Retrospective observational study of ethnicity-gender pay gaps among hospital and community health service doctors in England

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ABSTRACT

Objective To identify differences in average basic pay between groups of National Health Service (NHS) doctors cross-classified by ethnicity and gender. Analyse the extent to which characteristics (grade, specialty, age, hours, etc.) can explain these differences.

Design Retrospective observational study using repeated cross-section design.

Setting Hospital and Community Health Service (HCHS) in England.

Participants All HCHS doctors in England employed by the NHS between 2016 and 2020 appearing in the Digital Electronic Staff Record dataset (average N=99,953 per year).

Main outcome measures Hours-adjusted full-time equivalent pay gaps; given as raw data and further adjusted for demographic, job, and workplace characteristics (such as grade, specialty, age, whether British nationality, region) using multivariable regression and statistical decomposition techniques.

Results Pay gaps relative to white men vary with the ethnicity-gender combination. Indian men slightly out-earn white men and Bangladeshi women have a 40% pay gap. In most cases, pay gaps can largely be explained by characteristics that can be measured, especially grade, with the extent varying by specific ethnicity-gender group. However, a portion of pay gaps cannot be explained by characteristics that can be measured.

Conclusions This study presents new evidence on ethnicity-gender pay gaps among NHS doctors in England using high quality administrative and payroll data. The findings indicate all ethnicity-gender groups earn less than white men on average, except for Indian men. In some cases, these differences cannot be explained giving rise to discussions about the role of discrimination.

INTRODUCTION

Differences in earnings between men and women doctors and differences between white and non-white doctors are well known. An analysis of ethnicity pay gaps by the Nuffield Trust revealed a ‘small’ pay gap generally favouring white doctors in England.1 Additionally, the Independent Review into Gender Pay Gaps in Medicine in England revealed that there was a gender pay gap in mean annual pay of 24.4% favouring men among Hospital and Community Health Service (HCHS) doctors in National Health Service (NHS) trusts in England, which reduces to 18.9% when expressed on a pro-rata basis.2

The Review argued that this is about twice as large as the gender pay gap for professional employees in the UK.3 It is disappointing that the Review represented the most detailed analysis ever of gender pay gaps in a profession. However, it did not explore the interaction between ethnicity and gender in detail. This paper aims to do so for HCHS doctors in NHS trusts in England.

Prior analyses have only considered gender and ethnic pay gaps additively. The different question of how combinations of ethnicity by gender groups compare to each other (ie, multiplicatively) has very rarely been explored in medicine.4-6
from defining groups in this way are often different from analysing ethnicity or gender separately.7 In the social science literature, this is referred to as ‘intersectionality’origi-
nated in the USA to describe how viewing social advan-
tage and disadvantage along single and discrete axes was misleading.10 For instance, a recent analysis by the Bank of England on a nationally representative sample of all workers in the UK found that women earned less than men on average, although non-white men earned less than white men and non-white women earned more than white women.7 The main contribution we seek to make is to explore ethnicity-gender pay gaps among medics through the lens of ‘intersectionality’. In foreshadowing the findings, given the exploratory nature of this paper we, have only general expectations. Following on from the intersectional perspective, we expect that stereotypes about cross-cutting categories of difference such as gender and ethnicities may advantage some groups and disadvantage others.6 11 We therefore expect the average pay disadvantage of women to vary according to ethnicity, and white men to generally be the highest earning group of all groups.

A second contribution we seek to make is to move beyond the simplistic two-factor view of ethnicity identity, which collapses non-white ethnicity groups into a single category of Black, Asian and Minority Ethnic (‘BAME’). This is a pertinent issue in terms of pay gaps, because some non-white ethnicity subgroups experience slight pay advantages over white doctors.12 It is also especially pertinent to NHS HCHS doctors in England, where 44% are ‘BAME’,13 four times more than the proportion in the UK labour market.14

An innovation of the Independent Review into Gender Pay Gaps in Medicine in England was to ‘decompose’ pay gaps into explanations that account for the differing composition of groups across pay bands on the one hand, and the differing wage structures of groups on the other. Our final contribution is to apply these techniques to the analysis of intersecting ethnicity-gender pay gaps to draw conclusions. We expect that the majority proportion of pay gaps of groups relative to white men will be explained by factors we can observe, for example, that groups are unequally distributed across grades and levels of seniority.

METHODS
The following is an analysis of ethnicity-gender pay gaps using data from all NHS trust doctors in the Electronic Staff Record (ESR). The study pooled cross-sectional data for 51 months from January 2016 to March 2020. Our data access included all doctors working in NHS trusts in England. We used ordinary least squares regression analysis combined with Oaxaca-Blinder decomposition techniques to delineate the causes of pay gaps.

The ESR
The ESR is an administrative monthly payroll dataset. It records a rich set of information about each doctor including earnings, demographic, job, and workplace characteristics. We used this data within a pooled repeated cross-section design to determine differences in average pay between ethnicity-gender groups. We excluded those with a non-medical primary area of work (eg, corporate, estates, dental/oral and facilities) from the sample. We included all grades (foundation years 1 and 2, staff and local grades, core trainees and specialty registrars, consultants, associate specialists and specialty doctors). We excluded cases where basic pay was zero or negative, where monthly hours worked were zero or exceed 320, and those with an inactive contract. Our final sample consisted of 5 097 897 doctor-month observations generated from 164 820 individual doctors. The average number of doctors in the sample each year grew from 95 636 in 2016 to 108 408 in 2020. To maintain representativeness of the sample, all doctors are included, irrespective of their length of service.

Analysis
To identify pay gaps, we followed the government’s advice on gender pay gap reporting which defines the gender pay gap as the percentage difference in women’s relative to men’s mean earnings using hours-adjusted measures. Given our focus is on ethnicity-gender gaps, we took white men as the reference group, as they are usually the most advantaged.

To explore the extent to which average pay can be explained by observed characteristics (ie, characteristics that we have information about, see under the ‘Measures’ section), we used an ordinary least squares (OLS) regression decomposition technique known as Oaxaca-Blinder decomposition—see online supplemental appendix A—(henceforth OBD),15 16 which is widely used in the econometric analysis of pay gaps17 18 and in health studies.19 20 The OBD statistical technique decomposes a pay gap into two elements by deploying OLS regression techniques to illuminate how patterns in the composition of groups influence the gap. For example, one reason a pay gap emerges between two groups is because they are differently composed across, for example, grade, that is, one group is more likely to be found in higher-paid senior grades than the other. Pay also tends to increase with age (through accumulated experience and tenure) and surgery is the highest-paid specialty. All can be considered, on the face of it, legitimate reasons for pay gaps. The extent to which these compositional factors account for observed pay gaps is captured by ‘endowment effects’. Pay gaps may also emerge because one group may get paid more on average for attaining a grade, holding other factors constant. The extent to which these wage structure factors adjust observed pay gaps are captured within ‘coefficient effects’. Differing wage structures for a given set of characteristics can arise because, for example, groups consistently occupy lower
points on pay scales for a given grade, or specialty, or age etc. Coefficient effects may therefore indicate discrimination or wage bias.

Given we use population and not sample data, we only report 95% CIs in the text for estimated parameters (ie, those arising from the multivariable analysis focusing on the pooled sample).

Measures
Ethnicity is self-assigned in the ESR. Seventy-six ethnicity categories in the ESR were reduced to seven for the purposes of analysis: white, black, Indian, Pakistani, Bangladeshi, Chinese, including the South East Asian (SEA) group and mixed race. These were then cross-classified with gender to create 14 ethnicity-gender groups. The labels for these groups have been abbreviated in the figures to improve their legibility where W indicates women and M men. So, for example, blackW refers to black women and IndianM to Indian men.

The main dependent variable was basic monthly pay; that is the element of pay before overtime, bonuses and tax are applied. Because one reason why pay may vary between individuals and groups is differences in contracted hours, our finding adjusted basic monthly pay by contracted hours to create full-time equivalent pay. We also controlled for month and year fixed effects to account for inflation. Given the left skew in pay, we transformed pay using the natural logarithm. This transformation also had the added benefit for our purposes in exploring pay gaps because the coefficients in the multivariable analyses were roughly equivalent to percentage point differences.

A variety of explanatory factors for pay gaps were available in the ESR and were also included. Explanatory variables were classified into following categories: age (a proxy for career experience) and age squared (to account for non-linearities in age-earnings profiles). Grade was indicated in a set of binary indicators for senior doctors (consisting of consultants, associate specialists and specialty doctors) junior doctors (comprising of specialty registrars, core trainees and foundation year students) plus staff and local grade doctor. Specialty included binary indicators for primary area of work that is, clinical oncology, clinical support, general acute, imaging, medicine, surgery, obstetrics and gynaecology, psychiatry, pathology, public health and occupational health. Region is indicated by dividing the population into 10 strategic health authorities. Personal characteristics consisted of a comprehensive set of variables on nationality, religion, sexual orientation and disability status. Work-related characteristics included variables on whether each doctor held a fixed term contract or not, and whether there were multiple NHS assignments in a given person for the month.

Patient and public involvement
No patients were involved.

RESULTS
Descriptive overview
In our population of 164 820 total doctors, we found missing data either in the form of actual missing or refusal to divulge in the case of ethnicity for 11 318 doctors (6.9%) and nationality (4812, 2.9%). We also found substantial missing data in sexual orientation (57 752, 34.0%) and religious identity (62 881, 38.1%) categories. All other variables in the data had less than 0.1% missing observations. In the analysis, non-white doctors constituted 49.1%. Non-white men and women comprised 31.1% and 18.0% of all doctors, respectively.

White doctors constituted the majority (50.9%) followed by Indians (15.2%), others/unknown (6.9%), Chinese/SEA (6.5%), Pakistanis (6.2%), other Asians (5.1%), blacks (4.9%), mixed race (3.3%) and Bangladeshis (1.0%). In terms of ethnicity gender, white women constituted the largest group (25.8%) followed by white men (25.1%), Indian men (9.2%), Indian women (6.1%), Chinese/SEA men (3.8%) and Pakistani men (3.6%).

To commence the analysis in relation to the first study objective, we explored differences in basic monthly pay among detailed intersectional ethnicity-gender categories (figure 1) and tracked them over 5 years (figure 2). Both figures illustrate the importance of disassembling the ‘BAME’ category by showing considerable pay gap heterogeneity by ethnic and gender group. All ethnicity-gender groups earned less than white men, except for Indian male doctors who earned slightly more. There were especially large gaps for Bangladeshi, Pakistani, black, mixed race and Chinese/SEA women doctors, who experienced on average between 25% and 40% lower pay than white men, amounting to a monthly basic pay gap of around £1500–£2000. White women and Indian women received lower pay relative to white men in similar magnitude to the disadvantages experienced by Chinese/SEA men, Pakistani men and black men, with these groups experiencing around 10%–20% lower pay on average. Bangladeshi men were noticeably the most disadvantaged male category, earning on average approximately 20%
less than white men. Over the 5 years of measurement, pay gaps for most groups remained high, but did not increase. The exception to this was pay gaps for Chinese/SEA men, black men and Pakistani men which grew steadily by 3.4%, 10.3% and 3.5%, respectively.

**Decomposition analysis: explaining pay gaps**

To address the second study objective and determine the extent to which differences in the composition of ethnicity-gender groups and their wage structures accounted for pay gaps relative to those of white men, we undertook an OBD decomposition using all explanatory factors in the ESR. The results are presented in figures 3 and 4. Figure 3 presents a decomposition of pay gaps relative to white men in absolute terms, summing to the mean pay gap for each group with a percentage point interpretation. Figure 4 presents the same but in relative terms, where the contribution of endowments and coefficient effects can be compared across groups and sum to 100%.

The main finding from this figure is that for almost all groups, differences in endowments accounted for about two-thirds or more of the pay gap with white men. This means that, in most cases, the major proportion of each pay gap was explained by known factors (figure 5). These factors were especially important in the case of Bangladeshi, Pakistani, black, mixed race and Chinese/SEA women doctors where the pay gap was large but can be explained by differences in endowments. The exception here was black men doctors, where endowment effects accounted for only two-fifths of their pay gap.

The coefficient effect, which could be held to be evidence of direct pay discrimination, was a feature for all groups, and notably large for black men and women. In relative terms, the coefficient effect tended to be larger for the male categories, especially black men where it explained 57.6% of the pay gap (56.0%–59.1%), and for Indian men (who earn slightly more than white men) where it offset their seeming endowment advantage by 48.2% (50.6%–45.9%). However, a wide array of other unmeasured characteristics such as productivity, performance or work histories can also determine basic pay, so we are cautious in using this interpretation. For instance, although we included nationality as a control in our regression models, the ESR dataset does not include information on routes into training and employment. International medical graduates (IMG) and European Economic Area doctors tend to have longer training routes than UK-trained doctors to reach senior grades and this can be misattributed as ethnicity effects. However, we highlight that the wage structure effects uncovered here constitute an area for further, and urgent, investigation.

Finally, using decomposition techniques again, we disaggregated the most influential factors within the endowment element of the pay gap. In figure 5, we show that, all else being equal, grade consistently stood out as being the single most important factor, accounting for 40%–60% of pay gaps for all ethnicity-gender groups relative to white men.
men (see also table 1 below). Age was also important in explaining pay gaps with older doctors earning more on average. In the male ethnic doctor categories, its contribution to the pay gap was minor for many, for example, 7.8% (7.2%–8.4%) for Pakistani men. Since black men were, on average, older than white men (43.5 vs 42.5) this reduced their pay gap by 11.9% (13%–10.9%). For all female categories, the contribution of age to the pay gap varied between 20% and 30%. Specialty played a very minor role in the pay gap, explaining less than 1% for any group. Region and work-related variables had almost no influence.

The contribution of personal features to pay gaps is worth noting. As a group, Indian men were distributed across personal feature measures (nationality, religion, sexual orientation and disability status) in ways that were well-rewarded and reduced their pay gap by 24.1% (25.4%–22.9%). Personal features, however, disadvantaged other groups of ethnic minority men, explaining 12.1% (11.4%–12.8%), 17% (16.2%–17.8%) and 19.2% (18.5%–19.9%) of the pay gap for Bangladeshi, Pakistani and Chinese and SEA men, respectively. Personal features had a smaller, but still statistically significant, impact on the pay gaps of ethnic minority women, accounting for 5.9% (5.8%–6.1%), 7.6% (7.1%–8.1%), 8.6% (8.2%–9%) and 9% (8.6%–9.4%) of the pay gap for Indian, Bangladeshi, Pakistani and Chinese and SEA women, respectively. Work-related variables which include contract type (permanent vs fixed-term) and number of assignments played a minor, but also statistically significant, role in pay gaps, explaining between 3% and 4% of the total pay gap for most men and women with the exception of Indian men where they explained 7.4% (7.1%–7.6%).

To reinforce our understanding of why grade and age are important to the observable (endowment) pay gaps between ethnicity-gender categories relative to white men, we show the distributions of these factors by ethnicity-gender.

White men had among the highest mean age and were disproportionately found in senior ranks. Indian men were also favourably distributed across grade and age, but their advantage was mitigated by the coefficient effect explored above. The lowest-paid ethnicity-gender groups were less likely to be observed in senior grades. For instance, Chinese/SEA women, mixed race women, black women, Pakistani women and Bangladeshi women
can be accounted for by observed factors, with grade and gender demonstrating that most gaps (except for Black men) are accounted for by ethnicity-gender -that account for pay gaps between groups. Applying these variable decomposition techniques to explain the factors that cause these pay gaps, we also applied multivariable decomposition techniques widely used in pay gaps research and identifying these gaps, we also applied multivariable decomposition techniques widely used in pay gaps research and the recent Independent Review into Gender Pay Gaps in Medicine in England. Extending the findings in the Review, we find evidence of pay disadvantage particularly for all groups of women doctors, but with white men as the reference category, we find pay disadvantage for most non-white men too.

Importantly, we find that there is much heterogeneity in pay gaps across groups. Large gaps are found for Bangladeshi, Pakistani, black, mixed race and Chinese/SEA women. White women and Indian women suffer from lower pay relative to white men in a magnitude similar to the disadvantages experienced by Chinese/SEA men, Pakistani men and black men. Bangladeshi men are noticeably the most disadvantaged male category. Only Indian men earn similarly to white men on average.

One innovation of the Review was employing multivariable decomposition techniques to explain the factors that account for pay gaps between groups. Applying these techniques to ethnicity-gender pay gaps relative to white men, demonstrates that most gaps (except for Black men) can be accounted for by observed factors, with grade differences explaining the majority. The predominance of white men in senior ranks perpetuates the gap and this will need to be overcome if pay gaps are to reduce. White men are also, on average, older, which goes some way towards explaining their higher earnings. However, for most groups, the disadvantage role of personal features such as religion, sexuality and nationality in pay gaps requires investigation. Plus, a disadvantage if small, in absolute terms, the coefficient effect is found for all groups, indicating that they are paid less than white men for the same characteristics that we observe in the ESR (grade, specialty, age, etc.). This accounts for a larger share for male category doctors, especially black, Pakistani and Indian men, explaining more than a third of their pay gap. There is a possibility that these effects might, all else being equal, result from the differentiated allocation of opportunities for progression or value-enhancing experience, such as committee memberships, among homophilous networks especially in senior ranks. The differing wage structures between ethnicity-gender groups for a given set of characteristics certainly warrants urgent further investigation.

**Implications of the findings**

The Review highlighted the gender dimension of unsympathetic career structures as an important factor in understanding the gender pay gap. This study used the same methodology to understand differences in full-time equivalent mean pay for intersectional groups. Implications of this study highlight, first, the importance of disaggregating the workforce into intersectional groups as the success of one group (Indian male doctors) can obscure the extraordinary disadvantage of others (eg, Bangladeshi women). Second, via decomposition analysis we have highlighted the importance of overcoming pay gaps for all intersectional groups by working towards the goal of equalising grade via equal progression and building equal workplace experience by ensuring improved workforce retention relative to white men. A focus of workforce policy directed towards alleviating pay gaps should be to better understand why there is uneven progression through the grades between ethnicity-gender groups, especially the female groups plus targeting them with retention strategies. We find that specialty choice was not particularly important in understanding pay gaps, however, strategies of making high prestige specialties a more attractive medical career option for women and especially minority ethnic women would also pay dividends where they are needed.

Findings also revealed that even when accounting for grade and other compositional differences, certain groups appear to be paid less, on average, for these characteristics. Coefficient effects account for the majority of the pay gap for black men and more than one-third of the pay gap relative to white men in other cases. As we have stressed, the evidence presented here cannot be straightforwardly interpreted as pay discrimination, but the findings do demand further research into why certain groups

**Table 1** Composition of ethnicity-gender groups

<table>
<thead>
<tr>
<th>Ethnicity-Gender</th>
<th>Age (mean)</th>
<th>% Ethnicity Gender</th>
</tr>
</thead>
<tbody>
<tr>
<td>WhiteM</td>
<td>42.5</td>
<td>60.7</td>
</tr>
<tr>
<td>BlackM</td>
<td>43.5</td>
<td>49.5</td>
</tr>
<tr>
<td>IndianM</td>
<td>43.5</td>
<td>67.7</td>
</tr>
<tr>
<td>PakistaniM</td>
<td>41.6</td>
<td>51.5</td>
</tr>
<tr>
<td>BangladeshiM</td>
<td>39.0</td>
<td>42.3</td>
</tr>
<tr>
<td>ChineseSEAM</td>
<td>41.7</td>
<td>52.6</td>
</tr>
<tr>
<td>MixedM</td>
<td>38.9</td>
<td>44.3</td>
</tr>
<tr>
<td>OtherAsianM</td>
<td>41.1</td>
<td>49.5</td>
</tr>
<tr>
<td>WhiteW</td>
<td>38.3</td>
<td>45.0</td>
</tr>
<tr>
<td>BlackW</td>
<td>37.5</td>
<td>31.1</td>
</tr>
<tr>
<td>IndianW</td>
<td>39.5</td>
<td>50.8</td>
</tr>
<tr>
<td>PakistaniW</td>
<td>36.4</td>
<td>29.0</td>
</tr>
<tr>
<td>BangladeshiW</td>
<td>35.8</td>
<td>28.5</td>
</tr>
<tr>
<td>ChineseSEAW</td>
<td>36.4</td>
<td>34.3</td>
</tr>
<tr>
<td>MixedW</td>
<td>35.4</td>
<td>29.9</td>
</tr>
<tr>
<td>OtherAsianW</td>
<td>37.7</td>
<td>37.3</td>
</tr>
</tbody>
</table>

SEAM, South East Asian men; SEAW, South East Asian women.
are paid differently for the same characteristics, such as grade and specialty. For all the detail we achieve, we still only present a broad portrait here.

Limitations

There are several limitations to this study. First, we only focused on basic pay. Basic pay measures do not include overtime, CEAs, shift work premia, etc. The Review identified these non-basic components of pay as important contributing factors to the total gender pay gap among HCHS doctors, with gender gaps being larger for total pay than for basic pay. Moreover, in the Review, observed characteristics were even less able to explain non-basic components of pay. It is likely therefore that we have underestimated the full extent of intersectional disparities.

Second, we only focused on monthly pay, not annual pay, meaning gaps may be understated for groups that are more affected by absences or short contract working. Third, while we included a rich set of explanatory variables in our analyses, there are always other factors that determine pay and so differences in average pay between groups. Statistically, the balance between endowment and coefficient effects in accounting for pay gaps depends on the variables included in the model, with, in general, the inclusion of more variables increasing the proportion of pay gaps explained. There are a variety of work history factors including IMG doctor status that could be explored, if captured in the ESR. There are other unmeasured factors that may affect pay but have no straightforward means to be robustly explored, such as within-grade salary point, productivity and performance. Other unmeasured factors could include pay discrimination. Factors such as these may well account for the disadvantaging wage structure effects that were observed, but further research is clearly needed to understand them.

CONCLUSIONS

Differences in pay between men and women and white and non-white doctors are well known. Pay gaps in medicine have been raised as a concern across other national contexts. Although the demographic make-up of ‘minority’ and disadvantaged groups will alter across national contexts, it is likely that similar compositional differences will create similar intersectional pay gaps.

Previous analysis of doctors in England has not explored the interaction between ethnicity identity and gender in understanding pay gaps. This study considered pay gaps in mean basic monthly pay between detailed ethnicity-gender groups relative to white men for HCHS doctors in England. It applied multivariable decomposition techniques to explain pay gaps. Findings reveal non-white doctors earn less than white doctors on average, but there is much heterogeneity in the magnitude of pay gaps, with certain female ethnic groups being particularly disadvantaged. Much of the pay gap relative to white men can be explained by differing composition of groups, especially in terms of age grade, but all groups suffer at least a small disadvantaging pay penalty not accounted for by observed characteristics.

Finally, we are keen to stress that explaining pay gaps between groups using compositional differences does not justify them. Statistical models employed here cannot account for structural barriers and discrimination that led to the differences in composition between groups in the first place (eg, achieving a certain level of seniority for a given age). Supplemeting statistical analysis with robust qualitative evidence will help to elucidate how this occurs.

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REFERENCES


Appendix A

The Oaxaca-Blinder Decomposition.

The Oaxaca-Blinder decomposition (OBD), is a standard econometric technique used in wage gap decomposition analysis plus a range of health outcomes. It can be used to reveal the driving factors behind pay gaps. If we took two groups as a starting point, an OBD procedure estimates the earnings structures of group one and group two using separate ordinary least squares regression estimations. This produces estimates of the meaningfulness of differences in personal and job characteristics for group one and group two in terms of earnings, alongside the average values of those characteristics for each group (referred to as “endowments”). Second, the decomposition produces an estimate of rewards, referred to as the ‘coefficients’ effect. Here the estimate measures differences in financial returns for each group holding equal measures of the same characteristic. For example, in this context it estimates different returns for men and women of differing ethnic groups relating to, for example, being a surgeon, or being 40. This could be considered discrimination. The coefficient effect also includes the constant in the model, which captures unobserved attributes associated with the pay gap.

The original Oaxaca-Blinder decomposition takes the following form:

\[
\ln(W_{WM}) = X_{WM}\beta_{WM} + \epsilon_{WM}
\]

\[
\ln(W_{CG}) = X_{CG}\beta_{CG} + \epsilon_{CG}
\]

\[
R = \ln(W_{WM}) - \ln(W_{CG}) = (\bar{X}_{WM} - \bar{X}_{CG})\beta_M + \bar{X}_F(\beta_M - \beta_{CG})
\] (1)

Where WM = white male, CG = comparison group and R is the ethnic pay gap in percentage points. The first term on the right-hand side reflects the difference in endowments between white men and the comparison group and the second term reflects the difference in the slope in the regression term of white male and comparison group wages i.e. differences in the structure of rewards to these endowments.

In this paper, we follow an alternative specification of the OBD decomposition commonly used in labour economics. It is assumed that there is a non-discriminatory coefficient vector derived from the data which is used to determine the contribution of the differences in the predictors outcomes. Let \(\beta^*\) be such a nondiscriminatory coefficient vector. The outcome difference from (1) can then be rewritten as

\[
R = \{E(X_{WM}) - E(X_{CG})\}'\beta^* + \{E(X_{WM})'(\beta_{WM} - \beta^*) + E(X_{CG})'(\beta^* - \beta_{CG})\}
\] (2)

1 Sen B. Using the Oaxaca–Blinder decomposition as an empirical tool to analyze racial disparities in obesity. Obesity. 2014 Jul;22(7):1750-5.
This results in a “twofold” decomposition, 

\[ R = Q + U \]

where the first component, 

\[ Q = \{E(X_{WM}) - E(X_{CG})\}\beta^* \]

is the part of the wage gap that is explained by group differences in the explanatory variables (the “endowment effect”), and the second component;

\[ U = E(X_{WM})'(\beta_{WM} - \beta^*) + E(X_{CG})'(\beta^* - \beta_{CG}) \]

is the unexplained part, often known as the “coefficient effect”. The coefficient effect is sometimes attributed to discrimination or disadvantage stemming from the wage structure, but it is important to highlight that this can also potentially capture unmeasured and unobserved variables. We derive \( \beta^* \) by using a pooled regression for both white males and the comparison group and derive the robust variance covariance matrix of the estimates using the delta method as outlined by Oaxaca and Ransom.