

Original article

Educational attainment, health outcomes and mortality: a within-sibship Mendelian randomization study

Laurence J Howe ^{1,2,*} Humaira Rasheed,^{1,3,4} Paul R Jones,⁵ Dorret I Boomsma ^{6,7} David M Evans ^{1,8,9} Alexandros Giannelis,¹⁰ Caroline Hayward ¹¹ John L Hopper,¹² Amanda Hughes,^{1,2} Hannu Lahtinen ¹³ Shuai Li ^{12,14,15} Penelope A Lind,^{16,17,18} Nicholas G Martin,¹⁹ Pekka Martikainen ^{13,20,21} Sarah E Medland,^{16,18,22} Tim T Morris ^{1,2} Michel G Nivard,⁶ Jean-Baptiste Pingault,^{23,24} Karri Silventoinen,¹³ Jennifer A Smith ^{25,26} Emily A Willoughby,¹⁰ James F Wilson,^{11,27} Within Family Consortium, Bjørn Olav Åsvold ^{3,28} Øyvind E Næss,^{5,29} George Davey Smith ^{1,2} Jaakko Kaprio,³⁰ Ben Brumpton ^{1,3,†} and Neil M Davies ^{1,3,31,32,†}

¹Medical Research Council Integrative Epidemiology Unit, Population Health Sciences, University of Bristol, UK, ²Population Health Sciences, Bristol Medical School, University of Bristol, Bristol, UK, ³K.G. Jebsen Center for Genetic Epidemiology, Department of Public Health and Nursing, NTNU, Norwegian University of Science and Technology, Trondheim, Norway, ⁴Department of Medicine and Laboratory Sciences, University of Oslo, Oslo, Norway, ⁵Department of Community Medicine and Global Health, Institute of Health and Society, Faculty of Medicine, University of Oslo, Oslo, Norway, ⁶Department of Biological Psychology, Netherlands Twin Registry, Vrije Universiteit, Amsterdam, Netherlands, ⁷Amsterdam Public Health (APH) and Amsterdam Reproduction and Development (AR&D), ⁸University of Queensland Diamantina Institute, University of Queensland, Brisbane, Australia, ⁹Institute for Molecular Bioscience, University of Queensland, Brisbane, Australia, ¹⁰Department of Psychology, University of Minnesota, Minneapolis, MN, USA, ¹¹MRC Human Genetics Unit, Institute of Genetics and Cancer, University of Edinburgh, Western General Hospital, Edinburgh, UK, ¹²Centre for Epidemiology and Biostatistics, Melbourne School of Population and Global Health, The University of Melbourne, Parkville, Victoria, Australia, ¹³Population Research Unit, University of Helsinki, Helsinki, Finland, ¹⁴Centre for Cancer Genetic Epidemiology, Department of Public Health and Primary Care, University of Cambridge, Cambridge, UK, ¹⁵Precision Medicine, School of Clinical Sciences at Monash Health, Monash University, Clayton, Victoria, Australia, ¹⁶Psychiatric Genetics, QIMR Berghofer Medical Research Institute, Brisbane, Australia, ¹⁷School of Biomedical Sciences, Queensland University of Technology, Brisbane, Australia, ¹⁸Faculty of Medicine, University of Queensland, Brisbane, Australia, ¹⁹Department of Genetics and Computational Biology, QIMR Berghofer Medical Research Institute, Brisbane, Queensland, Australia, ²⁰The Max Planck Institute for Demographic Research, Germany, ²¹Department of Public Health Sciences, Stockholm University, Sweden, ²²School of Psychology, University of Queensland, Brisbane, Australia, ²³Social Genetic & Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, London,

UK, ²⁴Department of Clinical, Educational and Health Psychology, University College London, London, UK, ²⁵Department of Epidemiology, School of Public Health, University of Michigan, Ann Arbor, MI, USA, ²⁶Survey Research Center, Institute for Social Research, University of Michigan, Ann Arbor, MI, USA, ²⁷Centre for Global Health Research, Usher Institute, University of Edinburgh, Teviot Place, Edinburgh, UK, ²⁸Department of Endocrinology, Clinic of Medicine, St Olavs Hospital, Trondheim University Hospital, Trondheim, Norway, ²⁹Norwegian Institute of Public Health, Oslo, Norway, ³⁰Institute for Molecular Medicine FIMM, University of Helsinki, Helsinki, Finland, ³¹Division of Psychiatry, University College London, London, UK and ³²Department of Statistical Sciences, University College London, London, UK

*Corresponding author. Population Health Sciences, Bristol Medical School, University of Bristol, Oakfield House, Oakfield Grove, Bristol, BS8 2BN, UK. E-mail: laurence.howe@bristol.ac.uk

[†]Jointly supervised the work.

Received 28 January 2022; Editorial decision 15 March 2023; Accepted 12 May 2023

Abstract

Background: Previous Mendelian randomization (MR) studies using population samples (population MR) have provided evidence for beneficial effects of educational attainment on health outcomes in adulthood. However, estimates from these studies may have been susceptible to bias from population stratification, assortative mating and indirect genetic effects due to unadjusted parental genotypes. MR using genetic association estimates derived from within-sibship models (within-sibship MR) can avoid these potential biases because genetic differences between siblings are due to random segregation at meiosis.

Methods: Applying both population and within-sibship MR, we estimated the effects of genetic liability to educational attainment on body mass index (BMI), cigarette smoking, systolic blood pressure (SBP) and all-cause mortality. MR analyses used individual-level data on 72 932 siblings from UK Biobank and the Norwegian HUNT study, and summary-level data from a within-sibship Genome-wide Association Study including >140 000 individuals.

Results: Both population and within-sibship MR estimates provided evidence that educational attainment decreased BMI, cigarette smoking and SBP. Genetic variant–outcome associations attenuated in the within-sibship model, but genetic variant–educational attainment associations also attenuated to a similar extent. Thus, within-sibship and population MR estimates were largely consistent. The within-sibship MR estimate of education on mortality was imprecise but consistent with a putative effect.

Conclusions: These results provide evidence of beneficial individual-level effects of education (or liability to education) on adulthood health, independently of potential demographic and family-level confounders.

Key words: Within-sibship, Mendelian randomization, educational attainment, mortality

Key Messages

- Within-sibship Mendelian randomization can reduce bias from demographic and familial factors that may particularly impact analyses of social and behavioural phenotypes.
- Within-sibship Mendelian randomization indicated that higher educational attainment decreased body mass index, smoking behaviour and blood pressure.
- These findings are consistent with beneficial individual-level health effects of higher (liability to) educational attainment.

Introduction

Higher educational attainment is strongly associated with better adulthood health and reduced mortality.^{1,2} However, whether these associations are causal remains unclear because of inconclusive evidence from previous quasi-experimental designs such as the raising of the school leaving age and co-twin control or discordant-twin studies.^{3–11} Another source of evidence on effects of educational attainment on health are Mendelian randomization (MR) studies,¹² which have used genetic variants associated with educational attainment as instrumental variables to provide consistent evidence for beneficial effects of educational attainment on adulthood health outcomes.^{13–17} A caveat is that educational attainment as measured by years in full-time education is a categorical exposure and so it may be more appropriate to interpret effects in terms of liability to educational attainment.¹⁶

A key assumption of MR analyses is that the genetic variant–exposure (here: educational attainment) and genetic variant–outcome (here: health outcomes) associations represent downstream effects of inheriting the genetic variant (or a correlated variant).^{12,18–20} However, there is growing evidence that genotype–phenotype associations derived from samples of unrelated individuals can reflect other sources of variation^{19–21} (Figure 1). Previous studies have illustrated that Genome-wide Association Study (GWAS) estimates for educational attainment from unrelated individuals reflect population stratification,²² assortative mating^{18,23–25} and indirect genetic effects.^{18,26–29} Educational attainment is particularly distinctive amongst complex traits because of the large magnitude of indirect genetic effects, the high degree of assortative mating, strong correlations with geographical features and widespread genetic correlations with many phenotypes including health outcomes.³⁰ It follows that MR analyses of educational attainment may be biased if using genetic association estimates from unrelated individuals.

Genetic association estimates from within-sibship models are largely robust against population stratification, assortative mating and indirect genetic effects because genetic differences between full siblings are due to random segregation at meiosis.^{12,18–20,31} Within-sibship MR using individual-level sibling or within-sibship GWAS data¹⁸ can control for these sources of genetic association and so can be used to derive less-biased MR estimates.

Here, using individual-level data from UK Biobank and the Trøndelag Health Study, Norway (HUNT) and summary data from a recent within-sibship GWAS,¹⁸ we generate population and within-sibship MR estimates of the effects of liability to educational attainment on health outcomes and mortality. For health outcomes, we included

body mass index (BMI), pack years of cigarette smoking and systolic blood pressure (SBP), which were investigated in previous MR studies and are measured in the majority of UK Biobank and HUNT study participants. We also performed phenotypic analyses using self-reported educational attainment including a twin-based analysis using Finnish twin cohort data.

Methods

UK Biobank

UK Biobank is a large-scale prospective cohort study that has been described in detail previously.^{32,33} In total, 503 325 individuals aged between 38 and 73 years were recruited between 2006 and 2010 from across the UK and attended an assessment centre where they were interviewed, completed a touch-screen questionnaire, and provided various measurements (e.g. height) and biological samples (e.g. blood).

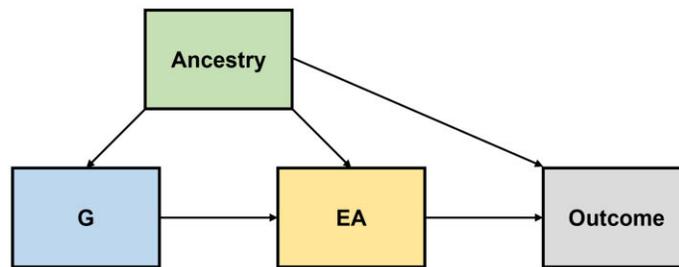
The UK Biobank study sample incidentally includes many related individuals. In our analyses we included individuals with one or more full siblings in the study sample. Siblings were identified in a previous study using the UK Biobank-derived estimates of pairwise identical by state kinships and the proportion of unshared loci (IBS0).^{19,34} After restricting the sample to sibships with two or more individuals with educational attainment data, our analysis sample included 40 734 individuals from 19 773 sibships. Further detail on the derivation of UK Biobank siblings is contained in the [Supplementary methods](#) (available as [Supplementary data](#) at *IJE* online).

Educational attainment was defined as in a previous study³⁵ using the self-reported qualifications from questionnaire data (field ID: 6138–0.0) to estimate the number of years each individual spent in full-time education. For example, ‘College or University degree’ was mapped to 17 years whereas ‘A levels/AS levels or equivalent’ was mapped to 14 years. Where individuals reported multiple qualifications, the highest qualification in terms of years in education was used. Information on health outcome phenotypes (BMI, smoking, SBP, mortality) and genotyping for UK Biobank study participants is contained in the [Supplementary methods](#) (available as [Supplementary data](#) at *IJE* online).

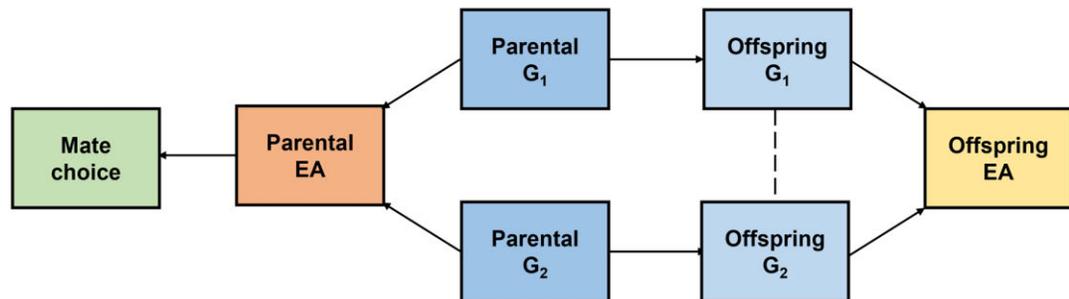
HUNT

HUNT is a series of general health surveys of the adult population of the Trøndelag region, Norway, as detailed in previous publications.^{36–38} Every 10 years, the adult population of this region (~90 000 adults at the start of

Population stratification



Assortative mating



Indirect genetic effects (parental)

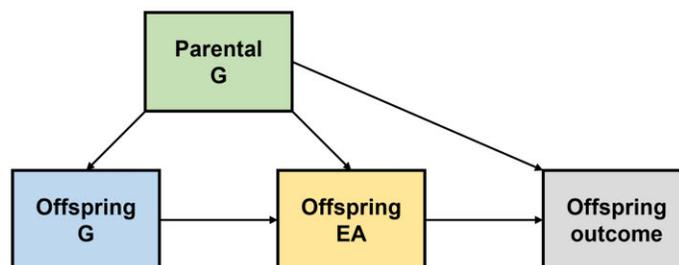


Figure 1 Population stratification, assortative mating and indirect genetic effects. ‘Population stratification’ occurs when ancestry is associated with the allele frequency of the genetic variant (G) and the phenotype of interest, distorting the association between G and the phenotype. In the context of MR, population stratification could distort the association between G and educational attainment (EA) and/or the association between G and the outcome, both of which could lead to bias in MR. ‘Assortative mating’ occurs when a heritable phenotype influences mate choice, e.g. if individuals are more likely to select a partner with a similar EA. Assortative mating leads to correlations for parental genotypes related to assorted phenotypes, which in turn leads to correlations between otherwise independent genotypes in the offspring. For example, if two genetic variants G₁ and G₂ influence EA then assortative mating on EA will lead to correlations in offspring for the EA-increasing alleles of G₁ and G₂ even if the two alleles are unlinked (linkage disequilibrium = 0). ‘Indirect genetic effects’ occur when the genotypes of relatives (e.g. parents, siblings) influence the phenotypes of the index individual. For example, parents with a higher EA polygenic score may produce an environment for their offspring that is more conducive to learning than parents with a lower EA polygenic score. This has been previously illustrated by evidence that non-transmitted parental EA polygenic scores also associate with offspring phenotypes

HUNT2 in 1995) is invited to attend a health survey (including comprehensive questionnaires, an interview, clinical examination and detailed phenotypic measurements). To date, four health surveys have been conducted, namely HUNT1 (1984–86), HUNT2 (1995–97), HUNT3 (2006–08) and HUNT4 (2017–19), and all surveys have had a >50% participation rate.³⁹ In this study, we used data

from 32 198 individuals from 12 578 sibships who reported their educational attainment in the HUNT2 survey. Siblings were identified using KING software,⁴⁰ with sibling-pairs identified based on the following criteria: kinship coefficient of between 0.177 and 0.355, the proportion of the genomes that share two alleles identical by descent (IBD) > 0.08 and the proportion of the genome

that share zero alleles IBD > 0.04 . Sibships of two or more siblings were constructed based on the identified sibling-pairs.

Educational attainment was measured using the following question: ‘What is your highest level of education?’ Participants answered one of five categories: (i) primary school, (ii) high school for 1 or 2 years, (iii) complete high school, (iv) college or university for <4 years and (v) college or university for ≥ 4 years. Participants with university degrees were assigned to 16 years of education, those who completed high school were assigned to 13 years, those who attended high school for 1 or 2 years were assigned to 12 years and those who only attended primary school were assigned to 10 years. Information on health outcome phenotypes (BMI, smoking, SBP, mortality) and genotyping for HUNT study participants is contained in the [Supplementary methods](#) (available as [Supplementary data](#) at *IJE* online).

Finnish twin cohort

Overview

The older part of the Finnish twin cohort was established in 1974 by identifying pairs of persons born on the same day, in the same local community, of the same sex and with the same surname at birth from the population registers of Finland. The selection was restricted to twin pairs born before 1958 and the baseline analysis cohort consists of 16 282 pairs (32 564 twins). A baseline questionnaire was mailed in the autumn of 1975 with some data collection in early 1976. It contained questions relating to the assignment of zygosity as well as questions on various phenotypes including smoking behaviour, weight and height.⁴¹ All twins in the cohort were asked to participate in a second survey in 1981.

Data on educational attainment were collected in both the 1975 and 1981 questionnaires using the following questions: ‘What kind of education have you had, and what courses have you taken?’ The 1975 information was updated by the 1981 response if additional educational attainment was reported. Eight response categories ranging from less than primary school (4 years) to university education (17 years) provided by study participants were converted into years of education. The ninth response alternative was ‘Other’ and coded as missing ($n = 587$, 2.1% of participants). Years of education were then standardized to a mean of 0 and standard deviation of 1. Further information on Finnish twin cohort phenotypes is contained in the [Supplementary methods](#) (available as [Supplementary data](#) at *IJE* online).

A validated algorithm classified respondent pairs as monozygotic (MZ), dizygotic (DZ) or of unknown zygosity (XZ) (excluded from all analyses).⁴² Data on

educational attainment and mortality were available for 27 229 individual twins living in Finland, which included 2779 individual twins (co-twin did not reply), 989 pairs of uncertain zygosity, 3518 MZ pairs and 7718 DZ pairs.

Statistical analysis

Population and within-sibship models

The population model is a standard regression model in which the outcome is regressed (e.g. linear) on the exposure (educational attainment or educational attainment polygenic score: PGS). The within-sibship model is an extension to the population model including the mean sibship exposure value in the model, e.g. mean educational attainment value of each sibship. Each sibling’s exposure value is centred on the mean sibship exposure value. To account for relatedness between siblings, standard errors are clustered by sibship in both models using a sandwich estimator. More information on these models is contained in previous publications.^{18,19}

Using individual-level data on 72 932 individuals from 32 351 sibships of European ancestry from UK Biobank ($n = 40 734$) and HUNT ($n = 32 198$) we estimated the association between measured educational attainment and outcomes (BMI, pack years of smoking, SBP, mortality) using population and within-sibship models. In population models, the outcome was regressed on educational attainment including relevant covariates. In within-sibship models, the mean educational attainment of each sibship was included as a covariate to account for variation in educational attainment between sibships. The continuous outcomes were chosen because they have previously been shown to be associated with educational attainment and because they were measured in the majority of UK Biobank and HUNT study participants. Linear regression models were used for BMI, pack years and SBP. Cox-proportional hazards models were used for mortality using date of birth as the baseline in UK Biobank and HUNT. Educational attainment, BMI, pack years and SBP were standardized after residualizing on birth year and sex.

We performed population and within-sibship PGS analyses using the UK Biobank and HUNT sibship data. The educational attainment PGS was constructed using weightings and directions of effect of independent variants identified at genome-wide significance ($P < 5 \times 10^{-8}$) in a BOLT-LMM⁴³ GWAS of educational attainment in UK Biobank with the siblings excluded, as in a previous publication.¹⁸ The summary data were linkage disequilibrium clumped ($r^2 < 0.001$, physical distance threshold = 10 000 kb, $P < 5 \times 10^{-8}$) in PLINK⁴⁴ to generate 350 independent genetic variants. We regressed the resulting PGS on age and sex, and the

standardized residuals (mean 0, SD=1) were used in the analysis. In the population model, the outcome was regressed on the PGS. In the within-sibship model, the mean sibship PGS was included as a covariate to account for variation in parental genotypes. The PGS approach is equivalent to an inverse-variance weighted estimator from a summary-based two-sample MR analysis.⁴⁵

UK Biobank and HUNT meta-analyses

Population and within-sibship models were fitted separately in UK Biobank and HUNT, and the estimates were meta-analysed using a fixed-effects model in the metafor R package. Shrinkage in estimates from the population to the within-sibship model was estimated as follows, with standard errors estimated using the delta method:

$$\frac{\beta^{Pop} - \beta^{WS}}{\beta^{Pop}}$$

UK Biobank and HUNT MR

MR estimates of the effects of educational attainment on the outcomes (BMI, pack years, SBP, mortality) were derived from the meta-analysis PGS association estimates (i.e. PGS–educational attainment, PGS–outcome). The point estimate was calculated using the Wald ratio of the PGS–outcome and PGS–educational attainment associations. Wald ratio standard errors were estimated using the delta method.

MR analyses require three core assumptions. First, genetic variants are strongly associated with the exposure (relevance); second, no unmeasured confounders of the association between the genetic variants and the outcome (independence); and third, genetic variants only influence the outcome via the exposure (exclusion–restriction).^{46–48} As discussed in previous work,¹⁷ MR estimates of categorical exposures such as educational attainment should generally be interpreted in terms of liability (e.g. liability to educational attainment) rather than effects of the categorical phenotype (e.g. years of schooling).

Within-sibship meta-analysis GWAS MR

We also performed two-sample MR analyses using within-sibship GWAS summary data from a recent within-sibship meta-analysis GWAS of 25 phenotypes.¹⁸ This study included data from UK Biobank, HUNT and an additional 16 cohorts (each with between $N=618$ and $13\,856$). GWAS data were available for educational attainment ($N=128\,777$) as well as BMI ($N=140\,883$), SBP ($N=109\,588$), ever smoking ($N=124\,791$) and cigarettes per day in ever smokers ($N=28\,134$). These GWASs were conducted in the same studies so there is near-complete sample overlap between the different GWASs.

As genetic instruments, we used the same 350 genetic variants as in the UK Biobank and HUNT analyses described above, which were derived from a BOLT-LMM GWAS of educational attainment in UK Biobank with the siblings excluded ($P < 5 \times 10^{-8}$, $r^2 < 0.001$, physical distance threshold = 10 000 kb). Using the within-sibship meta-analysis GWAS data we then derived MR effect estimates (β_{MR}) of educational attainment on the four health outcomes in both population and within-sibship models using an inverse-variance weighted approach¹⁸ as follows:

$$\beta_{MR} = \sum_1^n \frac{\beta_{EA}^k * \beta_{OUT}^k}{(\sigma_{OUT}^k)^2} / \sum_1^n \frac{(\beta_{EA}^k)^2}{(\sigma_{OUT}^k)^2}$$

where β_{EA} represents the association estimate from educational attainment GWAS, β_{OUT} represents the association estimate from outcome GWAS, σ_{OUT} represents the standard error from outcome GWAS, n represents the number of genetic variants and k represents the k -th variant.

The standard error of β_{MR} was estimated as follows:

$$SE(\beta_{MR}) = \sqrt{\frac{1}{\sum_1^n \frac{(\beta_{EA}^k)^2}{(\sigma_{OUT}^k)^2}}}$$

where n represents the number of genetic variants and k represents the k -th variant.

Finnish twin cohort analysis

UK Biobank and HUNT analyses included both twin and non-twin siblings, with the vast majority (>95%) being non-twin siblings of different ages—a potential concern with educational attainment trends changing over time. We investigated whether the observed inverse association between educational attainment and mortality persisted in twin-only analyses using data on 27 229 individuals from the Finnish twin cohort that included 3518 MZ and 7718 DZ twin pairs.

The association between educational attainment and mortality was estimated in the whole sample ($N=27\,229$) using Cox-proportional hazards models with adjustment for sex and smoking (population model). Stratified Cox-proportional hazard models were applied to MZ and DZ twins separately with baseline hazards stratified by twin pair, with adjustment for smoking. All twin pairs included were of the same sex. All analyses were performed in Stata using the stcox package.

Results

Phenotypic educational attainment, health outcomes and mortality in UK Biobank and HUNT

Higher educational attainment was strongly associated with lower BMI, pack years of smoking, SBP and mortality

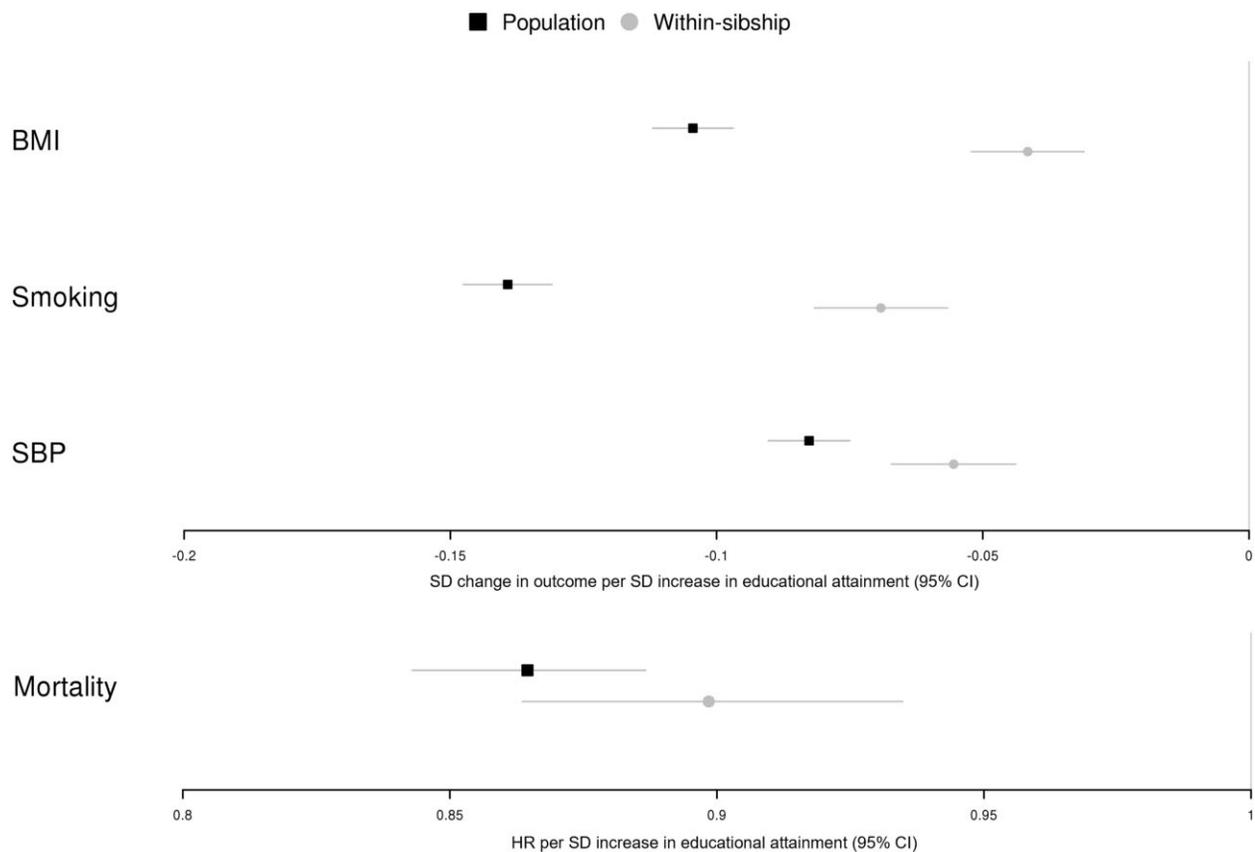


Figure 2 Phenotypic educational attainment and health outcomes. **Figure 2** shows associations between phenotypic educational attainment (years in full-time education derived using qualifications) and body mass index, smoking (cigarettes, measured in pack years), systolic blood pressure and mortality in the population and within-sibship models in UK Biobank and HUNT. Estimates for mortality are presented as hazard ratios with the rest of the estimates presented in standard deviation units. BMI, body mass index; SBP, systolic blood pressure

in both population and within-sibship models. In within-sibship models, a 1-SD higher educational attainment (corresponding to an additional 2.3 years in UK Biobank and 1.2 years in HUNT) was associated with lower BMI (0.04 SD; 95% CI 0.03, 0.05), fewer pack years of cigarette smoking (0.10 SD; 95% CI 0.08, 0.12), lower SBP (0.06 SD; 95% CI 0.04, 0.07) and lower mortality [hazard ratio (HR) 0.90; 95% CI 0.86, 0.93]. Population estimates that did not account for family-level confounding were in the same direction but substantially larger (34% for mortality to 146% for BMI) than the within-sibship point estimates (**Figure 2** and **Supplementary Table S1**, available as **Supplementary data** at *IJE* online).

Phenotypic educational attainment and mortality in the Finnish twin cohort

In non-twin population regression models, using data from the whole sample, 1-SD higher educational attainment was strongly associated with lower mortality (HR 0.95; 95% CI 0.93, 0.97) after adjusting for sex and smoking. Estimates from the within DZ twin pair (HR 0.91; 95% CI 0.83, 1.01)

and within MZ twin pair (HR 0.87; 95% CI 0.70, 1.08) analyses were broadly consistent with the Finnish twin cohort population estimate as well as the UK Biobank and HUNT within-sibship estimate (HR 0.90; 95% CI 0.86, 0.93) but confidence intervals overlap with the null hypothesis. In sex-stratified twin analyses, point estimates were larger in magnitude for males than for females (**Supplementary Table S2**, available as **Supplementary data** at *IJE* online).

Educational attainment PGS, educational attainment, health outcomes and mortality in UK Biobank and HUNT

The educational attainment PGS was strongly associated with educational attainment in both population and within-sibship models. Consistently with previous studies,^{27,29,49} the population PGS association estimate attenuated by 49% (95% CI 40%, 58%) in the within-sibship model. In the population model, a higher educational attainment PGS was associated with lower BMI, fewer pack years of cigarette smoking, lower SBP and lower mortality. In the within-sibship model, the PGS was associated with BMI, cigarette

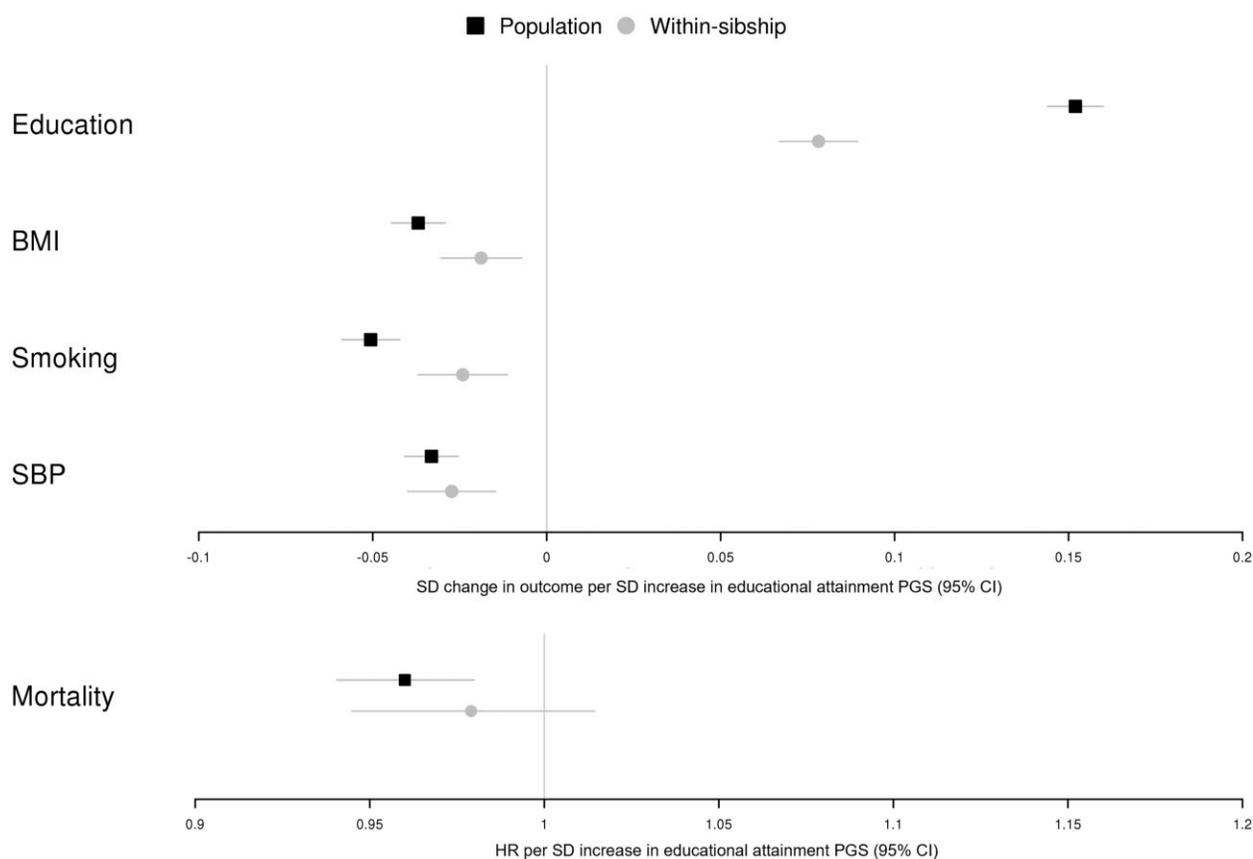


Figure 3 Educational attainment polygenic score (PGS), educational attainment and health outcomes. **Figure 3** shows associations between an educational attainment PGS and education (measured educational attainment), body mass index, cigarette smoking (pack years), systolic blood pressure and mortality in the population and within-sibship models in UK Biobank and HUNT. Estimates for mortality are presented as hazard ratios per standard deviation increase in the polygenic score with the rest of the outcome estimates presented in standard deviation units

smoking and SBP in the same direction but there was limited evidence for an association with mortality, likely because of lower statistical power. The within-sibship PGS association estimates for BMI and cigarette smoking were 49% (95% CI 16%, 82%) and 52% (95% CI 26%, 79%) smaller than the population PGS estimates, respectively, consistently with the within-sibship attenuations for educational attainment (**Figure 3** and **Supplementary Tables S3 and S4**, available as **Supplementary data** at *IJE* online).

MR of educational attainment on health outcomes and mortality using UK Biobank and HUNT PGS estimates

Population MR estimates indicated that a 1-SD increase in liability to educational attainment reduced BMI by 0.24 SD units (95% CI 0.19, 0.29), pack years of cigarette smoking by 0.33 SD (95% CI 0.28, 0.39) and SBP by 0.22 SD (95% CI 0.17, 0.27) (**Figure 4**). Within-sibship MR estimates were consistent with the population MR estimates for BMI (0.24; 95% CI 0.09, 0.39), cigarette smoking (0.31; 95% CI 0.14, 0.48) and SBP (0.35; 95% CI

0.18, 0.52). The population and within-sibship MR point estimates for mortality were also consistent (HR per SD increase in educational attainment; population 0.76; 95% CI 0.67, 0.88; within-sibship 0.76; 95% CI 0.48, 1.20) but the imprecision of the within-sibship estimate prevented stronger conclusions.

Differences between population and within-sibship MR estimates are a function of differences in the PGS–outcome and PGS–educational attainment estimates. If the PGS estimates change by the same proportion from the population model to the within-sibship model, then the population and within-sibship MR estimates will be consistent (**Figure 4** and **Supplementary Tables S4 and S5**, available as **Supplementary data** at *IJE* online).

MR of educational attainment on health outcomes using within-sibship GWAS summary data

Population and within-sibship MR estimates based on the within-sibship GWAS summary data provided further evidence that higher educational attainment lowers BMI, risk of ever smoking and SBP. Confidence intervals for

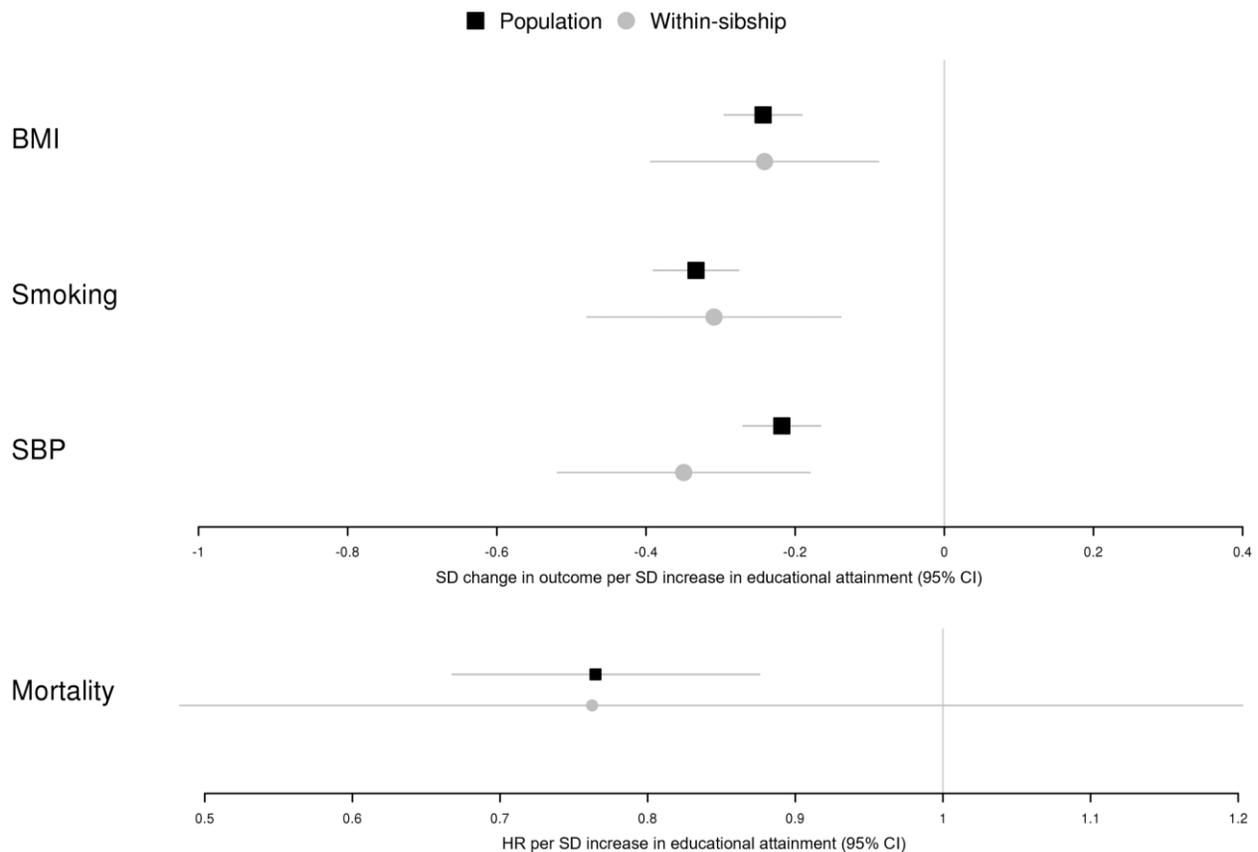


Figure 4 Mendelian randomization (MR) estimates of educational attainment on health outcomes from UK Biobank and HUNT. **Figure 4** shows population and within-sibship MR estimates of the effect of educational attainment on body mass index, smoking (pack years of cigarette smoking), systolic blood pressure and mortality. These estimates were derived using the polygenic score association estimates in **Figure 3** from the UK Biobank and HUNT studies. Estimates are presented in standard deviation units for body mass index, smoking and systolic blood pressure, and as hazard ratios for mortality

cigarettes per day overlapped the null in both models but statistical power was limited because data were only collected in ever smokers. There was some evidence that the within-sibship MR estimates were smaller than the population estimates for BMI (51%; 95% CI 19%, 83%) and SBP (52%; 95% CI 4%, 100%), which was not observed in the UK Biobank and HUNT analyses. The standard errors for the within-sibship MR estimates for BMI and SBP were 47% and 49% smaller, respectively, in the two-sample MR analyses compared with the UK Biobank and HUNT analyses because of the larger sample size of the within-sibship GWAS (**Figure 5** and **Supplementary Table S6**, available as **Supplementary data** at *IJE* online).

Discussion

We used within-sibship MR to provide evidence that higher liability to educational attainment reduces BMI, cigarette smoking and SBP. These findings strengthen evidence for beneficial effects of educational attainment, or closely correlated trait, on adulthood health by illustrating that previously observed effects persist when population stratification,

assortative mating and indirect genetic effects from parents are controlled for in within-sibship MR.

An important consideration for MR analyses of educational attainment is that genetic variants instrument liability to educational attainment, a latent measure of educational attainment, rather than specific measured educational attainment phenotypes (e.g. having a university degree)^{17,50} or other related traits (e.g. cognition). Genetic variants influence measured education phenotypes via their effect on liability to educational attainment and so genetic association estimates will capture all effects of liability, which may or may not act via changes in the measured education phenotype. For example, an educational-attainment-increasing genetic variant may influence an outcome because the variant increases the probability that an individual attains a measured qualification such as a university degree, but also because the variant influences related unmeasured characteristics such as choice of educational track, a personality phenotype or cognitive ability.

The distinction between measured education phenotypes and liability to educational attainment is particularly relevant in the within-sibship model because siblings often have

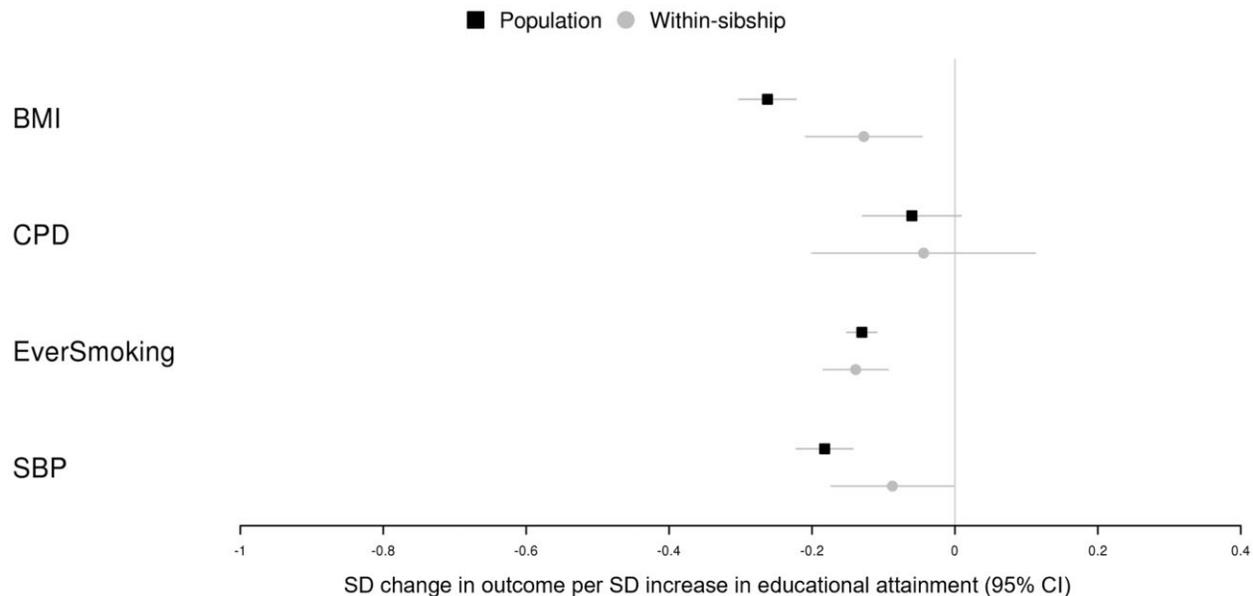


Figure 5 Mendelian randomization estimates of educational attainment on health outcomes using summary data from the within-sibship Genome-wide Association Study (GWAS). **Figure 5** shows population and within-sibship Mendelian randomization estimates (inverse-variance weighted) of the effect of educational attainment on body mass index, cigarettes per day measured in ever smokers only, ever smoking and systolic blood pressure. These estimates were derived using GWAS summary statistics from a within-sibship meta-analysis GWAS of ≤ 18 studies. Estimates are presented in standard deviation units except for ever smoking (binary) where estimates are in terms of risk difference %. BMI, body mass index; CPD, cigarettes per day; SBP, systolic blood pressure

genotype differences for educational attainment variants but often have the same measured educational attainment values (e.g. both siblings attended university). A conventional MR analysis assumes that genetic differences between siblings do not affect the outcome if the siblings have the same value for the measured education phenotype. This is implausible as genetic variants are likely to influence unmeasured differences between siblings such as the choice of degree. Therefore, we interpret the within-sibship MR results as providing evidence that an underlying liability to education has beneficial effects on health, rather than specific educational attainment qualifications.^{17,51}

Consistently with previous studies,^{18,26,52} the association estimate between the educational attainment PGS and educational attainment attenuated on average by around a half from the population model to the within-sibship model when using weights from a GWAS of unrelated individuals. However, the association estimates of educational attainment PGS on health outcomes were also attenuated by a similar degree. As the attenuation was balanced, the population and within-sibship MR effect estimates, which are a ratio of single nucleotide polymorphism (SNP) (or PGS)–outcome and SNP–exposure associations, were consistent. These results illustrate how population stratification, assortative mating and indirect genetic effects can distort genetic association estimates but will not necessarily affect MR estimates if the gene–exposure and gene–outcome association estimates are affected proportionally.

A caveat of our work is that MR estimates are sensitive to the assumption that genetic variants only influence the outcome via their effect on (liability to) educational attainment (exclusion–restriction), which could be violated by directional pleiotropy. Previous work has illustrated that MR with measured educational attainment is unlikely to satisfy this assumption, despite use of pleiotropy-robust methods.¹⁶ We presented estimates in terms of liability to educational attainment to acknowledge the likelihood of such effects but note that, in the context of interventions, effects of liability to educational attainment are less useful than effects of specific education phenotypes. Future studies could use multivariable MR and structural equation modelling approaches to disentangle mechanisms underlying our results by exploring potential pleiotropic pathways potentially relating to both cognitive and non-cognitive phenotypes.^{53,54}

Our work has further limitations. First, our within-sibship MR estimate for mortality was imprecise because mortality data were only available in UK Biobank and HUNT. Second, educational attainment is known to influence participation in biobanks so our study may have been susceptible to selection bias.⁵⁵ Third, the gene–exposure and gene–outcome estimates in the within-sibship GWAS two-sample MR analyses were from largely overlapping samples, which could have potentially induced modest bias.⁵⁶ Fourth, there is evidence that within-sibship models using PGS based on weights from population GWAS could

introduce bias.⁵⁷ This could have potentially affected our MR estimates from the individual-level PGS approach in UK Biobank and HUNT but not our MR estimates using the within-sibship meta-analysis GWAS data where the gene–exposure and gene–outcome estimates were derived from within-sibship models. Estimates from the two different MR approaches provided consistent qualitative evidence of effects of educational attainment on the tested outcomes suggesting that this potential limitation is unlikely to have affected our overall conclusions. There were some quantitative differences between the two approaches, with evidence of within-sibship shrinkage from the summary-based MR estimates of educational attainment on BMI and SBP but not from the PGS approach. Fifth, within-sibship models do not control for indirect genetic effects of siblings. Previous studies have indicated that sibling indirect genetic effects are likely to be small, suggesting that they are unlikely to have impacted our findings.

We found compelling evidence that educational attainment (or liability to educational attainment) influences BMI, smoking and SBP, even after accounting for population stratification, assortative mating and indirect genetic effects of parents. Within-family MR more closely emulates a randomized experiment because of random variation in meiotic segregation within families³¹ but has been historically limited by data availability. The emerging availability of within-family GWAS data will enable researchers to better disentangle the effects of social and behavioural phenotypes on health outcomes.

Ethics approval

This research has been conducted using the UK Biobank resource under Application Number 8786. UK Biobank obtained ethics approval from the North West Multi-centre Research Ethics Committee and obtained informed consent from all study participants. The use of HUNT data in this study was approved by the Regional Committee for Ethics in Medical Research, Central Norway (2017/2479). All participants signed informed consent for participation and the use of data in research. Register linkages and use of the questionnaire data were approved by the Ethics Committee at the Finnish Institute of Health and Welfare (THL 220/6.02.04/2021).

Data availability

UK Biobank individual-level participant data are available via enquiry to access@ukbiobank.ac.uk. Researchers associated with Norwegian research institutes can apply for the use of HUNT data and samples with approval by the Regional Committee for Medical and Health Research Ethics. Researchers from other countries may apply if collaborating with a Norwegian Principal Investigator. Information for data access can be found at <https://www.ntnu.edu/hunt/data>. The HUNT variables are available for browsing on the HUNT databank at <https://hunt-db.medisin.ntnu.no/hunt-db/>. Use

of the full genetic data set requires the use of an approved secure computing solution such as the HUNT Cloud (<https://docs.hdc.ntnu.no/>). Example scripts for population and within-sibship models are available on GitHub <https://github.com/LaurenceHowe/EducationSiblingMR/>. Summary data from the within-sibship meta-analysis GWAS are publicly available for download on OpenGWAS (<https://gwas.mrcieu.ac.uk/>) via the TwoSampleMR R package. Note that the summary data include both ‘population’ and ‘within-sibship’ estimates for each phenotype, with the model detailed in the metadata notes.

Supplementary data

Supplementary data are available at *IJE* online.

Author contributions

L.J.H., B.M.B. and N.M.D. conceptualized the project. L.J.H. performed UK Biobank and HUNT analyses under the supervision of G.D.S., B.M.B. and N.M.D. H.R. and B.M.B. contributed extensively to the HUNT analyses. Within-family consortium authors contributed data, expertise and scientific advice to the meta-analysis GWAS summary data used in MR analyses. J.K. proposed and performed the Finnish twin cohort analyses. L.J.H. drafted the original manuscript. All co-authors contributed to the interpretation of results and writing of the manuscript.

Funding

The University of Bristol support the MRC Integrative Epidemiology Unit (MC_UU_00011/1). The Trøndelag Health Study (HUNT) is a collaboration between HUNT Research Centre (Faculty of Medicine and Health Sciences, NTNU, Norwegian University of Science and Technology), Trøndelag County Council, Central Norway Regional Health Authority and the Norwegian Institute of Public Health. The genotyping in HUNT was financed by the National Institutes of Health; University of Michigan; the Research Council of Norway; the Liaison Committee for Education, Research and Innovation in Central Norway; and the Joint Research Committee between St Olavs hospital and the Faculty of Medicine and Health Sciences, NTNU. The K.G. Jebsen Center for Genetic Epidemiology is financed by Stiftelsen Kristian Gerhard Jebsen; Faculty of Medicine and Health Sciences, NTNU, Norway. N.M.D. is supported by a Norwegian Research Council Grant number 295989. J.F.W. acknowledges support from the MRC Human Genetics Unit programme grant, ‘Quantitative traits in health and disease’ (U. MC_UU_00007/10). J.B.P. is funded by the European Research Council (ERC) under the European Union’s Horizon 2020 research and innovation programme (grant agreement No. 863981). S.L. is a Victorian Cancer Agency Early Career Research Fellow (ECRF19020). J.L.H. is a National Health and Medical Research Council Senior Principal Research Fellow. P.R.J. and O.E.N. are funded by the Research Council of Norway (#28743). J.K. is funded by the Academy of Finland (grants 312073 and 336823) and the Sigrid Juselius Foundation. P.M. was supported by the Academy of Finland (#308247, # 345219) and the ERC under the European Union’s Horizon 2020 research and innovation programme (grant agreement No. 101019329). The funders had no role in study design, data collection and analysis, decision to publish or preparation

of the manuscript. This publication is the work of the authors, who serve as the guarantors for the contents of this paper.

Acknowledgements

Quality Control filtering of the UK Biobank data was conducted by R. Mitchell, G. Hemani, T. Dudding and L. Paternoster as described in the published protocol (doi: 10.5523/bris.3074krb6t2frj29yh2b03x3wxj).

Conflict of interest

None declared.

References

- Cutler DM, Lleras-Muney A. Education and health: evaluating theories and evidence. National Bureau of Economic Research. 2006. Report No. 0898–2937. <https://www.nber.org/papers/w12352> (June 2021, date last accessed).
- Mackenbach JP, Rubio Valverde J, Bopp M *et al.* Progress against inequalities in mortality: register-based study of 15 European countries between 1990 and 2015. *Eur J Epidemiol* 2019;**34**:1131–42.
- Galama TJ, Lleras-Muney A, van Kippersluis H. The effect of education on health and mortality: a review of experimental and quasi-experimental evidence. National Bureau of Economic Research. 2018. Report No. 0898-2937. <https://www.nber.org/papers/w24225> (June 2021, date last accessed).
- Davies NM, Dickson M, Davey Smith G, van den Berg GJ, Windmeijer F. The causal effects of education on health outcomes in the UK Biobank. *Nat Hum Behav* 2018;**2**:117–25.
- van Kippersluis H, O'Donnell O, van Doorslaer E. Long run returns to education: does schooling lead to an extended old age? *J Hum Resour* 2009;**4**:1–33.
- Clark D, Royer H. The effect of education on adult mortality and health: evidence from Britain. *Am Econ Rev* 2013;**103**:2087–120.
- Lager AC, Torssander J. Causal effect of education on mortality in a quasi-experiment on 1.2 million Swedes. *Proc Natl Acad Sci USA* 2012;**109**:8461–66.
- Meghir C, Palme M, Simeonova E. Education and mortality: evidence from a social experiment. *Am Econ J Appl Econ* 2018;**10**:234–56.
- Lundborg P, Lyttkens CH, Nystedt P. The effect of schooling on mortality: new evidence from 50,000 Swedish Twins. *Demography* 2016;**53**:1135–68.
- Behrman JR, Kohler HP, Jensen VM *et al.* Does more schooling reduce hospitalization and delay mortality? New evidence based on Danish twins. *Demography* 2011;**48**:1347–75.
- Silventoinen K, Piirtola M, Jelenkovic A *et al.* Smoking remains associated with education after controlling for social background and genetic factors in a study of 18 twin cohorts. *Sci Rep* 2022;**12**:13148.
- Davey Smith G, Ebrahim S. 'Mendelian randomization': can genetic epidemiology contribute to understanding environmental determinants of disease? *Int J Epidemiol* 2003;**32**:1–22.
- Tillmann T, Vaucher J, Okbay A *et al.* Education and coronary heart disease: Mendelian randomisation study. *BMJ* 2017;**358**:j3542.
- Gage SH, Bowden J, Smith GD, Munafo MR. Investigating causality in associations between education and smoking: a two-sample Mendelian randomization study. *Int J Epidemiol* 2018;**47**:1131–40.
- Zhou H, Zhang Y, Liu J *et al.* Education and lung cancer: a Mendelian randomization study. *Int J Epidemiol* 2019;**48**:743–50.
- Zeng L, Ntalla I, Kessler T *et al.*; UK Biobank CardioMetabolic Consortium CHD Working Group. Genetically modulated educational attainment and coronary disease risk. *Eur Heart J* 2019;**40**:2413–20.
- Howe LJ, Tudball M, Smith GD, Davies NM. Interpreting Mendelian randomization estimates of the effects of categorical exposures such as disease status and educational attainment. *Int J Epidemiol* 2021; 2020.12.14.20248168.
- Howe LJ, Nivard MG, Morris TT *et al.*; Within Family Consortium. Within-sibship genome-wide association analyses decrease bias in estimates of direct genetic effects. *Nat Genet* 2022;**54**:581–92.
- Brumpton B, Sanderson E, Heilbron K *et al.* Avoiding dynastic, assortative mating, and population stratification biases in Mendelian randomization through within-family analyses. *Nat Commun* 2020;**11**:3519. <https://doi.org/10.1038/s41467-020-17117-4>.
- Davies NM, Howe LJ, Brumpton B, Havdahl A, Evans DM, Davey Smith G. Within family Mendelian randomization studies. *Hum Mol Genet* 2019;**28**:R170–79.
- Hartwig FP, Davies NM, Davey Smith G. Bias in Mendelian randomization due to assortative mating. *Genet Epidemiol* 2018;**42**:608–20.
- Haworth S, Mitchell R, Corbin L *et al.* Apparent latent structure within the UK Biobank sample has implications for epidemiological analysis. *Nat Commun* 2019;**10**:333.
- Yengo L, Robinson MR, Keller MC *et al.* Imprint of assortative mating on the human genome. *Nat Hum Behav* 2018;**2**:948–54.
- Robinson MR, Kleinman A, Graff M *et al.*; The LifeLines Cohort Study. Genetic evidence of assortative mating in humans. *Nat Hum Behav* 2017;**1**:0016.
- Domingue BW, Fletcher J, Conley D, Boardman JD. Genetic and educational assortative mating among US adults. *Proc Natl Acad Sci USA* 2014;**111**:7996–8000.
- Kong A, Thorleifsson G, Frigge ML *et al.* The nature of nurture: effects of parental genotypes. *Science* 2018;**359**:424–28.
- Lee JJ, Wedow R, Okbay A *et al.*; Social Science Genetic Association Consortium. Gene discovery and polygenic prediction from a genome-wide association study of educational attainment in 1.1 million individuals. *Nat Genet* 2018;**50**:1112–21.
- Howe LJ, Evans DM, Hemani G, Davey Smith G, Davies NM. Evaluating indirect genetic effects of siblings using singletons. *PLoS Genet* 2022;**18**:e1010247.
- Wang B, Baldwin JR, Schoeler T *et al.* Robust genetic nurture effects on education: A systematic review and meta-analysis based on 38,654 families across 8 cohorts. *Am J Hum Genet* 2021;**108**:1780–91.
- Okbay A, Wu Y, Wang N *et al.*; Social Science Genetic Association Consortium. Polygenic prediction of educational attainment within and between families from genome-wide association analyses in 3 million individuals. *Nat Genet* 2022;**54**:437–49.

31. Davey Smith G, Holmes MV, Davies NM, Ebrahim S. Mendel's laws, Mendelian randomization and causal inference in observational data: substantive and nomenclatural issues. *Eur J Epidemiol* 2020;**35**:99–111.
32. Sudlow C, Gallacher J, Allen N *et al*. UK Biobank: an open access resource for identifying the causes of a wide range of complex diseases of middle and old age. *PLoS Med* 2015;**12**:e1001779.
33. Bycroft C, Freeman C, Petkova D *et al*. The UK Biobank resource with deep phenotyping and genomic data. *Nature* 2018;**562**:203–09.
34. Hill WG, Weir BS. Variation in actual relationship as a consequence of Mendelian sampling and linkage. *Genet Res (Camb)* 2011;**93**:47–64.
35. Okbay A, Beauchamp JP, Fontana MA *et al.*; LifeLines Cohort Study. Genome-wide association study identifies 74 loci associated with educational attainment. *Nature* 2016;**533**:539–42.
36. Holmen J, Midthjell K, Krüger Ø *et al*. The Nord-Trøndelag Health Study 1995–97 (HUNT 2): objectives, contents, methods and participation. *Norsk Epidemiologi* 2003;**13**:19–32.
37. Åsvold BO, Langhammer A, Rehn TA *et al*. Cohort Profile Update: The HUNT Study, Norway. *Int J Epidemiol* 2023;**52**:e80–e91.
38. Brumpton BM, Graham S, Surakka I *et al*. The HUNT Study: a population-based cohort for genetic research. *Cell Genom* 2022;**2**:100193. <https://doi.org/10.1016/j.xgen.2022.100193>.
39. Krokstad S, Langhammer A, Hveem K *et al*. Cohort profile: the HUNT study, Norway. *Int J Epidemiol* 2013;**42**:968–77.
40. Manichaikul A, Mychaleckyj JC, Rich SS, Daly K, Sale M, Chen WM. Robust relationship inference in genome-wide association studies. *Bioinformatics* 2010;**26**:2867–73.
41. Kaprio J, Bollepalli S, Buchwald J *et al*. The older Finnish Twin Cohort—45 years of follow-up. *Twin Res Hum Genet* 2019;**22**:240–54.
42. Kaprio J, Sarna S, Koskenvuo M, Rantasalo I. The Finnish Twin Registry: formation and compilation, questionnaire study, zygosity determination procedures, and research program. *Progr Clin Biol Res* 1978;**24**:179–84.
43. Loh P-R, Tucker G, Bulik-Sullivan BK *et al*. Efficient Bayesian mixed-model analysis increases association power in large cohorts. *Nat Genet* 2015;**47**:284–90.
44. Purcell S, Neale B, Todd-Brown K *et al*. PLINK: a tool set for whole-genome association and population-based linkage analyses. *Am J Hum Genet* 2007;**81**:559–75.
45. Dudbridge F. Polygenic Mendelian randomization. *Cold Spring Harb Perspect Med* 2021;**11**:a039586.
46. Haycock PC, Burgess S, Wade KH, Bowden J, Relton C, Davey Smith G. Best (but oft-forgotten) practices: the design, analysis, and interpretation of Mendelian randomization studies. *Am J Clin Nutr* 2016;**103**:965–78.
47. Didelez V, Sheehan N. Mendelian randomization as an instrumental variable approach to causal inference. *Stat Methods Med Res* 2007;**16**:309–30.
48. Lawlor DA, Harbord RM, Sterne JA, Timpson N, Davey Smith G. Mendelian randomization: using genes as instruments for making causal inferences in epidemiology. *Stat Med* 2008;**27**:1133–63.
49. Morris TT, Davies NM, Hemani G, Davey Smith G. Population phenomena inflate genetic associations of complex social traits. *Sci Adv* 2020;**6**:eaay0328.
50. Frisell T, Oberg S, Kuja-Halkola R, Sjölander A. Sibling comparison designs: bias from non-shared confounders and measurement error. *Epidemiology* 2012;**23**:713–20.
51. Morris TT, Heron J, Sanderson ECM, Davey Smith G, Didelez V, Tilling K. Interpretation of Mendelian randomization using a single measure of an exposure that varies over time. *Int J Epidemiol* 2022;**51**:1899–1909.
52. Mostafavi H, Harpak A, Agarwal I, Conley D, Pritchard JK, Przeworski M. Variable prediction accuracy of polygenic scores within an ancestry group. *eLife* 2020;**9**:e48376.
53. Demange PA, Hottenga JJ, Abdellaoui A *et al*. Estimating effects of parents' cognitive and non-cognitive skills on offspring education using polygenic scores. *Nat Commun* 2022;**13**:4801. <https://doi.org/10.1038/s41467-022-32003-x>.
54. Davies NM, Hill WD, Anderson EL, Sanderson E, Deary IJ, Davey Smith G. Multivariable two-sample Mendelian randomization estimates of the effects of intelligence and education on health. *eLife* 2019;**8**:e43990.
55. Munafò MR, Tilling K, Taylor AE, Evans DM, Davey Smith G. Collider scope: when selection bias can substantially influence observed associations. *Int J Epidemiol* 2018;**47**:226–35.
56. Mounier N, Kutalik Z. Bias correction for inverse variance weighting Mendelian randomization. *Genet Epidemiol* 2023;**47**:314–31.
57. Fletcher J, Wu Y, Li T, Lu Q. Interpreting Polygenic Score Effects in Sibling Analysis. *bioRxiv* 2021: 2021.07.16.452740, preprint: not peer reviewed.