Mathematical performance in childhood and early adult outcomes after very preterm birth: an individual participant data meta-analysis

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AIM To investigate the strength of the independent associations of mathematics performance in children born very preterm (<32wks’ gestation or <1500g birthweight) with attending postsecondary education and their current employment status in young adulthood.

METHOD We harmonized data from six very preterm birth cohorts from five different countries and carried out one-stage individual participant data meta-analyses (n=954, 52% female) using mixed effects logistic regression models. Mathematics scores at 8 to 11 years of age were standardized using contemporary cohort-specific controls. Outcomes included any postsecondary education, and employment/education status in young adulthood. All models were adjusted for year of birth, gestational age, sex, maternal education, and IQ in childhood.

RESULTS Higher mathematics performance in childhood was independently associated with having attended any postsecondary education (odds ratio [OR] per SD increase in mathematics z-score: 1.36 [95% confidence interval (CI): 1.03, 1.79]) but not with current employment/education status (OR 1.14 per SD increase [95% CI: 0.87, 1.48]).

INTERPRETATION Among populations born very preterm, childhood mathematics performance is important for adult educational attainment, but not for employment status.

Preterm birth has life-course consequences, in particular for the smallest infants and those born very preterm. Globally, about 1% to 2% of all infants are born at less than 32 weeks’ gestation every year.1 Improved neonatal care has led to increased survival rates for these infants,2 but the high prevalence of cognitive and learning problems in surviving infants has not substantially changed over the last two decades.3,4 Children born very preterm (<32wks’ gestation or <1500g birthweight) have an increased risk for poorer academic attainment than their term-born peers, especially in mathematics.5–8 This risk persists into adolescence and may diminish life-long opportunities, contributing to the frequently observed lower rates of postsecondary education and higher rates of unemployment in preterm cohorts.9 Longitudinal outcome studies in adulthood consistently show that problems of cognition, neurodevelopment, and academic achievement continue over the life-course.11–15 In addition to single cohort studies and literature-based reviews, two studies have documented considerable temporal and cross-national consistency in cognitive and mathematics performance among children born very preterm.16,17 However, it is unclear whether difficulties in mathematics among children born very preterm are accounted for by low IQ or not.7,18

Among general population studies, mathematics skills in childhood have independent positive effects on adult economic success.19,20 However, individuals born very preterm often experience severe medical complications and subsequent neurocognitive difficulties.21–23 Accordingly, studies in children and adults born very preterm have documented a higher prevalence of deficits in general cognitive performance in conjunction with poorer mathematics attainment than in term-born controls.24 One prospective longitudinal study of individuals born very preterm to 26 years of age...
investigated the extent to which mathematics performance in childhood explained adult economic success, but found no direct nor indirect effects after controlling for IQ. However, limitations of many follow-up studies into adulthood include attrition and relatively small samples, reducing statistical power.

Over and above IQ and mathematics performance, maternal education/family socioeconomic status is associated with long-term educational attainment, employment success, and wealth in adulthood. In addition, general population studies consistently report that females outperform males in school. For example, the life-course Panel Study of Income Dynamics found that females, on average, completed more years of schooling while males, on average, had higher annual incomes. Evidence for sex differences in life-course trajectories among individuals born very preterm is still limited, but there is emerging evidence of males being more affected by preterm birth than females with regard to educational attainment.

In this study we aimed to determine the strength of the independent associations between mathematics performance in children born very preterm and two adult outcomes—postsecondary education and being employed or in education—while adjusting for year of birth, gestational age, sex, maternal education, and child IQ.

METHOD

This study was registered as an individual participant data (IPD) meta-analysis with the International Prospective Register of Systematic Reviews PROSPERO (protocol #CRD42020175032).

Search strategy and study selection criteria

A two-step process was applied to identify and select studies for inclusion. First, members of the Adults born Preterm International Collaboration (APIC; www.apic-preterm.org) were presented with the study protocol. APIC is a collaborative group of international researchers who study life-course outcomes of individuals born very preterm. All APIC members were invited to contribute data to this study. Prospective longitudinal studies that had assessed participants born very preterm with standardized intelligence and mathematics tests in middle childhood and with adult data on postsecondary education and employment/education status were eligible for inclusion. Only studies with data available from at least three timepoints (i.e., birth, childhood, and adulthood) that included a contemporaneous comparison group of children born at term for the childhood assessments were included in the data harmonization and pooling process. There was no restriction for year of birth. A total of six APIC cohort studies were identified through this process by June 2019. Second, a systematic literature search was conducted in English in PubMed and PsycINFO using the keywords ‘very low birth weight’ OR (‘very preterm’ OR ‘premature’ OR ‘gestation’) AND (math OR ‘arithmetic’) AND (‘intelligence’ OR ‘IQ’ OR ‘cognitive’) AND (‘adult’).

Screening and review of articles did not identify additional data sets that fulfilled the eligibility criteria. Accordingly, data from the following six birth cohorts were included in this study (ordered alphabetically by country).

**Australia**

The Victorian Infant Collaborative Study included all livebirths less than 1000g and/or less than 28 weeks’ gestational age in the state of Victoria in 1991 to 1992 and a normal birthweight (>2500g) term comparison group recruited in the newborn period. At 8 years of age, mathematics performance was assessed using the Wide Range Achievement Test, Third Edition, Arithmetic scale, and IQ using the Wechsler Intelligence Scale for Children, Third Edition. Adult assessments were carried out at 18 and 25 years of age.

**Canada**

The McMaster University Study included all livebirths weighing 1000g or less in the Central-West Ontario region from 1977 to 1982, and a normal birthweight term comparison group recruited at school age. At 8 years of age, children were evaluated using the Wide Range Achievement Test Arithmetic scale and the Wechsler Intelligence Scale for Children – Revised. Adult assessments were carried out at 29 years of age.

**Germany**

The Bavarian Longitudinal Study included all livebirths born very preterm in Southern Bavaria in 1985 and 1986 and a normal birthweight term comparison group. At 8 years of age, children were evaluated using a standardized Mathematische Test and the Kaufman Assessment Battery for Children Mental Processing Composite scale (IQ-equivalent). Adult assessments were carried out at 26 years of age.

**UK and Republic of Ireland**

The EPIcure study is a nationwide study comprising all infants born at 22 to 25 completed weeks of gestation in the UK and Republic of Ireland between March and December 1995, and a term-born comparison group of peers recruited at school age. At 11 years of age, children were evaluated using the Wechsler Individual Achievement Test Second UK Edition Mathematics Composite scale and the Kaufman Assessment Battery for Children Mental Processing Composite scale (IQ). Adult assessments were carried out at 19 years of age.
USA
The Case Western Reserve University Cohort Study included all infants with birthweights less than 1500g admitted to Rainbow Babies and Children’s Hospital in Cleveland from 1977 to 1979 and a normal birthweight comparison group recruited at age 8 years.44 At 8 years of age, children were evaluated using the Woodcock-Johnson Achievement Test Mathematics Scale45 and the Wechsler Intelligence Scale for Children – Revised.33 Adult assessments were carried out at 20 years of age.44

USA
The Rhode Island Cohort of Adults Born Preterm Study included livebirths weighing less than 1500g in the speciality hospital from 1985 to 1989 and a ‘healthy’ term-born comparison group.46,47 At 8 years of age, children were evaluated using the Wide Range Achievement Test, Third Edition, Arithmetic scale30 and the Wechsler Intelligence Scale for Children, Third Edition.31 Adult assessments were carried out at 23 years of age.48

Data extraction, harmonization, and analysis
Participating groups (n=6) obtained ethical permission and completed a data sharing agreement with the University of Tennessee Knoxville to transfer non-identifiable individual level data for the analysis. Groups were requested to provide neonatal data (i.e., infant gestational age at birth [wks], birthweight [g], sex [binary coded], maternal education), childhood IQ and mathematics test results (standard scores), and adult employment and educational attainment information.

Data were harmonized in SPSS v24 (IBM Corp., Armonk, NY, USA), including a variable for each cohort’s year of birth. The content of the IQ and mathematics tests (e.g., dimensions assessed) were compared to confirm that these generally targeted very similar domains (see Table 1, online supporting information, for details). Test scores were z-standardized according to each cohort-specific contemporary comparison group before pooling. In line with an adaptive approach to scoring test performance, participants with IQ scores more than 2 SD below the control participants’ mean (i.e., ≤−2 SD) were coded as having intellectual impairment. Information about maternal education was recoded into an interval-scaled 9-level variable according to the International Standard Classification of Education.49 The two dependent variables in young adulthood (i.e., 19–29y) were binary-coded as ‘any postsecondary education’ (1=yes, ≥International Standard Classification of Education level 5 and 0=no, <International Standard Classification of Education level 5) and ‘currently employed or in education’ (1=yes/0=no).

One-stage IPD meta-analyses were performed as mixed effects logistic regressions in Stata 16 (StataCorp, College Station, TX, USA). All models were controlled for fixed effects of cohort year of birth, infant gestational age at birth, sex, maternal education, and childhood IQ, as well as the nestedness of data in cohorts (i.e., including a random effect for study site). The fixed and random effects of childhood mathematics scores on postsecondary education and being employed/in education in young adulthood are reported. Including random effects allowed for different relationships with adult outcomes within cohorts. As part of a stepwise process, random effects of other covariates were additionally included, and model fit was evaluated using log-likelihood goodness of fit tests. Random effects were retained if their inclusion significantly improved overall model fit. Specifically, with regard to attending postsecondary education, the additional inclusion of random effects for gestation and IQ improved model fit, indicated by a significant log-likelihood ratio χ² test (χ²[4]=26.30, p<0.001), fixed effects remained stable. With regard to being employed/in education, the additional inclusion of other covariates’ random effects did not improve model fit, indicated by a non-significant log-likelihood ratio χ² test (χ²[4]=9.11, p=0.058).

RESULTS
Individual cohort sample sizes ranged from 117 to 231 participants (Table 1). Sex was equally distributed across study cohorts but there was substantial variation with regard to gestational age and birthweight, because of different cohort inclusion criteria, and in distributions of maternal education levels. Mean childhood IQ and mathematics z-scores ranged from −1.57 to −0.36 and −1.55 to −0.51, respectively. Across cohorts there was variation in the proportion of adults who had attended postsecondary education (18–57%) and in those employed or in education at the time of assessment (71–88%).

One-stage IPD meta-analyses
Mixed effects logistic regression indicated substantial heterogeneity between study cohorts (i.e., significant log-likelihood χ² tests), supporting the use of mixed effects models as opposed to logistic models without random effects.

Table 2 shows that adults born very preterm with higher mathematics test scores in childhood had higher odds to have attended any postsecondary education (fixed effect odds ratio [OR] per SD increase in mathematics test scores: 1.36 [95% confidence interval {CI}: 1.03, 1.79]). The random effect of mathematics was estimated at SD 0.21 (95% CI: 0.05, 0.82), indicating significant variation between cohorts in the relationship between these two variables. In addition, adults born very preterm with higher childhood IQ (OR per SD increase in IQ: 1.58 [95% CI: 1.26, 1.97]), higher educated mothers (OR per SD increase in maternal education: 1.29 [95% CI: 1.09, 1.53]), and females (OR: 2.00 [95% CI: 1.43, 2.78]) had higher odds to have attended postsecondary education.

The fixed effect of childhood mathematics test scores (OR=1.14 [95% CI: 0.87, 1.48]) on current employment/education status was not significant, while the random effect was estimated at SD 0.21 (95% CI: 0.09, 0.51), indicating significant variation in this relationship between cohorts. Adults with higher childhood IQ were more likely
Table 1: Descriptive characteristics by included cohort

<table>
<thead>
<tr>
<th>Recruitment criteria</th>
<th>Australia (Victoria; n=166)</th>
<th>Canada (McMaster; n=122)</th>
<th>Germany (BLS; n=201)</th>
<th>UK (EPIcure; n=117)</th>
<th>USA (Cleveland; n=231)</th>
<th>USA (Rhode Island; n=117)</th>
</tr>
</thead>
<tbody>
<tr>
<td>&lt;28wks GA/&lt;1000g BW</td>
<td>143 (86.6%)</td>
<td>117 (96.7%)</td>
<td>213 (106%)</td>
<td>97 (83.3%)</td>
<td>226 (99.5%)</td>
<td>101 (91.1%)</td>
</tr>
<tr>
<td>&lt;32wks GA/&lt;1500g BW</td>
<td>231 (38.8%)</td>
<td>116 (95.8%)</td>
<td>231 (115%)</td>
<td>117 (99.2%)</td>
<td>231 (100%)</td>
<td>117 (100%)</td>
</tr>
<tr>
<td>Female</td>
<td>55.4%</td>
<td>53.3%</td>
<td>46.8%</td>
<td>55.6%</td>
<td>51.5%</td>
<td>55.6%</td>
</tr>
<tr>
<td>Birthweight (g)</td>
<td>885 (155)</td>
<td>841 (121)</td>
<td>1341 (324)</td>
<td>742 (126)</td>
<td>1177 (219)</td>
<td>1251 (329)</td>
</tr>
<tr>
<td>Gestation (wks)</td>
<td>26.7 (2.0)</td>
<td>27.0 (2.2)</td>
<td>30.6 (2.2)</td>
<td>24.4 (0.76)</td>
<td>30.2 (2.3)</td>
<td>30.0 (2.55)</td>
</tr>
<tr>
<td>Mother’s education</td>
<td>3.16 (0.87)</td>
<td>3.12 (1.26)</td>
<td>3.02 (1.08)</td>
<td>2.99 (0.99)</td>
<td>2.92 (0.74)</td>
<td>4.03 (1.12)</td>
</tr>
<tr>
<td>Child assessment age</td>
<td>8y</td>
<td>8y</td>
<td>8y</td>
<td>11y</td>
<td>8y</td>
<td>8y</td>
</tr>
<tr>
<td>Childhood mathb</td>
<td>-0.62 (1.05)</td>
<td>-0.90 (1.23)</td>
<td>-0.74 (1.15)</td>
<td>-1.55 (1.35)</td>
<td>-0.51 (1.46)</td>
<td>-0.57 (1.03)</td>
</tr>
<tr>
<td>Childhod IQb</td>
<td>-0.51 (1.14)</td>
<td>-0.95 (1.24)</td>
<td>-0.72 (1.33)</td>
<td>-1.57 (1.44)</td>
<td>-0.36 (1.15)</td>
<td>-0.59 (1.28)</td>
</tr>
<tr>
<td>IQ &lt; -2SDb</td>
<td>8.4%</td>
<td>18.9%</td>
<td>12.9%</td>
<td>35.0%</td>
<td>0.0%</td>
<td>12.0%</td>
</tr>
<tr>
<td>Adult assessment age</td>
<td>25y</td>
<td>29y</td>
<td>28y</td>
<td>19y</td>
<td>20y</td>
<td>23y</td>
</tr>
<tr>
<td>Employed/in education</td>
<td>78.3%</td>
<td>77.0%</td>
<td>87.6%</td>
<td>85.5%</td>
<td>80.5%</td>
<td>70.9%</td>
</tr>
<tr>
<td>Postsecondary education</td>
<td>56.6% of n=143</td>
<td>39.1% of n=64</td>
<td>17.9% of n=201</td>
<td>25.2% of n=111</td>
<td>36.4% of n=231</td>
<td>38.8% of n=116</td>
</tr>
</tbody>
</table>

aHarmonized according to the International Standard Classification of Education (0-8 point scale). bValues z-standardized according to cohort-specific controls. cProportion of participants who attended any postsecondary education (International Standard Classification of Education levels 5-8), binary coded yes/no; n=88 participants across different cohorts were still attending secondary education and/or did not have information about postsecondary education; the resulting total number of participants for this part of the analyses was 866. BLS, Bavarian Longitudinal Study; GA, gestational age at birth; BW, birthweight.

Table 2: Multilevel mixed-effects model results (one-stage individual participant data) showing associations of childhood mathematics scores and other covariates with attending postsecondary education and being employed/in education in young adulthood in six very preterm cohorts

<table>
<thead>
<tr>
<th>Dependent variable</th>
<th>Postsecondary education (n=866)</th>
<th>Employed/in education (n=954)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Fixed effects</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mathematics (per SD)</td>
<td>1.36 (1.03, 1.79)</td>
<td>1.14 (0.87, 1.48)</td>
</tr>
<tr>
<td>Year of birth (per year)</td>
<td>1.05 (0.97, 1.13)</td>
<td>1.00 (0.96, 1.03)</td>
</tr>
<tr>
<td>Gestation (per week)</td>
<td>1.06 (0.98, 1.15)</td>
<td>0.97 (0.90, 1.03)</td>
</tr>
<tr>
<td>Female</td>
<td>2.00 (1.43, 2.78)</td>
<td>0.83 (0.59, 1.16)</td>
</tr>
<tr>
<td>Mother’s education (per point)</td>
<td>1.29 (1.09, 1.53)</td>
<td>1.14 (0.96, 1.35)</td>
</tr>
<tr>
<td>Participant IQ (per SD)</td>
<td>1.58 (1.26, 1.97)</td>
<td>1.28 (1.06, 1.56)</td>
</tr>
<tr>
<td>Random effects by cohort</td>
<td></td>
<td></td>
</tr>
<tr>
<td>SD (Mathematics)</td>
<td>0.21 (0.05, 0.82)</td>
<td>0.21 (0.09, 0.51)</td>
</tr>
<tr>
<td>SD (Gestation)</td>
<td>0.02 (0.01, 0.04)</td>
<td>-</td>
</tr>
<tr>
<td>SD (IQ)</td>
<td>0.10 (0.00, 4.11)</td>
<td>-</td>
</tr>
<tr>
<td>Log-likelihood</td>
<td>-455.04</td>
<td>-450.93</td>
</tr>
</tbody>
</table>

For fixed effects: *p<0.05; **p=0.01; ***p<0.001. OR, odds ratio; CI, confidence interval; SD, standard deviation.

to be employed or in education at follow-up (OR=1.28 [95% CI: 1.06, 1.56]). Figure 1 displays forest plots of the effects of childhood mathematics scores on postsecondary education and current employment/education status by cohort.

As sensitivity analyses, calculations were repeated three times. First, to test the effects of mathematics test scores on adult outcomes without including control variables. These showed similar but stronger associations of mathematics performance with both postsecondary education and current employment/education status (see Table S2, online supporting information). Second, because 118 participants had intellectual impairment, models were rerun after their exclusion, again with similar findings (see Table S3, online supporting information). Finally, considering potential outlier effects of the EPIcure study data indicated in Figure 1, analyses were rerun without these 117 EPIcure participants. In these models, the fixed effect of childhood mathematics test scores on attending postsecondary education (OR=1.18 [95% CI: 0.87, 1.61]) was not significant, nor was the effect on current employment/education status (OR=1.11 [95% CI: 0.90, 1.38]).

**DISCUSSION**

In this IPD meta-analysis of six very preterm cohort studies from five different countries in Europe, North America, and Australia, born between 1977 and 1995, we found evidence that higher mathematics scores in childhood were independently associated with attending any postsecondary education. Evidence was weak for an independent association.
between mathematics performance and being employed or in education in young adulthood. The random effects of mathematics on attending postsecondary education and on being employed or in education were both significant, indicating substantial heterogeneity between study cohorts with regard to the relationship of mathematics with adult
outcomes. Over and above the effects of mathematics, childhood IQ was independently associated with attending any postsecondary education and being employed or in education. Moreover, females born very preterm and those with higher educated mothers at birth were more likely to attend postsecondary education.

It is important to highlight that this study did not investigate effects of very preterm birth as has traditionally been done in similar meta-analyses. The cohort-specific contemporaneous control data were only used to standardize childhood test scores. With regard to the other variables in the mixed effects models, adults born very preterm who had been born to higher educated mothers had 29% higher odds of attending postsecondary education, as previously indicated in other studies. The current findings confirm that educational and employment opportunities are enhanced with birth to higher educated mothers, no matter whether in Australia, Europe, or North America. Females born very preterm were twice as likely than males born very preterm to attend postsecondary education. This confirmation that males born very preterm have lower educational attainment than females across these different industrialized societies is important, and, in keeping with previous findings, suggests that males born very preterm may benefit from screening and targeted educational support. More research is needed to investigate why males born very preterm may do particularly poorly in education. Although not novel findings, the cross-national and temporal consistency of effects of maternal education and child sex found here may warrant renewed consideration for how early follow-up services and educational policies could support equal opportunities for all members of society. Moreover, our finding of heterogeneity between study cohorts with regard to the relationship of mathematics with adult outcomes may be further investigated by future studies. For instance, it may be worth assessing whether specific educational supports or therapies provided to some cohorts, or individuals within cohorts, might explain this variation and thereby point towards novel interventions.

Our models did not show fixed effects of gestational age at birth. However, variations in adult outcomes according to gestation may have been restricted by the limited range in gestation, mediated by functional childhood indicators such as IQ and mathematics performance, and they may also have been masked by the random effects of cohort membership, since each study cohort had applied different recruitment and sampling criteria, causing wide variations in gestation and birthweight between cohorts (Table 1).

Accordingly, Table 2 shows a significant random effect of gestation, indicating variation between cohorts in the relationship between gestational age and postsecondary education. Sensitivity analyses excluding EPICure study data also suggested that the association between mathematics performance and postsecondary education varied by gestational age, being stronger among participants born extremely preterm, whose average childhood performance may also have been comparably lower compared with term-born peers due to their birth on the lower extreme of the gestational age spectrum (see Table 1).

Strengths and limitations
Our IPD meta-analysis confirms our ability to predict life-course outcomes into early adulthood across different very preterm cohorts sampled across different birth epochs, neonatal health care services, and education systems. Using pooled IPD provided a large sample size and allowed better adjustment for confounding bias than a standard meta-analysis. Data harmonization and analysis protocols returned stable and reliable models that accounted for the heterogeneity between cohorts. In the future, we support recommendations to harmonize follow-up instruments to further facilitate comparative studies.

For reasons of ecological validity and data distribution we decided to use postsecondary education (i.e., having attended college after graduation from secondary school) as the dependent variable for educational attainment. Access to higher education has been found to represent an excellent indicator of attainment and economic success in adults born preterm. Although education data were recoded using International Standard Classification of Education levels, some structural differences between education systems may not be fully eliminated. For instance, graduating from high school and choosing the respective pathway into higher education may have a higher threshold in Germany and the UK than in North America, in part because alternative professional apprenticeship qualification pathways may be more typical for large parts of the population in Europe. For instance, according to 2020 OECD data, population rates of postsecondary educational attainment among 25- to 34-years-olds differ between Germany (35%) and Australia (54%) and the USA (52%). Respectively, the values in Table 1 should be interpreted with these country-specific differences in mind. Nevertheless, our main results focus on associations between variables, not descriptive rates of outcomes. Moreover, national differences in elementary school education (e.g., regular age at formal school entry, rates of delayed school entry, and rates of inclusion of children with disabilities [often very preterm] into mainstream schooling) and differences in ages at assessments may have confounded the distribution of variables included in our models. In addition, the childhood assessment was administered at age 11 years in the EPICure cohort, and at age 8 years in the other five cohorts, which may have affected cohort participants’ average performance in comparison to their age-matched controls. Indeed, our sensitivity analysis confirmed a slight variation in findings when excluding EPICure participants. At the time of the adult assessment, some cohorts’ participants were still in secondary education and were excluded, therefore we had a slightly lower sample size for that dependent variable. In addition, data collection for the young adult assessments was conducted between ages 19 to 29 years, before some participants had fully completed their postsecondary education and established career trajectories. Accordingly, our assessment of the role of...
CONCLUSIONS
Among infants born very preterm, performance in mathematics in childhood represents an independent functional indicator of attending postsecondary education, but it is not associated with being employed or in full-time education in early adulthood. IQ was strongly and consistently associated with both postsecondary education and employment/education in young adulthood. Despite substantial variations between cohorts (i.e., world regions and birth epochs), males born very preterm were consistently at higher risk than females born very preterm of not attending postsecondary education, as were infants born very preterm to mothers of lower educational background. These findings have implications for the long-term follow-up and support after very preterm birth.

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DATA AVAILABILITY STATEMENT
The aggregated data that support the findings of this study are available on request from the corresponding author. The individual data points are not available due to privacy or ethical restrictions of individual study cohorts.

SUPPORTING INFORMATION
The following additional material may be found online:

Table S1: Childhood standardized mathematics and IQ tests, and control group criteria by cohort

Table S2: Associations of childhood mathematics scores with attending postsecondary education and being employed/in education in young adulthood in six very preterm cohorts, without adjusting for confounding variables

Table S3: Associations of childhood mathematics scores and other covariates with attending postsecondary education and being employed/in education in young adulthood in six very preterm cohorts, excluding participants with intellectual impairment

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