

**Association between meditative capacities and cognitive functions in healthy older adults
naïve to meditation practice**

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Abstract

Objectives Aging people experience a slight decrease in their cognitive efficiency, even in the absence of brain pathology. Concurrently, several studies have reported positive effects of meditation practice on older adults' cognitive functioning. This study aimed to assess if dispositional mindfulness (or more generally trait meditation capacities) was associated with better cognition during aging. **Method** We analyzed cross-sectional data from 134 healthy elderly participants enrolled in the Age-Well trial (Age: 69.0 ± 3.8 , 61.2% female) using a series of linear regressions. Participants were naïve to meditation practice before inclusion in the study. Three core meditation capacities were assessed: *attentional* related to metacognition/regulation of attention, *constructive* assessing attitudes towards others or towards ourselves, and *deconstructive* focusing on cognitive defusion. Cognitive abilities were assessed through four composite measures of attention, executive function, episodic memory, and a global composite sensitive to subtle age-related cognitive changes linked to dementia risk (Preclinical Alzheimer's Cognitive Composite [PACC5]). **Results** There was a positive relationship between PACC5 and deconstructive capacity ($R^2=.03$; $p=.04$). However, no association was observed between the three meditation capacities and the three specific cognitive scores ($p>.05$). **Conclusions** We propose that deconstructive capacity, associated with self-inquiry and down-regulation of maladaptive affective schemes, could be a cognitive factor important for global cognition in healthy aging. It remains to be determined to what extent explicit training in meditation positively influences these capacities and whether these changes also contribute to better cognition in aging. **Preregistration** NCT02977819 (<https://ClinicalTrials.gov>).

Keywords: meditation; aging; cognition; personality; affective state

Based on United Nations predictions (2019), by 2050, 1/6 people in the world will be over 65, compared to 1/11 in 2019. This increase in life expectancy requires the promotion of healthy aging at physical, psychological, and cognitive levels. It is now well established that even in the absence of brain pathology, aging is associated with subtle declines in cognitive functioning and efficiency in the domains of memory, attention and executive processes (Craik & Salthouse, 2000). However, cognitive decline rates and trajectories are characterized by great variability (Olaya et al., 2017; Rönnlund et al., 2005) and do not necessarily lead to dementia (Qiu & Fratiglioni, 2018). The heterogeneity in cognitive aging profiles can be explained by the coexistence of adverse and protective factors at the biological and psychological levels (Park & Reuter-Lorenz, 2009; Reuter-Lorenz & Park, 2014). Cognitive decline might be influenced by education (Fratiglioni & Wang, 2007; Meng & D'Arcy, 2012), employment complexity and autonomy (Then et al., 2014), physical activity, engagement in cognitively demanding leisure activities and/or sustained social interactions (Wang et al., 2012). Similarly, negative affect including depression, anxiety, rumination, and worry are associated with cognition in aging even at subclinical levels (Dotson, 2014; Marchant et al., 2020).

In this context, it has been proposed that mindfulness meditation might be a useful approach to prevent or delay age-related cognitive decline (Gard et al., 2014a; Reuter-Lorenz & Park, 2014). Indeed, meditation practice positively affects brain functions and structures relevant to cognition, which could, therefore, enhance various cognitive processes, including attention, memory, and executive function (Chiesa et al., 2011; Tang et al., 2015). However, the benefit can sometimes be specific to subprocesses, for example regarding executive functioning (Gallant, 2016). With regard to the aging effects, cross-sectional studies on expert meditators revealed that they have higher levels of fluid intelligence (Gard et al., 2014b), as well as attention and executive function compared with non-meditators (Prakash et al., 2012; Sperduti et al., 2016), and also a younger brain age (Luders et al., 2016). Moreover, intervention

studies reported improvements in memory, attention, and executive function after short-term training of older people naïve (i.e., without previous experience) to meditation practice (Gallant, 2016; Gard et al., 2014a) and in patients at risk of or diagnosed with dementia (Innes et al., 2017; Newberg et al., 2010). A recent meta-analysis supported these results by showing a positive effect of mindfulness-based programs on executive function in older adults (≥ 60 years) (Whitfield et al., 2021).

Even if participants do not formally have a meditation practice, dispositional mindfulness could also be a protective factor for cognitive aging. Dispositional mindfulness is defined as the tendency to engage in receptive attention to and awareness of current experiences with non-judgment and acceptance (Baer et al., 2006; Kabat-Zinn, 2003). Several studies reported that dispositional mindfulness seems to increase with age (Boekel & Hsieh, 2018; Prakash et al., 2015), and this rise in mindfulness could partially account for lower negative affect levels (Raes et al., 2015) also seen to decrease with age (Jensen et al., 2020). However, at this time, very few studies have investigated the relationship between mindfulness dispositions and cognition in older adults.

Prakash et al. (2015) assessed the associations between interference resolution in various cognitive tasks (working memory, inhibition, flexibility) and dispositional mindfulness assessed via the Mindful Attention and Awareness Scale score (MAAS; Brown & Ryan, 2003). They failed to observe a positive link between dispositional mindfulness and their attentional control / interference resolution measures among either young or older adults. However, in another study using the same assessment of dispositional mindfulness, Fiocco and Mallya (2015) showed that cognitive flexibility (assessed by part B of the Trail Making Test) was positively associated with the MAAS score in older adults, while no association was found with episodic memory. Finally, Fountain-Zaragoza et al. (2016) observed that the MAAS score was

positively associated with reactive attentional control during a continuous performance task, but not during a Go/No-go task, and was not associated with proactive control or with variability in response times. Dispositional mindfulness was measured in these studies with the MAAS, a scale referring to mindfulness as a unidimensional construct based mainly on attentional capacities (Brown & Ryan, 2003). However, the term mindfulness, and more generally meditation, refers to a wide variety of contemplative practices, and meditation can be considered a family of complex emotional and attentional regulatory strategies which allow the cultivation of well-being and emotional balance (Lutz et al., 2008).

A recent taxonomy has been proposed to categorize meditation practices into attentional, constructive and deconstructive families (Dahl et al., 2015). These categories are theoretically informed by their primary cognitive mechanism, consequently the main mechanism of one family can also be cultivated by and/or be necessary for meditation practices in other families, leading to some overlap in processes between families. In short, practices in the attentional family (for example, mindfulness-based practices) systematically train the ability to initiate, direct, and sustain attentional processes and to be aware of thoughts, feelings, and perceptions (i.e., meta-awareness). These abilities are systematically trained in mindfulness-based practices. Practices in the constructive family (for example, compassion and loving-kindness practices) cultivate and strengthen cognitive and affective patterns that positively impact well-being. Perspective-taking and cognitive reappraisal are the core cognitive mechanisms of constructive practices. The deconstructive family (for example, Vipassana or insight practices) consists of analytical and object-oriented meditation. The aim is to weaken and dissolve maladaptive cognitive patterns by investigating the emotional, cognitive, and perceptual dynamics of lived experience (or self-inquiry) and to cultivate insights into the nature of all phenomena. For instance, a standard self-inquiry/ insight consists of dissecting the experience of pain into its component parts while noticing how the associated thoughts,

feelings, and physical sensations are constantly changing (Zorn et al., 2020). Empirical evidence in support of the three types of meditation practice comes notably from neuroimaging data showing that meditation practices related to each family induced distinct spatial neuroplasticity in brain areas associated with the core cognitive processes trained by these practices (Valk et al., 2017).

According to this recent taxonomy proposed by Dahl et al. (2015), the use of global scales such as the MAAS (Brown & Ryan, 2003) might not capture specific processes associated with different meditation practices. To go beyond this limitation, we recently defined a three-dimensional dispositional meditation trait (Schlosser et al., 2022). Based on a series of psychometric questionnaires, we computed three dispositional meditation scores corresponding to the attentional, constructive and deconstructive types of practice described in Dahl et al. (2015). These dispositional meditation traits, when assessed in participants naïve to meditation, may reflect their cognitive predisposition to engage with each of these three families of meditation. Determining the influence of such dispositional meditation traits on cognition in aging is particularly relevant in the context of models proposing that multiple psychological factors influence the rate of cognitive changes (Reuter-Lorenz & Park, 2014).

In that context, attentional practices could be associated with a better cognition regarding a recent meta-analysis reporting a positive impact of mindfulness-based interventions on attentional and executive functions in aging (Whitfield et al., 2021). Moreover, we recently argued that constructive and deconstructive practices can indirectly impact cognitive reserve by enhancing positive emotions and by down-regulating detrimental age-related factors, such as stress, or depression (Lutz et al., 2021). For instance, constructive practices such as loving-kindness and compassion were shown to alleviate social stress with a different mechanism than attentional-based meditation (Engert et al., 2017) and deconstructive practice seemed to reduce

depressive episodes by learning to cognitively decenter from ruminative thoughts (Segal et al., 2019). These results are particularly interesting as previous studies have shown that anxiety, depression and neuroticism levels are associated with cognitive performance in aging (Bunce et al., 2012; Luchetti et al., 2016)

We capitalized on cross-sectional data from the Age-Well trial in the “Silver Santé Study” (<https://silversantestudy.eu/>) to test whether dispositional meditation traits are associated with better cognitive functioning in older healthy participants naïve to formal meditation practice along different cognitive domains (memory, attention, executive function) as well as on a global composite considered as a relevant measure to assess risk of dementia in aging (Preclinical Alzheimer’s Cognitive Composite [PACC5]; Lu et al., 2019; Mormino et al., 2017). We predicted that the attentional practice would be most associated with cognitive performance. We further expected that constructive and deconstructive practices could show similar, but weaker, associations with cognitive performance. In a second exploratory analysis, we included measures of affective state and personality using self-report scales as mediators to observe if they could have an influence on the relationship between meditation scores and cognitive performance.

Method

Participants

According to the power analysis on G*Power (Faul et al., 2007), with a power level of 80% and an alpha at $p=.05$ on a linear regression including three predictors, a medium effect size ($f=0.25$) is expected with 48 participants, and a small effect size ($f=0.15$) with 77 participants. The final sample consisted of 134 cognitively unimpaired older adults (age range: 65 to 84 years old; one participant was excluded for missing data, flow diagram is in Figure 1)

from baseline visit of the Age-Well randomized clinical trial of the European Union's Horizon 2020-funded Medit-Ageing European project (public name: “Silver Santé Study”; Poisnel et al., 2018). The study has been registered on Clinicaltrials.gov (NCT02977819). A lack of previous experience with meditation training was a requirement for enrollment in the study. Participants were community-dwelling native French speakers (61.19% women), retired for at least one year and with at least seven years of education (demographic information is presented in Table 1). The Montgomery and Asberg Depression Rating Scale (MADRS; Montgomery & Asberg, 1979) was used as an eligibility criterion, and participants who obtained a score above 7 were excluded. None of the participants reported recent history of neurological or psychiatric disease (including alcohol or drug abuse) or taking medication likely to affect the central nervous system. Finally, all had normal or corrected-to-normal vision and hearing, and obtained a score superior to 27 on the Mini Mental State Examination (Folstein et al., 2001). All participants gave their written informed consent to enter the study.

< Table 1 and Figure 1 about here >

Procedure

This study was a monocentric (location at Caen; France) randomized, controlled clinical trial. Participants were included in one of the three groups: meditation training, non-native language training and no-intervention. After a prescreening interview, eligible participants were enrolled for a baseline pre-intervention visit. The mid and last intervention visits were performed 9 and 18 months after the start of the intervention, respectively. Behavioral, biological and neuroimaging data were collected at the three visits. Clinicians were blinded to group intervention and all the data were pseudonymized. Data analyzed here are scores obtained on cognitive tests and on meditation questionnaires at the baseline visit. The cognitive tasks

were administered by neuropsychologists and the questionnaires were filled in by participants themselves.

Measures

Cognitive assessment. The cognitive assessment consisted of a comprehensive battery of the French versions of neuropsychological tests assessing short-term and episodic memory, attentional and executive function. Composite scores were created and used for statistical analyses because, compared to individual tests, they possess higher sensitivity, might be less susceptible to floor and ceiling effects and reduce statistical multiplicity. Three composite scores corresponding to the memory, executive function and attentional domains were created based on previous categorizations made in the literature (Bangen et al., 2010; Ossenkoppele et al., 2012). We also created a global score (the Preclinical Alzheimer's Cognitive Composite 5 score; PACC5 score) sensitive to subtle cognitive changes linked to dementia risk (Donohue et al., 2014; Papp et al., 2017). The scores were computed with the same method as previously reported (e.g., Papp et al., 2017).

All cognitive composite scores were computed by averaging scores of each neuropsychological test (the tests were previously standardized) using the mean and standard deviation of the entire group (Z-scores). Thereafter, the composite scores were re-standardized (divided by the standard deviation of the composite). The episodic memory score (Cronbach's $\alpha = .82$ and McDonald's $\omega = .84$) was computed by averaging the standardized scores of the California Verbal Learning Test (CVLT-II; Poitrenaud et al., 2007; three raw scores: (a) sum of five free recalls from learning trials, (b) short-term free recall, (c) long-term free recall) and Logical Memory (Wechsler, 2008; two raw scores: (a) immediate recall (b) delayed recall). The executive function score (Cronbach's $\alpha = .67$ and McDonald's $\omega = .67$) was computed by averaging the standardized scores of Digit Span backward (Wechsler, 2008; raw score), Trail

Making Test part B (Godefroy, 2008; response time), Stroop interference index (Godefroy, 2008; response time for interfering-neutral items), and Letter Fluency (Godefroy, 2008; raw score). The attention composite score (Cronbach's $\alpha = .58$ and McDonald's $\omega = .61$) was computed by averaging the standardized scores of the Trail Making Test part A (Godefroy, 2008; response time), Stroop naming (Godefroy, 2008; response time), Digit Span forward (Wechsler, 2008; raw score), and Coding (Wechsler, 2008; raw score). Trail Making Test and Stroop scores were reversed so that higher individual scores and therefore higher total compound score indicate better performance. The PACC5 score (Cronbach's $\alpha = .60$ and McDonald's $\omega = .61$) (Donohue et al., 2014; Papp et al., 2017) was computed as the average of z-scores of the following cognitive measures: long-term free recall of the CVLT-II (raw score), delayed recall of the Logical Memory Test (raw score), total score of the Coding (raw score), Verbal Fluency raw score for the animal category and total raw score of the Mattis Dementia Rating Scale (GRECO, 1994). Raw data are presented in Supplementary Table S1.

Dispositional meditation assessment. A series of psychological self-report scales aimed to assess behavioral aspects and cognitive processes associated with meditation practice was administered to the participants: Multidimensional Assessment of Interoceptive Awareness (MAIA; Mehling et al., 2012), 15-item Five Facet Mindfulness Questionnaire (FFMQ-15; Baer et al., 2012), Drexel Defusion Scale (Forman et al., 2012), Interpersonal Reactivity Index (IRI; Davis, 1983), Compassionate Love Scale (Raes et al., 2011) and Prosocialness (Caprara et al., 2005).

The development of the meditation composite scores was based on a theoretical framework categorizing meditation practices into attentional, constructive, and deconstructive types based on the cognitive mechanisms these practices primarily target or necessitate (Dahl et al., 2015). A high value on a composite score is postulated to be associated with a frequent

recruitment of the psychological processes associated with that meditation type. A detailed description of the composite development process has previously been published (Schlosser et al., 2022). In short, three of the authors with scientific expertise in the domain (AL, MS, OK) independently assigned items from the questionnaires to the attentional, constructive or deconstructive families based on the cognitive mechanisms they were judged to primarily capture. Inadequate scales or subscales without clear links to the three families were excluded and any disagreement was resolved via group discussion. The meditation composites displayed adequate psychometric properties in three independent samples: meditation-naïve healthy older adults from the Age-Well trial ($n = 135$; Poisnel et al., 2018), meditation-naïve older adults with subjective cognitive decline from the SCD-Well trial ($n = 147$; Marchant et al., 2018), and healthy long-term meditators ($\geq 10,000$ h of practice including one 3-year meditation retreat) from the Brain & Mindfulness project ($n = 29$; Abdoun et al., 2018). Confirmatory factor analyses on the Age-Well sample corroborated the theoretically predicted distinction between the three meditation families.

The attentional capacity of meditation (corresponding to metacognition / regulation of attention processes) was primarily captured by five items from the MAIA (Noticing, Attention Regulation, Emotional Awareness, Self-Regulation, Body Listening) and two items from the FFMQ (Observing, Acting with Awareness) (Cronbach's $\alpha = .72$ and McDonald's $\omega = .75$). The constructive capacity (corresponding to positive attitudes toward oneself and others) was primarily captured by two items from the IRI (Perspective Taking and Empathic Concern) and scores on the Compassionate Love Scale and Prosocialness scale (Cronbach's $\alpha = .46$ and McDonald's $\omega = .49$). The deconstructive capacity (namely self-inquiry and cognitive defusion) was judged to be primarily captured by Personal Distress item from the IRI, two items from the FFMQ items (Non-Judging, Non-Reactivity), and the total score on the Drexel Defusion Scale (Cronbach's $\alpha = .36$ and McDonald's $\omega = .39$).

Each composite score was computed following the same procedure as was used to create the cognitive composite scores. Raw data are presented in Supplementary Table S2.

Affective and personality assessment. We also assessed affective state and personality using self-report scales, as previous studies have shown that anxiety, depression and neuroticism levels are associated with cognitive performance in older adults (Bunce et al., 2012; Luchetti et al., 2016). Depressive symptoms and trait anxiety were assessed with the Geriatric Depression Scale (Yasavage & Sheikh, 1986) and the part B of the State-Trait Inventory form (Spielberger et al., 1970). Neuroticism was assessed by the Big Five Inventory (Plaisant et al., 2010). Raw data are presented in Supplementary Table S3.

Data Analyses

We used linear regression analyses (PROC GLM) to model the relationship between meditation capacities and each cognitive score. Data were normally distributed, and collinearity was assessed using Tolerance ($TOL > .10$) and Variance Inflation Factors ($VIF < 10$). We employed separate models for each cognitive score (attention, executive function, episodic memory, and PACC5). In these models, the predictors were the three capacities of meditation, and the dependent variable was each cognitive score. We added the semi-partial R-square (R_{sp}^2) when a significant effect was found. Indeed, semi-partial R-square is useful to explain an effect of one predictor specified in a model with many predictors. No correction for multiple comparisons were applied as this was considered an exploratory study because there is scant data in this research field.

Next, we performed an exploratory mediation analysis (Structural Equation Model [SEM] using the Lavaan package in R) to determine the respective contribution of the predictors identified in the regression analyses to cognitive performance. We included in the model

affective/personality and demographic variables that were correlated with the meditation scores (PROC CORR).

Statistical analyses were performed with SAS 9.4 for Windows (SAS Institute Inc., 2013, <https://www.sas.com>). The exploratory analysis and the graphics were computed with R version 3.6.3 (R Core Team, 2020). The statistical threshold was set at $p < .05$.

Results

Associations between Demographic Variables, Affective State and Personality Trait and Meditation Capacities

We assessed the presence of significant correlations between demographic variables (age and education), affective state, neuroticism level, and dispositional meditation traits associated with the attentional, constructive, and deconstructive families (Table 2). No significant associations were observed between age and meditation scores. Regarding the association with affective state and personality (depression, anxiety, and neuroticism), we found a negative association between anxiety and some meditation capacities (attentional and deconstructive families), while depression and neuroticism were negatively associated with the deconstructive capacity only. Correlations between affective measures are presented in supplemental material.

< Table 2 about here >

Relationships between Cognition and the Three Meditative Capacities

No associations with attentional, constructive, and deconstructive meditation capacity were found for the episodic memory, executive function, and attentional cognitive domains (all p 's $> .05$, Table 3). However, a positive association between global cognition (PACC5) and

deconstructive capacity was observed ($R_{sp^2}=.03$; $p=.04$), but no association with attentional and constructive capacity (Table 3, Supplemental Figure S4).

< Table 3 about here >

Affective State and Personality Traits as mediators of the Relationship between Global Cognition and Meditative Capacities

We performed exploratory analyses to determine if the effect of deconstructive meditation we observed on global cognition (PACC5) was mediated by affective state and personality. We included in the model the depression, anxiety and neuroticism scores, and the three meditation capacities. The model (Figure 2, Table 4) had a good fit (CFI=0.98, TLI=0.94, RMSEA=0.07, SRMR=0.05). Positive associations were observed between all affective measures. There was also a significant positive association between deconstructive and attentional meditation capacities. Moreover, the deconstructive capacity had a large negative effect on depression, anxiety and neuroticism values. As expected from our previous regression analyses, we observed a significant positive effect of deconstructive capacity on global cognition (direct effect, in purple in the Figure 2). However, there was no direct effect of attentional capacity, constructive capacity, anxiety or depression on global cognition score, while a significant positive effect of neuroticism was observed. Among the three indirect effects (the effect of deconstructive mediated by the anxiety, depression and neuroticism, respectively in blue, green and red in the Figure 2), only the indirect effect of neuroticism was significant. To summarize, the model is appropriate to use and the mediation analysis showed that neuroticism and deconstructive scores have an impact on global cognitive performance (assessed with PACC5), with neuroticism partly mediating the effect of deconstructive practice.

< Figure 2 about here >

< Table 4 about here >

Discussion

Our study aimed to investigate the relationship between trait meditation reflecting mechanisms involved in attentional, constructive and deconstructive practices (Schlosser et al., 2022) and cognition in aging. In partial agreement with our hypotheses, we observed a positive association between deconstructive capacity and PACC5 score, assessing risk of dementia in aging. In the exploratory mediation analysis, when affective and personality variables were simultaneously considered, our model showed that neuroticism influenced PACC5, and partly mediated the effect of the deconstructive capacity. No associations between any of the meditation capacities and cognitive scores of attention, episodic memory and executive function were observed.

These results suggest that trait meditation capacities, and more particularly deconstructive mechanisms, are associated with cognitive performance in healthy aging. This relationship was observed with the PACC5, a measure sensitive to early amyloid-beta-related cognitive decline (Donohue et al., 2014; Papp et al., 2017), rather than the cognitive composites measuring memory, attention or executive function. The PACC5 is associated with age-related changes and possible dementia prodromes, such as amyloid deposition (Donohue et al., 2014), tau accumulation (Hanseeuw et al., 2019), and lower functional connectivity (Buckley et al., 2017). It is worth mentioning that most studies do not find cross-sectional differences in PACC5 between at-risk/lower-risk groups (Donohue et al., 2014), so we might also expect to observe no associations. However, our finding indicates that deconstructive meditation capacity may be an important protective factor. In line with our results, Strikwerda-Brown et al. (2022) showed that trait mindfulness is associated with less decline on a global measure of cognition and less tau and amyloid-beta brain deposits. They further reported that the nonjudging and non-

reactivity components of their mindfulness measure (the FFMQ) were significant predictors of global cognitive decline. These findings strongly align with our own, considering that nonjudging and non-reactivity formed part of our deconstructive composite.

The deconstructive family of meditation practice implicates meta-cognitive processes and self-reflection (Dahl et al., 2015; Schlosser et al., 2022). Deconstructive practices involve inquiries into the nature of thought, perception, and other cognitive and affective processes. Interestingly, self-reflection has been associated with better global cognition in older age (Demnitz-King et al., 2022). Furthermore, individuals with greater deconstructive capacity may be more able to distance themselves from negative thoughts and emotions. Accordingly, several studies have shown a positive effect of meditation practice on rumination (Wolkin, 2015). In a study that included 1281 meditators with 2 to 13 years of practice, a greater meditation experience was associated with less repetitive negative thinking (Schlosser et al., 2020); and separately, higher levels of repetitive negative thinking have been associated with cognitive decline in older age (Marchant et al., 2020). Together these findings suggest that deconstructive meditation capacity may be associated with better cognition via engagement in different thinking styles (higher self-reflection, lower rumination), however this remains speculative because no study has yet examined these specific associations. More generally, Wolkin (2015) proposed that the positive effect of meditation on rumination is a two-step process entailing first distraction and secondly decentering. Decentering is enabled through enhanced metacognition to allow a thought to become a mental event rather than an intrinsic part of the self (Wallace & Shapiro, 2006).

We did not find a mediating effect of anxiety or depression on the relationship between deconstructive trait capacities and performance on the PACC5. We posit that this absence of effect may be due to our participant population. To be eligible for the study, participants could

not have significant anxious or depressive symptoms, which could lead to reduced variance on these affective measures. However, the relationship we observe between deconstructive capacity and better performance on PACC5 is mediated by neuroticism level. Namely, lower deconstructive capacity was associated with greater neuroticism, and greater neuroticism with better performance. Interestingly, our data hint that reinforcing a disposition to deconstructive meditation practice could have a positive influence on a personality trait associated with increased dementia risk. However, the observation that higher neuroticism was associated with better cognitive performance contradicts previous data showing that neuroticism is a risk factor for dementia (Neupert et al., 2008; Terracciano et al., 2017). One possible explanation would be that a high cognitive functioning can be linked to perfectionism and therefore negative affect or neuroticism (Chang, 2000). Further, self-reflection has been positively associated with neuroticism (Creed & Funder, 1998; Trapnell & Campbell, 1999), yet it has also been associated with better global cognition and brain function (Demnitz-King et al., 2022). Therefore, there may be a complex interplay between certain correlates of neuroticism (e.g., self-reflection) and cognition.

However, our main prediction was that high scores on measures associated with the attentional family would be related to better cognition, but the association was not observed with our sample. This prediction was based on previous studies that showed that meditation programs centered on attentional processes such as Mindfulness-Based Stress-Reduction led to increases in some aspects of cognition, even after a short training. Indeed, a systematic review and meta-analysis showed a greater improvement in executive functioning and working memory performance in mindfulness-based practice groups, compared to control groups, while no difference was found concerning attention or episodic memory outcomes (Whitfield et al., 2021). Additionally, expert meditators have been shown to recruit attentional processes very efficiently (Brefczynski-Lewis et al., 2007). Here, we explored the correlates of meditation

traits in older adults with no experience of formal meditation practices. It thus seems plausible that the dynamics between these informal meditation capacities and cognition differ from those observed after explicit and formal practice. Moreover, we used a compound attentional score while in previous studies, such associations were observed between very specific attentional measures and a global measure of meditation (the MAAS scale) (Fiocco & Mallya, 2015; Fountain-Zaragoza et al., 2016; Prakash et al., 2015). Finally, whilst we did not observe an association between the attention meditation composite and cognition in this study, it is possible that a relationship would be observed after meditation training and in meditators. We argue this on the basis that, in the context of meditation practice, psychological processes reflecting different composites could strengthen one another (Dahl et al., 2015; Lutz et al., 2021). These evolving inter-relationships might then result in a different correlation structure with the meditation composites and cognition (Schlosser et al., 2022).

Limitations and Future Research

A limitation of this study is that it is cross-sectional and therefore causality cannot be determined. We report possible cognitive dispositions for psychological processes important in meditation, and not meditative skills *per se*. Consequently, it remains to be determined to what extent explicit training in meditation practice is linked to changes in dimensional dispositional meditation traits (Schlosser et al., 2022) and cognitive performance. However, the factors that were identified as contributing to the cognitive reserve capacity (and are consequently protective factors against dementia; Dekhtyar et al., 2015) exert their influence over long time periods (e.g., education or occupation), and it is reasonable to expect that a long-term engagement in meditation (that is by formal or informal practice) will be required to yield lasting effects. Furthermore, the responses given by the participants on questionnaires of meditation could be biased by participant's or examiner's characteristics. Indeed, Podsakoff

et al. (2012) reported different bias linked to subjective measures such as social desirability, response styles and priming effect. Another potential limitation is the absence of correction for multiple comparisons in the regression models; our approach was to transparently report the number of analyses performed particularly given that this was an exploratory study. Future replication of these results is thus required.

In conclusion, the current study supports the idea that the informal development of trait meditation capacities related to the deconstructive family is associated with a higher level of global cognitive functioning during late life. Moreover, it shows that the benefit of such an attitude could partly depend on personality traits (e.g., neuroticism). The interest of meditation practice as a non-pharmacological intervention to promote healthy cognitive aging is developing. Future studies should determine if the positive effect of an informal meditation attitude on cognition in aging observed here is magnified through intervention implementing formal meditation practice. If a positive effect is confirmed, it should be determined which intervention programs (possibly combining different families of practice appear the most appropriate).

Declarations

Conflicts of Interest

GC has received research support from the EU's Horizon 2020 research and innovation programme (grant agreement number 667696), Fondation d'entreprise MMA des Entrepreneurs du Futur, Fondation Alzheimer, Programme Hospitalier de Recherche Clinique, Région Normandie, Association France Alzheimer et maladies apparentées and Fondation Vaincre Alzheimer (all to Inserm), and personal fees from Inserm, Fondation Alzheimer and Fondation d'entreprise MMA des Entrepreneurs du Futur.

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Ethics Statement

After a description of the study’s procedures to the participants and before any data acquisition, written informed consent were signed by participants. The Age-Well trial received ethics approval from the Comité de Protection des Personnes CPP Nord-Ouest III in Caen (trial registration number: EudraCT 2016-002441-36).

Author’s Contributions

Florence Requier: wrote the first draft of the manuscript; statistical analysis; interpretation of data; and incorporation of manuscript feedback. Anne Sophia Hendy: statistical analysis; interpretation of data. Marco Schlosser: substantial contributions to the analysis of data; and revision of the manuscript for important intellectual content. Harriet Demnitz-King: substantial contributions to the analysis of data; and revision of the manuscript for important intellectual content. Tim Whitfield: substantial contributions to the analysis of data; and revision of the manuscript for important intellectual content. Gael Chételat: substantial contributions to the conception of the work; and revision of the manuscript for important intellectual content. Olga Klimecki: substantial contributions to the design of the work; and revision of the manuscript for important intellectual content. Antoine Lutz: substantial contributions to the conception and

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Data Availability

Data used in the Medit-Ageing project are available upon request.

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Table 1. Descriptive data of the demographics of 134 healthy old participants

Demographic information	Mean (<i>n</i>=134)
Age (<i>years, mean ± SD</i>)	69.3 ± 3.8
Education (<i>number of years, mean ± SD</i>)	13.1 ± 3.1
Gender (<i>women, %</i>)	61.2
MMSE (<i>raw score, mean ± SD</i>)	29.0 ± 1.0

Note. SD: Standard Deviation MMSE: Mini Mental State Examination.

Table 2. Correlations between demographic information, affective state, personality and meditation capacities

	<i>Education (years)</i>	<i>Age (years)</i>	<i>STAI-B</i>	<i>GDS</i>	<i>Neuroticism</i>
<i>Meditation capacities</i>					
<i>Attentional</i>	<i>r = .12</i> <i>p = .16</i>	<i>r = -.04</i> <i>p = .61</i>	<i>r = -.17</i> <i>p < .05</i>	<i>r = .004</i> <i>p = .96</i>	<i>r = -.07</i> <i>p = .40</i>
<i>Deconstructive</i>	<i>r = .14</i> <i>p = .10</i>	<i>r = .10</i> <i>p = .26</i>	<i>r = -.61</i> <i>p < .001</i>	<i>r = -.18</i> <i>p < .05</i>	<i>r = -.69</i> <i>p < .001</i>
<i>Constructive</i>	<i>r = .05</i> <i>p = .58</i>	<i>r = .10</i> <i>p = .26</i>	<i>r = .11</i> <i>p = .20</i>	<i>r = -.02</i> <i>p = .79</i>	<i>r = .13</i> <i>p = .14</i>

Note. *n*=134, GDS: Geriatric Depression Scale, STAI: State-Trait Anxiety Inventory.

Table 3. Statistical outcomes of the regression analyses seeking for the association between cognitive performance and meditation capacities.

<i>Capacities</i>	Estimate ± SE	F value (df)	P
(A) Cognitive composite score of attention			
<i>Attentional</i>	-0.02 ± 0.09	0.04 (1,130)	.85
<i>Deconstructive</i>	0.07 ± 0.09	0.56 (1,130)	.45
<i>Constructive</i>	-0.07 ± 0.09	0.57 (1,130)	.45
(B) Cognitive composite score of executive function			
<i>Attentional</i>	0.07 ± 0.09	0.54 (1,130)	.46
<i>Deconstructive</i>	0.12 ± 0.09	1.97 (1,130)	.16
<i>Constructive</i>	0.04 ± 0.09	0.27 (1,130)	.61
(C) Cognitive composite score of episodic memory			
<i>Attentional</i>	-0.08 ± 0.09	0.77 (1,130)	.38
<i>Deconstructive</i>	0.12 ± 0.09	1.98 (1,130)	.16
<i>Constructive</i>	0.15 ± 0.09	2.98 (1,130)	.09
(D) PACC5 score			
<i>Attentional</i>	-0.12 ± 0.09	1.81 (1,130)	.18
<i>Deconstructive</i>	0.18 ± 0.09	4.34 (1,130)	.04 (R_{sp}²=.03)
<i>Constructive</i>	0.11 ± 0.09	1.61 (1,130)	.21

Note. SE: Standard Error, df: degrees of freedom, R_{sp}²: semi-partial R², n=134.

Table 4. Mediation analysis: meditative capacities and affective data as mediators on PACC5 performance.

Parameters	Estimate ± SE	Z	P	CI
Attentional ~ Deconstructive	0.20 ± 0.09	2.30	.02	[0.03:0.37]
PACC5 ~ Attentional	-0.14 ± 0.08	-1.70	.09	[-0.31:0.02]
PACC5 ~ Constructive	0.10 ± 0.08	1.16	.25	[-0.07:0.26]
PACC5 ~ Deconstructive	0.30 ± 0.12	2.49	.01	[0.06:0.53]
GDS ~ STAI-B	3.62 ± 0.87	4.16	<.001	[1.91:5.33]
STAI-B ~ Neuroticism	1.52 ± 0.29	5.28	<.001	[0.96:2.08]
Neuroticism ~ GDS	0.34 ± 0.08	4.07	<.001	[0.18:0.51]
GDS ~ Deconstructive	-0.32 ± 0.15	-2.14	.03	[-0.60:-0.03]
PACC5 ~ GDS	0.01 ± 0.05	0.11	.91	[-0.10:0.11]
STAI-B ~ Deconstructive	-4.27 ± 0.48	-8.89	<.001	[-5.21:-3.33]
PACC5 ~ STAI-B	-0.02 ± 0.02	-1.10	.27	[-0.05:0.02]
Neuroticism ~ Deconstructive	-0.51 ± 0.05	-10.93	<.001	[-0.60:-0.42]
PACC5 ~ Neuroticism	0.38 ± 0.18	2.09	.04	[0.02:0.74]
Indirect effect with STAI-B	0.08 ± 0.08	1.09	.28	[-0.07:0.23]
Total effect with STAI-B	0.10 ± 0.11	0.92	.36	[-0.12:0.33]
Indirect effect with GDS	-0.002 ± 0.02	-0.11	.91	[-0.04:0.03]
Total effect with GDS	0.30 ± 0.12	2.39	.02	[0.05:0.54]
Indirect effect with neuroticism	-0.19 ± 0.09	-2.05	.04	[-0.38:-0.01]
Total effect with neuroticism	0.10 ± 0.11	0.92	.36	[-0.12:0.33]

Note. SE: Standard Error, STAI: State Trait Anxiety Inventory, CI: Confidence Interval, $n=134$,

~~ association/correlation, ~ regression, indirect effect: mediating effect, total effect: direct

and indirect effect of one mediation model.

Figure 1. Flow diagram of the inclusion process

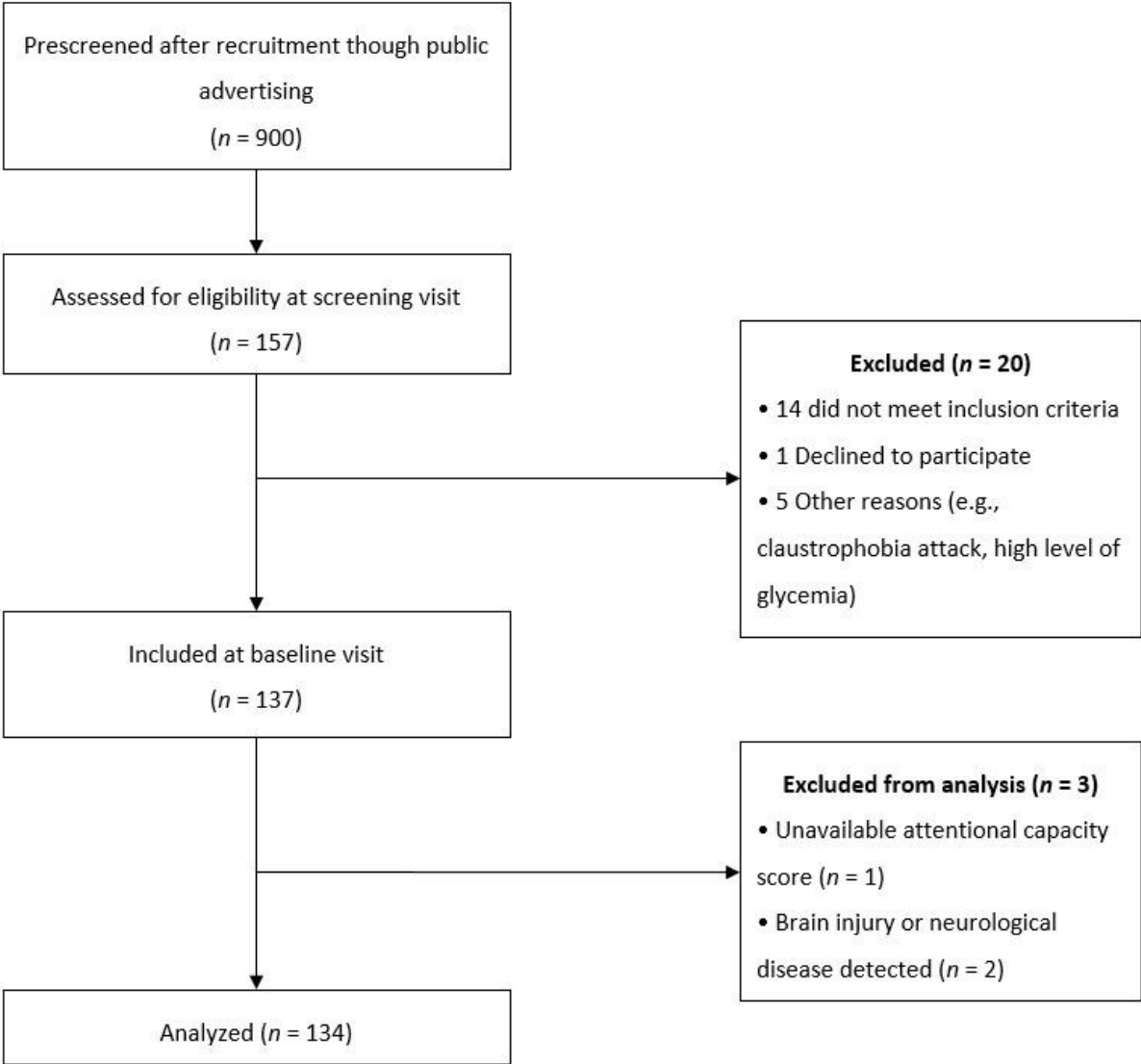
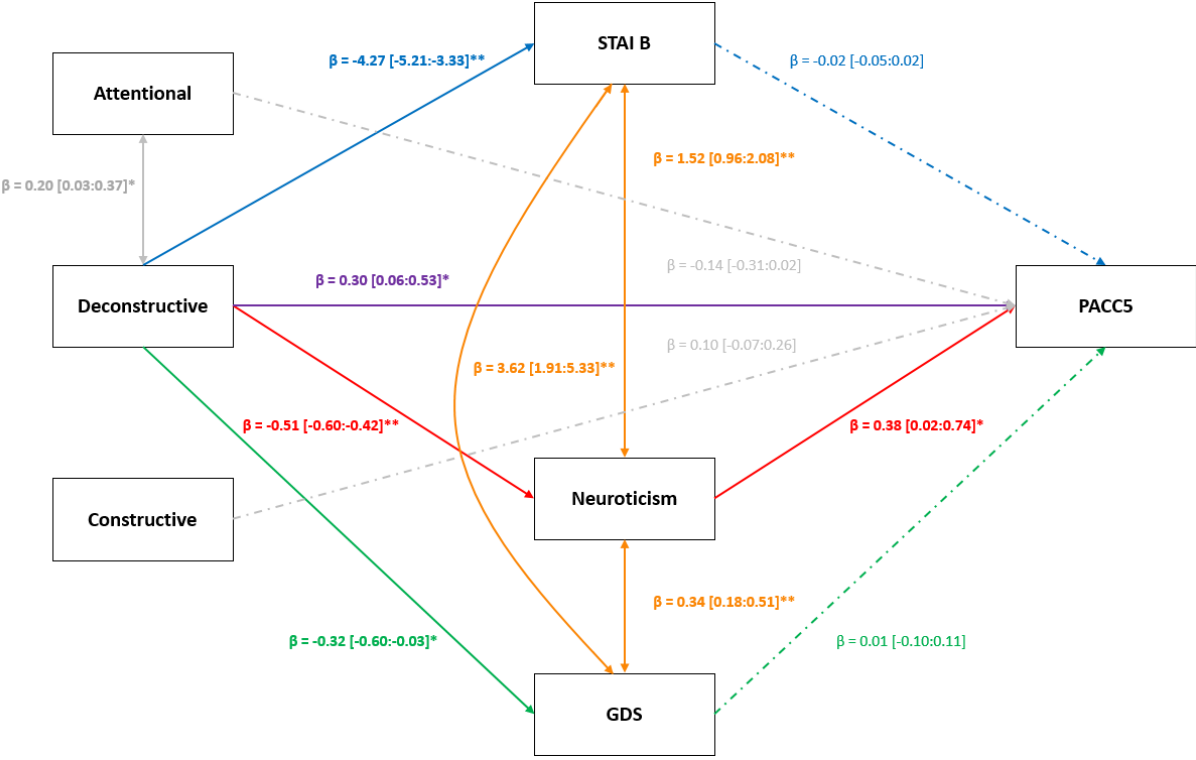


Figure 2. Mediation analysis: relationship between deconstructive capacity and global cognition (PACC5) with STAI-B, GDS and neuroticism as mediators



Note. * $p < .05$, ** $p < .001$. Bidirectional arrows represent associations between variables, and unidirectional arrows the direct effects. Statistical values for indirect and total effects are presented in Table 4.