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### Publication bias casts doubt on implicit processing in inattentional blindness

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#### Abstract

Two recent meta-analyses on inattentional blindness (Kreitz, Pugnaghi, & Memmert, 2020; Nobre et al., 2020) concluded that objects can be processed implicitly even when attention is directed elsewhere. However, signs of publication bias are evident in both of these metaanalyses. Here, we employed multiple tools to correct for publication bias in the data aggregated

in those meta-analyses. Analyses using the Precision-Effect Test (PET) and robust Bayesian meta-analysis (RoBMA) suggest that the estimates in the original meta-analyses were inflated, together with strong evidence of publication bias. Indeed, the data are consistent with no overall implicit effects. We suggest that more evidence, particularly from well-powered pre-registered experiments, is needed before solid conclusions can be drawn regarding implicit processing during inattentional blindness.

### Keyword

meta-analysis; Inattentional blindness; implicit processing; publication bias.

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Research on inattentional blindness has established the remarkable finding that salient objects in plain view can be unseen if attention is engaged elsewhere. Two recent meta-analyses (Kreitz, Pugnaghi, & Memmert, 2020; Nobre et al., 2020) reviewed evidence that such unseen objects are nonetheless sometimes processed implicitly – for instance, affecting response times or accuracy in a separate task. Across k = 59 experiments, Nobre et al. found a medium-sized average meta-analytic implicit effect of r = .33, 95 % confidence interval (CI) [.21, .45], while Kreitz et al. (who analyzed data from their own research group) found a smaller effect, k = 16, Cohen's d = 0.211 [0.106, 0.316]. Nobre et al.'s (2020) dataset included a subset (k = 14 studies employing visual stimuli) of those aggregated by Kreitz et al. (2020).

At face value these meta-analyses, combining data across 3,464 participants, conclusively demonstrate that unseen objects are processed implicitly. We argue here, however, that there are signs of severe publication bias in this field and that when cutting-edge techniques are employed

to correct the above estimates for publication bias, the true average effect size is approximately zero.

While Kreitz et al. (2020) did not discuss publication bias in their meta-analysis, Nobre et al. (2020) examined a funnel-plot for the effect sizes (see Figure 1) and duly noted the presence of publication bias, which they corrected for using the Trim-and-Fill method (Duval & Tweedie, 2000). After this correction, they still obtained a significant, albeit smaller, overall effect (r = .18), indicating that implicit processing occurred during inattentional blindness. However, recent simulation studies (Carter et al., 2019) suggest that, in the prevailing conditions of the meta-analysis, the Trim-and-Fill method has a very high false positive rate (the rate at which a method falsely concludes that there is a non-zero effect when the true effect is zero). Therefore, we conducted alternative analyses (using methods that were not available when the original meta-analyses were conducted) to obtain a better estimate of the overall effect size after correcting for publication bias.

## Assessing publication bias in the published meta-analyses

Using Carter et al.'s (2019) MetaExplorer application (http://shinyapps.org/apps/metaExplorer/), we estimated false positive rates to assess the performance of five different methods: standard random-effects meta-analysis, Trim-and-Fill, the Precision-Effect (PET) and Precision-Effect Estimate with Standard Errors (PEESE) tests, and a 3-parameter selection model (3PSM; see Carter et al. for further details about these methods). We adopted a maximum acceptable false positive rate of 20% as the threshold to decide whether each method was acceptable or not.

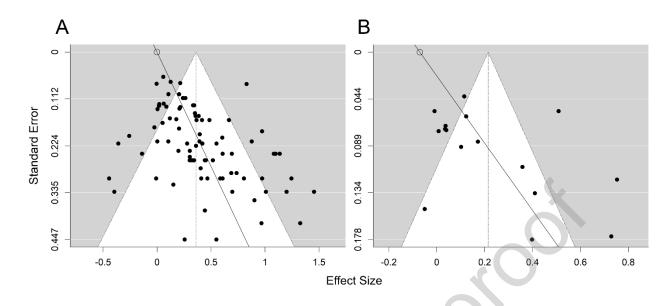


Figure 1. Funnel plots for the datasets analyzed. A: Funnel plot for the data of Nobre et al. (2020), in correlation values (r). B: Funnel plot for the data of Kreitz et al. (2020), in Cohen's d values. Open circles are PET estimates.

The parameters used in the simulations were selected as follows. For Nobre et al.'s (2020) dataset, we assumed medium severity of publication bias and a medium level of questionable research practices (QRPs) environment (see Carter et al., 2019, for definitions of these parameters). We examined results for two levels of heterogeneity: medium ( $\tau = 0.2$ ) and high ( $\tau = 0.4$ ). Under these conditions the Trim-and-Fill method employed by Nobre et al. (2020) has a false positive rate over 80% and is clearly inadequate as a correction method. However, whereas the simulation showed lower false positive rates for both PET and 3PSM (see Tables S1 and S2 in the Supplementary Information), for PEESE we obtained an unacceptably high false positive rate (72% for medium heterogeneity, 53% for high heterogeneity), so we removed PEESE from further analyses.

Table 1 shows the estimates and confidence intervals for the analysis reported by Nobre et al. (2020) as well as PET and 3PSM, using correlation coefficients as effect sizes. The table also includes the Trim-and-Fill estimate, even though the above analysis suggests that it is inappropriate in this context. The estimate is included to confirm that we reproduce the same value as Nobre et al. (2020). PET showed an influence of standard error on effect size (p < .001). The bias-corrected estimate computed by PET is close to 0 (Table 1) and is not significant in the model (p = .99). Thus, when a bias-correcting method is employed that is less prone to false positives, the overall effect disappears. On the other hand, 3PSM yielded an estimate close to the one obtained using a random-effects model. A likelihood ratio test comparing the fit of the adjusted model to that of the random-effects model was not significant,  $\chi^2$  (1) = 0.09, p = .76, indicating that 3PSM does not improve the fit compared to the original model, revealing no publication bias.

Tat	ble 1. Estimates, confidence intervals and outcome of publication bias assessment by
method.	

Bias correction method	Estimate	95% CI	Publication bias	
Nobre et al. (2020)				
RE	.33	[.21, .45]	-	
Trim-and-Fill	.18	[.08, .28]	Yes	
PET	01	[17, .17]	Yes	
3PSM	.35	[.22, .46]	No	
RoBMA	.037	[09, .42]	Yes	

Kreitz et al. (2020)						
RE	.21	[.10, .33]	-			
Trim-and-Fill	.21	[.10, .33]	No			
PET	07	[34, .20]	Yes			
3PSM	.20	[.02, .38]	No			
RoBMA	.13	[20, .68]	Yes			

Note: RE = Random-Effects model; PET = Precision-Effect Test; 3PSM = 3-Parameter Selection Model; RoBMA = Robust Bayesian Meta-Analysis. Effect sizes estimates are correlation coefficients for Nobre et al. (2020) and Cohen's *d* values for Kreitz et al. (2020).

Lastly, we employed the recently developed robust Bayesian meta-analysis approach (RoBMA; Bartoš, Maier, Quintana, & Wagenmakers, 2022; Maier, Bartoš, & Wagenmakers, 2022). RoBMA uses Bayesian model-averaging to combine estimates from multiple models — including PET, PEESE and selection models — both with and without publication bias. Each method is fit to the data and then the estimated effect is computed by weighting each of them by its likelihood, given the data. This method computes Bayes factors to quantify the evidence for the presence or absence of an effect as well as of publication bias. RoBMA has been shown to be superior to other bias-correction methods via several simulation studies (Bartoš, Maier, Quintana, & Wagenmakers, 2022; Maier et al., 2022). We fit RoBMA assuming equal prior probabilities across model types. Bayes factors indicated moderate evidence for publication bias ( $BF_{10} = 3303.90$ ). These results suggest that the overall effect Nobre et al. (2020) reported might be distorted by publication bias in the inattentional blindness literature.

For Kreitz et al.'s (2020) dataset, we ran the Carter et al. (2019) simulation assuming high heterogeneity ( $\tau = 0.4$ ) and a medium level of QRPs. We simulated two levels of publication bias, medium and high. Again, Trim-and-Fill shows a high false positive rate (30%, above our cutoff point of 20%). PEESE also returned a high false positive rate, so we removed it from further analysis.

Table 1 displays the results of those analyses. Whereas RE, Trim-and-Fill, and 3PSM provide confidence intervals that do not include zero, PET shows a wide confidence interval consistent with an effect size of zero. For publication bias assessment, results were mixed: among frequentist models, only PET found publication bias. RoBMA also showed evidence for publication bias ( $BF_{10} = 10.11$ ), as well as moderate evidence against an effect ( $BF_{10} = 0.32$ ). Thus, the results again suggest it is possible that the true average effect is close to zero when corrected for publication bias.

Turning to heterogeneity, Nobre et al.'s (2020) meta-analysis revealed considerable heterogeneity among effect sizes, which was partly explained by the presence of moderators. For this reason, we conducted the same analyses within subgroups defined by those moderators (see Supplemental Information). Although these analyses found that the evidence for publication bias differs among subgroups, in none was the effect size significantly different from zero.

### Discussion

In sum, the analyses reported here provide considerable evidence that the results of both Nobre et al. (2020) and Kreitz et al. (2020) were distorted by publication bias. In both metaanalyses, correcting for this bias with PET and RoBMA yields effects sizes that are smaller than those originally reported, and not significantly greater than zero. Moreover, the high

heterogeneity in both datasets, as well as the characteristics of several of the studies included in those meta-analyses (small *N*s, use of multiple measures of implicit processing, selective reporting of results) make the results even more difficult to interpret.

It is important to emphasize that the bias-correction methods employed here only perform satisfactorily under certain assumptions. As Carter et al.'s (2019) simulations demonstrated, these methods may underperform in some settings, and we selected a subset of methods for our analysis — for example, including PET but not PEESE — based on how they are expected to perform according to the simulations. Nevertheless, evaluating the performance of these methods is a complex issue, and our results should be read as suggestive rather than conclusive.

We should note that two of us are authors of one the meta-analysis examined here (Nobre et al., 2020). Research on bias correction is rapidly developing, and important improvements in this field have occurred since the publication of our previous review. The difference between our conclusions here and those of Nobre et al. (2020) is due to the availability of cutting-edge bias-detection methods and growing recognition of the need to employ them in interpreting meta-analyses (e.g., Bartoš, Maier, Quintana, & Wagenmakers, 2022; Carter et al., 2019; Maier et al., 2022; Stanley et al., 2021).

Importantly, both Nobre et al. (2020) and Kreitz et al. (2020) included only a few preregistered experiments (10 out of 59 and 6 out of 16, respectively). In each review, a little over half of the pre-registered studies (6/10 for Nobre et al., 2020; 4/7 for Kreitz et al., 2020) provided results that support the hypothesis of implicit processing during inattentional blindness. Fitting a random-effects model to the subset of pre-registered experiments in Nobre et al. (2020) shows an overall effect size of r = .22 that is significantly different from zero (t = 2.46, p = .03). For preregistered experiments in Kreitz et al. (2020), an effect size of d = .32 was found, which was also

significant (z = 2.11, p = .03). However, the resulting sample of pre-registered experiments in each meta-analysis is small, so these results should be interpreted with care.

Implicit processing is a field where the lack of reproducible results and publication bias is widely found, as in the case of priming studies which could not be reproduced in the last decade (e.g., O'Donnell et al., 2018; Rohrer, Pashler, & Harris, 2015; Shanks, et al., 2015). Our results demonstrate how the application of bias-detection methods may substantially change the conclusions of a meta-analysis. This approach has implications for other areas within psychology, since more than half of all meta-analyses in psychology probably overestimate the evidence for the presence of effects (Bartoš, Maier, Shanks, et al., 2022).

Our results suggest that the existence of implicit processing during inattentional blindness is unproven and questionable. Thus, although further meta-analyses may provide powerful conclusions on this question, the field may benefit from experiments that conform to more recent methodological guidelines for reproducibility and transparency, such as multi-lab approaches. In particular, to reach more solid conclusions on effects of implicit processing during inattentional blindness, we emphasize a need for pre-registered experiments, to minimize the possibility of publication bias; and further research aimed at understanding the roots of the heterogeneity evident amongst these effects.

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