avoid those that lead to punishments (Skinner,

1966). Accordingly, they accumulate evidence to make decisions which result in positive outcomes. One such value-based decision is the decision to share information (Globig et al., 2023; Lin et al., 2022; Scholz et al., 2020).

#### External Incentives

Prior research shows that people share information with others to optimize external outcomes (Lindström et al., 2021; Scholz et al., 2017, 2020; Scissors et al., 2016). Often, these outcomes are intangible, in the form of social feedback (Barasch, 2020). This feedback can either be positive ('social rewards'), through social approval, reputational gains, or increased sense of social connection (Delgado et al., 2023) or negative ('social punishments'), through criticisms, reputational losses, or isolation (Brudner et al., 2023). Since the brain processes social feedback analogously to monetary outcomes (Bhanji & Delgado, 2014; Gu et al., 2020; Meshi et al., 2013; Ruff & Fehr, 2014; Wake & Izuma, 2017), and the receipt of social rewards activates reward-regions of the brain including the ventral striatum and the ventromedial prefrontal cortex (VMPFC, Davey et al., 2010; Klucharev et al., 2009; Morelli et al., 2014), it is no surprise that social feedback has been shown to be just as, if not more, effective in directing human action (for a review see Tamir & Hughes, 2018). For instance, humans are more likely to share information with others if they believe it will elicit positive feedback, and choose to abstain from doing so if they fear it will elicit negative feedback (Brudner et al., 2023). Social media platforms have capitalized on this sensitivity to incentives, by implementing social rewards, such as 'likes' and 'reposts' to maximize user engagement. More recently, they have also implemented monetization schemes, which financially reward some users whose sharing behaviour generates particularly high engagement (Alizadeh et al., 2023). In these cases, intangible rewards, in the form of social feedback, bring about tangible rewards in the form of monetary gains. Thus, social media users are motivated to share information with others in order to maximize the rewards they receive. The more responsive they perceive their social network to be, the more they share (Walsh et al., 2020). Over time, users learn what elicits positive or negative reactions and thus adjust their sharing behaviour to increase the likelihood of positive

feedback (Brady et al., 2021; Lindström et al., 2021; Scissors et al., 2016). As such, information-sharing is motivated by external incentives (**see Figure 1.2**).

#### Internal Incentives

On top of these external incentives, it has also been suggested that sharing may be rewarding in itself (Baek et al., 2017; Tamir et al., 2015; Tamir & Mitchell, 2012). For example, consider sharing your experience about a recent unsuccessful job interview. Disclosing this experience may elicit reactions of social support, but it can also help you regulate your emotions by allowing you to reflect on the interview (Berger, 2014; Rimé, 2009). This may help you make sense of the experience. For instance, through self-reflection you may realize that the unsuccessful interview was not a result of your performance, but rather because the job did not match your interests, thereby bolstering your perceived self-efficacy and mitigating negative emotions (Niederhoffer & Pennebaker, 2009). Indeed, a recent study shows that information-sharing processes involve the activation of self-related-processing regions of the medial prefrontal cortex (MPFC) and the posterior cingulate cortex (PCC, Baek et al., 2017), suggesting that self-reflection is a key motivation for information-sharing. It then follows that sharing decisions are likely motivated not only by external incentives but also by internal incentives (Chen et al., 2019; Fu et al., 2017; Rode, 2016, **see Figure 1.2**).

### <span id="page-22-0"></span>**1.2.2 Sharing as a Value-Based Decision**

Building on this work, sharing has been characterized as a value-based decision, that involves both self-related and other-related considerations (Scholz et al., 2020, 2023). Sharing decisions elicit activity in regions associated with positive valuation, including the ventral striatum and VMPFC (Berger, 2014; De Angelis et al., 2012; Lee & Ma, 2012; Wien & Olsen, 2014). Activity in these regions scales with how enthusiastically messages are shared (Falk et al., 2012). Echoing these findings, this thesis conceptualizes sharing decisions, as the result of reward-oriented evidence accumulation, in which sharers adjudicate between, sometimes competing, internal and external

incentives (**see Figure 1.2**), weighing the costs against the benefits throughout. In some cases, they favour internal incentives, as evidenced by research reporting that participants choose to forgo monetary rewards in order to disclose information about themselves (Tamir & Mitchell, 2012). In others, the prospect of external rewards, such as social capital may override internal rewards (Fu et al., 2017). This is especially pertinent in online contexts, where shared information can reach a larger audience and thus elicits a greater volume of feedback, i.e., social incentives (Bodaghi & Oliveira, 2020). Over time users learn what type of information elicits the greatest amount of rewards and thus adaptively adjust their behaviour to maximize overall utility (Brady et al., 2021; Lindström et al., 2021; Scissors et al., 2016). Sharing that elicits rewards is reinforced, sharing that elicits punishments is suppressed. As, external rewards tend to be easier to quantify than internal rewards and given that social incentives are already a core feature of existing social media platforms, the reinforcing feedback loop between sharing and social incentives provides a bullseye for measures designed to improve sharing behaviour online.



<span id="page-23-1"></span>**Figure 1.2 Reward-Oriented Information-Sharing.** Individuals share information to obtain external rewards (e.g., positive social feedback and/or monetary rewards). However, they are also motivated to share information which gives rise to internal rewards (e.g., positive emotions, self-efficacy). Informed by the outcomes they obtain and weighing the costs against the benefits throughout, they adaptively adjust their behaviour to maximize overall utility.

### <span id="page-23-0"></span>**1.2.3 Shaping Sharing Decisions with Social Feedback**

The reinforcement-model like pattern of online behaviour, in which users respond to social rewards and punishments (Brady et al., 2021; Lindström et al., 2021; Rosenthal-von der Pütten et al., 2019; Scissors et al., 2016) may also

explain the apparent disconnect between what people believe themselves and what they share online (Pennycook et al., 2021; Ren et al., 2021). For instance, users may choose to abstain from sharing content they believe will elicit negative reactions and instead post content which they believe will maximize engagement (Brady et al., 2021; Lindström et al., 2021; Scissors et al., 2016). Notably, it has been shown that misinformation often generates more engagement than reliable posts (Lazer et al., 2018). As such, users have little reason to take into account the veracity of information when deciding what type of content to share. Instead, they are incentivized to share the type of content that they anticipate will elicit the largest amount of positive engagement, which sometimes translates to monetary rewards. Thus, both social (e.g., 'likes', 'shares') and monetary incentives (e.g., monetization schemes) contribute to spread of misinformation online (Alizadeh et al., 2023).

With time, this not only vastly accelerates the spread of misinformation, but may also alter users' beliefs. Empirical evidence suggests that when people learn that others share their belief, their confidence in those beliefs increases (Kappes et al., 2020). Moreover, it has been shown that even a single repetition of misinformation leads to it being perceived as more accurate, irrespective of whether it aligns with the user's ideology (Murray et al., 2020; Pennycook et al., 2018). This 'illusory truth effect' (Hasher et al., 1977) then in turn influences sharing, suggesting a bi-directional relationship of sharing and beliefs (Van Bavel et al., 2021; Vellani et al., 2023). Thus, the current incentive structure of social media platforms, in which incentives, such as rewards in the form of 'likes' and punishments in the form of 'dislikes' are dissociated from the veracity of the information shared, facilitates the dissemination of misinformation, and may also fuel the spread of erroneous beliefs (see also Epstein et al., 2023). Users strive to maximize external rewards, without being motivated to consider whether the information they are sharing is accurate.

In Chapter 4, I suggest that to mitigate the spread of misinformation, a modified incentive structure is needed in which rewards and punishments are directly contingent on the veracity of the information shared. To test this assumption, I examine the impact of slightly altering the engagement options offered to users. Specifically, I add an option to react to posts using 'trust' and 'distrust' buttons, which are, by definition, related to veracity. I then examine how this change affects discernment in sharing behaviour between true and false information and if it can help in correcting false beliefs. Using computational modelling I dissect the mechanisms underlying this change.

## <span id="page-25-0"></span>**1.3 Drift-Diffusion Modelling: A Window into the Mechanisms of Decision-Making**

A common issue of behavioural studies is the difficulty to obtain insights into the mechanisms underlying observed behaviour (Roberts & Hutcherson, 2019). In many studies, directly observable variables are limited to choices and response times. Notably, these variables are aggregate outcomes and may mask the underlying processes (Clithero, 2018; Stafford et al., 2020). For example, slower response times can be the result of more tenuous cognitive processing, and/or a higher threshold for decision certainty, reflecting a more cautious decision-making process (Forstmann et al., 2016; Voss et al., 2013). Unsurprisingly therefore, a growing number of researchers are shifting focus towards computational models (Calder et al., 2018). By formalizing observed behaviours as mathematical frameworks, these models allow researchers to examine the underlying mechanisms of decision-making (Wilson & Collins, 2019).

For instance, Drift-Diffusion Models (DDM), a subtype of sequential sampling models, describe the process by which individuals accumulate evidence to make a decision (Ratcliff et al., 2016), for example to share information (Globig et al., 2023; Lin et al., 2023). They model the decision process between two (or more) alternatives as the noisy accumulation of evidence over time (Ratcliff, 1978; Ratcliff & Rouder, 1998; Voss et al., 2013, see Roxin, 2019 for multialternative DDMs). In doing so, they assume evidence accumulation is driven by random fluctuations, i.e., diffusion, to mimic the gradual integration of sensory information in the brain. More specifically, DDMs assume that agents

accumulate evidence at a given rate until the evidence for one alternative relative to the other alternative(s) is large enough to reach a pre-determined decision threshold and a decision is made (Forstmann et al., 2016; Ratcliff et al., 2016). This evidence can be external, e.g., from the stimuli observed, or internal, based on internal representations of the alternatives (Krajbich, 2019).

Conventionally, DDMs include the following parameters: (1) the drift rate (v) – which is the rate at which evidence is accumulated; (2) the distance between decision thresholds  $(\alpha)$  — which captures the amount of evidence required to form a decision; (3) the starting point (z) of the accumulation process; and (4) the amount of non-decision time (t0)—which includes stimulus encoding as well as response preparation and execution. Each of these parameters can be fixed based on prior assumptions or can be allowed to vary freely, for example as a function of different variables, such as trial condition or trial difficulty. These parameters have been found to correspond with specific neural and physiological patterns (Basten et al., 2010; Cavanagh, Wiecki, et al., 2011; Krajbich et al., 2010; Mulder, 2014). This underscores the potential of DDMs to successfully capture and provide insights into neural processes.

Taken together, DDMs therefore offer a viable, cost-efficient, and complementary solution to gather insights into decision-making processes (Berlinghieri et al., 2023). They can help identify potential target mechanisms for measures designed to reduce biases (Arkes, 1991; Krajbich, 2022). In this thesis I capitalize on the potential of DDM to dissect the mechanisms underlying the influence of incentives on evidence accumulation and sharing decisions.

### <span id="page-26-0"></span>**1.4 Summary**

When navigating today's information ecosystem, individuals are perpetually faced with opportunities to accumulate evidence and share information. This behaviour is motivated by (1) external incentives, such as financial or social rewards and punishments (Gold & Shadlen, 2002; Rosenthal-von der Pütten et al., 2019), and (2) internal incentives, such as positive and negative emotions

(Baek et al., 2017; Gesiarz et al., 2019; Leong et al., 2019; Tamir et al., 2015; Tamir & Mitchell, 2012). Sometimes, these incentives can steer individuals towards decisions that yield adverse consequences, such as overly optimistic investment decisions (Barber & Odean, 1999; Shefrin, 2015), or sharing of misinformation (Pennycook et al., 2021). While prior research has sought to understand how different motivations influence evidence accumulation (Gesiarz et al., 2019; Hausmann & Läge, 2008; Kelly & O'Connell, 2013; Leong et al., 2019) and information-sharing decisions (Baek et al., 2017; Rode, 2016; Scholz et al., 2020), how and when changing these incentives can reduce biased evidence accumulation and increase the sharing of accurate information has been largely unexplored and is the focus of this thesis.

Individuals accumulate evidence to arrive at desirable, internally rewarding conclusions, even at the expense of external outcomes (Gesiarz et al., 2019; Leong et al., 2019). By prioritizing internal incentives, they discount undesirable, negative evidence. As a result, they are biased towards internally rewarding conclusions (Gesiarz et al., 2019; Leong et al., 2019). This bias can contribute to negligent decision-making, as for instance observed during the financial collapse of 2008 (Shefrin, 2015). To prevent such negative ramifications, it is imperative to understand how and when biased evidence accumulation can be mitigated.

It has been suggested that individuals adaptively prioritize internal and external incentives depending on their environment (Dunning, 2009; Haselton & Nettle, 2005; Johnson & Fowler, 2011). In Chapter 2, I test the hypothesis that when individuals find themselves in a high-threat environment, in which the costs of false beliefs are high, they are less biased towards desirable conclusions. To that end, I combine a social-threat-manipulation with a sequential sampling task and DDM to understand how threat alters evidence accumulation.

Intuitively, we assume that financially incentivizing individuals to provide correct responses will enhance the accuracy of their decisions. In Chapter 3, I test an explanation for why financial accuracy incentives may fail to mitigate the

influence of internal incentives on evidence accumulation. I tease apart the underlying mechanisms by combining an incentivized perceptual evidence accumulation task with DDM.

Together, these findings will enhance our understanding of the role of internal and external incentives in evidence accumulation and can aid in the development of interventions to prevent the potential negative ramifications of biased evidence accumulation.

One example of a decision based on evidence accumulation is the decision to share information with others. Over five billion people globally are now active on social media platforms (Statista, 2022), making them a vital source of information for many (Pew Research Center, 2021). Ensuring this information is reliable is therefore of utmost importance. Much of the success of social media platforms has been attributed to the human need for rewards (Rosenthalvon der Pütten et al., 2019). Over time, users learn what type of information elicits the greatest amount of positive feedback (e.g., 'likes' and 'shares') and adjust their sharing behaviour accordingly (Brady et al., 2021; Lindström et al., 2021; Scissors et al., 2016). But as engagement is not tied to the veracity of information, and misinformation often elicits more engagement than reliable information, users are not motivated to discern between the two. I thus argue that the spread of misinformation is facilitated by the existing incentive structure of social media platforms. In Chapter 4, I explore whether a slight change to this structure, such that social rewards and punishments are contingent on the veracity of information can overturn the adverse effects of incentives on sharing behaviour and thereby reduce the spread of misinformation online. I provide a mechanistic account of how this change affects discernment between true and false posts, by modelling sharing decisions as a drift-diffusion process.

In conclusion, my research will shed new light on how incentives to accumulate evidence and share information can be modulated to enhance discernment in decision-making. The results not only expand our understanding of how incentives impact evidence accumulation and sharing but also highlight the need for innovative approaches in policy and decision-making strategies. Ultimately, this work paves the way for more informed and nuanced applications of behavioural interventions in various domains ranging from public policy to individual decision-making processes.

## <span id="page-30-0"></span>**Chapter 2: Under Threat Weaker Evidence is Required to Reach Undesirable Conclusions**

### <span id="page-30-1"></span>**2.1 Abstract**

Critical decisions, such as in domains ranging from medicine to finance, are often made under threatening circumstances that elicit stress and anxiety. The negative effects of such reactions on learning and decision-making have been repeatedly underscored. In contrast, here we show that perceived threat alters the process by which evidence is accumulated in a way that may be adaptive. Participants ( $N = 91$ ) completed a sequential sampling task in which they were incentivized to accurately determine whether they were in a desirable state, which was associated with greater rewards than losses, or an undesirable state, which was associated with greater losses than rewards. Prior to the task participants in the 'threat group' experienced a social-threat manipulation. Results show that perceived threat led to a reduction in the strength of evidence required to reach an undesirable conclusion. Computational modelling revealed this was due to an increase in the relative rate by which negative evidence was accumulated. The effect of the threat manipulation was global, as the alteration to evidence accumulation was observed for evidence which was not directly related to the cause of the threat. Requiring weaker evidence to reach undesirable conclusions in threatening environments may be adaptive as it can lead to increased precautionary action.

### <span id="page-30-2"></span>**2.2 Introduction**

Many important decisions are made when people feel stressed and anxious (Beilock, 2010). Consider a doctor in the operating theatre who needs to decide on the best course of action, a soldier on the battlefield who must decide whether to attack, or a driver stuck in traffic selecting which route to take. Whether calm or stressed, to make good decisions people need to accumulate evidence over time (Forstmann et al., 2016; Platt & Glimcher, 1999; Ratcliff,

1978; Usher & McClelland, 2001). For example, a doctor may decide to consult multiple colleagues before deciding to amputate. Because evidence can be unlimited, an agent needs to determine when the available data is strong enough to reach a conclusion (Gluth et al., 2012, 2013). Here, we examine how perceived threat impacts the process by which evidence is accumulated to reach a conclusion.

A feature of threatening environments is that the potential for adverse outcomes is high. In these instances, it is adaptive to err on the side of caution. For example, imagine you are walking through a dark alley and hear a 'pop'. The sound may be a gunshot or perhaps the uncorking of a champagne bottle. Interpreting the sound as the former will cause you to escape and mitigate potential risk. Thus, under perceived threat it may be adaptive to interpret a stimulus as undesirable even if the strength of the evidence supporting this conclusion is only limited. The psychophysiological reaction induced by threat can provide a global, rather than specific, danger signal. We thus hypothesized that the effects of threat on evidence accumulation may be observed even when the source of the threat is unrelated to the decision at hand (e.g., a psychophysiological reaction triggered by a professional conflict may impact how the 'pop' is interpreted).

Computationally, this process may occur in at least two ways. First, under perceived threat people may be predisposed towards undesirable conclusions before attaining any evidence (e.g., you may believe the road you are walking down is dangerous before observing any evidence to that effect). A second, not mutually exclusive possibility, is that under perceived threat an undesirable piece of evidence (e.g., an anxious looking man walking down the road) drives beliefs towards an undesirable conclusion ('this road is dangerous'), more so than a desirable piece of evidence (e.g., people are walking past you relaxed and happy) towards a desirable conclusion ('this road is safe'). These two distinct mechanisms will result in the same observable behaviour. Specifically, weaker evidence will be needed to support undesirable conclusions under perceived threat.

To tease apart these mechanisms, we used a sequential sampling model to model noisy evidence accumulation towards either of two decision thresholds (Ratcliff, 1978; Ratcliff & Rouder, 1998; Voss et al., 2013). The model allows estimation of both (1) starting point, and (2) rate of evidence accumulation, reflecting the quality of information processing. We can then measure whether either of these factors are influenced by the desirability of a decision and how this is influenced by perceived threat.

We exposed participants to an acute threat manipulation in the lab (Garrett et al., 2018) or a control condition, and then asked them to complete an evidence accumulation task (Gesiarz et al., 2019) that was unrelated to the cause of the threat. In the task, participants witnessed various stimuli that were contingent upon which one of two hidden states they were in. One state was associated with greater rewards than losses (desirable state) and the other with greater losses than rewards (undesirable state). Participants had no control over which state they were in; their task was simply to determine the state, gaining additional rewards for accurate conclusions and losing rewards for inaccurate conclusions. Thus, it was in participants' interest to be as accurate as possible. They were allowed to accumulate as much evidence as they wished before making a decision. Here, we examined if and how perceived threat impacts the accumulation of evidence towards a decision.

### <span id="page-32-0"></span>**2.3 Methods**

#### **Experimental Design.**

**Participants**. A total of 91 individuals participated in this study at two sites: University College London (UCL,  $N = 51$ ) and Massachusetts Institute of Technology (MIT,  $N = 40$ ). They were recruited via the participant pools of UCL and MIT. All analyses were repeated separately for participants tested in the two different locations (MIT, UCL). There were no differences between locations in any of our results including model-free analysis, psychometric equations and DDM.

Participants gave written, informed consent and were remunerated £7.50/\$15 for their participation plus an unspecified performance-related bonus. Ethical approval was provided by the Research Ethics Committees at UCL and MIT. Both experiments were performed in accordance with the principles expressed in the Declaration of Helsinki. One participant who terminated the experiment early and another who failed all comprehension checks were excluded from the analysis. In addition, we followed the exclusion criteria previously published for this task (Gesiarz et al., 2019): we excluded two participants whose accuracy was below chance (50%) and four who provided responses based only on the first stimulus in over half the trials. Thus, data of 83 participants was included in the analysis (M<sub>age</sub> = 30.29,  $SD<sub>age</sub> ± 12.20$ ; female = 37, male = 46, UCL = 43,  $MIT = 40$ ). Each participant was randomly assigned to either the threat manipulation group ( $N = 40$ ,  $M_{age} = 28.98$ ,  $SD_{age} \pm 11$ ; female = 14, male = 26, UCL = 21 and MIT = 19) or the control group (N = 43, M<sub>age</sub> = 31.51, SD<sub>age</sub>  $\pm$ 13.23; female = 23, male = 20; UCL = 22, MIT = 21).

**Manipulation Procedure and Manipulation Check**. We followed the exact same threat manipulation as in Garrett et al., (2018). Participants assigned to the 'threat group' were informed that at the end of the experiment they would be required to deliver a speech on a surprise topic, which would be recorded on video and judged live by a panel of staff members. They were shown an adjacent room where chairs and tables were already organized for the panel. This manipulation is a variation of the Trier Social Stress Test (Birkett, 2011) with the key difference being that participants in this task were threatened by the possibility of a stressful social event and completed the main task under anticipation of the threat, but the threat was never executed. Having participants believe the threatening event would take place at the end of the task, rather than before, increased the likelihood that participants' anxiety levels remained high throughout the task. In addition, participants were presented with six difficult mathematical problems that they were asked to try and solve in 30 seconds (s). This exact manipulation procedure was previously executed in our

lab, and has been shown to significantly heighten cortisol levels, skin conductance and self-reported state anxiety (Garrett et al., 2018). Our lab has also shown that the manipulation-induced changes in self-reported state anxiety (measured using the Spielberger State Trait Anxiety Inventory, STAI, Marteau & Bekker, 1992) correlated across participants with physiological indicators of stress (Garrett et al., 2018).

Participants assigned to the 'control group' were informed that at the end of the experiment they would be required to write a short essay on a surprise topic, which would not be judged. They were then presented with six elementary mathematical problems to solve in 30s. This control manipulation has been shown not to heighten cortisol levels, skin conductance and self-reported state anxiety (Garrett et al., 2018). As a manipulation check, before and after the induction procedure, we asked participants to complete the STAI (Marteau & Bekker, 1992) as a measure of anxiety.

**Behavioural Task.** After completing the threat/control manipulation, participants played 80 trials of the 'Factory Task' (Gesiarz et al., 2019, **see Figure 2.1**). On each trial participants witnessed an animated sequence of televisions (TVs) and telephones passing along a conveyor belt. There were two types of trials: telephone factory trials and TV factory trials. In telephone factory trials, the probability of each item in the animated sequence being a telephone was 0.6, and the probability of each item in the animated sequence being a TV was 0.4. For TV factory trials the proportions were reversed. The trial type was randomly determined with replacement on every trial with an equal probability for each trial type. Participants were tasked with determining whether they were in a telephone factory trial or a TV factory trial. Since the trial type was not directly observable, their means of doing this was through reverse inference over the sequence of objects they were seeing. Participants were free to respond as soon as they wished after initiating the trial and the sequence would continue until they made their decision.

Participants began the game with an endowment of 5000 points. Each 100 points was worth 1 pence/1 cent. One of the two factory types was randomly assigned per participant to be the 'desirable' factory type and the other to be the 'undesirable' factory type. Participants were informed that each time they visited the desirable factory (desirable state), they would win points, and each time they visited the undesirable factory (undesirable state), they would lose points. We did not specify the exact number of points they would win or lose. Crucially, this bonus was entirely outside of the participants' control, i.e., it was not affected by the conclusions participants made. Separately, participants were informed that they would earn an unspecified number of points for making a correct conclusion and lose an unspecified number of points for making an incorrect conclusion. We informed participants that the magnitudes of each unspecified bonus/loss were independent of each other, potentially unequal and varied randomly on each trial.

The task was the same as published previously (Gesiarz et al., 2019), except that we jittered the presentation time of the stimuli, so that participants were less likely to have a clear expectation of when the next stimulus would be observed. Due to a technical error this jitter was slightly different across sites (average stimuli presentation time at UCL:  $M = 657.28$  milliseconds (ms), SD  $\pm$ 1060.73ms; MIT:  $M = 373.65$ ms,  $SD \pm 49.69$ ms). The lag between stimuli was  $~150$ ms.

Trials in which participants made their decision before observing the second object were removed. In cases where a participant did this in over half their trials, we assumed that the participant was not appropriately engaging with the task and eliminated the entirety of their trials. Following Gesiarz et al., (2019), we excluded four participants for this reason, as well as a further 72 responses made before seeing the second item.

**Training.** Prior to playing the task, participants received extensive instructions and were required to answer multiple-choice comprehension check questions on the key points of the task, with the question repeated until they either chose
correctly or failed three times, upon which the correct answer was displayed. The comprehension check questions addressed the following key points of how the game worked: that telephone factories mostly produced telephones, but sometimes produced TVs; that the bonus for visiting desirable factories was independent of the decision they made; which factory was their desirable factory; and that trial types (i.e., if they were in a TV or telephone factory) were randomly determined and it was not guaranteed that they would see exactly the same amount of each type of factory. Participants then played a practice session of 20 trials, where the trial type was visibly displayed to them (i.e., if they were in a TV or telephone factory), so they could have prior experience of the outcome contingencies and the trial type distribution.



**Figure 2.1 The 'Factory Task'.** In each trial, participants saw an animated sequence of TVs and telephones passing along a conveyor belt. Their task was to accurately determine whether they were in (a) a telephone factory, i.e., a factory that produces telephones most of the time or (b) a TV factory, i.e., a factory that produces TVs most of the time. They were incentivized for accuracy and could enter their decision whenever they liked. Each participant had 'invested' in one factory. On trials where they happened to be in that (desirable) factory they gained points, on trials in which they happened to be in the other (undesirable) factory they lost points. Notably, this bonus was beyond participants' control and was not affected by the actual decision made. Stimulus presentation time was jittered, so that participants were less likely to have a clear expectation of when the next stimulus would be observed. Stimulus presentation time on average was ~521ms. The lag between stimuli was on average ~150ms.

#### **Statistical Analysis.**

**Manipulation Check.** An independent two-tailed t-test was computed to assess the difference in percentage change in STAI ((post STAI - pre STAI)/pre STAI) between the threat and control group. One-sample t-tests were computed to assess percentage change against zero within each group. All statistical tests conducted in the present article are two-sided. Analysis was conducted using IBM SPSS 27 and R Studio (Version 1.3.1056).

**Psychometric Function.** We followed the same analysis as in Gesiarz et al., (2019) to relate participants' responses to the strength of evidence they observed. We fitted a psychometric function, using a generalized mixed effects equivalent of a logistic regression, with fixed and random effects for all independent variables. We fitted these functions separately for participants for whom the TV factory was desirable and for whom the TV factory was undesirable, and separately for each group (control, threat).

$$
P(TV) \; = \; \frac{1}{1 + e^{-(\beta_1 X - \beta_0)}}
$$

Where *P(TV)* is the probability of a participant indicating they were in a TV factory; *X* is the proportion of TV stimuli out of all stimuli observed in a trial. This variable was centred, thus ranging from 0.5 when all samples were TVs to -0.5 when all samples were phones; *β<sup>0</sup>* is the indifference point – reflecting the proportion of TVs required to respond TV 50% of the time. If  $\beta_0 = 0$ , participants would indicate they were in a TV factory half the time when half the samples were TVs. When  $β_0$  is low, the function will move left and vice versa.  $β1$  is the slope, reflecting by how much the probability of a participant indicating they were in a TV factory increases when the proportion of TVs increases by one unit.

**Pieces of evidence gathered.** We examined whether the total number of pieces of evidence (TVs + telephones) differed when participants reached an undesirable or desirable conclusion (within-subject variable) and/or depending on group (between-subject variable). Number of pieces of evidence gathered before responding was entered into a mixed 2 (group: control/threat) by 2 (valence of conclusion: desirable/undesirable) ANOVA. We also allowed for an interaction of group and valence of conclusion.

**Drift-Diffusion Modelling.** Our aim in modelling our task using the Drift-Diffusion framework was to assess how perceived threat impacted the evidence accumulation process. In particular, we wanted to assess (1) whether the evidence accumulation process in the threat and control groups was best represented by the same model or a different model, and (2) whether perceived threat impacted the parameters of the evidence accumulation process in our data.

We implemented and compared four different specifications of a DDM (**see Table 2.1**). The models included the following parameters: (1) t0—amount of non-decision time; (2) α—distance between decision thresholds; (3) z—starting point of the accumulation process; and (4) v—drift rate - the rate of evidence accumulation. Crucially, in models 1 and 3 the starting point was fixed to 0.5, while in models 2 and 4, we allowed the starting point to vary towards one threshold (its value could vary between 0 and 1, thus allowing a valencedependent starting point bias). In models 1 and 2 with an unbiased drift rate, the parameter was symmetric for desirable and undesirable factories (v and v). In models 3 and 4 we allowed the drift rate to vary (which we call a valencedependent drift rate bias) depending upon whether the participant was visiting a desirable factory or an undesirable factory (thus allowing a process bias). In these models we included a term reflecting the difference between drift rates for desirable and undesirable factories (β1factory desirability). 'Factory desirability'—is the true factory visited coded as 1 for desirable factories and 0 for undesirable factories. Positive values indicated a bias towards desirable conclusions, and negative values indicated a bias towards undesirable conclusions. β0 is a constant for the drift rate.





**Table 2.1. DDM Specification.** For each group we ran four models which differed in whether we allowed the starting point to vary (model 2 & 4), whether we included a valence-dependent drift rate bias (model 3 & 4), or neither (model 1).

We used the HDDM software toolbox (Wiecki et al., 2013) to estimate the parameters of our models. The HDDM package employs hierarchical Bayesian parameter estimation, using Markov Chain Monte Carlo (MCMC) methods to sample the posterior probability density distributions for the estimated parameter values. We estimated both group-level parameters as well as parameters for each individual participant. Parameters for individual participants were assumed to be randomly drawn from a group-level distribution. Participants' parameters both contributed to and were constrained by the estimates of group-level parameters. In fitting the models, we used priors that assigned equal probability to all possible values of the parameters. Models were fit to log-transformed response times (RTs) as done previously (Gesiarz et al., 2019), because RTs were non-normally distributed and had a heavy positive skew. Also, since our 'error' RT distribution included relatively fast errors we included an inter-trial starting point parameter (sz) for both models to improve model fit (Ratcliff & Rouder, 1998). We sampled 20,000 times from the posteriors, discarding the first 5,000 as burn in and thinning set at 5. MCMC are guaranteed to reliably approximate the target posterior density as the number of samples approaches infinity. To test if the MCMC converged within the allotted time, we used Gelman-Rubin statistic (Rubin & Gelman, 1992) on 5 chains of our sampling procedure. The Gelman–Rubin diagnostic evaluates MCMC convergence by analysing the difference between multiple Markov chains. The convergence is assessed by comparing the estimated betweenchains and within-chain variances for each model parameter. In each case, the Gelman-Rubin statistic was close to one (<1.1), suggesting that MCMC were able to converge. To assess if the parameters describing the bias in prior and drift rate are significantly different in the control and threat group, we compared

95% Confidence Intervals (CIs) of the parameters' values. Specifically, for each parameter in each group we calculated the 95% CIs. If the 95% CIs for a parameter between groups did not overlap, we consider there to be a significant difference. We also calculated the difference in the posterior distributions and reported the 95% Highest Density Interval (HDI) of the difference. If this HDI did not overlap zero, we considered there to be a meaningful difference between the two groups. HDI testing was conducted in R using *HDInterval* (Meredith & Kruschke, 2016).

In addition, model fits were compared using the Deviance Information Criterion (DIC, Spiegelhalter et al., 2002), which is a generalization of the Akaike Information Criterion (AIC) for hierarchical models. The DIC is commonly used when the posterior distributions of the models have been obtained by MCMC simulation (Gamerman & Lopes, 2006). It allows one to assess the goodness of fit, while penalizing for model complexity.

To validate the winning model, we used each group's parameters obtained from participants' data to simulate log-transformed RTs and responses separately for the threat and control group. We used the exact number of participants, total number of trials and trial structure as in the experiment. Simulated data was then used to (1) perform model recovery analysis and (2) to compare the pattern of participants' response to the pattern of simulated responses, separately for each group. We sampled 2,000 times from the posteriors, discarding the first 500 as burn in. Simulation and model recovery analysis were performed using the HDDM software toolbox (Wiecki et al., 2013).

**Proportion of correctly identified factories.** We examined whether group (between-subject variable) and valence of factory visited (within-subject variable) affected the proportion of correctly identified factories as desirable or undesirable. Proportions were calculated for each participant and then entered into a mixed 2 (group: control/threat) by 2 (valence of factory: desirable/undesirable) ANOVA. We also allowed for an interaction of group and

valence of factory. We compared the pattern of results obtained from participants' real data to those obtained from the simulated data.

## **2.4 Results**

**Threat manipulation was successful.** The manipulation was successful in inducing perceived threat. Participants in the threat group reported a significantly larger increase in anxiety as a result of the manipulation (increase in STAI score after the manipulation relative to before  $M = 40.815\%$ , SE = 5.928,  $t(39) = 6.885$ ,  $p < 0.001$ , Cohen's  $d = 1.089$ ), compared to those in the control group, who in fact showed a reduction in anxiety ( $M = -5.742\%$ , SE = 1.257,  $t(42) = 4.568$ ,  $p < 0.001$ , Cohen's d = 0.697, difference between the two groups: t(81) = 7.943, p < 0.001, Cohen's d = 1.715, **Figure 2.2**) an effect often observed in control participants, who tend to relax as they learn more about the task at hand (Garrett et al., 2018).



**Figure 2.2. Threat manipulation was successful.** Participants in the threat group became significantly more anxious after the manipulation than in the control group. Data are plotted as boxplots for each condition, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5 x interquartile range. Diamond shape indicates

the mean percentage change in STAI per experimental group. Individuals' percentage STAI change are shown separately as grey dots. \*\*\*p < 0.001

**Under threat participants required weaker evidence to conclude they are in an undesirable factory.** We first examined whether perceived threat alters the strength of evidence participants required to reach desirable and undesirable conclusions. To that end, we fit a psychometric function to the data which relates the percentage of TVs observed on a trial (i.e., the strength of the evidence to judge a factory as a TV factory) to participants' conclusion about whether they were visiting a TV or telephone factory. This was done separately for participants for whom the TV factory was desirable and for whom it was undesirable in the threat and control group.

As observed in **Figure 2.3a**, under perceived threat the psychometric function of participants for whom the TV factory was undesirable (solid orange line) was shifted left compared to controls (dotted orange line). This means that compared to control, under threat participants required a smaller proportion of TVs to be observed before reaching the conclusion that they were in a TV factory when the TV factory was undesirable (indifference parameter was higher for the threat group:  $β0 = 0.113$ ,  $95%$  CI [-0.185, 0.411], than controls: β0 = -0.582, 95% CI [-0.964, -0.20], Cohen's d = 0.61, **Figure 2.3a**). No such difference is observed when the TV factory is desirable. Participants in both groups required an equal proportion of TVs to be observed before reaching the conclusion that they were in a TV factory. This can be seen in **Figure 2.3b** where the psychometric functions for threat and control participants overlap (indifference parameter was not different for the threat group:  $\beta$ 0 = 0.277, 95% CI [-0.068, 0.622] and control group: β0 = 0.23, 95% CI [-0.079, 0.53], Cohen's d = 0.05, **Figure 2.3b**).

While participants in the control group required weaker evidence to conclude they were in a desirable factory than an undesirable factory (replicating previous findings from Gesiarz et al., 2019), this difference was abolished under perceived threat. This can be observed where the psychometric function of control participants for whom the TV factory was desirable (dotted blue line, **Figure 2.3c**) is shifted to the left of control participants for whom the TV factory was undesirable (dotted orange line, **Figure 2.3c**; indifference parameter was greater for desirable factory:  $β0 = 0.23$ ,  $95%$  CI [-0.079, 0.53], than undesirable: β0 = -0.582, 95% CI [-0.964, -0.20], Cohen's d = 0.694, **Figure 2.3c**), while for participants in the threat group they overlap (indifference parameter when the TV factory was desirable β0 = 0.277, 95% CI [-0.068, 0.622] and undesirable β0 = 0.113, 95% CI [-0.185, 0.411], Cohen's d = 0.16, **Figure 2.3d**).

As expected, both in the threat and control group the greater the proportion of TVs in a trial the more likely participants were to judge the factory as a TV factory (control: TV factory desirable:  $β1 = 25.55$ , 95% CI [23.20, 27.90], TV factory undesirable: β1 = 24.943 [15.623, 34.262], Cohen's d = 0.023, **Figure 2.3c**; threat: TV factory desirable: β1 = 27.79, 95% CI [19.169, 36.411], TV factory undesirable:  $β1 = 23.043$ ,  $95\%$  CI [17.308, 28.778], Cohen's d = 0.2, **Figure 2.3d**).



**Figure 2.3. Under threat weaker evidence is required to reach undesirable conclusions.** Fitted psychometric functions for data of the threat group (solid line) and control group (dotted line). Y axis shows the proportion of times participants indicated they were in a TV factory as a function of the proportion of TV stimuli they observed in a trial prior to making a decision (x axis). In blue is the data of participants for whom the TV factory was the desirable factory. In orange is the data of participants for whom the TV factory was the undesirable factory. **(a)** Compared to the control group, under perceived threat participants required a smaller proportion of TVs to be observed before reaching the conclusion that they were in a TV factory, when the TV factory was undesirable. This can be seen as the solid line (threat group) is shifted left relative to the dotted line (control group). **(b)** No such difference is observed when the TV factory is desirable. **(c)** Participants in the control group required a smaller proportion of TVs to be observed before reaching the conclusion that they were in a TV factory, when the TV factory was desirable than undesirable. This is seen as the blue line (desirable) is shifted left relative to the orange line (undesirable). **(d)** This difference is abolished under perceived threat. Error bars show SE at given level of proportion of TVs observed (error bars for threat group are indicated by 'x' at the centre of the error bar). Grey dashed line indicates point of indifference – i.e., how much evidence is needed for participants to say 'TV' half the time.

Note, that the total number of pieces of evidence (TVs + telephones) did not differ when participants reached an undesirable or desirable conclusion  $(F(1,81) = 1.363, p = 0.247,$  partial  $p = 0.017$ , nor did it differ as a function of perceived threat  $(F(1,81) = 0.376, p = 0.542,$  partial  $p2 = 0.005$ ), neither was there an interaction between these two factors  $(F(1,81) = 0.023, p = 0.879,$ partial  $n^2$  = 0.00). Rather, as shown above, it is the proportion of evidence (which signifies the strength of the evidence) needed to reach a conclusion that differed as a function of perceived threat and valence.

Thus far our analysis suggests that perceived threat led to a reduction in the strength of the evidence needed to reach undesirable conclusions, even though the cause of the threat (anticipating a negative social situation) had nothing to do with the task at hand. We next sought to identify the precise computational factor(s) affected by perceived threat during evidence accumulation.

**Under threat the drift rate towards undesirable conclusions is greater.** Computationally, there are at least two different ways by which perceived threat can lower the strength of evidence needed to reach undesirable conclusions. First, threat may alter the starting point of the accumulation process. That is, if under perceived threat participants are a priori more likely to believe they are in an undesirable state relative to controls, then weaker evidence will be needed to reach that conclusion. Alternatively, perceived threat can enhance the weight given to each piece of undesirable evidence relative to control. This again will lead to weaker evidence needed to reach an undesirable conclusion.

To tease apart these possibilities we modelled the responses as a drift-diffusion process (Ratcliff, 1978; Ratcliff & McKoon, 2008; Voss et al., 2013) with the following parameters: (1) t0—amount of non-decision time; (2)  $\alpha$ —distance between decision thresholds; (3) z—starting point of the accumulation process; and (4) v—drift rate – i.e., the rate of evidence accumulation (**see Methods for details**). Crucially, in models 1 and 3 the starting point was fixed to 0.5, while in models 2 and 4 we allowed the starting point to vary towards one threshold (thus allowing a starting point bias). In models 3 and 4 we allowed the drift rate to vary (which we call a valence-dependent drift rate bias) depending upon whether the participant was visiting a desirable factory or an undesirable factory (thus allowing a process bias).

The DIC, a generalization of the AIC for hierarchical models, was calculated for each model **(Table 2.2)**. The DIC scores indicated that Model 4 (the valencedependent model), which included a valence-dependent starting point and drift rate, outperformed all other models for both the threat and the control group. As can be observed in **Figure 2.4,** while for the control group the valencedependent model was clearly a better fit than the valence-independent model (replicating previous findings from Gesiarz et al., 2019), for the threat group the advantage in terms of fit was modest.



**Table 2.2 DDM Model Fits.** For each group we ran four models which differed in whether we allowed the starting point to vary (model 2 & 4), whether we included a valence-dependent drift rate bias (model 3 & 4), or neither (model 1). DIC scores show goodness of fit, with lower numbers indicating better fit.



**Figure 2.4. Difference in fit between winning valence-dependent model and valence-independent model as a function of perceived threat.** The Y axis shows the difference in DIC scores between the valence-independent model and the winning valence-dependent models for the control group (dark grey) and threat group (light grey).

We next examined which of the accumulation parameters were affected by perceived threat. To that end, we calculated 95% CIs of each parameter for each group. If the 95% CIs do not overlap, we infer a significant difference between the two groups.

As observed in **Table 2.3** and **Figure 2.5**, only one element in the accumulation process was significantly altered by perceived threat: the valence-dependent drift rate bias. The drift rate bias is the difference in drift rates between desirable and undesirable factories, the greater the bias the greater the drift rate for desirable factories relative to undesirable ones. As can be observed in **Figure 2.5e** the valence-dependent bias in drift rate in the control group was significantly greater than in the threat group (control:  $\beta$ 1 = 0.17 [0.07, 0.27]; threat:  $\beta$ 1 = -0.08 [-0.20, 0.04]). For controls the bias in drift was significantly *positive* (95% CI does not include zero: β1 = 0.17 [0.07, 0.27]), leading to a drift rate that was more than double when participants were in the desirable factory ( $V_{\text{desirable}} = 0.63$ ) than undesirable factory ( $V_{\text{undesirable}} = 0.46$ ). In contrast, under perceived threat the bias in drift rate was numerically negative and not

significantly different from zero (95% CI includes zero:  $\beta$ 1 = -0.08 [-0.20, 0.04]), leading to a drift rate that was numerically and non-significantly larger when participants were in the undesirable factory ( $v_{\text{undesirable}} = 0.63$ ) than when they were in the desirable factory ( $v_{\text{desirable}} = 0.55$ ). We also corroborate these results using 95% HDI comparisons (**see Appendix 7.1 Supplementary Table 2.1 for HDI Comparisons**).



**Table 2.3. Parameter estimates of the evidence accumulation process.**  Displayed are the model estimates from the winning model for the control and threat groups. These include distance between decision thresholds (α), nondecision time (t0), starting point (0<z<1), inter-trial starting point parameter (sz), constant drift rate (β0) and drift rate bias (β1). The latter is the term reflecting the additional weight added to the drift rate as a function of factory desirability. Positive values indicate a bias towards desirable conclusions, and negative values indicate a bias towards undesirable conclusions. [CI].



**Figure 2.5. Under threat the valence-dependent drift rate bias is abolished.** Displayed are the posterior distributions of parameter estimates for the threat group (light grey) and the control group (black). No significant difference is observed between groups for estimates of **(a)** distance between decision thresholds, **(b)** non-decision time, **(c)** starting point, and **(d)** drift-rate constant. **(e)** In contrast, a significant difference is observed for the valence-dependent drift rate bias. In the control group the bias indicates a significantly larger drift rate towards desirable than undesirable conclusions. This bias is corrected for under perceived threat and is numerically inverse (that is the bias is nonsignificantly negative under perceived threat but significantly positive in the control group). \*indicates significant difference between parameters in the threat and the control group (i.e., CIs do not overlap). Dashed vertical lines indicate group mean.

We simulated data using group parameters from the threat and control group separately (**see Methods for details**). We first examined if the model parameters could be successfully recovered based on the simulated data. To do so the valence-dependent model was fit to simulated data, in the same way as for the experimental data. We sampled 2,000 times from the posteriors, discarding the first 500 as burn in. As shown in **Table 2.4** model parameters could be successfully recovered based on the simulated data. Additionally, we examined if the simulated data reproduced the same behavioural pattern of results as the participants' data. This was indeed the case (see **Figure 2.6d,**  detailed explanation below).



**Table 2.4. Recovered Parameter estimates of the evidence accumulation process based on simulated data.** Displayed are the winning model estimates recovered from the simulated data for the control and threat groups. These include distance between decision thresholds  $(\alpha)$ , non-decision time (t0), starting point (0<z<1), inter-trial starting point parameter (sz), constant drift rate (β0) and drift rate bias (β1). The latter is the term reflecting the additional weight added to the drift rate as a function of factory desirability. Positive values indicate a bias towards desirable conclusions, and negative values indicate a bias towards undesirable conclusions. [CI].

As DDM parameters are computed partially based on participants' responses we expected the model-based valence-dependent drift rate bias to correlate across individuals with a valence-dependent bias in judgments. Indeed, across participants there was a strong positive correlation between valence-dependent drift rate bias and the proportion of correctly identified desirable factories minus the proportion of correctly identified undesirable factories (threat group:  $r =$ 0.802, p < 0.001, **Figure 2.6a,** control: r = 0.918, p < 0.001, **Figure 2.6a)**, which we term 'valence-dependent judgement bias'. Individuals with a greater drift rate towards desirable than undesirable conclusions were more likely to correctly identify desirable factories as desirable when they observed them, than undesirable factories when they observed them. In contrast, the starting point bias did not correlate with the valence-dependent bias in judgements in the threat group ( $r = 0.223$ ,  $p = 0.191$ ), but did in the control group ( $r = 0.517$ , p < 0.001). In the latter, a larger starting point bias was related to the proportion of correctly identified desirable factories minus the proportion of correctly identified undesirable factories (**Figure 2.6b**).

As we have already shown that participants in the control group had a greater drift rate bias than those under perceived threat, it follows that they would also show a greater valence-dependent judgement bias. This is indeed what we observed. Entering the proportion of correctly identified factories as desirable or undesirable into a mixed 2 by 2 ANOVA with valence of factory (withinsubject), group (between-subject) and their interaction, revealed a group by valence interaction (F(1,81) = 3.868, p = 0.049, partial η2 = 0.05, **Figure 2.6c**) as well as a main effect of factory valence  $(F(1,81) = 5.045, p = 0.027,$  partial  $n^2$  = 0.06) and no main effect of group (F(1,81) = 0.002, p = 0.985, partial  $n^2$  = 0.00). To tease apart the interaction we followed up with pairwise comparisons. This revealed that participants in the control group were less likely to correctly categorize undesirable factories ( $M = 0.727$ ,  $SE = 0.022$ ) than desirable factories (M = 0.80, SE = 0.015;  $t(42) = 3.099$ , p = 0.003, Cohen's d = 0.473). In contrast, under perceived threat the effect of valence dissappeared  $(t(39) =$ 0.19,  $p = 0.85$ , Cohen's  $d = 0.03$ ); participants were just as likely to correctly categorize undesirable factories ( $M = 0.76$ ,  $SE = 0.02$ ) as they were desirable factories ( $M = 0.765$ ,  $SE = 0.019$ ). This suggests that under perceived threat the valence-dependent judgement bias is abolished.

We conducted the same analysis on our simulated data and find that it nicely reproduced the behavioural pattern of results (**Figure 2.6d, see Appendix 7.1 Supplementary Table 2.2**).



**Figure 2.6. Threat-induced change in valence-dependent drift rate bias is expressed as a valence-dependent change in the proportion of correctly identified factories. (a)** A positive relationship is observed between the valence-dependent drift rate bias (Y axis) and the valence-dependent bias in the proportion of correctly identified factories (X axis). Individuals with a greater drift rate towards desirable than undesirable conclusions are more likely to correctly categorize desirable than undesirable factories. This is true both for both participants in the control group (dark grey) and participants in the threat group (light grey). For those in the control group, the regression line is above that of participants in the threat group, which is due to the fact that their drift rate bias is significantly greater. The regression line for controls is also shifted to the right which indicates a significantly greater valence-dependent judgement bias. **(b)** By contrast we did not observe a relationship between the starting point bias (Y axis) and the valence-dependent bias in the proportion of correctly identified factories (X axis) in the threat group (light grey). A positive correlation was observed in the control group (dark grey). In the control group individuals with a large starting point bias were more likely to correctly identify desirable than undesirable factories. While the line for the threat group is above that of the control group this is not a significant difference. **(c)** Controls are less likely to correctly categorize undesirable factories (orange) than desirable factories (blue), while this is not the case for participants in the threat group. (d) Simulated data based on model parameters reproduced these findings. Data are plotted as boxplots for each condition, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers

indicate 1.5 x interquartile range. Diamond shape indicates the mean.  $*$ p < 0.01, ns = not significant. Clouds represent CIs.

# **2.5 Discussion**

The findings show that perceived threat has a profound effect on the process by which evidence is accumulated. In particular, it leads to a reduction in the strength of the evidence needed to reach undesirable conclusions. Relative to controls, participants under perceived threat required a smaller proportion of negative stimuli to be observed before reaching an undesirable conclusion. In contrast, there was no between-group difference in the strength of evidence accumulated before reaching a desirable conclusion. We found this to be true despite the fact that the cause of the threat (anticipating a socially stressful event) was unrelated to the task performed (judging whether more phones or more TVs were observed).

Computationally, there are different mechanisms by which perceived threat can lower the strength of evidence needed to reach undesirable conclusions. First, under threat participants may be a priori more likely to believe they are in an undesirable state relative to controls leading to weaker evidence needed to reach that conclusion. Another possibility is that perceived threat can selectively increase the rate of negative evidence accumulation (drift rate) relative to control. This again will lead to weaker evidence required to reach an undesirable conclusion. To tease apart these possibilities we modelled responses as a drift-diffusion process (Ratcliff, 1978; Ratcliff & McKoon, 2008; Voss et al., 2013). We found support for the latter. Specifically, perceived threat altered only one feature of the accumulation process: the relative drift rate towards desirable and undesirable conclusions (the 'valence-dependent drift rate bias'). For controls the bias in drift rate was positive – the rate of evidence accumulation was greater towards desirable than undesirable conclusions (as observed before in Gesiarz et al., 2019). Under threat, however, the bias disappeared due to the drift rate towards undesirable conclusions increasing.

These results fit with previous suggestions that perceived threat shifts neural valuation from desirable to potential aversive outcomes in risky choice (Engelmann et al., 2015), directs attention towards negative stimuli (Macatee et al., 2017) and leads to greater impact of such stimuli on belief updating (Garrett et al., 2018). Indeed, it is possible that the effect of perceived threat on the rate of negative evidence accumulation is partially due to increased attention towards negative stimuli. The current findings go beyond these previous demonstrations to illuminate the effects of perceived threat on the process of sequential evidence accumulation and show that weaker evidence is needed to reach undesirable conclusions under threat.

Here, we show a causal link between perceived threat and evidence accumulation in healthy individuals. It is interesting, however, to consider how our findings may be related to evidence accumulation in individuals with affective disorders, as these are often triggered by stressful events and/or characterized by high anxiety. With regards to individuals with high trait anxiety, a processing advantage for threatening words has been previously reported (White et al., 2010). While that study was correlational and thus could not determine whether anxiety caused the changes to the drift rate and/or vice versa, our results support the notion that anxiety can in fact alter the drift rate towards undesirable conclusions, even if the anxiety is short-lived rather than chronic. With regards to individuals with anxiety and mood disorders, one study (Aylward et al., 2019) found a lower drift rate towards desirable conclusions compared to healthy individuals. Interestingly, the latter study did not detect any effects of induced threat, which may be due to the fact that the task used in that study (as well as in all the above-mentioned studies) unlike ours, was a nonsequential perceptual decision-making task. The process by which pieces of evidence are accumulated over time may be especially impacted by perceived threat.

Our study suggests that evidence accumulation is a flexible process which quickly adjusts to the environment. In particular, the findings show that perceived threat leads to a valence-dependent change to the accumulation

process, even when the evidence is not directly related to the cause of the threat. An increased rate of negative evidence accumulation can then enhance the probability of taking precautionary action to avoid aversive consequences. As aversive outcomes can be more severe and frequent in threatening environments, such generalization can be, on average, adaptive. However, in individuals who are hypersensitive to threat and/or falsely perceive situations as threatening, such as those suffering from anxiety and depression, an increased rate of negative evidence accumulation could be maladaptive. This is because such an increased rate can produce overly pessimistic predictions, which induce stress and anxiety, further worsening symptoms.

# **Chapter 3: Futile Rewards: Why Accuracy Incentives Fail to Reduce Biased Evidence Accumulation**

# **3.1 Abstract**

Intuitively, we assume that financially incentivizing individuals to provide a correct response will enhance the accuracy of their decisions. Here, we show that accuracy incentives fail to reduce a well-known bias in which people reach desirable (over undesirable) conclusions (= desirability bias), because the two operate on orthogonal aspects of the evidence accumulation process. Over three experiments, participants ( $n = 235$ ) completed a perceptual evidence accumulation task in which they had to determine whether they were in a desirable state, which was associated with greater rewards, or an undesirable state. In some trials they were also financially incentivized for correct responses. Results show that while accuracy incentives led participants to take more time in reaching a conclusion, they did not impact participants' desirability bias. Fitting the data to an evidence accumulation model revealed that while accuracy incentives led to an increase in the distance between the decision thresholds, the desirability bias was associated with greater weight on desirable relative to undesirable evidence (drift rate bias). These results suggest that the desirability bias is likely unconscious.

# **3.2 Introduction**

Humans demonstrate an impressive ability to understand the world around them. At the same time, however, people also exhibit a host of systematic errors in judgement (known as biases and heuristics, Gigerenzer & Gaissmaier, 2011). These errors can lead to poor and costly decisions in domains ranging from finance to health (Dunning et al., 2004; Shefrin, 2015). There is thus a clear need to identify ways to mitigate such systematic errors. One obvious solution is to financially incentivize people to form more accurate beliefs in situations where biases and heuristics are common. But would such incentives work? The answer may depend on the nature of the error.

One category of systematic errors, known as heuristics, are cognitive shortcuts used to make efficient and effortless decisions (Gigerenzer & Gaissmaier, 2011). These are not errors that people are motivated to make, but rather the result of 'thoughtless' processing. Past studies suggest financial incentives are generally successful at reducing such errors. For instance, financial incentives have been shown to reduce the tendency to falsely believe large-number ratios (e.g., 30/100) convey a higher probability than equivalent small-number ratios (3/10; Dale et al., 2007; Lefebvre et al., 2011). Incentives also reduce the anchoring effect, that is the tendency for decisions to be disproportionally influenced by an initial piece of observed information (Meub et al., 2013; Wright & Anderson, 1989; but see Epley & Gilovich, 2005; Enke et al., 2023), and the conjunction fallacy, which describes the tendency to incorrectly consider multiattribute hypotheticals that are more specific (e.g., "The dog surfs and wears a hat") to be more probable than singular hypotheticals (e.g., "The dog surfs", Zizzo et al., 2000; for a meta-analysis see Yechiam & Zeif, 2023; but see Charness et al., 2010). Thus, there is some evidence that incentives can increase ('slow') rational thinking and reduce the reliance on cognitive shortcuts.

Here, we focus on a distinct, second, category of systematic errors motivational biases. These are errors that occur not due to 'cognitive laziness', but because the individual prefers one belief over the other (Montibeller & Von Winterfeldt, 2015). Only a handful of studies have examined whether incentives can overcome such motivational biases. Almost all have focused on the partisan bias, which is the tendency to process, interpret, and favour information in a way that aligns with one's political ideology (Van Bavel & Pereira, 2018). Indeed, monetary incentives have been shown to reduce the impact of the partisan bias on judging the accuracy of political statements (Prior et al., 2015; Rathje et al., 2023), however not on non-political advice-taking (Zhang and Rand, 2023). Beyond the partisan bias, financial incentives also

failed to mitigate the bias to interpret a perceptual stimulus as predicting the absence of an upcoming shock (Engelmann et al., 2019).

Thus, while muddy, the literature seems to suggest that monetary incentives are relatively unreliable in reducing motivational biases. Here, we propose and test a possible explanation. We hypothesize that financial incentives do indeed lead individuals to invest cognitive effort in reaching accurate conclusions, by for example accumulating more evidence. These efforts may indeed be fruitful when the cause of the error is fast, careless, processing. However, when the cause of the error is an unconscious bias in how information itself is processed (for example as when people put greater weight on desirable than undesirable evidence resulting in a desirability bias, as shown in Gesiarz et al., 2019; Globig et al., 2021), they will have little effect. In other words, if financial incentives alter a feature of the decision process that is orthogonal to the one the bias works on, they will fail.

To test this hypothesis we conducted three experiments, where participants (total  $N = 236$ ) played a modified random-dot task. In the task, participants observed a cloud of moving dots. Their goal was to determine whether most of the dots were moving left or right. On some trials, they were rewarded for correct responses, while on other trials there was no reward for correct responses. Importantly, participants were told that one of the two directions was desirable. That is whenever the dots moved in that direction, they could receive an additional bonus. Crucially, participants had no control over the direction of the dots; their task was simply to detect it. In Experiment 1 and its replication, the reward for correct responses on incentivized trials was equal to this bonus, and in Experiment 2 it was five times greater. Thus, in all experiments, to maximize reward across the task, it was in participants' interest to be as accurate as possible.

We fit a DDM to our data. This models the process of noisy evidence accumulation toward either of two decision thresholds from a given starting point (Ratcliff & Rouder, 1998; Voss et al., 2013) and enables us to determine which elements are altered by accuracy incentives. Previous work from our lab shows, that the motivation to form desirable conclusions alters the rate at which desirable evidence is accumulated (Gesiarz et al., 2019; Globig et al., 2021; see also Leong et al., 2019). We speculate that if accuracy incentives fail to mitigate biased evidence accumulation, this could be because they affect a different element of the accumulation process. For example, incentives may increase the distance between decision thresholds, thus making participants more cautious, but may not impact the relative rate at which desirable evidence is accumulated. Knowledge of the precise decision-making features that financial accuracy incentives impact, can shed light on when and why they may fail to reduce motivational biases.

## **3.3 Methods**

#### **Experimental Design**

#### **Participants (Experiment 1):**

Seventy participants residing in the United States (US) completed the task on *Prolific Academic* (www.prolific.com). Participants received £7.50 per hour for their participation in addition to a performance-related bonus. Ethical approval was provided by the Research Ethics Committee at University College London and all participants gave informed consent. All experiments were performed in accordance with the principles expressed in the Declaration of Helsinki. Participants who failed the comprehension checks at the beginning of the experiment more than twice were not allowed to participate. We also adopted the following exclusion criteria: response times (RTs) faster than 200ms were discarded from further analysis, as recommended in previous literature (Ratcliff & Tuerlinckx, 2002; Rollwage et al., 2020; Wiecki et al., 2013). In cases where a participant did this in over half their trials, we assumed that the participant was not appropriately engaging with the task and eliminated the entirety of their trials (as in Gesiarz et al., 2019; Globig et al., 2021). Based on this we removed 1 participant and a further 64 trials from Experiment 1. Thus, data of 69 participants were analysed (Experiment 1: N = 69 M<sub>age</sub> = 35.319, SD<sub>age</sub>  $\pm$ 10.222; female  $= 31$ , male  $= 38$ ). Sample size was calculated based on a pilot study. Calculations were performed using g\*Power (Faul et al., 2009) to achieve power of 0.8.

#### **Participants (Replication):**

Seventy-three participants residing in the US completed the task on *Prolific Academic*. Participants received £7.50 per hour for their participation in addition to a performance-related bonus. RTs faster than 200ms were discarded from further analysis (**see Participants Experiment 1 for details**). Forty-nine trials were discarded for this reason. No participants were removed. Thus, data of 73 participants were analysed (Replication:  $N = 73$ ; M<sub>age</sub> = 35.699, SD<sub>age</sub>  $\pm$  9.253;  $female = 37$ , male  $= 36$ ).

#### **Participants (Experiment 2):**

Ninety-three participants residing in the US completed the task on *Prolific Academic*. Participants received £7.50 per hour for their participation in addition to a performance-related bonus. RTs faster than 200ms were discarded from further analysis (**see Participants Experiment 1 for details**). In cases where a participant did this in over half their trials, we assumed that the participant was not appropriately engaging with the task and eliminated the entirety of their trials. Based on this we removed 1 participant from Experiment 2 and a further 134 trials. Thus, data of ninety-two participants were analysed (Experiment 2:  $N = 92$ ;  $M_{\text{aoe}} = 29.396$ ,  $SD_{\text{aoe}} \pm 5.702$ ; female = 44 male = 48).

#### **Behavioural Task (Experiment 1 & Replication):**

In each experiment, participants played a modified forced choice random dot kinetogram (RDK) task (Newsome & Pare, 1988; **see Figure 3.1**). The task consisted of four blocks with 60 trials each. On each trial participants observed a cloud of white dots (number of dots = 100, dot radius =  $2^\circ$ , dot life = 5 frames) moving within a circular aperture on a black background. The direction of motion of the clouds was either right or left along the horizontal meridian. Each set was replotted one aperture later in which a subset of dots, determined by the noise level, was offset from their previous location towards the target movement direction, and another subset was offset in the opposite direction,

whereas the rest was replotted randomly. Participants had to judge whether the majority of the dots were moving in the right or left direction. They were free to respond as soon as they wished. The dot motion would either continue until they made their choice or up to 5s. Across trials we varied whether participants were incentivized for accuracy (i.e., paid for correct responses; reward = \$5 vs no reward =  $$0$ ) and the level of noise of the dot motion (i.e., low noise = 0.256 motion-coherence vs high noise = 0.064 motion-coherence). Participants were informed that at the end of the experiment, one trial would randomly be selected for pay off. If in this trial participants were incentivized for accuracy, and they responded correctly they would receive \$5. In addition to this, in each block, one direction was assigned to be the desirable direction, and the other to be the undesirable direction. Participants were told that if the dots in the trial selected for pay off moved in the desirable direction, they would receive an additional \$5 bonus. Crucially, this bonus was entirely outside of their control, i.e., it was not affected by their responses. Desirability was counterbalanced across participants and varied across blocks, such that in two blocks participants would receive a bonus if the dots moved to the left, and in the remaining two blocks participants would receive the bonus if the dots moved to the right. Trial type was randomly determined with replacement on every trial with an equal probability for each trial type.



**Figure 3.1. The Random Dot Kinetogram Task (Experiment 1 & Replication).** In each trial, participants saw an animated cloud of dots moving

either rightward or leftward within a circular aperture. Their task was to determine the direction of the majority of the dots. Across trials, we varied the level of noise of the dot motion (i.e., % of dots moving in the same direction) and whether participants were incentivized for accuracy. The reward for correct responses on incentivized trials was \$5. Participants were informed that at the end of the experiment one trial would randomly be selected for pay off. In each block, one direction was assigned to be desirable. On trials where the dots happened to move in that direction, participants could receive an additional \$5 bonus, if this trial was selected for pay off. Notably, this bonus was beyond participants' control and was not affected by the actual decision made. Participants were told which direction was desirable at the start of each block. In each trial, they were first informed about whether the trial was incentivized for accuracy. Next, they were presented with a reminder about which direction was desirable. This was followed by a fixation cross. Each of these was presented for 1000ms. Afterwards the random-dot stimulus was displayed for a maximum of 5000ms or until button press.

For example, imagine that at the end of the experiment, a trial is selected in which there is no reward for correct responses. The trial is in a block where the right side is considered desirable. If a participant correctly deduces that the dots move to the right, they receive no performance-related reward despite answering correctly, because there is no reward for correct responses in this trial. However, they do receive a bonus (\$5) because the dots are in fact moving in the desirable direction. The task was coded using *JsPsych* and *Javascript*. The task for the replication was identical to Experiment 1.

**Training.** Prior to playing the task, participants received extensive instructions and were required to answer multiple-choice comprehension check questions on the key points of the task, with the question repeated until they either chose correctly or failed twice, upon which the experiment was terminated. The questions addressed the following key points of how the task worked: that participants' task was to correctly identify the direction of the dot motion; that on some trials they could receive a reward for correct responses; and that the bonus for dots moving in the desirable direction was independent of the responses they gave. Participants then played two practice sessions of 10 trials each where they received feedback. During the first practice session, they were given the opportunity to familiarize themselves with RDKs before the bonus

structure was introduced. The second practice session was included to give participants prior experience with the trial types. If their choice accuracy was below 50% in practice session 1 and 40% in practice session 2 respectively, they had to repeat the training procedures. If they failed the training procedure twice, the experiment was terminated.

#### **Statistical Analysis (Experiment 1 & Replication):**

**Response Bias.** To determine whether participants were biased towards reaching desirable conclusions we calculated each participant's response bias. We define this as their tendency to overestimate the proportion of desirable trials encountered. We calculated response bias separately for each participant as follows:

 $Response Bias = Prop.\n<sub>desirable judgments made</sub> - Prop.\n<sub>desirable trials encountered</sub>$ 

Such that positive values indicate a bias towards desirable conclusions while negative values indicate a bias towards undesirable conclusions. Response bias values were then entered into a one-sample t-test against zero to assess whether participants showed a significant response bias in their behaviour for each incentive level separately. We then compared the average response bias for each incentive level (\$0 vs \$5) using paired t-tests. Bayes tests were calculated to corroborate non-significant findings (Ly et al., 2016). Finally, onesample t-tests against zero were performed to assess whether participants showed a significant response bias in their behaviour for each incentive level separately. We compared the pattern of results obtained from participants' real data to those obtained from simulated data (**see Drift-Diffusion Modelling**). All statistical tests conducted in the present article are two-sided. Analysis was conducted using IBM SPSS 27 and R Studio (Version 1.3.1056). All results of interest hold when controlling for noise (**see Appendix 7.2 Supplementary Tables 3.9-3.16**).

**Response Times.** To determine whether participants cared about accuracy incentives, we calculated their mean log-transformed RTs for trials in which they were incentivized for accuracy (\$5) and trials in which they were not incentivized (\$0). We then computed a paired t-test to assess how accuracy incentives (\$0 vs \$5) affected log-transformed RTs.

**Drift-Diffusion Modelling.** Our aim in modelling our task using DDM was to gain a better understanding of the underlying mechanisms of how accuracy incentives and desirability may alter the evidence accumulation process.

To that end, we implemented and compared 65 different specifications of a DDM. The models included the following parameters: (1) t0—amount of nondecision time; (2) α—distance between decision thresholds; (3) z—starting point of the accumulation process; and (4) v—drift rate - the rate of evidence accumulation. To reduce computational load, we adopted a theory-driven approach: First, a baseline model was estimated where we allowed all parameters to vary - but did not include dependencies on any of the motives. Then, informed by prior research (Gesiarz et al., 2019; Globig et al., 2021; Leong et al., 2019), we reasoned that desirability might (1) reflect as a shift in the starting point to be closer to the bound associated with the desirable conclusion; and/or (2) induce selective accumulation of evidence in line with the desirable conclusion, as evidenced by a valence-dependent drift-rate bias. Furthermore, we speculated that when individuals are motivated to make accurate decisions, (1) they may be more cautious about making a decision, thereby increasing the distance between decision thresholds; and/or (2) they might weigh evidence more carefully thus slowing the rate of evidence accumulation (Shevlin et al., 2022). We also controlled for noise. We hypothesized that noise could (1) increase the distance between decision thresholds; such that participants are more cautious when the evidence is noisy; and/or (2) slow the rate of evidence accumulation, by making it harder to separate the evidence for each response option.

We first compared 64 models in which we tested these possibilities as main effects; and then assessed whether adding interactions would improve model fit of the winning model (**see Appendix 7.2 Supplementary Tables 3.1 & 3.2**

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**for list of all models**). RTs faster than 200 ms were discarded from the model fits and further analysis, as recommended in previous literature (Ratcliff & Tuerlinckx, 2002; Rollwage et al., 2020; Wiecki et al., 2013). In cases where a participant did this in over half their trials, we assumed that the participant was not appropriately engaging with the task and eliminated the entirety of their trials. Based on this, from Experiment 1 we removed 1 participant and a further 64 trials, and from its replication, we removed 49 trials.

We used the HDDM software toolbox (Wiecki et al., 2013) to estimate the parameters of our models. The HDDM package employs hierarchical Bayesian parameter estimation, using MCMC methods to sample the posterior probability density distributions for the estimated parameter values. We estimated both group-level parameters as well as parameters for each individual participant. Parameters for individual participants were assumed to be randomly drawn from a group-level distribution. Participants' parameters both contributed to and were constrained by the estimates of group-level parameters. In fitting the models, we used priors that assigned equal probability to all possible values of the parameters. Models were fit to log-transformed RTs because RTs were non-normally distributed and had a heavy positive skew. Also, since our 'error' RT distribution included relatively fast errors we included an inter-trial starting point parameter (sz) to improve model fit (Ratcliff & Rouder, 1998). We sampled 10,000 times from the posteriors, discarding the first 5000 as burn in and thinning set at 5. MCMC are guaranteed to reliably approximate the target posterior density as the number of samples approaches infinity. To test if the MCMC converged within the allotted time, we used Gelman-Rubin statistic (Rubin & Gelman, 1992) on 5 chains of our sampling procedure. The Gelman– Rubin diagnostic evaluates MCMC convergence by analysing the difference between multiple Markov chains. The convergence is assessed by comparing the estimated between-chains and within-chain variances for each model parameter. In each case, the Gelman-Rubin statistic was close to one (<1.1), suggesting that MCMC were able to converge.

Model fits were then compared using the DIC (Spiegelhalter et al., 2002), which is a generalization of the AIC for hierarchical models. The DIC is commonly used when the posterior distributions of the models have been obtained by MCMC simulation (Gamerman & Lopes, 2006). It allows one to assess the goodness of fit, while penalizing for model complexity. To examine the robustness of these results, we performed a complementary analysis and calculated Bayesian Predictive Information Criterion (BPIC) for each model (Ando, 2007). This corrects for over-fitting by adjusting for the asymptotic bias of the posterior mean of the log-likelihood as an estimator for its expected loglikelihood.

To validate the winning model, we used group parameters obtained from participants' data to simulate log-transformed RTs and responses. We used the exact number of participants, total number of trials and trial structure as in the experiment. Simulated data was then used to (1) perform model recovery analysis and (2) to compare the pattern of participants' responses to the pattern of simulated responses. We sampled 2000 times from the posteriors, discarding the first 500 as burn in. Simulation and model recovery analysis were performed using the HDDM software toolbox (Wiecki et al., 2013).

To assess if desirability, accuracy incentives and noise significantly altered the model parameters, we calculated 95% CIs of the parameters' values. If the 95% CI for a parameter does not overlap zero (or 0.5 for the starting point), we consider there to be a significant effect. We also computed 95% HDIs of the parameter's posterior distributions. If this HDI did not overlap zero, we considered there to be a meaningful effect. HDI testing was conducted in R using *HDInterval* (Meredith & Kruschke, 2016).

#### **Behavioural Task and Analysis (Experiment 2):**

The task and analysis in Experiment 2 was identical to that used in Experiment 1 and its replication except for the following differences:

1) As before, we varied whether participants were incentivized for accuracy across trials. But this time, we increased the potential reward for correct

responses five-fold. In some trials participants could receive \$25 for correct responses and in others \$0 **(see Figure 3.2)**. The bonus for dots moving in the desirable direction was \$5. Thus, in Experiment 2, we compared \$0 vs \$25 accuracy incentives, instead of \$0 vs \$5 accuracy incentives in Experiment 1 and its replication.

- 2) We also increased the range of noise levels. Noise varied across trials between 0.03 (high noise) to 0.3 (low noise) in increments of 0.03.
- 3) RTs faster than 200ms were discarded from the DDM fits (**see Appendix 7.2 Supplementary Table 3.3 for list of all models**) and further analysis, as recommended in previous literature (Ratcliff & Tuerlinckx, 2002; Rollwage et al., 2020; Wiecki et al., 2013). In cases where a participant did this in over half their trials, we assumed that the participant was not appropriately engaging with the task and eliminated the entirety of their trials. Based on this, from Experiment 2 we removed 1 participant and a further 134 trials.



**Figure 3.2. The Random Dot Kinetogram Task (Experiment 2).** The task in Experiment 2 was identical to the task in Experiment 1 and its replication except for the following: Across trials, we varied the level of noise of the dot motion (0.03-0.3 in increments of 0.03) and whether participants were incentivized for accuracy. In Experiment 2 we increased the reward on incentivized trials to \$25. On trials which were not incentivized the reward for correct responses was \$0.

We maintained the bonus associated with the dots moving in a desirable direction at \$5.

# **3.4 Results**

In this study, we assessed whether rewarding participants for correct responses reduces the bias towards desirable conclusions in evidence accumulation. To test this, we ran two identical experiments ( $N_{\text{Experiment 1}} = 70$ ;  $N_{\text{Replication}} = 73$ ), in which participants completed a modified forced choice RDK task (Newsome & Pare, 1988, **see Figure 3.1**). In this task they judged whether most dots were moving right or left. Across trials we varied whether participants were incentivized for accuracy (i.e., paid for correct responses; reward = \$5 vs no reward = \$0) and the level of noise (low noise vs high noise). We also told participants that in each block, one direction was assigned to be the desirable direction, and the other the undesirable direction. Participants were told that each time the dots moved in the desirable direction they would win an additional \$5 bonus. Crucially, participants had no control over which type of trial they were in (i.e., which direction the dots were moving), their task was simply to determine the direction, gaining rewards for accurate decisions on incentivized trials. We anticipated that this manipulation would cause participants to systematically err towards believing the dots were moving in the desirable direction even when they were not.

**The manipulation successfully biased participants towards desirable conclusions.** We first assessed whether our manipulation successfully induced a bias towards desirable conclusions. To that end, we calculated each participant's response bias as follows:

 $Response~Bias = Prop.$  desirable judgments made  $- Prop.$  desirable trials encountered

such that positive values indicate a bias towards desirable conclusions while negative values indicate a bias towards undesirable conclusions.

Indeed, participants judged a larger proportion of trials as desirable than the proportion of desirable trials they actually encountered (Experiment 1: MResponse  $Bias = 0.032$ ,  $SE = 0.014$ ,  $t(68) = 2.361$ ,  $p = 0.021$ , Cohen's d = 0.284, Replication: MResponse Bias =  $0.047$ , SE =  $0.014$ , t(72) = 3.323, p = 0.001, Cohen's  $d = 0.389$ . Thus, our manipulation was successful at inducing a bias towards desirable conclusions (replicating previous findings from Gesiarz et al., 2019; Globig et al., 2021; Leong et al., 2019).

**Participants become more cautious when incentivized for accuracy.** We next examined whether accuracy incentives altered participants' behaviour. We speculated, that if participants were sensitive to accuracy incentives, they should be more cautious on incentivized trials (\$5 reward) than on trials in which there was no reward for correct responses (\$0 reward). Indeed, we found that participants were slower to respond when they were incentivized for accuracy (Experiment 1: \$5 accuracy incentives –  $M_{\text{loaRT}} = 7.25$ , SE = 0.048; \$0 accuracy incentives:  $M_{\text{logRT}} = 7.219$ ,  $SE = 0.05$ ;  $t(68) = 3.478$ ,  $p < 0.001$ ; Cohen's d = 0.07; Replication: \$5 accuracy incentive –  $M_{\text{loaRT}} = 7.282$ , SE = 0.043; \$0 accuracy incentives:  $M_{logRT} = 7.26$ ,  $SE = 0.044$ ;  $t(72) = 2.755$ ,  $p = 0.007$ ; Cohen's d = 0.059**; see Figure 3.3**). Adding the level of noise to the analysis does not alter the results (**see Appendix 7.2 Supplementary Tables 3.4 & 3.5**). While the effect of incentives is small, these results suggests that participants are sensitive to accuracy incentives and are slower about their decisions when incentivized.



**Figure 3.3. Participants are more cautious when incentivized for accuracy (Experiment 1 & Replication).** In both **(a)** Experiment 1, and **(b)** its replication, participants took longer to reach a conclusion when incentivized for accuracy (\$5), compared to when there was no reward for correct responses (\$0). Y axis shows log-transformed RT. X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5  $\times$  interquartile range. Diamond shape indicates the mean logtransformed RT per incentive level. Individuals' mean response time is shown separately as dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p <  $0.001$ , \*\*p <  $0.01$ .

### **Accuracy incentives do not reduce the bias towards desirable conclusions.**

We next turned to our primary question - do accuracy incentives reduce the bias towards desirable conclusions? We found that despite participants taking longer to reach a conclusion when incentivized for accuracy, their bias towards desirable conclusions remained unchanged. In particular, the magnitude of the response bias, measured as the proportion of trials judged as desirable minus the proportion of desirable trials actually encountered, remained the same regardless of whether accuracy was incentivized (\$5 accuracy incentives – Experiment 1: MResponse Bias =  $0.035$ , SE =  $0.014$ , Replication: MResponse Bias = 0.042,  $SE = 0.015$ ) or not (\$0 accuracy incentives – Experiment 1: MResponse Bias  $= 0.03$ , SE = 0.001, Replication: MResponse Bias = 0.052, SE = 0.017; comparison

between the two: Experiment 1:  $t(68) = -0.532$ ,  $p = 0.96$ , Cohen's  $d = 0.064$ , Replication: t(72) = 0.765, p = 0.447, Cohen's d = 0.089, **see Figure 3.4**). Adding the level of noise to the analysis does not alter the results (**see Appendix 7.2 Supplementary Tables 3.6 & 3.7**). Bayes tests further provide moderate to strong support in favour of the null hypothesis (Experiment 1: BF01)  $= 9.181$ , Replication: BF $_{01} = 8.139$ ).

One-sample t-tests against zero revealed that for each incentive level participants concluded they were in a desirable trial significantly more often than an undesirable trial and thus overestimated the proportion of desirable trials they encountered (\$0 accuracy incentives – Experiment 1:  $t(68) = 2.051$ ,  $p = 0.044$ , Cohen's d = 0.247; Replication: M = 0.052, SE = 0.017, t(72) = 3.146,  $p = 0.002$ , Cohen's  $d = 0.368$ ; \$5 accuracy incentives – Experiment 1:  $M = 0.035$ ,  $SE = 0.014$ ,  $t(68) = 2.429$ ,  $p = 0.018$ , Cohen's d = 0.119; Replication: M = 0.042, SE = 0.015,  $t(72) = 2.79$ , p = 0.007, Cohen's d = 0.327). Thus, participants are biased towards desirable responses irrespective of whether they are incentivized for accuracy. In line with this, we did not observe an improvement in overall discernment between signal and noise when incentivized vs when not incentivized (**see Appendix 7.2 Supplementary Results & Supplementary Tables 3.8 & 3.9 for analysis of dPrime**).


**Figure 3.4. Accuracy incentives do not reduce the response bias (Experiment 1 & Replication).** In both **(a)** Experiment 1, and **(b)** its replication participants were biased towards desirable responses. Accuracy incentives did not reduce this bias. For each accuracy incentive level, the proportion of trials they judged as desirable was significantly larger than the proportion of desirable trials they encountered. Y axis shows response bias, i.e., proportion of desirable judgements made minus proportion of desirable trials encountered. X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5  $\times$  interquartile range. Diamond shape indicates the mean response bias per incentive level. Individuals' response bias is shown separately as dots. Symbols above each boxplot indicate significance level compared to 0.  $p < 0.05$ , ns = not significant.

# **Specifying the impact of accuracy incentives and desirability on the stages of evidence accumulation.**

Thus far, our results suggest that although participants are sensitive to accuracy incentives, the way in which they accumulate evidence remains biased. We next considered different potential mechanisms through which this behaviour could be explained. To that end, we considered how accuracy incentives, desirability and noise may influence the evidence accumulation process.

Informed by prior research (Gesiarz et al., 2019; Globig et al., 2021; Leong et al., 2019), we speculated that when participants are motivated to reach a desirable conclusion they may (1) selectively accumulate evidence in line with this desirable conclusion (drift rate); and/or (2) may be a priori more likely to believe they are in the desirable state even before seeing any evidence (starting point). Rewarding individuals for accurate responses makes them more cautious, potentially (1) increasing the amount of evidence they require before reaching a conclusion (distance between decision thresholds); and/or (2) causing them to weigh the evidence they gather in a less biased manner (drift rate, Shevlin et al., 2022). When the evidence is noisy, participants may (1) increase the amount of evidence they require before reaching a conclusion (distance between decision thresholds); and/or (2) slow the rate at which they accumulate evidence (drift rate).

To determine which, if any, of the above holds, we fit a DDM to the data. Specifically, we modelled participants' responses as a drift-diffusion process (Ratcliff, 1978; Ratcliff & McKoon, 2008; Voss et al., 2013) with the following parameters: (1) α—distance between decision thresholds; (2) t0—amount of non-decision time; (3) z—starting point of the accumulation process; and (4) v—drift rate, i.e., the rate of evidence accumulation.

In a first step, we set out to identify the model that best captures the data. To that end, we compared different model specifications, each allowing for different dependencies of each parameter on accuracy incentives, desirability, and noise. To reduce computational load, we first compared 64 different model specifications in which we tested all the above possibilities as main effects. We then added interactions to the winning model to assess whether this would improve model fit (**see Methods for further details**).

The DIC, a generalization of the AIC for hierarchical models, was calculated for each model (**see Appendix 7.2 Supplementary Tables 3.1 & 3.2**). DIC scores indicated that for both experiments the model that outperformed all other models allowed: (1) accuracy incentives to influence the distance between decision thresholds (2) desirability to influence starting point and drift rate and (3) noise to influence drift rate. Allowing for interaction effects did not improve model fit (**see Appendix 7.2 Supplementary Tables 3.1 & 3.2**). This suggests that the effect of desirability is not modulated by accuracy incentives, rather accuracy incentives and desirability influence different elements of the accumulation process. While desirability and noise alter the quality of evidence accumulation through the drift rate, accuracy incentives do not alter the quality of accumulation but instead influence the decision-making process by adjusting the caution with which decisions are made, i.e., the distance between decision thresholds. BPIC results corroborate these findings (**see Appendix 7.2 Supplementary Tables 3.1 & 3.2**).

Thus far, we have therefore established that a model which incorporates desirability, accuracy incentives and noise, without interactions among these variables fits the data best. To reveal the direction, magnitude, and significance of each effect, we, we calculated the 95% CIs for each parameter. If the 95% CI of the parameter distribution does not overlap with zero (or 0.5 for the starting point), we infer a significant effect.

This revealed that accuracy incentives increased the distance between decision thresholds (Experiment 1:  $\alpha=0.059$ ; 95% CI [0.021, 0.089], Replication: α=0.032; 95% CI [0.001, 0.072], **Table 3.1 & 3.2, Figure 3.5 a&b**). When rewarded for correct responses, participants were more cautious and needed more evidence before reaching a conclusion. In addition to this, desirability had a statistically meaningful effect on the drift rate (Experiment 1: v=0.125, 95% CI [0.002, 0.257]; Replication: v=0.278, 95% CI [0.12, 0.431], **Figure 3.5 c&d** replicating previous findings from Gesiarz et al., 2019; Globig et al., 2021; Leong et al., 2019). Participants selectively accumulated evidence towards desirable conclusions, and thus had a larger drift rate when the dots moved in the desirable direction. By contrast, desirability did not significantly modulate the starting point (Experiment 1: z=0.504, 95% CI [0.494, 0.515], Replication: z=0.49, 95% CI [0.479, 0.50], replicating previous findings from Globig et al., 2021). Finally, we observed a significant effect of noise on the drift rate (Experiment 1: v=0.383; 95% CI [0.299, 0.461], Replication: v=0.396; 95% CI [0.298, 0.492], **Figure 3.5 e&f**). That is, when the evidence was noisy, it was harder for participants to separate the evidence for each response option, thereby slowing the rate of evidence accumulation. 95% HDI comparisons corroborate this result **(see Appendix 7.2 Supplementary Table 3.10 for HDI Comparisons**).



**Figure 3.5. Different motives affect different aspects of the evidence accumulation process (Experiment 1 & Replication).** When participants were incentivized for accuracy, **(a&b)** the distance between decision thresholds increased. **(c&d)** The rate of evidence accumulation was greater when the dots were moving in the desirable direction and **(e&f)** when noise was high. Displayed are the posterior distributions of the parameter estimates for each feature for the parameter it significantly modulates in Experiment 1 (left) and its

replication (right). Coloured dashed vertical lines indicate group means. Black line represents 95% CI. \*indicates significant effect.

These findings lend support to some of our hypothesized mechanisms. Specifically, we find that (i) when individuals are motivated to make accurate decisions, they are more cautious - increasing the distance between decision thresholds; (ii) desirability induces selective accumulation of evidence towards desirable conclusions, and (iii) high levels of noise slow the rate of evidence accumulation. In line with the behavioural results, we do not find evidence for accuracy incentives modulating the effect of desirability on behaviour. This suggests that accuracy incentives fail to mitigate the influence of desirability on evidence accumulation, because the two act on different elements of the accumulation process. While participants required more information to make a decision, the way in which they accumulated evidence to reach that decision remained biased.

# **Model recovery is successful and simulated data reproduces experimental results**.

To determine whether our winning model accurately captured the data we examined whether the model parameters could be successfully recovered based on simulated data. To that end, we first simulated data using the group parameters (**see Methods for details**). We then fit the winning model to the simulated data, in the same way as for the experimental data. We sampled 2000 times from the posteriors, discarding the first 500 as burn in. As shown in **Table 3.1**-**3.2** model parameters could be successfully recovered based on the simulated data.



**Table 3.1. Real and Recovered Parameter estimates of the evidence accumulation process (Experiment 1).** Displayed are the real (left) and recovered (right) model estimates from the winning model. These include a constant for Distance Between Decision Thresholds (α), βAccuracy Incentives Distance between Decision Thresholds, Non-decision Time (t0), Starting Point (0<z<1), Inter-trial Starting Point (sz), constant Drift Rate (β0), βDesirability Drift Rate (β1), and βNoise Drift Rate. [CI].



**Table 3.2. Real and Recovered Parameter estimates of the evidence accumulation process (Replication).** Displayed are the real (left) and recovered (right) model estimates from the winning model. These include a constant for Distance between Decision Thresholds (α), βAccuracy Incentives Distance between Decision Thresholds, Non-decision Time (t0), Starting Point (0<z<1), Inter-Trial Starting Point (sz), constant Drift Rate (β0), βDesirability Drift Rate (β1), and βNoise Drift Rate. [CI].

Additionally, we examined whether the simulated data reproduced the same behavioural pattern of results as participants' data. This was indeed the case. In both the simulated data for Experiment 1 and its replication (**see Figure 5**) the response bias was not reduced by accuracy incentives (\$5 accuracy incentives – Simulated Data Experiment 1: MResponse Bias =  $0.027$ , SE =  $0.009$ , Simulated Data Replication: MResponse Bias =  $0.06$ , SE =  $0.007$ ; \$0 accuracy incentives – Simulated Data Experiment 1: MResponse Bias =  $0.032$ , SE =  $0.009$ , Simulated Data Replication: MResponse Bias =  $0.057$ , SE =  $0.008$ ; comparison between the two: Simulated Data Experiment 1:  $t(68) = 0.433$ ,  $p = 0.666$ , Cohen's  $d = 0.052$ , Simulated Data Replication:  $t(72) = 0.291$ ,  $p = 0.772$ , Cohen's d = 0.034, **see Figure 3.6**). Adding the level of noise to the analysis does not alter the results (**see Appendix 7.2 Supplementary Tables 3.11 & 3.12**). Bayes tests further provide moderate to strong support in favour of the null hypothesis (Simulated Data Experiment 1:  $BF_0 = 9.625$ , Simulated Data Replication:  $BF<sub>01</sub> = 10.407$ ).

One-sample t-tests against zero show that the response bias was significant in each incentive level (\$0 accuracy incentives – Simulated Data Experiment 1:  $t(68) = 3.582$ ,  $p < 0.001$ , Cohen's  $d = 0.431$ ; Simulated Data Replication:  $t(72) = 6.959$ ,  $p < 0.001$ , Cohen's  $d = 0.814$ ; \$5 accuracy incentives – Simulated Data Experiment 1:  $t(68) = 3.089$ ,  $p = 0.003$ , Cohen's  $d = 0.372$ Simulated Data Replication:  $t(72) = 7.993$ ,  $p < 0.001$ , Cohen's  $d = 0.935$ ). Thus, both the simulated data and the experimental data show a bias towards desirable conclusions irrespective of accuracy incentives.



**Figure 3.6. Simulated Data reproduces experimental data (Experiment 1 & Replication).** As in the experimental data, simulated data from DDM parameters for both **(a)** Experiment 1, and **(b)** its replication shows a bias towards desirable responses. Accuracy incentives did not reduce this bias. Y axis shows response bias, i.e., the proportion of trials judged as desirable minus proportion of desirable trials encountered. X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate  $1.5 \times$  interquartile range. Diamond shape indicates the mean response bias per incentive level. Simulated individuals' response bias is shown separately as dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p $<$  0.001, \*\*p $<$ 0.01, ns = not significant.

# **Participants remain biased despite a five-fold increase in the reward for correct responses.**

Up to this point, we show that modest accuracy incentives (\$5), which match the magnitude of the bonus amount (\$5) participants get when the dots move in the desirable direction, fail to mitigate biased evidence accumulation. Could a larger accuracy incentive, however, succeed in mitigating the bias? To test this, we increased the reward for correct responses five-fold, offering \$25 for correct responses on some trials, and \$0 in others. We maintained the bonus for desirable trials at \$5 **(see Figure 3.7)**. As before, participants had no control over the trial type and were incentivized to give accurate response to maximize financial gain in the task.

We repeated the analysis as in Experiment 1 and its replication. Once again, participants judged a larger proportion of trials as desirable than the proportion of desirable they encountered (Experiment 2:  $M = 0.035$ ,  $SE = 0.01$ , t(91) = 3.568,  $p < 0.001$ , Cohen's  $d = 0.372$ ). Thus, our manipulation was successful. Consistent with our previous results, participants were also slower to respond when they were incentivized for accuracy (Experiment 2: \$25 accuracy incentives – M = 7.211, SE = 0.043; \$0 accuracy incentives – M = 7.157, SE = 0.044; t(91) =  2.649, p < 0.001; Cohen's d = 0.132, **see Figure 3.7).** Adding the level of noise to the analysis does not alter the results (**see Appendix 7.2 Supplementary Table 3.13**).



**Figure 3.7. Participants are more cautious when incentivized for accuracy (Experiment 2).** Participants in Experiment 2 took longer to reach a conclusion when incentivized (\$25) for correct responses, compared to when they were not (\$0). Y axis shows log-transformed RT. X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate  $1.5 \times$  interquartile range. Diamond shape indicates the mean log-transformed RT per incentive level. Individuals' mean response time is

shown separately as dots. Symbols above boxplot indicate significance level compared to 0. \*\*\*p  $< 0.001$ .

Crucially, despite the large incentive for accuracy, the magnitude of participants' response bias was not significantly reduced (\$25 accuracy incentives: MResponse Bias =  $0.023$ ,  $SE = 0.012$ ; \$0 accuracy incentives: MResponse  $B_{\text{bias}} = 0.047$ , SE = 0.013; comparison between the two: t(91) = 1.508, p = 0.135, Cohen's d = 0.157, **see Figure 3.8**). Adding the level of noise to the analysis does not alter the results (**see Appendix 7.2 Supplementary Table 3.14**). A Bayes test further provides moderate to strong support in favour of the null hypothesis (Experiment 2:  $BF_{01} = 3.998$ ).

One-sample t-tests against zero showed that participants were biased towards desirable conclusions in trials in which there was no reward for correct responses (\$0 accuracy incentives:  $M = 0.047$ ,  $SE = 0.013$ ,  $t(91) = 3.567$ ,  $p <$ 0.001, Cohen's  $d = 0.372$ ) and in trials in which there was a \$25 reward for correct responses (\$25 accuracy incentives:  $M = 0.023$ ,  $SE = 0.012$ ,  $t(91) =$ 1.976, p = 0.05, Cohen's d = 0.206; **see Appendix 7.2 Supplementary Material & Supplementary Table 3.15 for analysis of dPrime**).



**Figure 3.8. Accuracy incentives that are five times larger than the bonus do not reduce the response bias (Experiment 2).** The proportion of trials participants judged as desirable was significantly larger than the proportion of desirable trials they encountered irrespective of whether the accuracy incentive was \$0 or \$25. Y axis shows response bias, i.e., proportion of desirable judgements made minus proportion of desirable trials encountered. X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate  $1.5 \times$  interquartile range. Diamond shape indicates the mean response bias per incentive level. Individuals' response bias is shown separately as dots. Symbols above each boxplot indicate significance level compared to 0.  $p$  < 0.05, ns = not significant.

As before, we then fit the DDM to our data to examine the mechanisms underlying these results. As in Experiment 1 and its replication, we found that the model that best explained participants' behaviour allowed (1) accuracy incentives to influence the distance between decision threshold; (2) desirability to influence starting point and drift rate; and (3) noise to influence drift rate. Allowing for interaction effects once again did not improve model fit (**see Appendix 7.2 Supplementary Table 3.3**).

As before, accuracy incentives increased the distance between decision thresholds such that participants were more cautious when incentivized for accuracy (Experiment 2: α=0.058; 95% CI [0.024, 0.093, **Table 3.3, Figure 3.9a**). The drift rate was larger when the majority of the dots moved in the desirable direction (Experiment 2: v=0.124; 95% CI [0.041, 0.209], **Figure 3.9b)**  and decreased when the evidence was noisy, thus making it harder for participants to separate the evidence for each response option (Experiment 2: v=1.68; 95% CI [1.263, 2.105]; **Figure 3.9c**). Desirability did not significantly modulate the starting point (Experiment 2: z=0.503, 95% CI [0.493, 0.513]). We therefore replicate the results from Experiment 1 and its replication. 95% HDI comparisons corroborate this result **(see Appendix 7.2 Supplementary Table 3.16 for HDI Comparisons**).



**Figure 3.9. Different motives affect different aspects of the evidence accumulation process (Experiment 2).** When participants were incentivized for accuracy, **(a)** the distance between decision thresholds increased**. (b)** The drift rate was greater when the dots were moving in the desirable direction and **(c)** when noise was low. Displayed are the posterior distributions of the parameter estimates for each motive for the parameter it significantly modulates in Experiment 2. Coloured dashed vertical lines indicate group means. Black line represents 95% CI%. \*indicates significant effect.

Finally, we successfully recovered the model parameters based on simulated data (**see Table 3.3; see Methods for details**) and found that the simulated data reproduced the same behavioural pattern of results as participants' data. Accuracy incentives did not alter the bias towards desirable conclusions (Simulated Data Experiment 2:  $t(91) = 0.032$ ,  $p = 0.975$ , Cohen's  $d = 0.003$ , **Figure 3.10).** The response bias across trials in which there was no reward for correct responses (\$0 accuracy incentives – Simulated Data Experiment 2:  $0.034$ ,  $SE = 0.004$ ) did not differ from trials in which there was a reward, even when the reward was five times larger than the bonus (\$25 accuracy incentives – Simulated Data Experiment 2:  $M = 0.034$ ,  $SE = 0.004$ ). These results hold when controlling for noise (**see Appendix 7.2 Supplementary Table 3.17**). Bayes tests further provide moderate to strong support in favour of the null hypothesis (Simulated Data Experiment 2:  $BF_{01} = 12.143$ ).





**Table 3.3. Real and Recovered Parameter estimates of the evidence accumulation process (Experiment 2).** Displayed are the real (left) and recovered (right) model estimates from the winning model. These include a constant for distance between decision thresholds (α), βAccuracy Incentives distance between decision thresholds, non-decision time (t0), starting point (0<z<1), inter-trial starting point parameter (sz), constant drift rate (β0), βDesirability drift rate (β1), and βNoise Drift Rate. [CI].

One-sample t-tests against zero show that the response bias was significant irrespective of incentive level (\$0 accuracy incentives – Simulated Data: t(91)  $= 9.201$ ,  $p < 0.001$ , Cohen's d = 0.959, \$25 accuracy incentives – Simulated Data:  $t(91) = 9.551$ ,  $p < 0.001$ , Cohen's  $d = 0.996$ ). Thus, we observed a bias towards desirable responses, in both the simulated data and the experimental data, irrespective of the potential reward for correct responses. The results from Experiment 2 therefore illustrate that the desire to form positive beliefs and thus reach desirable conclusions, are not mitigated by rewarding participants for correct responses, even when the magnitude of the reward is relatively large.



**Figure 3.10. Simulated Data reproduces experimental data (Experiment 2).**  As in the experimental data, simulated data from DDM parameters shows a bias towards desirable responses. Accuracy incentives did not reduce this bias. For each accuracy incentive level, the proportion of trials judged as desirable was significantly larger than the proportion of desirable trials encountered. Y axis shows response bias, i.e., proportion of desirable judgements made minus proportion of desirable trials encountered. X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate  $1.5 \times$  interquartile range. Diamond shape indicates the mean response bias per incentive level. Individuals' response bias is shown separately as dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p <  $0.001$ , ns = not significant.

# **3.5 Discussion**

Over three experiments we find that although participants are more cautious in reaching conclusions when incentivized for accuracy, the magnitude of their bias is unaffected. Specifically, when incentivized participants take longer to reach a conclusion, but their tendency to falsely believe they are in a desirable state (a state where rewards are greater) remains unchanged. This was true even when the reward for correct responses was increased five-fold.

Importantly, we reveal that this failure is due to accuracy incentives impacting a feature of the accumulation process that is orthogonal to the one influenced by wishful thinking. Specifically, when incentivized for accuracy participants required more evidence before making a decision, which is signified as a larger distance between decision thresholds in the evidence accumulation model (the DDM). However, the desire to hold a preferred belief was associated with greater weight assigned to desirable evidence, signified in the model as a larger drift rate when in a desirable state (replicating previous findings from Gesiarz et al., 2019; Globig et al., 2021; Leong et al., 2019). Thus, while participants accumulate more evidence before reaching a conclusion, the accumulation process remains biased.

These results fit with previous suggestions that even substantial accuracy incentives fail to alleviate certain biases, including the partisan bias in advice taking (Zhang & Rand), the repeated truth effect (Speckmann & Unkelbach, 2022), and anxiety-induced wishful thinking (Engelmann et al., 2019). The current findings go beyond these previous demonstrations and provide an explanation for why accuracy incentives may fail to reduce biased evidence accumulation.

Interestingly, our results suggest that participants consciously strive to maximize financial gains, slowing down when incentivized. Yet, the fact that this attempt is unfruitful may suggest they are not consciously aware of their bias towards forming desirable beliefs. This idea, that a desirability bias in evidence accumulation is automatic and unconscious, is in accord with findings showing that the magnitude of the bias is constant under time constraints and different degrees of cognitive load (Kappes & Sharot, 2019). These findings lend support to previous suggestions that motivational biases are not merely the product of overt preferences but instead beyond individuals' awareness (Chen & Krajbich, 2018; Desai & Krajbich, 2022; Krajbich, 2022; but see Sánchez-Fuenzalida et al., 2023).

A crucial feature of the task is that participants did not receive feedback. Thus, they did not have an opportunity to reflect on their errors and become aware of their bias. We speculate that to correct motivational biases, accuracy incentives should be coupled with feedback. Chapter 4 tests this assumption, by providing participants with direct feedback about the accuracy of the content they shared in simulated social media environments. We find that providing participants with accuracy-related feedback reduced the tendency to share misinformation in domains such as politics, culture, and health (Globig et al., 2023).

In summary, the study shows that accuracy incentives fail to mitigate the influence of the motivational bias on evidence accumulation. We suggest that this is because the desirability bias and accuracy incentives alter distinct elements of the evidence accumulation process. The former alters the relative weight assigned to desirable and undesirable information, while the latter increases the amount of evidence required to reach a conclusion. These findings are particularly relevant for policymakers and industry leaders as they may explain why financial bonuses may be ineffective in improving decisionmaking and performance.

# **Chapter 4: Changing the Incentive Structure of Social Media Platforms to Halt the Spread of Misinformation**

# **4.1 Abstract**

The powerful allure of social media platforms has been attributed to the human need for social rewards. Here we demonstrate that the spread of misinformation on such platforms is facilitated by existing social 'carrots' (e.g., 'likes') and 'sticks' (e.g., 'dislikes') that are dissociated from the veracity of the information shared. Testing 951 participants over six experiments, we show that a slight change to the incentive structure of social media platforms, such that social rewards and punishments are contingent on information veracity, produces a considerable increase in the discernment of shared information. Namely, an increase in the proportion of true information shared relative to the proportion of false information shared. Computational modelling revealed that the underlying mechanism of this effect is associated with an increase in the weight participants assign to evidence consistent with discerning behaviour. The results offer evidence for an intervention that could be adopted to reduce misinformation spread, which in turn could reduce violence, vaccine hesitancy and political polarization, without reducing engagement.

# **4.2 Introduction**

In recent years, the spread of misinformation online has skyrocketed, increasing polarization, racism and resistance to climate action and vaccines (Barreto et al., 2021; Rapp & Salovich, 2018; Tsfati et al., 2020; Van Bavel et al., 2021). Existing measures to halt the spread, such as flagging posts, have had limited impact (Chan et al., 2017; Grady et al., 2021; Lees et al., 2022).

We hypothesize that the spread of misinformation on social media platforms is facilitated by the existing incentive structure of those platforms, where social rewards (in the form of 'likes' and 'shares') are dissociated from the veracity of the information (Sharot, 2021). The rationale for this hypothesis is as follows; users can discern true from false content to a reasonable degree (Allen et al., 2021; Pennycook & Rand, 2019). Yet, because misinformation generates no less retweets and 'likes' than reliable information (Lazer et al., 2018; Vosoughi et al., 2018), and online behaviour conforms to a reinforcement-learning model by which users are reacting to social rewards, users have little reason to use their discernment to guide sharing behaviour. Thus, people will share misinformation even when they do not trust it (Pennycook et al., 2021; Ren et al., 2021).

To halt the spread, an incentive structure is needed where 'carrots' and 'sticks' are directly associated with accuracy (**Figure 4.1a right panel**, (Sharot, 2021). Such a system will work with the natural human tendency to select actions that lead to the greatest reward and avoid those that lead to punishment (Skinner, 1966). Scientists have tested different strategies to reduce the spread of misinformation, including educating people about fake news (Guess et al., 2020; Traberg et al., 2022), using a prompt to direct attention to accuracy (Kozyreva et al., 2020; Pennycook et al., 2020, 2021) and limiting how widely a post can be shared (Jackson et al., 2022). Surprisingly, possible interventions in which the incentive structure of social media platforms is altered to reduce misinformation had previously been overlooked.

Here, we test the efficacy of such a structure by slightly altering the engagement options offered to users. Specifically, we add an option to react to posts using 'trust' and 'distrust' buttons (**Figure 4.1b**). We selected these buttons because trust by definition is related to veracity - it is defined as 'a firm belief in the reliability, truth, ability, or strength of someone or something' (Oxford Dictionary).

We hypothesize that (1) people will use the 'trust' and 'distrust' buttons to discern true from misinformation more so than the commonly existing engagement options (such as a 'like' button; **Figure 4.1b, top panel**). By 'discernment' we mean that true posts will receive more 'trust' reactions than

'distrust' reactions and vice versa for false posts. This will create an environment in which rewards ('trusts') and punishments ('distrusts') are more directly associated with the veracity of information. Thus, (2) when exposed to this environment, users will start sharing more true information and less false information in order to obtain more 'trust' carrots and fewer 'distrust' sticks (**Figure 4.1b, bottom panel**). The new feedback options could both reinforce user behaviour that generates trustworthy material and signal to others that the post is dependable.

We also test environments in which participants receive only 'trusts' (a different number of *trusts* for different posts) or only 'distrusts' (a different number of *distrusts* for different posts) to examine if and how the impact of small vs large positive feedback ('trust') on discernment differs from the impact of small vs large negative feedback ('distrust'). It has been proposed that the possibility of reward is more likely to reinforce action than the possibility of punishment, while the possibility of punishment is more likely to reinforce inaction (Guitart-Masip et al., 2011, 2012, 2014). This may translate to a large number of 'trusts' selectively increasing sharing of true information without decreasing sharing of misinformation and vice versa for a large number of 'distrusts'. Further, being mindful of potential differences in sharing behaviour across political parties (Grinberg et al., 2019; Guess et al., 2020) we test participants from both sides of the political divide (Republicans & Democrats).



**Figure 4.1. Theoretical Framework. (a)** The current incentive structure (blue) is such that the veracity of shared information is dissociated from rewards ('carrots') and punishments ('sticks'). That is, true information and misinformation may lead to roughly equal number of rewards and punishments. An optimal incentive structure (orange) is such that sharing true information is rewarded with more 'carrots' than sharing misinformation, which in turn is penalized with more 'sticks' than true information. To create an optimal environment an intervention is needed by which the number of rewards and

punishments are directly associated with the veracity of information. **(b)** We test one such possible intervention (Experiment 1). In particular, we allow people to engage with posts using 'trust' reaction buttons and 'distrust' reaction buttons (orange). The rationale is that they will use these reactions to discern true from false information more so than 'like' and 'dislike' reaction buttons. **(c)** As a result, to obtain a greater number of 'trust' carrots and a smaller number of 'distrust' sticks in response to a post, people in the optimal environment (orange) will share more true than misinformation compared to those in the suboptimal environment which includes no feedback at all (grey), and those in an environment where the association between veracity of information and number of carrots and sticks is weak (blue). This second step is tested in Experiment 2 & 3.

To that end, over six experiments 951 participants engaged in simulated social media platforms where they encountered true and false information. In Experiment 1 we examined whether participants would use 'trust' and 'distrust' buttons to discern true from false information more so than existing 'like' and 'dislike' buttons (**Figure 4.1b,** replication: Experiment 4). In Experiment 2 and 3 we tested whether new groups of participants would share more true than false information in social media platforms that introduce real 'trust' and 'distrust' feedback from other participants (**Figure 4.1c,** replication: Experiment 5 & 6). The intuition is that 'trust' and 'distrust' reactions will naturally be used to indicate veracity and thus provide a reward structure contingent on accuracy, thereby reducing the sharing of misinformation and generating a healthier information ecosystem. Using computational modelling we provide insights into the specific mechanism by which our intervention improves sharing discernment.

### **4.3 Methods**

#### **Experimental Design**

#### **Power Calculations**.

Sample sizes for all experiments were computed based on our pilot study (**see Appendix 7.3 Experiment 4-6**). Power calculations were performed using  $g^*$ Power (Faul et al., 2009) to achieve power of 0.8 (beta = 0.2, alpha = 0.05; Experiment 1: partial  $n^2 = 0.51$ ; Experiment 2: Cohen's d = 0.33; Experiment 3: Cohen's  $d = 0.327$ ).

**Participants (Experiment 1).** One-hundred and eleven participants residing in the US completed the task on *Prolific Academic* (www.prolific.com). Data of four participants who failed more than two memory checks were excluded from further analysis (**see Memory/Attention Check for details**). Thus, data of 107 participants were analysed (52 Democrats, 54 Republican, 1 Other,  $M_{\text{age}} =$ 40.579,  $SD_{ace} \pm 14.512$ ; female = 55, male = 52; Non-White = 20, White = 87). Participants received £7.50 per hour for their participation in addition to a memory test performance-related bonus. For all experiments presented in this study, ethical approval was provided by the Research Ethics Committee at University College London and all participants gave informed consent. All experiments were performed in accordance with the principles expressed in the Declaration of Helsinki. All samples were politically balanced for Democrats and Republicans. All experiments were replicated (**see Appendix 7.3 Experiment 4-6**).

#### **Participants (Experiment 2).**

Three-hundred and twenty participants completed the task on *Prolific Academic*. Data of four participants who failed more than two memory checks were excluded from further analysis (**see Memory/Attention Check for details**). Thus, data of three-hundred and sixteen participants were analysed (146 Democrats, 142 Republican, 28 Other,  $M_{age} = 37.598$ , SD<sub>age</sub>  $\pm$  13.60; female = 157, male = 157, other = 2, Non-White = 77, White = 239). Participants received £7.50 per hour for the participation in addition to a memory test performance-related bonus.

#### **Participants (Experiment 3).**

Four-hundred and nine participants completed the task on *Prolific Academic*. Data of three participants who failed more than two memory checks were excluded from further analysis (**see Participants Experiment 1 for details**). Further data of three participants who suspected that the feedback provided did not stem from real participants were excluded. Thus, data of four-hundred and three participants were analysed (194 Democrats, 197 Republican, 12 Other,

 $M_{\text{age}} = 35.179$ ,  $SD_{\text{age}} \pm 11.051$ ; female = 204, male = 194, other = 4, Non-White  $= 85$ , White  $= 218$ ). Participants received £7.50 per hour for their participation in addition to a memory test performance-related bonus.

#### **Task (Experiment 1).**

Participants engaged in a simulated social media platform where they saw 100 news posts, each consisting of an image and a headline (**see Figure 4.2 & see Appendix 7.3 Supplementary Table 4.1 for stimuli and ratings**). Half of the posts were true, and half were false. They covered a range of different topics including COVID-19, environmental issues, politics, health, and culture. They were all extracted from the fact-checking website Politifact (https://www.politifact.com). For each post, participants had the option to either 'like', 'dislike', 'trust' or 'distrust' the post, or they could choose to 'skip' the post. They could select as many options as they wished (e.g., 'like' and 'distrust') or none at all. Participants were informed that if they chose to react to a post other users would be able to see their reactions. They were asked to treat the platform as they would any other social media platform. The order in which reaction buttons appeared on screen was counterbalanced across participants. Participants also indicated their age, gender, ethnicity, and political orientation. The task was coded using the *Qualtrics* online platform [\(https://www.qualtrics.com\)](https://www.qualtrics.com/).



**Figure 4.2. Task (Experiment 1).** Participants observed a series of 100 posts in random order (50 true, 50 false). Their task was to react using one or more of the 'like', 'dislike', 'trust' or 'distrust' buttons or to skip the post. The task was self-paced.

**Memory/Attention check.** At the end of the experiment, participants were presented with five posts and had to indicate whether these were old or new. This is to ensure that participants were attentive during the experiment. Participants who failed more than two of the memory checks were excluded from the analysis.

#### **Task (Experiment 2).**

In Experiment 2 participants engaged in a simulated social media platform where they saw the same 100 posts (50 true, 50 false) shown to participants in Experiment 1. Participants had to either 'repost' or 'skip' each post (**see Figure 4.3**). They were told that if they decided to repost, then the post would be shared to their feed, and they would observe other participants' reactions to it. We used a between-subject design with five environments. Depending on the environment participants were randomly assigned to, they could either see (i) how many people *disliked* the post, (ii) how many people *liked* the post, (iii) how many people *distrusted* the post, or (iv) how many people *trusted* the post. We also included a *Baseline* environment, in which participants received no feedback. Due to logistic constraints, the feedback was not collected in real time but was instead taken from participants' reactions in Experiment 1. The participants, however, believed the reactions were provided in real time as indicated by a rotating cogwheel (1s). If participants selected to skip, they would also observe a rotating cogwheel (1s) and then a screen asking them to click continue. The average duration of the white screen  $(M = 2.351s; SE = 0.281)$ was not different from the average duration of feedback ( $M = 2.625$ s; SE = 0.245; t(233) = 0.853,  $p = 0.395$ , Cohen's d = 0.056). Though the duration of trials in which participants chose to skip ( $M = 9.046$ s,  $SE = 0.38$ ) was slightly shorter than those in which they chose to share  $(M = 9.834s, SE = 0.358; t(233))$  $= 2.044$ ,  $p = 0.042$ , Cohen's d = 0.134). Thereafter, participants were presented with all the posts again and asked to indicate if they believed the post was accurate or inaccurate on a continuous scale from  $0 =$  *inaccurate* to  $100 =$ *accurate*. Finally, participants completed a short demographic questionnaire assessing age, gender, ethnicity, and political orientation. The task was selfpaced. The task was coded using *JsPsych* and *Javascript*.



**Figure 4.3. Task (Experiment 2 and 3).** In Experiment 2 on each of 100 trials participants observed a post (50 true, 50 false). They then choose whether to share it or skip (self-paced). They were told that if they chose to share a post, it would be shared to their feed such that the other participants would be able to see the post and react to it in real time (feedback). Depending on the environment participants were in, they could either observe the number of (i) 'dislikes' (N = 45), (ii) 'likes' (N = 89), (iii) 'distrusts' (N = 49), or (iv) 'trusts' (N = 46) feedback. The feedback was in fact the number of reactions gathered from Experiment 1, though the participants believed the reactions were in real time as indicated by a rotating cogwheel (1s). Once the feedback appeared, participants could then click continue (self-paced). If participants selected to skip, they would observe a white screen asking them to click continue (selfpaced). In the Baseline environment ( $N = 59$ ) participants received no feedback. Experiment 3 was identical to Experiment 2 with two distinctions: (1) Depending on the environment participants were in, they could either observe the number of (i) both 'dislikes' and 'likes'  $(N = 128)$ , (iii) both 'distrusts' and 'trusts' (N = 137) or (iii) no feedback (Baseline, N = 126). (2) In Experiment 3 we selected 40 posts (20 true, 20 false) to which Republicans and Democrats had on average reacted to similarly using the 'trust' button in Experiment 1. Discernment was calculated for each participant by subtracting the proportion of sharing false information from the proportion of sharing true information. High discernment indicates greater sharing of true than false information.

#### **Task (Experiment 3).**

Experiment 3 (**see Figure 4.3**) was identical to the task used in Experiment 2 with three exceptions:

- (1) We selected 40 posts (20 true, 20 false), in which there was no significant difference in the way Republicans and Democrats reacted to them using the trust button during Experiment 1. This was done by entering participants' trust responses (0/1) into a vector for Democrats and Republicans for each post. We then performed Pearson Chi Square Tests for each of the 100 posts to identify whether Democrats and Republicans used the trust button differently. Posts where no significant difference was observed were included in Experiment 3.
- (2) Three environments were included: a *Baseline* environment, in which participants received no feedback, a 'Trust & Distrust' environment, in which participants received both *Trust* and *Distrust* feedback whenever they chose to share a post, and a 'Like & Dislike' environment, in which participants received *Like* and *Dislike* feedback whenever they chose to share a post.
- (3) At the end of the experiment, we asked participants: (1) "What do you think the purpose of this experiment is?"; and (2) "Did you, at any point throughout the experiment, think that the experimenter had deceived you in any way? If yes, please specify."

#### **Statistical Analysis**

#### **Statistical Analysis (Experiment 1).**

We examined whether participants used the different reaction buttons to discern true from false information. For positive reactions (e.g., *'likes'* and *'trusts'*) discernment is equal to the proportion of those reactions for true information minus false information, and vice versa for negative reactions (*'dislikes'* and *'distrusts'*). Proportions were calculated for each participant and then entered into a 2 (**type of reaction**: 'trust' and 'distrust' / 'like' and 'dislike') by 2 (**valence:** positive, i.e., 'like', 'trust'/ negative, i.e., 'dislike', 'distrust') withinsubject ANOVA. Political orientation was also added as a between-subject factor (Republican/Democrat), allowing for an interaction of political orientation and type of reaction to assess whether participants with differing political beliefs used the reaction buttons in different ways. We performed one-sample t-tests

to compare discernment (equal to the proportion of those reactions for true information minus false information, and vice versa for negative reactions) against zero to assess whether each reaction discerned between true and false information. To examine whether participants' frequency of use of each reaction option differed we again ran a within-subject ANOVA, but this time with percentage frequency of reaction option used as the dependent variable. We computed a Pearson's correlation across participants between frequency of skips and discernment.

One participant selected 'other' for political orientations. This participant was not included in the analysis because political orientation was included in the analysis, and such small group sizes could heavily skew results. Analysis was conducted using IBM SPSS 27 and R Studio (Version 1.3.1056). All statistical tests conducted in the present article are two-sided. All results of interest hold when controlling for demographics (age, gender and ethnicity, when not including political orientation in the analysis, and, if applicable, when allowing for an interaction between type of feedback and valence.

#### **Discernment Analysis (Experiment 2 and 3).**

Discernment is calculated for each participant by subtracting the proportion of sharing false information from the proportion of sharing true information. High discernment indicates greater sharing of true than false information. In Experiment 2 scores were submitted to an ANOVA with type of feedback ('(Dis)Trust' vs '(Dis)Like' vs Baseline), valence of feedback (positive, i.e., 'like', 'trust' vs negative, i.e., 'dislike', 'distrust'), political orientation and an interaction of political orientation and type of feedback. To assess whether frequency of posts shared differed we used the same ANOVA, this time with percentage of posts shared out of all trials as the dependent variable.

To test whether '(Dis)Trust' feedback improves belief accuracy, we transformed participants' belief ratings (which were given on a scale from 'post is accurate'  $=100$  to 'post is inaccurate' = 0) to indicate error. If the post was false (inaccurate) error was equal to the rating itself, if the post was true (accurate) error was equal to 100 minus the rating. Participants' average error scores were then entered into a between-subject ANOVA with type of feedback (Baseline, '(Dis)Trust', '(Dis)Like'), valence of feedback, political orientation and an interaction of political orientation and type of feedback.

Analysis of Experiment 3 followed that of Experiment 2 with the difference being that we had three types of feedback environments (Baseline, 'Like & Dislike', 'Trust & Distrust') and of course no valence of feedback (as all environments were mixed valence or no valence). Data of participants who selected 'other' for political orientations (Experiment  $2 = 28$ , Experiment  $3 = 12$ ) were not analysed, because political orientation was included in the analyses variable, and small group sizes of 'other' could heavily skew results.

#### **Drift-Diffusion Modelling (Experiment 2 and 3).**

To assess whether being exposed to an environment with '(Dis)Trust' feedback impacted the parameters of the evidence accumulation process in our data compared to Baseline and '(Dis)Like' feedback we modelled our data using DDM. To that end, we ran three separate models – one for each type of feedback and included the following parameters: (1) t0—amount of nondecision time; (2) α—distance between decision thresholds; (3) z—starting point of the accumulation process; and (4) v—drift rate, i.e., the rate of evidence accumulation.

We used the HDDM software toolbox (Wiecki et al., 2013) to estimate the parameters of our models. The HDDM package employs hierarchical Bayesian parameter estimation, using MCMC methods to sample the posterior probability density distributions for the estimated parameter values. We estimated both group-level and individual-level parameters. Parameters for individual participants were assumed to be randomly drawn from a group-level distribution. Participants' parameters both contributed to and were constrained by the estimates of group-level parameters. In fitting the models, we used priors that assigned equal probability to all possible values of the parameters. Models were fit to log-transformed RTs. We sampled 20,000 times from the posteriors, discarding the first 5000 as burn in and thinning set at 5. MCMCs are guaranteed to reliably approximate the target posterior density as the number of samples approaches infinity. To test whether the MCMC converged within the allotted time, we used Gelman–Rubin statistic (Rubin & Gelman, 1992) on five chains of our sampling procedure. The Gelman–Rubin diagnostic evaluates MCMC convergence by analysing the difference between multiple Markov chains. The convergence is assessed by comparing the estimated betweenchains and within-chain variances for each model parameter. In each case, the Gelman–Rubin statistic was close to one (<1.1), suggesting that MCMC were able to converge.

We then compared parameter estimates using 95% CIs. Specifically, for each parameter in each group ('(Dis)Trust' vs '(Dis)Like', '(Dis)Trust' vs Baseline, '(Dis)Like' vs Baseline) we calculated the 95% CI. If the 95% CI of two groups do not overlap, we consider there to be a significant difference between the two feedback types compared. We also calculated 95% HDIs. For each comparison ('(Dis)Trust' vs '(Dis)Like', '(Dis)Trust' vs Baseline, '(Dis)Like' vs Baseline) we calculated the difference in the posterior distributions and reported the 95% HDI of the difference. If this HDI did not overlap zero, we consider there to be a meaningful difference between the two feedback types compared**.** HDI testing was conducted in R using *HDInterval* (Meredith & Kruschke, 2016).

To validate the DDM, we used each group's parameters obtained from participants' data to simulate log-transformed RTs and responses separately for each feedback type. We used the exact number of participants and total number of trials as in the experiments. Simulated data were then used to (1) perform model recovery analysis and (2) to compare the pattern of participants' response to the pattern of simulated responses, separately for each group. We sampled 2000 times from the posteriors, discarding the first 500 as burn in. Simulation and model recovery analysis were performed using the HDDM software toolbox (Wiecki et al., 2013). One-way ANOVAs were computed to examine if simulated data reproduced the behavioural pattern from the experimental data. To that end, discernment was entered into a one-way ANOVA with type of feedback as the independent variable for Experiment 2 and 3 separately. Note, that we did not enter veracity of the post into our DDM and instead entered responses as either 'veracity-promoting' (true post shared or false post skipped) or 'veracity-obstructing' (false post shared or true post skipped). Thus, discernment was calculated as the proportion of true posts shared and false posts skipped minus the proportion of true posts skipped and false posts shared.

# **4.4 Results**

**Participants' use 'trust' and 'distrust' buttons to discern true from false information (Experiment 1).**

In a first step, we examined whether participants used 'trust' and 'distrust' reactions to discern true from false information more so than 'like' and 'dislike' reactions. In Experiment 1, participants saw 100 news posts taken from the fact-checking website Politifact (https://www.politifact.com; **see Figure 4.2**). Half of the posts were true, and half were false. Participants were given the opportunity to react to each post using 'like', 'dislike', 'trust' and 'distrust' reaction buttons. They could select as many buttons as they wished or none at all ('skip'). Five participants were excluded according to pre-determined criteria (**see Methods for details**). Thus, one-hundred and six participants (52 Democrats, 54 Republican, M<sub>age</sub> = 40.745, SD<sub>age</sub>  $\pm$  14.479; female = 54, male  $= 52$ ) were included in the analysis.

We then examined whether participants used the different reaction buttons to discern true from false information. Discernment was calculated as follows, such that high numbers always indicate better discernment:

For 'like':

 $Discernment = Prop_{likes true} - Prop_{likes false}$ 

For 'dislike':

Discernment =  $Prop_{dislikes false} - Prop_{dislikes true}$ 

For 'trust':

 $Discernment = Prop<sub>trusts true</sub> - Prop<sub>trusts false</sub>$ 

For 'distrust':

$$
Discernment = Prop_{distrusts false} - Prop_{distrusts true}
$$

With *Prop* indicating the proportion of that response out of all true posts, or out of all false posts, as indicated.

These discernment scores were calculated for each participant separately and then entered into a 2 (**type of reaction**: 'trust' and 'distrust' / 'like' and 'dislike') by 2 (**valence:** positive, i.e., 'like', 'trust'/ negative, i.e., 'dislike', 'distrust') withinsubject ANOVA. Political orientation was also added as a between-subject factor (Republican/Democrat), allowing for an interaction of political orientation and type of reaction to assess whether participants with differing political beliefs used the reaction buttons in different ways.

The results reveal that participants' use of '(Dis)Trust' reaction buttons ( $M =$ 0.127;  $SE = 0.007$ ) was more discerning than their use of '(Dis)Like' reaction buttons (M = 0.047; SE = 0.005; F(1,104) = 95.832, p < 0.001**,** partial η2 = 0.48 **Figure 4.4**). We additionally observed an effect of valence  $(F(1,105) = 17.33, p$  $<$  0.001, partial  $n^2$  = 0.14), with negatively valenced reaction buttons (e.g., 'dislike' and 'distrust'  $M = 0.095$ ,  $SE = 0.007$ ) being used in a more discerning manner than positively valenced reaction buttons (e.g., 'like' and 'trust',  $M =$ 0.087,  $SE = 0.005$ ) and an effect of political orientation (F(1,104) = 25.262, p < 0.001, partial  $n2 = 0.2$ ), with Democrats (M = 0.115, SE = 0.007) being more discerning than Republicans ( $M = 0.06$ ,  $SE = 0.005$ ). There was also an interaction of type of reaction and political orientation ( $F(1,104) = 24.084$ , p < 0.001, partial  $n^2 = 0.19$ ), which was characterized by Democrats showing greater discernment than Republicans in their use of '(Dis)Trust' reaction buttons (F(1,104) = 33.592, p < 0.001, partial  $n/2$  = 0.24), but not in their use of '(Dis)Like' reaction buttons (F(1,104) = 2.255, p = 0.136, partial η2 = 0.02).

Importantly, however, both Democrats  $(F(1,51) = 93.376, p < 0.001$ , partial  $n^2$  $= 0.65$ ) and Republicans (F(1,53) = 14.715, p < 0.001, partial  $\eta$ 2 = 0.22) used the '(Dis)Trust' reaction buttons in a more discerning manner than the '(Dis)Like' reaction buttons.

One-sample t-tests against zero further revealed that participants' use of each reaction button discerned true from false information ('like':  $M = 0.06$ , SE = 0.006,  $t(105) = 10.483$ ,  $p < 0.001$ , Cohen's d = 1.018; 'trust': M = 0.099; SE = 0.01,  $t(105) = 9.744$ ,  $p < 0.001$ , Cohen's d = 0.946; 'dislike': M = 0.034; SE = 0.007; t(105) = 4.76,  $p < 0.001$ , Cohen's d = 0.462; 'distrust': M = 0.156; SE = 0.01,  $t(105) = 15.872$ ,  $p < 0.001$ , Cohen's d = 1.542).



**Figure 4.4. Participants use 'trust' and 'distrust' reactions to discern true from false information.** 'Trust and 'distrust' reactions were used in a more discerning manner than 'like' and 'dislike' reactions. Y axis shows discernment between true and false posts. For positive reactions (e.g., 'likes' and 'trusts') discernment is equal to the proportion of positive reactions for true information minus false information, and vice versa for negative reactions ('dislikes' and 'distrusts'). X axis shows reaction options. Data are plotted as boxplots for each reaction button, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5  $\times$  interguartile range. Diamond shape indicates the mean discernment per reaction. Individuals' mean

discernment data are shown separately as grey dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p < 0.001.

Thus far, we have shown that participants use '(Dis)Trust' reaction buttons in a more discerning manner than '(Dis)Like' reaction buttons. As social media platforms care about overall engagement not only its quality, we examined how frequently participants used the different reaction buttons. An ANOVA with the same specifications as above was conducted, but this time submitting frequency of reaction as the dependent variable. We found that participants used '(Dis)Trust' reaction buttons more often than '(Dis)Like' reaction buttons (Percentage use of reaction out of all trials: 'trust':  $M = 28.057\%$ ; 'distrust':  $M =$ 34.085%; 'like': M = 18.604%; 'dislike': M = 23.745%; F(1,104) = 36.672, p < 0.001, partial  $n2 = 0.26$ ). In addition, negative reaction buttons ('distrust' and 'dislike':  $M = 28.915\%$ ,  $SE = 1.177$ ) were used more frequently than positive reaction buttons ('trust' and 'like':  $M = 23.33\%$ ,  $SE = 1.133$ :  $F(1.105) = 16.96$ , p  $<$  0.001, partial  $n/2$  = 0.07). No other effect was significant. Interestingly, we also found that participants who skipped more posts were less discerning  $(R = -$ 0.414,  $p < 0.001$ ). Together, the results show that the new reaction options increase engagement.

The results hold when controlling for demographics, when not including political orientation in the analysis, and allowing for an interaction between type of reaction and valence (**see Appendix 7.3 Supplementary Tables 4.2 & 4.3**). The results also replicate in an independent sample (**see Appendix 7.3 Experiment 4**)**.** 

# **'Trust' and 'distrust' incentives improve discernment in sharing behaviour (Experiment 2).**

Thus far, we have shown that participants use '(Dis)Trust' reaction buttons in a more discerning manner than '(Dis)Like' reaction buttons. Thus, an environment which offers '(Dis)Trust' feedback is one where the number of 'carrots' (in the form of 'trusts') and the number of 'sticks' (in the form of 'distrusts') are directly associated with the veracity of the posts. It then follows that submitting participants to such an environment will increase their sharing of true information (to receive 'trusts') and reduce their sharing of misinformation (to avoid 'distrusts').

To test this, we ran a second experiment. A new group of participants ( $N = 320$ ) were recruited to engage in a simulated social media platform. They observed the same 100 posts (50 true, 50 false) shown to the participants in Experiment 1, but this time instead of reacting to the posts they could either share the post or skip it (**see Figure 3**). They were told that if they chose to share a post, it would be shared to their feed such that the other participants would be able to see the post and would then be able to react to it in real time (*feedback*). Depending on the environment participants were in, which varied betweensubjects, they could receive feedback in the form of the number of users who (i) '*disliked'*, or (ii) '*liked'*, or (iii) '*distrusted'*, or (iv) '*trusted'* their posts. We also included a (v) baseline condition, in which participants received no feedback. If participants selected to skip, they would observe a white screen asking them to click continue. Data of 32 participants were not analysed according to predetermined criteria (**see Methods for details**). Two-hundred and eighty-eight participants (146 Democrats, 142 Republicans,  $M_{age} = 38.073$ ,  $SD_{age} \pm 13.683$ ; female = 147, male = 141) were included in the analysis (**see Methods for details**).

$$
Discernment = Prop_{reposts\ true} - Prop_{reposts\ false}
$$

These scores were submitted to a between-subject ANOVA with type of feedback ('trust' & 'distrust'/ 'like' & 'dislike'/ Baseline), valence (positive, i.e., 'like' & 'trust' vs negative, i.e., 'dislike', 'distrust' vs neutral i.e., no feedback) and political orientation (Republican/Democrat) as factors. We also allowed for an interaction of political orientation and type of feedback.

We observed an effect of type of feedback  $(F(1,281) = 15.2, p < 0.001,$  partial  $\eta$ 2 = 0.051), such that participants shared more true than false information in the '(Dis)Trust' environments (M = 0.18,  $SE = 0.018$ ) than the '(Dis)Like' environments (M = 0.085, SE = 0.019, F(1,225) = 14.249, p < 0.001, partial  $n^2$  $= 0.06$ ) and Baseline environment (M = 0.084, SE = 0.025; F(1,150) = 10.906, p = 0.001, partial η2 = 0.068, **Figure 5a**). Moreover, participants who received 'trust' feedback ( $M = 0.176$ ,  $SE = 0.026$ ) were more discerning in their sharing behaviour than those who received 'like' feedback  $(M = 0.081, SE = 0.021,$  $F(1,131) = 10.084$ ,  $p = 0.002$ , partial  $p = 0.071$ ). Those who received 'distrust' feedback ( $M = 0.175$ ,  $SE = 0.026$ ) were more discerning than those who received 'dislike' feedback (M =  $0.092$ , SE =  $0.039$ , F(1,90) =  $5.003$ , p =  $0.028$ , partial  $n^2$  = 0.053). We further observed a trend interaction between type of feedback and political orientation (F(1,281) = 2.939,  $p = 0.055$ , partial  $p =$ 0.02). While Democrats (M =  $0.213$ ; SE = 0.014) were generally more discerning than Republicans (M = 0.017; SE = 0.016; F(1,281) = 77.392, p < 0.001, partial  $n^2 = 0.216$ ), this difference was smaller in those who received  $'(Dis)$ Trust' feedback (M = 0.082, SE = 0.034) compared to those who received '(Dis)Like' feedback (M = 0.23, SE = 0.03; F(1,224) = 4.879, p = 0.028, partial  $n2 = 0.021$ ) and by trend smaller than those who received no feedback (M = 0.229, SE =  $0.045$ ; F(1,149) = 3.774, p = 0.054, partial  $n^2$  = 0.025). There was no difference between the latter two (F(1,188) = 0.00, p = 0.988, partial η2 = 0.00). No other effects were significant. Overall engagement, measured as percentage of posts shared out of all trials, did not differ across environments  $(F(1,281) = 1.218, p = 0.271,$  partial  $p = 0.004$ ; Mean % posts shared out of all trials: Baseline = 27.712%; Dislike = 35.889%; Like = 33.258%; Distrust = 32.51%; Trust = 30.435%; **see Appendix 7.3 Supplementary Table 4.4 for means for true and false posts**).

Results hold when controlling for demographics, when not including political orientation in the analysis, and allowing for an interaction between type of reaction and valence (**see Appendix 7.3 Supplementary Table 4.5 & 4.6**). Results replicate in an independent sample (**see Appendix 7.3 Experiment 5**)**.** 

To recap - participants in Experiment 2 decided whether to share content or skip. They then observed the reactions of other participants to their post (they believed this was happening in real-time, but for simplicity we fed them reactions of participants from Experiment 1). Each participant in Experiment 2 observed only one type of feedback. For example, only 'distrusts'. How is it that observing 'distrusts' alone increases discernment? The rationale behind this design is that for any given post, true or false, some users will distrust the post. However, true posts will receive fewer 'distrusts' than false posts. It is the number of 'distrusts' per post that matters. Participants are motivated to minimize the average number of 'distrusts' they receive. To achieve this, they should post more true posts and fewer false posts. Of course, if participants were simply trying to minimize the *total* number of distrusts, they would just skip on every trial. Participants do not do that, however. Potentially because sharing in and of itself is rewarding (Tamir & Mitchell, 2012). The results indicate that participants are sensitive to the number of 'distrusts' per posts not just to the total number of 'distrusts' over all posts.

The same rationale holds for the participants that only observe 'trusts'. They receive more 'trusts' for true than false posts. It is the magnitude of 'trusts' that is associated with veracity. This motivates participants to post more true posts and fewer false posts in order to maximize the average number of 'trusts' per post. Of course, if participants were simply trying to maximize the total number of 'trusts', they would just share on every trial. Participants do not do that, however. This indicates that they are sensitive to the number of 'trusts' per post not just to total number over all posts. Any user of social media platforms could relate to this; when posting a tweet, for example, many people will be disappointed with only a handful of 'hearts'. The user's goal is to maximize positive feedback per post. The same rationale as above holds for 'likes' and 'dislikes' except that those are less associated with veracity, thus impact discernment less.

The posts included in the experiment covered a range of topics including politics, science, health, environment, and society. As observed in **Figure 4.5b**, the effect of '(Dis)Trust' environment on discernment is observed regardless of content type.
Thus far, our results show that changing the incentive structure of social media platforms by coupling the number of 'carrots' and 'sticks' with information veracity could be a valuable tool to reduce the spread of misinformation. If feedback promotes discernment in sharing behaviour, it is plausible that it may in turn improve belief accuracy. To test this, we asked participants at the end of the experiment to indicate how accurate they thought a post was on a scale from *inaccurate* (0) to *accurate* (100). Participants' error in estimating whether a post was true or false was calculated as follows: for false posts error was equal to the participants' accuracy rating and for true posts it was equal to 100 minus their rating. Participants' average error scores were entered into a between-subject ANOVA with type of feedback and valence of feedback, as well as political orientation and its interaction with feedback type. We observed an effect of type of feedback (F(1,281) = 7.084, p = 0.008, partial  $n^2 = 0.025$ ), such that participants were more accurate (less errors) when they received '(Dis)Trust' feedback (M = 47.24,  $SE = 0.938$ ) compared to '(Dis)Like' feedback  $(M = 50.553, SE = 0.851, F(1,224) = 7.024, p = 0.009, partial n2 = 0.03)$ . We further observed an effect of political orientation  $(F(1,281) = 11.402, p < 0.001,$  $\eta$ 2 = 0.039), with Democrats (M = 47.264, SE = 0.773) being more accurate than Republicans ( $M = 51.117$ ,  $SE = 0.802$ ). No other effects were significant. All results hold when controlling for demographics, when not including political orientation in the analysis, and allowing for an interaction between type of feedback and valence (**see Appendix 7.3 Supplementary Table 4.7)**. Results are replicated in an independent sample (**see Appendix 7.3 Experiment 5**)**.**



**Figure 4.5. Altering the incentive structure of social media environments increases discernment of information shared. (a) Participants operating in** an environment where '(Dis)Trust' feedback was introduced shared more true information relative to false information than participants operating in an environment where only '(Dis)Like' feedback was available, or no feedback at all (Baseline). Y axis shows discernment, i.e., proportion of true posts shared minus proportion of false posts shared. X axis shows the group environment (type of feedback). **(b)** This was the case regardless of the topic of the post (politics = turquoise, science = blue, health = olive, environment = salmon, society = pink, other = green). Bubble size corresponds to number of the posts included in the study. Diagonal dashed line indicates point of equivalence, where discernment in equal across the '(Dis)Like' and '(Dis)Trust' environments. As can be seen, all bubbles are above the dashed line indicating that in all cases discernment is greater in an environment that offers '(Dis)Trust' feedback. Y axis shows discernment in the '(Dis)Trust' environment, X axis shows discernment in the '(Dis)Like' environment. **(c)** Experiment 3 showed the same results as Experiment 2. Data are plotted as boxplots for each reaction, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate  $1.5 \times$  interquartile range. Diamond shape indicates the mean discernment per reaction. Individuals' mean discernment data are shown separately as grey dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p <  $0.001$ , \*\*p <  $0.01$ .

#### **'Trust' and 'distrust' incentives together improve discernment in sharing behaviour (Experiment 3).**

Given that Experiment 2 revealed that receiving 'trust' or 'distrust' feedback separately improves discernment, it is likely that the coupled presentation of both will jointly also improve discernment. To test this, we ran a third experiment with a new group of participants. The task was identical to Experiment 2 (**see Figure 4.3**), but this time we included three between-subject environments: a *Baseline* environment, in which participants received no feedback, a 'Trust & Distrust' environment, in which participants observed both the number of *trust* and the number of *distrust* feedback, and a 'Like & Dislike' environment, in which participants observed both the number of *like* and the number of *dislike* feedback.

Additionally, to ensure posts align equally with Democratic and Republican beliefs, in Experiment 3 we selected 40 posts (20 true, 20 false) in response to which Republicans and Democrats utilized the 'trust' button in a similar manner in Experiment 1 (**see Methods**). Data of 18 participants were not analysed according to pre-determined criteria (**see Methods for details**). Analysis of Experiment 3 ( $N = 391$ , 194 Democrats, 197 Republican, M<sub>age</sub> = 35.304, SD<sub>age</sub>  $\pm$  11.089; female = 196, male = 192, other = 3) was the same as in Experiment 2 except that there were three environments (Baseline, 'Like & Dislike', 'Trust & Distrust') and no valence of feedback, because all environments either include both positive and negative feedback or no feedback.

Discernment was submitted to a between-subject ANOVA with **type of feedback** (Baseline/ 'Like & Dislike' / 'Trust & Distrust'), political orientation and their interaction as factors. Again, we observed an effect of type of feedback (F(1,385) = 11.009, p < 0.001, partial η2 = 0.054, **Figure 4.5c**), with participants in the 'Trust & Distrust' feedback group posting more true relative to false

information ( $M = 0.101$ ,  $SE = 0.015$ ) than those in the 'Like & Dislike' group (M  $= 0.042$ , SE = 0.013; F(1,261) = 8.478, p = 0.00, partial  $n^2 = 0.031$ ) or those who received no feedback at all (M = 0.008, SE = 0.014, F(1,259) = 20.142, p  $<$  0.001, partial  $n^2$  = 0.0724). By contrast there was no difference between the latter two groups (F(1,250) = 2.981,  $p = 0.085$ , partial  $p = 0.012$ ). As observed in Experiment 2, Democrats ( $M = 0.073$ ,  $SE = 0.011$ ) were more discerning than Republicans (M = 0.031, SE = 0.012; F(1,385) = 6.409, p = 0.012, partial  $n^2$  = 0.016). No other effects were significant.

Interestingly participants shared more frequently in the 'Trust & Distrust' environment compared to the other two environments (% of all trials: 'Trust & Distrust' = 36.2%, 'Like & Dislike' = 30.41%; Baseline = 25.853%; F(1,385) = 8.7692,  $p < 0.001$ , partial  $n = 0.044$ ). This illustrates that '(Dis)Trust' feedback improves discernment without reducing engagement. No other effects were significant.

All results hold when controlling for demographics, when not including political orientation in the analysis, and allowing for an interaction between type of reaction and valence (**see Appendix 7.3 Supplementary Tables 4.8 & 4.9**). Results replicate in an independent sample (**see Appendix 7.3 Experiment 6**)**.**

At the end of Experiment 3, we again asked participants to indicate how accurate they thought a post was. Participants' average error scores were calculated as in Experiment 2 and entered into a between-subject ANOVA with type of feedback, political orientation and their interaction as factors. Democrats  $(M = 40.591$ ; SE = 6.371) were more accurate than Republicans  $(M = 42.056)$ ; SE = 5.633; F(1,385) = 5.723, p = 0.017, partial  $n^2$  = 0.015). No other effects were significant (for results when controlling for demographics, when not including political orientation in the analysis **see Appendix 7.3 Supplementary Table 4.10)**. Note, that in the replication study (Experiment 6) we did observe an effect of type of feedback  $(F(1, 147) = 4.596, p = 0.012,$  partial  $n2 = 0.059$ ), with '(Dis)Trust' being most accurate. Thus, we see accuracy effects in three (Experiment 2, 5, 6) out of our four studies.

Taken together these findings suggest that changing the incentive structure of social media platforms, such that 'carrots' and 'sticks' are strongly associated with veracity promotes discernment in sharing behaviour, thereby reducing the spread of misinformation.

**'(Dis)Trust' incentives improve discernment in sharing behaviour by increasing the relative importance of evidence consistent with discerning behaviour.** Next, we set out to characterize the mechanism by which the new incentive structure increased discernment. Imagine you observe a post on social media, and you need to decide whether to share it – how do you make this decision? First, you examine the post. Second, you retrieve your existing knowledge. For example, you may think about what you already know about the topic, what you heard others say, you may try to estimate how others will react to the post if you share it, and so on. This process is called 'evidence accumulation' - you gradually accumulate and integrate external evidence and internal evidence (memories, preferences etc.) to decide. Some of the evidence you retrieve will push you towards a 'good' response that promotes veracity (that is posting a true post and skipping a false post) and some will push you towards a 'bad' response that obstructs veracity (that is posting a false post and skipping a true post). We can think of the evidence that pushes you toward a response that promotes veracity as 'signal'. Using computational modelling it is possible to estimate how much a participant is influenced ('pushed') by signal relative to noise, by calculating a parameter known as a 'drift rate' in a class of models known as DDM. One possibility then is that in the '(Dis)Trust' environment evidence towards responses that promote veracity is given more weight than towards responses that obstruct veracity (that is the drift rate is larger), thus people make more discerning decisions.

Another, non-exclusive possibility, is that in the '(Dis)Trust' environment participants are more careful about their decisions. They require more evidence before making a decision. For example, they may spend more time deliberating the post. In DDM this is estimated by calculating what is known as the distance

between the decision thresholds (that is how much total distance do I need to be 'pushed' in one direction or the other to finally make a choice).

To test the above possible mechanisms, we modelled our data using the DDM (Lin et al., 2023; Ratcliff, 1978; Ratcliff & McKoon, 2008). We modelled participants' responses ('veracity-promoting' vs 'veracity-obstructing' choice) separately for each type of feedback ('(Dis)Trust', '(Dis)Like', Baseline) and each Experiment (Experiment 2 and 3). The following parameters were included: (1) t0—the amount of non-decision time, capturing encoding and motor response time; (2)  $\alpha$ —the distance between decision thresholds ('veracity-promoting' response vs 'veracity-obstructing' response); (3) z starting point of the accumulation process; and (4) v—the drift rate (Ratcliff, 1978; Ratcliff & McKoon, 2008; Voss et al., 2013).

We next examined which of the parameters were different in the different environments (**see Table 4.1 & 4.2**). To that end, we calculated 95% CIs of each parameter for each pair of incentive environments ('(Dis)Trust' vs '(Dis)Like', '(Dis)Trust' vs Baseline, '(Dis)Like' vs Baseline). If the 95% CIs do not overlap, we infer a significant difference between the two incentive environments.

For both Experiment 2 (**see Figure 4.6a**) and Experiment 3 (**see Figure 4.6c**) we observed a significant difference in the drift rate. In particular, in the  $'(Dis)$ Trust' environments the drift rate was larger (Experiment 2:  $v = 0.216$ ; 95% CI [0.17, 0.262]; Experiment 3:  $v = 0.12$ ; 95% CI [0.086, 0.155]) than in the'(Dis)Like' environments (Experiment 2: v = 0.01; 95% CI [0.056, 0.145]; Experiment 3: 0.037; 95% CI [0.002, 0.069]) or no feedback environment (Experiment 2: v = 0.098; 95% CI [0.039, 0.158]; Experiment 3: v = 0.006; 95% CI [-0.027, 0.037]). The Baseline and '(Dis)Like' environments did not differ for drift rate. This suggests that relative to the other environments, in the '(Dis)Trust' environments evidence consistent with a 'veracity-promoting' response is weighted more than 'evidence' consistent with a 'veracityobstructing' response. 95% HDI comparisons corroborate this result **(see** 

**Appendix 7.3 Supplementary Tables 4.11 & 4.12 for HDI Comparisons**). We replicate these results in Experiment 5 and Experiment 6 **see Appendix 7.3 Experiment 5 & 6**)**.** 



**Figure 4.6. '(Dis)Trust' feedback increases the drift rate.** Displayed are the posterior distributions of parameter estimates for the Baseline environment, the '(Dis)Like' environment and the '(Dis)Trust' environment. Dashed vertical lines indicate respective group means. **(a)** In both Experiment 2 and **(c)** Experiment 3 95% CI comparison revealed that participants had a larger drift rate in the '(Dis)Trust' environments than in the other environments. No significant difference was observed between the latter two environments. Recovered model parameter estimates reproduced experimental results for both **(b)** Experiment 2 and **(d)** Experiment 3. \*indicates significant difference between parameters (i.e., CI do not overlap).



Non- <b>Decision</b> Time $(t0)$	7.025 [6.898, 7.154]	6.936 [6.802, 7.071]	6.681 [6.425, 6.94]
<b>Starting</b> Point (z)	0.497 [0.486, 0.508]	$0.491$ [0.483, 0.50]	$0.48$ [0.471, 0.49]
<b>Drift Rate</b> (v)	0.098 [0.039, 0.158]	$0.10$ [0.056, 0.145]	$0.216$ [0.17, 0.262]

**Table 4.1. Group estimates for DDM in Experiment 2.**



**Table 4.2. Group estimates for DDM in Experiment 3.**

While in Experiment 2 the distance between decision thresholds in the Baseline environment was lower than in the other two environments, and non-decision time (t0) higher than in the '(Dis)Trust' environment, these differences are not replicated in Experiment 3**.** More importantly, neither distance between decision thresholds nor non-decision time differed between the '(Dis)trust' and '(Dis)like' environments (see **Table 4.1 & 4.2 and Appendix 7.3 Supplementary Tables 4.11 & 4.12 for HDI Comparisons**).

Model parameters could be successfully recovered with data simulated using group-level parameters from Experiment 2 and Experiment 3 separately (**see Methods for details**, **see Figure 4.6 b, d, Appendix 7.3 Supplementary Tables 4.13 & 4.14)**. This was done by fitting the model to simulated data, in the same way as for the experimental data. We sampled 2000 times from the posteriors, discarding the first 500 as burn in. The same pattern of results was reproduced with the simulated data as with real participants' data (**Figure 4.7**). For each Experiment we ran two separate one-way ANOVAs to assess the effect of type of feedback on discernment: one for the real data and one for the simulated data. We remind the reader that we entered responses into our DDM as either 'veracity-promoting' (true post shared or false post skipped) or 'veracity-obstructing' (false post shared or true post skipped). Thus, discernment here is calculated as:

> $Discernment = Prop_{\text{veracity-promoting responses}}$ − Prop<sub>veracity</sub>-obstructing responses

Which is equal to:

 $Discernment = Prop_{reposts true\ posts+reposts\ false\ posts}$  $-$  Prop<sub>reposts</sub> false posts+skips true posts

As expected, we observed an effect of type of feedback in both the simulated  $(Experiment 2: F(1, 285) = 3.795, p = 0.024, p = 0.026; Experiment 3: F(1, 388)$  $= 7.843$ ,  $p = 0.001$ ,  $p = 0.039$ , **Figure 4.7 b, d**), and the experimental data  $(Experiment 2: F(1, 287) = 7.049, p = 0.001, p2 = 0.047; Experiment 3: F(1, 388)$  $= 11.166$ ,  $p < 0.001$ ,  $p = 0.054$ ). That is, discernment was higher in '(Dis)Trust' environments relative to '(Dis)Like' environments or no feedback environments **(Figure 4.7 a, c, see Appendix 7.3 Supplementary Tables 4.15 & 4.16 for pairwise comparisons** and **Supplementary Table 4.17 for correlations between real and recovered individual-level parameters**).



**Figure 4.7. Simulated Data reproduced experimental findings.** In both **(a)** Experiment 2 and **(c)** Experiment 3 participants who received '(Dis)Trust' feedback were more discerning than participants in the '(Dis)Like' and Baseline environments. Simulated data reproduced these findings **(b&d)**. Y axis shows discernment, i.e., proportion of true posts shared and false posts skipped minus the proportion of true posts skipped and false posts shared. X axis shows feedback group. Data are plotted as boxplots for each reaction, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5 × interquartile range. Diamond shape indicates the mean discernment per reaction. Individuals' mean discernment data are shown separately as grey dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p < 0.001, \*\*p < 0.01, \*p < 0.05.

#### **4.5 Discussion**

Here, we created a novel incentive structure that significantly reduced the spread of misinformation and provide insights into the cognitive mechanisms that make it work. This structure can be adopted by social media platforms at no cost. The key was to offer reaction buttons (social 'carrots' and 'sticks') that participants were likely to use in a way that discerned between true and false information. Users who found themselves in such an environment, shared more true than false posts in order to receive more 'carrots' and less 'sticks'.

In particular, we offered 'trust' and 'distrust' reaction buttons, which in contrast to 'likes' and 'dislikes', are by definition associated with veracity. For example, a person may dislike a post about Joe Biden winning the election, however this does not necessarily mean that they think it is untrue. Indeed, in our study participants used 'distrust' and 'trust' reaction buttons in a more discerning manner than 'dislike' and 'like' reaction buttons. This created an environment in which the number of social rewards ('carrots') and punishments ('sticks') were strongly associated with the veracity of the information shared. Participants who were submitted to this new environment were more discerning in their sharing behaviour compared to those in traditional environments who saw either no feedback or 'dislike' and/or 'like' feedback. The result was a reduction in sharing of misinformation without a reduction in overall engagement. All the effects were replicated, and effect sizes of misinformation reduction were large to medium.

Using computational modelling we were able to pin-point the changes to participants' decision-making process. In particular, DDM revealed that participants in the new environment assigned more weight to evidence consistent with discerning than non-discerning behaviour relative to traditional environments. In other words, the possibility of receiving rewards that are consistent with accuracy led to an increase in the weight participants assigned to accuracy-consistent evidence when making a decision. 'Evidence' likely includes external information that can influence the decision to share a post (such as the text and photo associated with the post) as well as internal information (e.g., retrieval of associated knowledge and memories).

Our results held when the potential feedback was only negative ('distrust'), only positive ('trust'), or both ('trust' and 'distrust'). While negative reaction buttons were used in a more discerning manner and more frequently than positive reaction buttons, we did not find evidence for a differential strength of positively or negatively valenced feedback on discernment of sharing behaviour itself.

The findings also held across a wide range of different topics (e.g., politics, health, science, etc.) and a diverse sample of participants, suggesting that the intervention is not limited to a set group of topics or users, but instead relies more broadly on the underlying mechanism of associating veracity and social rewards. Indeed, we speculate that these findings would hold for different 'carrots' and 'sticks' (beyond 'trust' and 'distrust'), as long as people use these 'carrots' and 'sticks' to reward true information and punish false information. However, we speculate that the incentives were especially powerful due to being provided by fellow users and easily quantifiable (just as existing buttons including 'like' and 'heart'). This may contrast with incentives which are either provided by the platform itself and/or not clearly quantified such as verification marks (Edgerly & Vraga, 2019) or flagging false news (Brashier et al., 2021; Chan et al., 2017; Grady et al., 2021; Lees et al., 2022). Interestingly, a trust button has also been shown to increase sharing of private information (Bălău & Utz, 2016).

Finally, we observed that feedback not only promotes discernment in sharing behaviour but may also increase the accuracy of beliefs. Though, while we see an increase in accuracy of beliefs in three of the four experiments, we did not observe this effect in Experiment 3. Thus, the new incentive structure reduces the spread of misinformation and may help in correcting false beliefs. It does so without drastically diverging from the existing incentive structure of social media networks by relying on user engagement. Thus, this intervention may be a powerful addition to existing intervention such as educating users on how to detect misinformation (Lewandowsky & van der Linden, 2021; Maertens et al., 2021; Pilditch et al., 2022; Roozenbeek & van der Linden, 2019; Traberg et al., 2022) or prompting users to think about accuracy before they engage in the platform (Capraro & Celadin, 2022; Fazio, 2020; Pennycook & Rand, 2022a). Over time, these incentives may help users build better habits online (Anderson & Wood, 2021; Ceylan et al., 2023).

As real-world platforms are in the hands of private entities, studying changes to existing platforms requires testing simulated platforms. The advantage of this approach is the ability to carefully isolate the effects of different factors. However, real-world networks are more complex and involve additional features which may interact with the tested factors. Our hope is that the science described here will eventually impact how privately owned platforms are designed, which will reveal whether the basic mechanisms reported here hold in more complex scenarios.

This study lays the groundwork for integration of the new incentive structure into existing (and future) social media platforms to further test the external validity of the findings. Rather than removing existing forms of engagement, we suggest an addition that complements the existing system and could be adopted by social media platforms at no cost. The new structure could subsequentially help reduce violence, vaccine hesitancy and political polarization, without reducing user engagement.

# **Chapter 5: General Discussion**

### **5.1 Synthesis**

The research detailed in this thesis sheds new light on whether and how, prominent incentives to gather and share information can be altered to reduce biased evidence accumulation and make sharing decisions more discerning.

In the first part of this thesis, I focus on evidence accumulation as the pathway of decision-making (Platt & Glimcher, 1999). Prior research illustrates that individuals selectively accumulate evidence to form positive beliefs (Gesiarz et al., 2019; Leong et al., 2019) from which they derive internal rewards, such as positive emotions, and increased sense of self-efficacy (Bromberg-Martin & Sharot, 2020; Loewenstein, 2006; Sharot et al., 2023). Often, they prioritize internal incentives over external incentives, such as financial gains. This results in a bias towards desirable, internally rewarding conclusions (Gesiarz et al., 2019; Leong et al., 2019). Such positively biased beliefs can have detrimental consequences, such as lack of preparation for natural disasters (Paton, 2003).

In Chapter 2 and 3, I investigate if and how incentives can be altered to make evidence accumulation less biased. I hypothesize that evidence accumulation becomes less biased towards desirable conclusions when false beliefs resulting in false conclusions are costly. I test this in two ways: In Chapter 2, I examine how perceived threat impacts evidence accumulation. I find that under threat the relative rate at which negative evidence is accumulated increases and the bias towards desirable conclusions disappears. This may be adaptive, as it can lead to increased precaution in environments in which the risk of adverse consequences is high. Building on this work, in Chapter 3, I then examine whether increasing the external incentives to reach accurate conclusions can also alter the influence of internal incentives on evidence accumulation. Results show that while accuracy incentives led participants to take more time to reach a conclusion, they did not impact participants' bias. I provide a mechanistic explanation for why this might be. In particular, DDM

reveals that accuracy incentives and the desirability bias act on orthogonal aspects of the accumulation process. While accuracy incentives led to an increase in the distance between decision thresholds, the bias was associated with greater weight on desirable relative to undesirable evidence (drift rate bias). These results suggest that participants may not be aware of their own bias. I suggest that when accuracy incentives are coupled with direct feedback, decision-making becomes more discerning.

One decision that is informed by evidence accumulation is the decision to share information (Globig et al., 2023; Huang et al., 2015; Lin et al., 2023). In recent years, the rising ease of sharing, brought about by the advent of the internet, has also facilitated the dissemination of misinformation (Kreps, 2020). This has been attributed to the existing incentive structure of social media platforms (Brady et al., 2021; Lindström et al., 2021; Scissors et al., 2016), in which social rewards ('likes') and punishments ('dislikes') are dissociated from the veracity of the information shared. In the Chapter 4, I therefore test the hypothesis that the role of this incentive structure in promoting misinformation spread can be curtailed by making social rewards and punishment contingent on the veracity of the information shared. Results show that when social incentives are contingent on the veracity of information, discernment in sharing behaviour increases which in turn reduces the spread of misinformation.

Taken together, this thesis provides insights into how incentive structures can be altered to improve the quality of evidence accumulation and sharing behaviour. This discussion will summarize the key findings of the studies in each chapter, and delve into their implications, limitations, and potential future pathways for this research.

### **5.2 Under threat weaker evidence is required to reach undesirable conclusions**

#### **5.2.1 Summary**

Many important decisions are made in threatening environments. Such settings, often characterized by heightened stress and anxiety, are known to have adverse effects on both learning and decision-making (FeldmanHall et al., 2015; Porcelli & Delgado, 2009, 2017; Raio et al., 2013; Starcke & Brand, 2012). More recent work, however, highlights that stress can also have an adaptive function (Akinola & Mendes, 2012; Garrett et al., 2018; Graybeal et al., 2011).

In Chapter 2, I demonstrate that stress induced by perceived threat also alters the process by which evidence is accumulated in a way that may be adaptive. Ninety-one participants completed a sequential sampling task in which they were incentivized to accurately judge whether they were in a desirable state, which was associated with greater rewards than losses, or an undesirable state, which was associated with greater losses than rewards (Gesiarz et al., 2019). Participants were assigned to one of two groups: a 'threat group' and a 'control group'. Prior to the task, participants in the threat group experienced a socialthreat manipulation. As expected, we found that participants in the control group were biased towards desirable conclusions. They weighed desirable evidence more than undesirable evidence (replicating previous findings from Gesiarz et al., 2019). Under threat this bias disappears. Relative to the control group, participants in the threat group required weaker evidence to reach an undesirable conclusion. DDM revealed that this was due to an increase in the relative rate at which negative evidence is accumulated. This is line with previous findings that stress increases attention to negative stimuli, resulting in fewer decision errors (Akinola & Mendes, 2012), and bolsters the integration of negative information, thereby moderating the bias for positive information in belief updating (Garrett et al., 2018). Such increased attention may also, in part, explain the observed effect of stress on the rate of evidence accumulation. The results reported in Chapter 2 extend previous research by shedding light on

how perceived threat alters the bias towards desirable conclusions in evidence accumulation and show that perceived threat reduces the strength of evidence required to reach an undesirable conclusion. As the source of the threat in this study was unrelated to the behavioural task this threat-induced change in evidence accumulation points to a global effect of stress. Taken together, these results lend support to the hypothesis that when participants are exposed to an environment in which the costs of false beliefs are high, such as a threatening environment, they are less biased towards internally rewarding conclusions. This study thus suggests that individuals adaptively prioritize internal and external incentives depending on the environment.

#### **5.2.2 Limitations and Future Directions**

Chapter 2 sheds light on how perceived threat alters evidence accumulation and provides support for the idea that when the costs of false beliefs are high, the motivational bias towards internally rewarding conclusions is mitigated. This is in line with prior research, showing that stress improves the ability to flexibly adjust previously learned behaviours to fit new task requirements (Graybeal et al., 2011) and triggers defence mechanisms (Baas et al., 2006; Cornwell et al., 2007), especially in the face of aversive stimuli (Blanchard et al., 2011; Davis et al., 2010; Grillon, 2008; Robinson et al., 2012). However, it is important to bear in mind some limitations of this study and future directions for this work.

Anecdotal evidence illustrates that psychophysiological stress-responses often provide a global rather than specific signal of threat. For example, stress elicited by personal conflict may impact professional performance (Piotrkowski, 1979). This fits with prior research reporting a general effect of stress on behaviour and neural responses, that goes beyond the source of the threat itself, in both humans (Cavanagh, Frank, et al., 2011; Lenow et al., 2017; Otto & Daw, 2019; Robinson, Overstreet, et al., 2013; Youssef et al., 2012) and non-human animals (Harding et al., 2004; Rygula et al., 2013). As such, in this study, the source of the threat (anticipating a stressful event) was intentionally disconnected from the behavioural task (the 'Factory Task'). As expected, the

results point to a global effect of perceived threat on evidence accumulation. This may be adaptive, as it enhances the likelihood of taking precautions against potential aversive outcomes, which tend to be more extreme and/or common in threatening environments.

However, there is some evidence to suggest that the effect of global stress may differ from the effect of task-specific stress. In some instances, the latter has even been found to induce biased cognition (for a review see Yu, 2016). For example, a study Engelmann et al., (2019) observed that participants who engaged in a visual pattern recognition task, in which some patterns were associated with receiving an electric shock, were worse at recognizing these patterns, relative to patterns that were not associated with a shock. This suggests that participants engaged in 'wishful thinking' to deceive themselves about the looming electric shock. This stands in contrast to the study described in Chapter 2, in which the threat-induction was decoupled from the behavioural task and participants were motivated to hold one belief over another using financial incentives. In Engelmann et al., (2019) the threat itself motivated participants to form a desirable conclusion. This in line with suggestions, that people at risk of terminal illnesses avoid negative information to remain optimistic and protect their emotional well-being (Ganguly & Tasoff, 2017; Lerman et al., 1998; Oster et al., 2013). As participants have no (perceived) control over the incoming threat, this response is likely adaptive. This may also explain reports of a reduction in attentional bias towards aversive stimuli under threat (Jiang et al., 2017). In the study outlined in Chapter 2, the threat response is unspecific, and thus participants are alerted to potential dangers in their environment. I thus speculate, that when accumulating evidence under threat, prioritizing external over internal incentives is an adaptive response to allow participants to adopt defensive mechanisms (see also Baas et al., 2006; Cornwell et al., 2007); especially in response to threatening stimuli (Blanchard et al., 2011; Davis et al., 2010; Grillon, 2008). Taken together, these studies suggest, that stress has an adaptive function designed to restore homeostasis (de Kloet et al., 2008) and therefore its' precise effect on decision-making depends on the specific context (Starcke & Brand, 2012). Future studies should

further investigate this hypothesis and could explore to what extent the findings from Chapter 2 hold in different contexts, using different threat induction methods, such as cold-pressor tasks (Lenow et al., 2017; Otto et al., 2013), and threat of electric shock (Robinson, Overstreet, et al., 2013), or different task designs. Moreover it would be interesting to explore how global threat as used in Chapter 2 interacts with a task-induced motivational manipulation which uses threat as adopted by Engelmann et al., (2019).

In Chapter 2, I establish a causal relationship between perceived threat and evidence accumulation in healthy individuals. An intriguing avenue for future research is to examine how these findings might apply to chronic stress and individuals with affective disorders, who often report heightened anxiety. It has been suggested that chronic stress has adverse effects on cognitive and neural functions (for a review see Sousa & Almeida, 2012) and may have a more pronounced effect on decision-making compared to acute stress (Hales et al., 2016). In individuals, who are hypersensitive to stress, enhanced accumulation of undesirable stimuli could therefore be maladaptive and result in overly pessimistic predictions, further aggravating distress. Indeed, Aylward et al. (2019) observed a diminished drift rate toward desirable conclusions in individuals with anxiety and mood disorders, when studying the effect of stress in a two-alternative-forced-choice task, in which participants had to discriminate between two types of auditory tones (high reward tone vs low reward tones). Future research should explore this effect in a sequential sampling task, like the 'Factory Task' (Gesiarz et al., 2019) used here. This would allow us to examine whether the mechanism by which stress impacts evidence accumulation identified in Chapter 2 could serve as a potential target mechanism for therapeutic interventions to restore healthy cognitive processing.

A further potential limitation of the study is the absence of physiological measures, such as salivary cortisol levels, to confirm the efficacy of the elicited stress response. In this experiment, anxiety was induced using a variation of the Trier Social Stress Test (Birkett, 2011; Marteau & Bekker, 1992; Spielberger et al., 1970) and assessed the success of the manipulation using a wellestablished state anxiety questionnaire (Allen et al., 2014). While this procedure was validated using salivary cortisol levels and skin conductance in a previous study from our lab (Garrett et al., 2018), recording cortisol as a biomarker for stress, may have provided a more objective assessment of the stress response, complementing the study's findings based on behavioural data (for a detailed discussion on this topic see Allen et al., 2014). Furthermore, the effect of perceived threat on decision-making varies as a function of sensitivity to cortisol response (Kudielka et al., 2009; Van den Bos et al., 2009). Thus, the precise effect of perceived threat on evidence accumulation may also be contingent on individual differences in sensitivities to threat. Future studies should record salivary cortisol changes to provide further insights into the impact of these and other differences on the desirability bias in evidence accumulation. Neuroimaging could further complement these findings. Additionally, it is plausible that a moderate level of stress may enhance behaviour adaptively, while minimal or excessive stress, as often seen in clinical disorders, might impair performance, thereby suggesting the presence of an inverted U-shaped relationship. This hypothesis aligns with existing literature, which reports an inverted U-shaped correlation between stress and memory performance (for a review see Lupien et al., 2007). Future research should aim to empirically test this hypothesis using cortisol levels to further elucidate the nuanced effects of stress on biased cognition.

In conclusion, despite its limitations, the present study supports the idea that evidence accumulation is a dynamic process in which individuals adaptively prioritize internal and external incentives depending on their costs and benefits in a given environment. This discussion has also highlighted potential directions for future research, which would contribute to a more comprehensive understanding of this process.

## **5.3 Futile Rewards: Why Accuracy Incentives Fail to Reduce Biased Evidence Accumulation**

#### **5.3.1 Summary**

When accumulating evidence to make a decision, individuals are biased by the desire to form positive beliefs, from which they derive internal rewards, such as positive emotions. They assign greater weight to evidence that aligns with their desired belief. As a result, they are more likely to reach a conclusion that confirms this belief, even if it is incorrect (Gesiarz et al., 2019; Globig et al., 2021; Leong et al., 2019). Such false beliefs can lead to suboptimal decisionmaking (for a review see Karayanni & Nelken, 2022) which can have detrimental consequences, such as falsely optimistic investment decisions preceding the financial crisis (Shefrin, 2015). One obvious solution is to financially incentivize people to form more accurate beliefs in situations where biases and heuristics are common. This idea is predicated on the understanding that biases and heuristics can be moderated through deliberate and effortful thinking (Bonner & Sprinkle, 2002; Botvinick & Braver, 2015; Smith & Walker, 1993). Yet, despite the intuitive appeal of this approach, the empirical evidence is mixed (Engelmann et al., 2019; Prior et al., 2015; Rathje et al., 2023; Zhang & Rand, 2023).

In Chapter 3, I demonstrate that accuracy incentives fail to counteract biased evidence accumulation and provide a mechanistic explanation for why this might be. Over three experiments, participants  $(N = 235)$  completed a perceptual evidence accumulation task in which they had to judge whether they were in a desirable state, which was associated with greater rewards, or an undesirable state, associated with no reward. In some trials they were also financially incentivized for correct responses. Crucially participants had no control over which type of trial they were in and had to be as accurate as possible to maximize their rewards. The results show that while participants were more cautious about their conclusions when incentivized for accuracy, they remained biased towards desirable conclusions. This was true even when the reward for correct responses was increased five-fold.

These results fit with previous suggestions that even substantial accuracy incentives fail to alleviate certain biases, including the partisan bias in advice taking (Zhang & Rand), and anxiety-induced wishful thinking (Engelmann et al., 2019). The current findings go beyond these previous demonstrations and provide a mechanistic explanation for why accuracy incentives may fail to reduce biased evidence accumulation. DDM revealed that accuracy incentives and the bias towards internally rewarding conclusions alter orthogonal elements of the accumulation process. While the bias selectively increases the rate of evidence accumulation towards desirable conclusions, accuracy incentives increase the distance between decision thresholds, thereby increasing the amount of evidence participants required to make a decision, but not changing the way in which this evidence is accumulated.

Taken together, these findings lend support to previous suggestions that internal incentives alter pre-conscious processing of information and the resultant motivational biases are not merely the product of overt preferences but instead beyond individuals' awareness (Chen & Krajbich, 2018; Desai & Krajbich, 2022; Krajbich, 2022; but see Sánchez-Fuenzalida et al., 2023). This study lends support to the hypothesis that the inherent subjective value of internal rewards exceeds the subjective value of external rewards in safe environments. The findings are particularly relevant for policymakers and industry leaders as they may explain why financial bonuses may be ineffective in improving decision-making.

#### **5.3.2 Limitations and Future Directions**

Chapter 3 provides a mechanistic explanation for why accuracy incentives may fail to mitigate the influence of the motivational bias on evidence accumulation. These results offer valuable insights for both private and public sectors and may help explain why financial bonuses do not always result in more rational decision-making. Nevertheless, it is essential to note the limitations inherent in this study and to outline future research directions.

The findings from Chapter 3 seemingly stand in contrast to those of Chapter 2. In theory, both perceived threat and accuracy incentives can increase the external costs of holding false beliefs. Participants assign greater weight to undesirable than desirable evidence under threat, but not when incentivized for accuracy. This could be because stress has a global effect on evidence accumulation. The stress response thus alters both conscious and unconscious processes. In contrast monetary incentives seemingly operate via more highlevel conscious processes.

Our results indicate that when incentivized for accuracy, participants are more cautious about their responses. Yet, despite their increased effort, they remain biased towards desirable conclusions. This suggests that participants' biased perception is sincere, and thus cannot be easily modulated through financial incentives. This fits with prior work on motivated perception positing that the influence of internal incentives extends to pre-conscious processing of information (Balcetis & Dunning, 2006,). By contrast, perceived threat elicits a global stress response, which can be both implicit and explicit (Verkuil et al., 2009) and thus also alters the implicit influence of internal incentives on evidence accumulation. I speculate, that the reason accuracy incentives fail to reduce biased evidence accumulation, is that participants are not aware of their mistakes. Their biased processing of information is genuine and not a result of deliberation. This is in line with suggestions that preferential belief updating for good news compared to bad news occurs even when there is a time limit or participants have reduced cognitive load (Kappes & Sharot, 2019). It is possible that allowing participants to learn about the inaccuracy of their beliefs, for instance by providing feedback would help mitigate biased processing (but see Engelmann et al., 2019). Indeed, in Chapter 4, I found that participants who received feedback about the accuracy of the content they shared in simulated social media environments, were better at discerning between true and false posts, than those who did not receive accuracy-related feedback. Similar results were also reported by Zhang and Rand (2023) who observed that while accuracy incentives failed to mitigate partisan bias in advice-taking, allowing participants to learn about their bias through feedback, successfully reduced the partisan bias and improved accuracy. I thus hypothesize that if participants learn that their beliefs are false, they adjust how they weigh the evidence they observe, reducing the bias towards desirable conclusions. Future studies, for instance using neuroimaging and eyetracking, should test this hypothesis and investigate whether the preferential sampling of desirable evidence is indeed sincere and not deliberate (for a discussion see Krajbich, 2022).

In addition to this, Chapter 2 and Chapter 3 are not entirely comparable, as they utilize different paradigms. While Chapter 2 utilizes a sequential sampling task, Chapter 3 uses a perceptual accumulation task. Although the bias in the drift rate towards desirable conclusions is consistently observed across both tasks, a reliable comparison between the two ways of increasing external incentives requires identical tasks designs. Future research should therefore test if accuracy incentives remain ineffective when a sequential sampling task, like the 'Factory Task', is used.

Finally, it is important to acknowledge, that although the effect of accuracy incentives on response times was observed consistently across experiments in Chapter 3, the effect sizes were small. Interestingly, the effect of accuracy incentives on response times increased when the reward was increased fivefold. While it is possible that further increasing the magnitude of accuracy incentives could augment caution in such a way that participants become less biased, I speculate that while caution would increase, a reduction of the bias is unlikely given that the two alter distinct elements of the accumulation process. Indeed, prior research using substantially larger financial accuracy incentives also failed to show a reduction of motivational biases (Enke et al., 2023; Zhang & Rand, 2023). Nevertheless, future studies should examine whether changing the type of accuracy incentive, e.g., using a non-monetary accuracy incentive, or using punishment instead of reward would alter these results.

In conclusion, Chapter 3 sheds light on why accuracy incentives might fail to debias evidence accumulation. This discussion has highlighted important limitations and potential directions for future research, which would contribute to a more comprehensive understanding of the influence of internal incentives on evidence accumulation and how to mitigate it.

## **5.4 Changing the Incentive Structure of Social Media Platforms to halt the Spread of Misinformation**

#### **5.4.1 Summary**

The wide-spread success of social media platforms has been attributed to the human need for social rewards (Nadkarni & Hofmann, 2012). People share information in order to receive rewards, in the form of engagement, and in turn this engagement reinforces sharing behaviour (Lindström et al., 2021; Rosenthal-von der Pütten et al., 2019; Scissors et al., 2016). In Chapter 4, I demonstrate that the existing incentive structure of social media platforms, in which social rewards (e.g., 'likes') and punishments (e.g., 'dislikes') are dissociated from the veracity of the information shared, contributes to the spread of misinformation online. I show that slightly altering this incentive structure, such that social rewards and punishments are contingent on the veracity of information, improves discernment in sharing behaviour. Recent studies corroborate this result (Alizadeh et al., 2023; Butler et al., 2023; Kapoor et al., 2023; Pretus et al., 2023; Rathje et al., 2023).

Over six experiments, 951 participants engaged in simulated social media platforms where they encountered both true and false information. In some platforms 'trust' and 'distrust' reaction buttons, which are, by definition, related to veracity, were added to the existing 'like' and 'dislike' reaction buttons. We found that participants used the buttons in a more discerning manner than the existing engagement options, thereby creating an environment in which the number of social rewards and punishments were strongly associated with the veracity of the information shared. Other participants who were then exposed to this environment subsequently shared more true relative to false posts than those in traditional environments. This improved the overall quality of the information shared thereby reducing the spread of misinformation without reducing overall engagement. DDM revealed that the underlying mechanism of this effect is associated with an increase in the weight participants assign to evidence consistent with discerning behaviour.

Taken together, these results offer evidence for an intervention that capitalizes on the reward-based engagement structures of social media platforms. By creating an environment in which engagement is contingent on the veracity of the information shared, platforms could create healthier information ecosystem without reducing overall engagement. This study therefore lays the groundwork for an intervention which could easily be incorporated into a variety of different existing social media platforms to reduce the spread of misinformation and thus mitigate its devastating consequences for society.

#### **5.4.2 Limitations and Future Directions**

The results of this study provide compelling evidence for an intervention that leverages the existing incentive structure of social media platforms and could easily be adopted in the real world. However, there are several limitations which bear consideration.

First, it is unclear to what extent the efficacy of the intervention generalizes to real-world contexts. The study utilizes a highly controlled simulated social media environment. Although such approximations of real-world situations are vital to establish proof-of-concept (Charness & Fehr, 2015; Falk & Heckman, 2009), they also bring about several limitations which constrain the ecological validity of the study. For instance, participants in this study were either given the opportunity to react to posts (Experiment 1 & 4), or to share them (Experiment 2-3 & 5-6), never both. Moreover, in the latter experiments, participants were assigned to one of three between-subject conditions. They either saw no feedback, 'like' and/or 'dislike' feedback, or 'trust' and/or 'distrust' feedback. By contrast, in real social media platforms users can react and share at the same time and are also able to see a range of social feedback simultaneously. Furthermore, the design of these experiments was such, that

participants first decided whether to repost a piece of content or not, and only then would they receive feedback. This was done to ensure that learning what others thought about a specific piece of content could not directly impact whether participants would repost that *specific* piece. Instead, the idea was that knowing that others would be able to 'trust' or 'distrust' the post would motivate participants to consider the veracity of the post when making their decision. However, when users decide to repost something in real social media platforms, they can often already see how others have reacted to the original post. Thus, future studies should examine whether the observed effects hold when users can react and share at the same time and are able to see how others have reacted to the original post.

The prevalence of misinformation online has been attributed to large rumour cascades which originate from so-called 'super-spreaders' with high-follower counts (Mosleh & Rand, 2022; Vosoughi et al., 2018). In line with this, 0.01% of Twitter (now X) users generated 80% of the misinformation surrounding the US election (Grinberg et al., 2019). As such, the reposting of other-generated content plays a pivotal role in the spread of misinformation. In Chapter 4, I therefore focused on sharing other-generated, rather than sharing selfgenerated content. This not only mimics prevalent real-world behaviour, but also ensures that all participants in the study observe the same content thereby increasing the reliability of the study. Nevertheless, in real platforms users are able to share their own content, as well as repost that of others. While I predict that anticipating others' '(Dis)Trust' feedback influences reposting and sharing behaviour equally, future studies should directly test this hypothesis. Furthermore, participants in this study only observed news-style content which, was either true or false. Future studies should investigate how users utilize the '(Dis)Trust' buttons on other types of content, such as personal interest updates or opinion pieces. As '(Dis)trust' feedback is intended to complement existing '(Dis)Like' feedback, I predict that participants will primarily use the latter when interacting with social content, and the former when interacting with verifiable content.

In addition to utilizing feedback to make sharing decisions, social media users may also (rightly or wrongly) use a range of other signals to infer the reliability of the content they encounter online, such as the source of the information (Epstein et al., 2019). To reduce confounds, this study did not provide information about the identity of the original poster or the source of the content. Future studies should investigate how these different signals interact. Furthermore, it is important to bear in mind that the sample in this study was politically balanced. As such, the feedback participants received was also politically balanced. Future studies should investigate how this intervention behaves in organic polarized networks. For instance, in this study Democrats were more discerning in their use of '(Dis)Trust' buttons than Republicans, as such it is possible that feedback obtained from a primarily Republican sample would be less discerning. In line with this, due to ethical considerations users remained anonymous in this study. However, it is possible that the use of the '(Dis)Trust' buttons as well as the sensitivity to such feedback may vary depending on the social proximity of the parties involved. Indeed it has been suggested that social media interactions differ depending on the proximity of the users (Stopczynski et al., 2018). Another limitation of this study is that it only investigated behaviour at one timepoint. It did not assess the longevity of the intervention. Future research should investigate whether the results hold long-term or whether users habituate to '(Dis)Trust' feedback over time and become less responsive to it.

To further aid in the adoption of novel interventions into real-world platforms, it is essential not only to consider whether they achieve their desired effect in reducing the spread of misinformation, but also to investigate how they influence holistic user behaviour. For instance, some platforms use algorithms which do not primarily rely on user interactions, but instead utilize more passive methods of engagement (Geers et al., 2024), such as dwell-time, that is the time spent looking at a piece of content (Epstein et al., 2022). As such, another limitation of this study is that it solely focusses on discernment in sharing behaviour. Future research should therefore also take into account other variables, such measures of attention, that can be utilized by algorithms to

determine the dissemination of content. Relevant insights could for instance be obtained using eye tracking to assess attention.

To systematically address these limitations and examine the holistic impact of this (and other) intervention(s) on user behaviour, future studies should test the efficacy of this intervention in real-world settings. However, as platforms are seemingly wary of directly implementing proposed interventions, possibly for fear of adverse effects, novel solutions are necessary. For instance, some researchers are trying to develop immersive social media simulations, that are modelled on existing platforms, such as Twitter (now X) or Facebook and have varying degrees of experimental control (for an overview see Jagayat & Choma, 2023). While the existing prototypes of these platforms are either not fully interactive or not scalable to larger audiences, they hold great promise for both academia and industry alike. By bridging the gap between highly controlled experimental settings and the real world, this type of hyper-realistic social media simulation could not only provide insights into how this intervention impacts user behaviour 'in the wild' but it would also aid in the development of more optimal solutions to combatting the spread of misinformation online.

Another limitation of the study is that it does not contain a baseline measure of participants' belief accuracy. This was due to a variety of reasons. First, it has been shown that directly asking users to indicate whether they think a post is true or false alters sharing behaviour (for a review see Pennycook & Rand, 2022b). To rule out that such accuracy prompts could confound the influence of social feedback, belief accuracy was only assessed once participants had already made their sharing decisions. Second, it has been shown that the repeated exposure of information alters its perceived accuracy and thus increases the likelihood of sharing. To avoid such interference, no initial (baseline) beliefs were recorded. Thus, it is possible that existing betweengroup differences prior to the intervention contributed to the observed effect on belief accuracy. Furthermore, it has been shown that participants asymmetrically update their beliefs when learning about others' opinions (Kappes et al., 2020). Finding out others agree with them significantly increases their confidence. By contrast, finding out that others disagree only marginally diminishes their confidence. It then follows that when participants observe social feedback they disagree with, they do not update their beliefs, as much as for positive feedback (Kappes et al., 2020). Future studies should strive to tease apart this effect of valence on belief updating in response to online feedback from others.

In recent years there are also growing concerns about non-genuine user activity online. It has been suggested that some social media users deliberately spread misinformation to pursue specific geopolitical objectives (Bradshaw & Howard, 2019). Moreover, many of the accounts disseminating misinformation have been identified as social bots (Polychronis & Kogan, 2023; Shao et al., 2018). Therefore, at a time where large language models are continuously gaining popularity and provide unprecedented access to this technology, there is growing concern that social media companies provide platforms which can easily be leveraged to maliciously shape public opinion (Hajli et al., 2022; Himelein-Wachowiak et al., 2021). It is thus crucial to examine whether the '(Dis)Trust' button is susceptible to this type of abuse. Thus far, I have studied the impact of genuine '(Dis)Trust' feedback. Future studies should assess whether the results generalize to environments in which the feedback is nonsensical, i.e., the buttons are used randomly and not in a manner that signals the veracity of the content, as well as whether the results generalize to environments in which those using the '(Dis)Trust' button pursue a specific misinformation narrative, and thus does not provide a genuine 'trust' signal.

Misinformation is thought to have contributed to increasing polarization, racism and resistance to climate action and vaccines (Barreto et al., 2021; Rapp & Salovich, 2018; Tsfati et al., 2020; Van Bavel et al., 2021). In light of these detrimental consequences, it is unsurprising that researchers are devoting significant amounts of time to developing new interventions (for a review see Ziemer & Rothmund, 2024). Until now however, few have compared the efficacy of these interventions (but see Hameleers, 2022; Heuer & Glassman, 2022). This is important to ensure that policy and industry resources are assigned to the most promising efforts. Future research should therefore systematically compare the '(Dis)Trust buttons to other interventions by determining their efficacy in the same environment. This would also help identify which (or if any) combination of interventions is most effective (Bak-Coleman et al., 2022)

It is also important to acknowledge that participation in our study was limited to the US. Emerging research indicates that cultural norms significantly influence sharing behaviour on social media (Hsu et al., 2021). While US users are more reactive to negative content that others' have generated, users in Japan respond more to positive content. This disparity might also influence how they use the '(Dis)Trust' buttons and underscores the importance of considering cultural variances in future research. Thus, cultural differences could significantly affect the generalizability and efficacy of the proposed intervention.

Finally, it is important to recognize that this study focussed solely on external incentives. However, it has been shown that sharing is also internally rewarding (Baek et al., 2017; Tamir et al., 2015; Tamir & Mitchell, 2012). Indeed, it is unclear to what extent external incentives provided on social media platforms affect internal incentives. It has been suggested that when external rewards are introduced, activities become less intrinsically rewarding (Deci, 1971). As yet, it is unclear how this relates to sharing decisions. Future studies should strive to explore how external and internal incentives interact to determine sharing behaviour and whether internal incentives could also be utilized to reduce the spread of misinformation online. Related to this, it is plausible that platforms are more likely to implement interventions which improve the user experience and increase their enjoyment of the platform. In the experiments outlined in Chapter 4, I found that users utilized the new trust and distrust buttons more frequently than existing options, and there was no reduction in frequency of sharing. As such, I speculate, that the '(Dis)Trust' buttons did not negatively affect the user experience. Nevertheless, it would be interesting to ask users about their subjective experience of using the buttons.

Taken together, this study presents an innovative approach to reducing misinformation spread on social media. Here I have outlined several limitations which constrain its generalizability to the real world and have outlined future research directions which will help determine the efficacy of this intervention in real-world contexts and will shed light on how social dynamics and network effects impact the efficacy of this (and other) interventions at a population-level.

#### **5.5 Conclusion**

Every day humans are faced with copious opportunities to accumulate evidence and sharing information. This behaviour is driven by (1) external incentives, such as financial or social gains (Gold & Shadlen, 2002; Rosenthal-von der Pütten et al., 2019), and (2) internal incentives, such as positive emotions and self-efficacy (Baek et al., 2017; Gesiarz et al., 2019; Leong et al., 2019; Tamir et al., 2015; Tamir & Mitchell, 2012). This thesis explores how the incentive structure people face can be altered to improve the quality of evidence accumulation and sharing behaviour.

First, I test the assumption that evidence accumulation should be less biased when false beliefs are costly. I test this assumption in two ways: In Chapter 2, I examine the impact of exposing participants to a threatening environment in which external punishments for holding negative beliefs may be high; and in Chapter 3, I examine the impact of increasing monetary incentives to form accurate beliefs. I find that while perceived threat induces a global shift towards negative evidence such that evidence accumulation becomes less biased, in a manner that may be adaptive, monetary incentives to form accurate conclusions fail to mitigate biased evidence accumulation. This is because external and internal incentives alter orthogonal elements of the accumulation process. I speculate that stress elicits a global psychophysiological response which elicits an automatic shift in evidence accumulation towards negative evidence. By contrast accuracy incentives alter deliberate processing. While participants try to be more cautious when incentivized for accuracy, this does not change their biased perception, as they believe it to be genuine.

I speculate that feedback will allow participants to become aware of their biases and thus will enable them to update their beliefs in order to maximize external rewards. I test this assumption in Chapter 4 by providing participants with social feedback about the accuracy of the information they share online. I demonstrate that the existing incentive structure of social media platforms, in which social rewards (e.g., 'likes') and punishments (e.g., 'dislikes') are dissociated from the veracity of the information shared, contributes to the spread of misinformation online. I show that slightly altering this incentive structure, such that social rewards and punishments are contingent on the veracity of information, improves discernment in sharing behaviour and reduces the spread of misinformation. These findings highlight the importance of considering both automatic and deliberate processes when designing interventions to increase the accuracy of decision-making.

In conclusion, this thesis not only sheds light on the mechanisms by which incentives alter evidence accumulation and sharing behaviour, but also bridges the gap between theory and practice, offering actionable strategies that hold the potential to mitigate biased decision-making across digital and physical realms.

## **Chapter 6: Bibliography**

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# **Chapter 7: Appendix**

### **7.1 Chapter 2: Under threat weaker evidence is required**

### **to reach an undesirable conclusion**

**Supplementary Table 2.1. Mean difference in posterior distributions and 95% HDI Comparison (Experimental Data)**



**Supplementary Table 2.2. Proportion of correctly identified factories (Simulated Data).**



As in the experimental data, we observed a significant group by valence interaction (F(1,78) = 13.113,  $p < 0.001$ , partial  $n^2 = 0.144$ ) in the simulated data. In the simulated control group data, the proportion of correctly categorized desirable factories ( $M = 0.812$ ,  $SE = 0.016$ ) was larger than the proportion of correctly categorized undesirable factories ( $M = 0.735$ ,  $SE = 0.017$ ; t(42) = 3.552,  $p < 0.001$ , Cohen's  $d = 0.542$ ). This difference was not observed in the simulated threat group data (proportion of correctly identified undesirable factories:  $M = 0.819$ ,  $SE = 0.019$ ; proportion of correctly identified desirable factories:  $M = 0.783$ ,  $SE = 0.02$ ;  $t(36) = 1.611$ ,  $p = 0.116$ , Cohen's  $d = 0.265$ ).

# **7.2 Chapter 3: Futile Rewards: Why Accuracy Incentives Fail to Reduce Biased Evidence Accumulation**

#### **7.2.1 Supplementary Methods**

**Sensitivity.** To assess whether accuracy incentives influenced participants' discernment between signal and noise, we calculated each individual's dPrime (d') for each incentive level separately using the *psycho* R package (Makowski, 2018). d' is as a measure of sensitivity from signal detection theory (Pallier, 2002) to distinguish between correct and incorrect trials. It reflects the distance between the signal and signal + noise distributions and is calculated as follows:

$$
d' = Z(Hit Rate) - Z(False Alarm Rate)
$$

Here, a higher d' value indicates greater sensitivity in differentiating correct from incorrect trials. We then compared the average d' between the different incentive levels (Experiment 1 & Replication: \$0 vs \$5; Experiment 2: \$0 vs \$25) using paired t-tests. To substantiate findings that were not statistically significant, Bayes tests were computed.

#### **7.2.2 Supplementary Results**

#### **Accuracy incentives do not alter discernment between signal and noise.**

Thus far we have shown that accuracy incentives do not reduce the bias towards desirable responses, it then follows that they may also be ineffective as a tool to increase accuracy and therefore might not improve participants' ability to discriminate between signal and noise. To test this, we calculated each participant's d', representing their sensitivity in distinguishing correct from incorrect trials, for each incentive level. We entered the d' for each incentive level into a paired t-tests. This revealed that providing incentives for correct responses did not significantly enhance participants' discrimination accuracy (Experiment 1:  $t(68) = 1.055$ ,  $p = 0.295$ , Cohen's  $d = 0.127$ , Replication:  $t(71) =$ 0.387,  $p = 0.7$ , Cohen's d = 0.046, Experiment 2: t(91) = 0.834,  $p = 0.407$ ,

Cohen's d = 0.087; **see Figure S3.1**). Specifically, the d' scores in trials without a reward for correct responses (\$0 accuracy incentives - Experiment 1:  $M =$ 0.81, SE = 0.078, Replication: M = 0.779, SE = 0.09, Experiment 2: M = 0.731,  $SE = 0.072$ ) were not significantly different from those in trials where participants were rewarded for accuracy, regardless of the reward amount (\$25 accuracy incentives - Experiment 1:  $M = 0.763$ ,  $SE = 0.078$ , Replication:  $M =$ 0.762, SE =  $0.095$ , \$25 accuracy incentives - Experiment 2: M =  $0.767$ , SE = 0.072). These results hold when controlling for noise (**see Supplementary Tables 3.8 & 3.9 & 3.15**). Again, Bayes tests provide moderate to strong support in favour of the null hypothesis (Experiment 1:  $BF_{01} = 6.125$ , Replication:  $BF_{01} = 10.009$ , Experiment 2:  $BF_{01} = 8.623$ ).

Taken together these results illustrate that accuracy incentives do not improve participants' discernment between signal and noise.



**Figure S3.1. Accuracy Incentives do not improve discernment between signal and noise.** In **(a)** Experiment 1, **(b) its** replication, and **(c)** Experiment 2 participants' sensitivity (d') did not differ between trials in which they were incentivized for accuracy and those in which they were not incentivized. Y axis shows Sensitivity (d'). X axis shows accuracy incentive level. Data are plotted as boxplots for each incentive level, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5  $\times$ interquartile range. Diamond shape indicates the mean d' score per incentive level. Individuals' d' scores are shown separately as dots. Symbols above each boxplot indicate significance level compared to  $0.$  ns = not significant.

# **7.2.3 Supplementary Tables**











## **Supplementary Table 3.2. Model Fits (Replication)**






### **Supplementary Table 3.3. Model Fits for Experiment 2**









#### **Supplementary Table 3.4. Repeated Measures ANOVA log Response Times (Experiment 1).**



**Supplementary Table 3.5. Repeated Measures ANOVA log Response Times (Replication).**



#### **Supplementary Table 3.6. Repeated Measures ANOVA Response Bias (Experiment 1).**





### **Supplementary Table 3.7. Repeated Measures ANOVA Response Bias (Replication).**



### **Supplementary Table 3.8. Repeated Measures ANOVA dPrime (Experiment 1).**



**Supplementary Table 3.9. Repeated Measures ANOVA dPrime (Replication).**



### **Supplementary Table 3.10. 95% HDI Comparisons (Experiment 1 & Replication).**





### **Supplementary Table 3.11. Repeated Measures ANOVA Response Bias based on simulated data (Experiment 1).**



### **Supplementary Table 3.12. Repeated Measures ANOVA Response Bias based on simulated data (Replication).**



### **Supplementary Table 3.13. Repeated Measures ANOVA log Response Times (Experiment 2).**



### **Supplementary Table 3.14. Repeated Measures ANOVA Response Bias (Experiment 2).**



dPrime	df	<b>F-value</b>	p-value	partial n2
<b>Accuracy Incentives</b>	(1.91)	1.398	0.240	0.015
<b>Noise</b>	(7.04, 640.272)	20.261	< 0.001	0.182
<b>Accuracy</b> <b>Noise</b> X	(9,819)	0.864	0.557	0.009
<b>Incentives</b>				

**Supplementary Table 3.15. Repeated Measures ANOVA dPrime (Experiment 2).**

**Supplementary Table 3.16. 95% HDI Comparison in Experiment 3.**



**Supplementary Table 3.17. Repeated Measures ANOVA Response Bias based on simulated data (Experiment 2).**



# **7.3 Chapter 4: Changing the Incentive Structure of Social Media Platforms to halt the Spread of Misinformation**

### **7.3.1 Replication Studies**

### **Methods (Experiment 4-6)**

**Participants (Experiment 4).** Fifty participants residing in the US completed the task on *Prolific Academic* (25 Democrats, 8 Republican, 17 Other, Mage = 33.16,  $SD_{age} \pm 9.804$ ; female = 24, male = 25, other = 1, Non-White = 15, White = 35). No participants failed the attention checks. Participants received £7.50 per hour for their participation in addition to a memory test performance-related bonus. For all experiments in this study, ethical approval has been provided by the Research Ethics Committee at University College London and all participants gave informed consent.

### **Participants (Experiment 5).**

Two-hundred and sixty-one participants completed the task on *Prolific Academic* (132 Democrats, 90 Republican, 39 Other, Mage = 34.824, SDage 12.632; female = 122, male = 131, others = 8, Non-White = 84, White = 177). Participants received £7.50 per hour for their participation in addition to a memory test performance-related bonus.

### **Participants (Experiment 6).**

One-hundred and fifty participants completed the task on *Prolific Academic* (74 Democrats, 14 Republican, 62 Other,  $M_{\text{age}} = 34.2$ ,  $SD_{\text{age}} \pm 12.489$ ; female = 70, male = 77, other = 3, Non-White = 39, White = 150). Participants received £7.50 per hour for their participation in addition to a memory test performance-related bonus.

### **Tasks and Statistical Analysis.**

The tasks and analysis in Experiment 4-6 were identical to those used in Experiment 1-3 except for the following differences:

- 1) In Experiment 4 a 'repost' button was included in addition to 'skip', '(Dis)Like' and '(Dis)Trust' options.
- 2) In Experiment 5 feedback symbols were coloured 'distrusts' and 'dislikes' in red and 'trusts' and 'likes' in green, instead of black and white.
- 3) Experiment 6 contained all 100 posts instead of a selection of 40 posts and did not contain final questions to assess whether participants believed the feedback was real.
- 4) The samples were not politically balanced (**see Participants Experiment 4-6**), as such analysis did not take into account political orientation.

### **Results (Experiments 4-6)**

### **Participants use '(Dis)Trust' buttons to discern true from false information (Experiment 4).**

Experiment 4 is a replication of Experiment 1, in which participants observe posts (half true half false) and could respond by clicking all, none or some of the following buttons: 'like', 'dislike', 'trust', 'distrust' (**see Methods for details**). Participant's use of '(Dis)Trust' ( $M = 0.111$ ; SE = 0.01) was more discerning than their use of '(Dis)Like' (M = 0.03; SE = 0.006; F(1,49) = 51.996, p < 0.001, partial  $n^2 = 0.51$ , Figure S4.1). They also used negative reactions (M =  $0.082$ ;  $SE = 0.011$ ) in a more discerning manner than positive reactions (M = 0.06, SE  $= 0.006$ ; F(1,49) = 7.147, p = 0.01, partial  $n^2 = 0.13$ ). Participants' used all reaction buttons, except 'dislike', to discern between true and false posts ('like':  $M = 0.053$ ; SE = 0.008; t(49) = 6.982, p < 0.001, Cohen's d = 0.987; 'trust': M  $= 0.066$ ; SE = 0.01; t(49) = 6.641, p < 0.001, Cohen's d = 0.939; 'dislike': M = 0.007; SE = 0.008;  $t(49)$  = 0.883, p = 0.381, Cohen's d = 0.125; 'distrust': M = 0.157; SE = 0.014;  $t(49)$  = 11.312, p < 0.001, Cohen's d = 1.6). These results hold when including an interaction of type of reaction and valence (**see Supplementary Table 4.18**).

Experiment 4 thus replicates the results of Experiment 1 to show that '(Dis)Trust' buttons are used to discern true from false information.



**Figure S4.1. Participants' use '(Dis)Trust' to discern true from false information. '**Distrust' and 'trust' reactions were more discerning than 'like' and 'dislike' reactions. Y axis shows discernment between true and false posts. For positive reactions (e.g., *'likes'* and *'trusts'*) discernment is equal to the proportion of positive reactions for true information minus false information, and vice versa for negative reactions (*'dislikes'* and *'distrusts'*). X axis shows reaction buttons. Data are plotted as boxplots for each reaction, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5 × interquartile range. Diamond shape indicates the mean discernment per reaction. Individuals' mean discernment data are shown separately as grey dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p < 0.001.

**'(Dis)Trust' incentives improve discernment in sharing behaviour (Experiment 5 and Experiment 6).** As in Experiment 2, we found an effect of type of feedback  $(F(1,257) = 8.112, p = 0.005,$  partial  $p2 = 0.031$ ). In particular, participants reposted more true relative to false information in the '(Dis)Trust' conditions (M =  $0.236$ , SE =  $0.019$ ) than the '(Dis)Like' conditions (M =  $0.111$ , SE = 0.022; F(1,212) = 7.682, p = 0.006, partial η2 = 0.035, **Figure S4.2a**) and Baseline condition (M =  $0.102$ , SE =  $0.026$ ; F(1,163) = 16.246, p <  $0.001$ , partial η2 = 0.087). '(Dis)Like' feedback did not improve discernment in sharing behaviour compared to Baseline (F(1,142) = 2.184,  $p = 0.142$ , partial  $p =$ 

0.015). No other effects were significant. These results hold when including an interaction of type of feedback and valence (**see Supplementary Table 4.19**).

In line with our results from Experiment 2, we observed an effect of feedback type  $(F(1.258) = 4.179, p = 0.042,$  partial  $n2 = 0.016$ ), such that participants were more accurate (less errors) when they received '(Dis)Trust' feedback (M  $= 35.717$ , SE = 0.65) compared to '(Dis)Like' feedback (M = 37.63, SE = 0.767;  $F(1,212) = 3.955$ ,  $p = 0.048$ , partial  $p = 0.018$ ) and also more accurate than those who received no feedback (Baseline,  $M = 39.73$ ,  $SE = 0.886$ ;  $F(1,162) =$ 11.759,  $p < 0.001$ , partial  $n = 0.068$ ). There was no difference in accuracy between participants in the '(Dis)Like' environment and those in the Baseline environment (F(1,143) = 2.746,  $p = 0.1$ , partial  $p = 0.019$ ). No other effects were significant. These results hold when allowing for an interaction between type of feedback and valence (**see Supplementary Table 4.20**).



**Figure S4.2. '(Dis)Trust' Feedback improves discernment in sharing behaviour.** In both **(a)** Experiment 2 and **(b)** Experiment 3 participants who received '(Dis)Trust feedback shared more true relative to false information than participants in the '(Dis)Like' and Baseline conditions. '(Dis)Like' reactions are indicated in blue, '(Dis)Trust in orange and Baseline in black. Y axis shows proportion of true posts shared minus proportion of false posts shared. X axis shows feedback type. Data are plotted as boxplots for each reaction, in which horizontal lines indicate median values, boxes indicate 25/75% interquartile range and whiskers indicate 1.5 × interquartile range. Diamond shape indicates the mean discernment per reaction. Individuals' mean discernment data are shown separately as grey dots. Symbols above each boxplot indicate significance level compared to 0. \*\*\*p <  $0.001$ , \*\*p <  $0.01$ .

In Experiment 6, we again observed an effect of type of feedback  $(F(1, 147) =$ 11.150,  $p < 0.001$ , partial  $p = 0.132$ , **Figure S2b**), with participants in the 'Trust and Distrust' feedback group posting more true relative to false information (M  $= 0.264$ , SE = 0.023) than those in the 'Like and Dislike' group (M = 0.147, SE  $= 0.027$ ; F(1,101) = 11.122, p = 0.001, partial  $n^2 = 0.099$ ) or those who received no feedback at all (M =  $0.106$ , SE =  $0.026$ ; F(1,101) =  $21.141$ , p <  $0.001$ , partial  $n2 = 0.173$ ). By contrast there was no difference between the latter two groups  $(F(1, 92) = 1.188, p = 0.279,$  partial  $p = 0.013$ ).

There was also an effect of type of feedback on accuracy  $(F(1, 147) = 4.596, p$  $= 0.012$ , partial  $n^2 = 0.059$ . Participants were more accurate when they received '(Dis)Trust' feedback (M = 35.464, SE =  $0.596$ ) compared to no feedback (Baseline, M = 39.7, SE = 0.89; F(1,101) = 10.694, p < 0.001, partial η2 = 0.096) with no difference between participants in the '(Dis)Like' environment and those in the Baseline environment  $(F(1,92) = 1.949, p = 0.11, p = 0.11)$ partial  $n2 = 0.021$ ) or (Dis)trust environment (M = 37.656, SE = 1.193; F(1.101)  $= 2.149$ ,  $p = 0.146$ , partial  $n2 = 0.021$ )

The findings replicate those of Experiment 2 and 3 in suggesting that changing the incentive structure of social media platforms, such that 'carrots' and 'sticks' are partially contingent on accuracy, promotes discernment in sharing behaviour.

#### **'(Dis)Trust' feedback increases the drift rate.**

DDM was conducted as for Experiment 2 and 3. For both Experiment 5 (**see Table S4.1**) and Experiment 6 (**see Table S4.2**) we observed a significant difference in the drift rate. That is to say, '(Dis)Trust' feedback (Experiment 5: v  $= 0.287$ ; 95% CI [0.242, 0.335]; Experiment 6: v = 0.321; 95% CI [0.265, 0.379]) increased the relative importance of the veracity of information when participants decided whether or not to share a post, compared to '(Dis)Like' feedback (Experiment 5: v = 0.198; 95% CI [0.147, 0.248], Experiment 6: 0.172; 95% CI [0.112, 0.233]) or no feedback at all (Experiment 5: v = 0.11; 95% CI [0.048, 0.171]; Experiment 6: v = 0.139; 95% CI [0.065, 0.215]). In Experiment

5 the drift rate in the '(Dis)Like' condition was also higher than in the Baseline condition. While in Experiment 6 there was no difference between the Baseline condition and the '(Dis)Like' condition. As shown in **Table S4.1** and **Table S4.2** the '(Dis)Like' and '(Dis)Trust' conditions did not differ on any other parameters, though some other parameters were different between the Baseline condition and the other conditions, but those were not consistent over all replications. 95% HDI comparisons corroborate this result **(see Supplementary Table 4.21 & 4.22 for HDI Comparisons).** In sum, Experiment 5 and 6 replicate the main results of interest in Experiment 2 and 3.

<b>Estimate</b>	<b>Baseline</b> [95% CI]	'(Dis)Like' [95% CI]	'(Dis)Trust' [95% CI]
<b>Distance</b> between <b>Decision</b> <b>Thresholds</b> (α)	2.2 [2.11, 2.298]	2.451 [2.326, 2.58]	2.458 [2.35, 2.573]
Non- <b>Decision</b> Time $(t0)$	6.973 [6.799, 7.139]	6.685 [6.429, 6.949]	6.757 [6.541, 6.973]
<b>Starting</b> Point (z)	0.496 [0.484, 0.504]	0.477 [0.467, 0.487]	0.472 [0.463, 0.482]
<b>Drift</b> Rate (v)	$0.11$ [0.048, 0.171]	$0.198$ [0.147, 0.248]	0.287 [0.242, 0.335]

**Table S4.1. Group estimates for DDM in Experiment 5.**







#### **Effects on True and False Posts**

Within the field, the gold standard is to measure discernment (that is endorsement of true relative to false posts) rather than measure endorsement of true and/or false posts separately (for a detailed explanation see Guay et al., 2023; Pennycook & Rand, 2021). In our manuscript we follow these recommendations and indeed find that the effects on discernment are very consistent across all experiments. Nevertheless, here we also report the effects of '(Dis)Trust' relative to the effects of '(Dis)Like' on true and false posts separately. We find that sometimes the effect is observed as a decrease in endorsing false posts, sometimes as an increase in endorsing true posts and sometimes as both, but always (as reported in the main text) as an increase in discernment. In particular, participants selected the 'trust' reaction button more for true (M = 37.906%, SE = 2.477) than false posts (M = 18.208%, SE = 1.568;  $t(105) = 9.744$ ,  $p < 0.001$ , Cohen's  $d = 0.946$ ), and the 'distrust' reaction button more for false (M = 49.66%,  $SE = 2.179$ ) than for true posts (M = 18.509%, SE  $= 1.476$ ; t(105) = 15.872, p < 0.001, Cohen's d = 1.542). Moreover, participants selected 'trust' (M = 37.906%, SE = 2.477) more than 'like' (M = 24.604%, SE  $= 1.601$ ) for true posts (t(105) = 4.843, p < 0.001, Cohen's d = 0.47), and selected 'distrust' ( $M = 49.66\%$ ,  $SE = 2.179$ ) more than 'dislike' ( $M = 27.132\%$ , SE = 1.897) for false posts (t(105) = 8.53,  $p < 0.001$ , Cohen's d = 0.829). This latter effect is also observed in Experiment 4, where 'distrust' (M = 45.32%, SE  $= 3.317$ ) is selected more often than 'dislike' (M = 13.88%, SE = 1.879) for false posts (t(49) = 8.637,  $p < 0.001$ , Cohen's d = 1.221). With regards to sharing, in Experiment 2 participants shared fewer false posts in the 'trust' condition ( $M =$ 21.217%,  $SE = 2.423$ ) than in the 'like' condition (M = 29.214%,  $SE = 2.233$ ;  $t(133) = 0.244$ ,  $p = 0.027$ , Cohen's  $d = 0.407$ ), and tended to do the same in the 'distrust' condition (M = 23.755%, SE = 2.7) than the 'dislike' condition (M =

31.289%, SE = 3.343;  $t(92) = 1.766$ ,  $p = 0.081$ , Cohen's d = 0.365). In Experiment 3, participants shared more true posts in the 'Trust & Distrust' condition (M = 41.241%, SE = 2.162) than the 'Like & Dislike' condition (M = 32.5%, SE = 2.161; t(263) = 2.857, p = 0.005, Cohen's = 0.351). This was repeated in Experiment 6, with true posts shared more in 'Trust & Distrust' condition ( $M = 46.607\%$ ,  $SE = 3.197$ ) compared to those in the 'Like & Dislike' condition (M = 33.702%, SE = 3.618; t(101) = 2.266, p = 0.026, Cohen's d = 0.448).

### **7.3.2 Supplementary Tables**

**Supplementary Tables for Experiment 1-3.**

**Supplementary Table 4.1. Individual Ratings per Stimulus in Experiment 1.**























The interaction of type and valence of reaction is characterized by participants using the 'distrust' reaction button ( $M = 0.157$ ,  $SE = 0.008$ ) in a more discerning manner than the 'trust' reaction button ( $M = 0.099$ ,  $SE = 0.008$ ; t(106) = 9.338,  $p < 0.001$ , Cohen's d = 0.903), but the 'like' reaction button (M = 0.06, SE = 0.008) in a more discerning manner than the 'dislike' button ( $M = 0.034$ , SE = 0.008;  $t(106) = 3.474$ ,  $p < 0.001$ , Cohen's d = 0.336).







**Supplementary Table 4.4. % true and false posts shared out of all true or false posts in that feedback condition (Experiment 2).**



### **Supplementary Table 4.5. Discernment of sharing behaviour (Experiment 2).**



% posts shared	df	<b>F-value</b>	p-value
including demographics			
<b>Intercept</b>	(1, 278)	78.161	< 0.001
<b>Type of Feedback</b>	(1,278)	1.031	0.311
<b>Valence of Feedback</b>	(1,278)	1.219	0.270
Gender	(1,278)	4.479	0.035
<b>Political Orientation</b>	(1,278)	1.518	0.219
<b>Ethnicity</b>	(1, 278)	2.341	0.127
Age	(1, 278)	0.032	0.858
Type of Feedback x Political	(1,278)	0.117	0.890
<b>Orientation</b>			
including valence x reaction			
<b>Intercept</b>	(1, 311)	701.419	< 0.001
<b>Type of Feedback</b>	(1, 311)	1.533	0.217
<b>Valence of Feedback</b>	(1, 311)	0.741	0.39
Type of Feedback x Valence of <b>Feedback</b>	(1, 311)	0	0.986

**Supplementary Table 4.6. % posts shared out of all posts (Experiment 2).**

### **Supplementary Table 4.7. Belief Accuracy (Experiment 2).**



**Supplementary Table 4.8. Discernment of sharing behaviour (Experiment 3).**





### **Supplementary Table 4.9. % posts shared out of all posts (Experiment 3).**



### **Supplementary Table 4.10. Belief Accuracy (Experiment 3).**





**Supplementary Table 4.11. Mean difference in posterior distributions and 95% HDI Comparison in Experiment 2.**

**Supplementary Table 4.12. Mean difference in posterior distributions and 95% HDI Comparison in Experiment 3.**



**Supplementary Table 4.13. Recovered Group estimates for DDM in Experiment 2 based on simulated data.**





### **Supplementary Table 4.14. Recovered Group estimates for DDM in Experiment 3 based on simulated data.**



### **Supplementary Table 4.15. Pairwise Comparisons for Discernment Experiment 2.**



### **Supplementary Table 4.16. Pairwise Comparisons for Discernment Experiment 3.**





**Supplementary Table 4.17. Correlations between participants' real and recovered DDM estimates in Experiment 2 and Experiment 3.**



We estimated both group-level and individual-level parameters. We then used the individual-level parameter estimates to simulate data for each participant respectively in the dataset. We used the same number of trials as in the experiments. Simulated data from each participant were then combined and used to perform model recovery analysis. We sampled 2000 times from the posteriors, discarding the first 500 as burn in. We then correlated the real and the recovered individual-level parameters.

**Supplementary Tables for Experiment 4-6.**

**Supplementary Table 4.18. Discernment of reactions (Experiment 4, including type x valence of reaction interaction).**



The interaction is characterized by participants using the 'distrust' reaction button ( $M = 0.157$ ,  $SE = 0.014$ ) in a more discerning manner than the 'trust' reaction button (M =  $0.066$ , SE =  $0.01$ ; t(49) =  $7.192$ ,  $p < 0.001$ , Cohen's d = 1.017), but the 'like' reaction button (M = 0.053,  $SE = 0.008$ ) in a more discerning manner than the 'dislike' button (M =  $0.007$ , SE =  $0.008$ ; t(49) = 4.407,  $p < 0.001$ , Cohen's d = 0.623).

### **Supplementary Table 4.19. Discernment of sharing behaviour (Experiment 5).**



### **Supplementary Table 4.20. Belief Accuracy (Experiment 5).**



<b>Estimate</b>	'(Dis)Trust' minus Baseline		'(Dis)Trust' minus '(Dis)Like'		'(Dis)Like' minus <b>Baseline</b>	
between <b>Distance</b> <b>Decision</b> Thresholds $(\alpha)$	0.258 0.415	[0.101]	0.007 0.186	$[-0.17,$	0.251 0.418	[0.085,
<b>Non-Decision Time</b> (10)	$-0.216$ $0.07$ ]	$[-0.502,$	0.072 0.432	$[-0.293,$	$-0.288$ 0.046	622, I-0.
<b>Starting Point (z)</b>	$-0.024$ $0.007$ ]	[-0.04, -	$-0.005$ 0.009	$[-0.019,$	$-0.019$ 0.002	[-0.036,
Drift Rate (v)	0.177 0.259	[0.095,	0.089 0.163	[0.018,	0.088 0.173	[0.005,

**Supplementary Table 4.21. Mean difference in posterior distributions and 95% HDI Comparison in Experiment 5.**

**Supplementary Table 4.21. Mean difference in posterior distributions and 95% HDI Comparison in Experiment 6.**



## **7.4 Bibliography**

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