
This paper is not the copy of record and may not exactly replicate the final, authoritative version of the article. The final article is available, upon publication, via the above journal.
University Lecturers’ Adaptability: Examining Links with Perceived Autonomy Support, Organisational Commitment, and Psychological Wellbeing

First submitted 25th May 2020. Accepted 24th July 2020.

Holliman, A. J.,¹ Keen, A.,² & Waldeck, D.³

¹Department of Psychology and Human Development, UCL Institute of Education

²Her Majesty’s Prison

³School of Psychological, Social and Behavioural Sciences, Coventry University

Author Note

Andrew J. Holliman. https://orcid.org/0000-0002-3132-6666

Daniel Waldeck. https://orcid.org/0000-0003-1201-7179

We have no known conflict of interest to disclose.

Due to the nature of this research, participants of this study did not agree for their data to be shared publicly, so supporting data is not available.

Correspondence concerning this article should be addressed to Andrew J. Holliman, Department of Psychology and Human Development, UCL Institute of Education, 25 Woburn Square, London, WC1H 0AA, United Kingdom. Email: a.holliman@ucl.ac.uk
Abstract

In this study, we examined associations between university lecturers’ perceived autonomy support (PAS), adaptability, organisational commitment, and psychological wellbeing. A sample of university lecturers ($N = 102$) from a single ex-polytechnic higher education institution in the United Kingdom completed validated scales for each construct in the spring term. Inspired by prior work in pre-tertiary education with schoolteachers, a conceptual model of predicted relations was developed and tested using structural equation modelling (SEM). Findings showed that PAS was positively associated with lecturers’ adaptability, organisational commitment, and psychological wellbeing; however, adaptability was unable to influence these outcomes independently of its association with PAS. The findings extend prior work with schoolteachers suggesting that, while adaptability is of importance, its influence may be more salient at pre-tertiary level—where there is typically heightened regulation and lower autonomy—and less salient when autonomy options are wider, as is the case in higher education.

Keywords: Adaptability; Autonomy support; Commitment; Wellbeing; University lecturers; Job demands-resources model.
Introduction

Over the last decade, the higher education (university) sector has experienced significant and wide-ranging change (Kinman & Johnson, 2019). This has likely impacted upon the mental health and wellbeing of teaching staff (referred to as ‘university lecturers’ from this point forward), which is reportedly of international concern (Morrish, 2019). Research is warranted therefore, that investigates factors that may influence healthy and effective workplace functioning in this ever-changing environment. In a recent study with high school mathematics teachers, Collie and Martin (2017) found that being able to adjust ones’ thoughts, emotions, and behaviours effectively under these conditions (referred to as adaptability; Martin et al., 2012) and feeling that those in positions of authority (i.e., the school principal in this case) are autonomy-supportive (referred to as perceived autonomy support, PAS; Deci & Ryan, 1987) has a positive influence on organisational commitment (i.e., the loyalty an employee feels towards the institution that they work for; McInerney et al., 2015) and psychological wellbeing. However, it remains to be tested whether these patterns of association would be replicated in higher education, where there are different occupational demands and where autonomy options are wider (that is, where there are greater actual/perceived options and fewer structural constraints). To fill this gap, the present study uses Collie and Martin’s (2017) model (see Figure 1)—that proved effective for pre-tertiary teachers—to examine the predictive roles of adaptability and PAS on the organisational commitment and psychological wellbeing of university lecturers.

A Changing Higher Education Environment

The higher education (university) sector has changed significantly over recent years (Kinman & Johnson, 2019). For example, there has been a rapid rise in student numbers, increased internationalization and the adoption of new learning technologies and alternative teaching methods (e.g., online and blended study), excessive workloads, increased auditing
and metrics (e.g., the Research Excellence Framework, REF, and introduction of the Teaching Excellence Framework, TEF), and more uncertain (precarious) contracts, to name just a few. Such is the magnitude of change, that university lecturers have often felt overloaded and under-resourced (Kinman & Wray, 2019), which has had mostly negative consequences. For example, in a recent report by Morrish (2019), involving 59 higher education institutions in the United Kingdom from 2009 to 2016, it was found that there were rising stress levels, occupational health referrals, and counselling visits, and declining levels of mental health and wellbeing, among university lecturers (see Morrish, 2019). It would seem timely then, to investigate the interplay between the ‘resources’ available to university lecturers and the influence these may have on their psychological functioning at work.

The Job Demands-Resources Model

The job demands-resources model (JD-R model; Schaufeli & Bakker, 2004) was developed to elucidate the relationship between work-related stress and negative outcomes such as job strain (e.g., burnout) and poor motivation (e.g., lack of engagement). Research adopting the JD-R model has been found useful in predicting a range of personal and organisational outcomes across different cultures and occupations (Bakker et al., 2003; Bakker et al., 2010; Brough et al., 2013; Huynh et al., 2014; Schaufeli & Bakker, 2004). It has also helped to explain teachers’ experiences of coping with demands at work (e.g., Bakker et al., 2007; Collie & Martin, 2017; Desrumaux et al., 2015; Hakanen et al., 2006). According to the JD-R model, there are two major components (resources) which can influence the capacity for university lecturers to positively function (and cope with demand) at work; that of job resources and personal resources. Job resources (e.g., performance feedback, participation in decision making, social support; Schaufeli & Bakker, 2004) refer to support (psychological, physical, social, or organisational) that stimulates personal and professional growth, enabling one to meet the demands of the job. Personal resources refer to
personal capacities (or aspects of the self) that enable individuals to have a sense of control and successfully contribute to their work environment (e.g., hope, optimism, self-efficacy, resilience; Grover et al., 2018).

The JD-R model proposes two key processes that help explain occupational outcomes (e.g., wellbeing, organisational commitment) from job and personal resources. First, job resources have a positive influence on an individual’s personal resources (Van den Broeck et al., 2013). For example, research has consistently shown that high levels of perceived principal support (i.e., a job resource) is associated with an increase in teachers’ personal resources such as optimism, work motivation, and emotion regulation ability (Brackett et al., 2010; Desrumaux et al., 2015; Dicke et al., 2017; Fernet et al., 2016). Second, both job and personal resources have been found to act as a buffer against health impairment such that chances of experiencing burnout are reduced. Such protective processes help enable positive outcomes to equip employees (e.g., university lecturers) in addressing their work demands (Van den Broeck et al., 2013).

It should be acknowledged at this point that there is a limited, but expanding literature focused on the impact of personal resources (see Van den Broeck et al., 2013). Indeed, several personal resources have been recently investigated (e.g., self-efficacy, optimism, resilience; see Grover et al., 2018). However, only one study to date has examined the influence of a relatively novel personal resource, that of adaptability (which will be elaborated on in due course). Indeed, Collie and Martin (2017) reported that PAS was positively related to teacher’s adaptability, and in turn, promote positive workplace outcomes (i.e., wellbeing, organisational commitment). With respect to the JD-R model—and partially replicating the model proposed by Collie and Martin (2017)—we positioned PAS as a job resource, adaptability as a personal resource, and well-being and organisational commitment as university lecturers’ outcomes. We introduce these constructs below.
**Perceived Autonomy Support**

Autonomy support is a social-contextual factor assumed to have an important role in enabling employee motivation, wellbeing, and performance (Baard et al., 2004). Autonomy support occurs when an authority figure (e.g., principal, manager) respects and acknowledges the perspective of their subordinate (e.g., teacher, university lecturers), as well as promoting choice and encouraging decision making (Baard et al., 2004; Ware & Kitsantas, 2011). For example, PAS might manifest in practice where a university lecturer feels their manager is listening to and conveying confidence in their ideas regarding course development. Research has consistently demonstrated that when teachers experience high levels of PAS, this helps to foster increased motivation, engagement, and feelings of trust towards their organisation (e.g., Collie et al., 2016; Collie & Martin, 2017; Fernet et al., 2015; Klassen et al., 2012; Lee & Nie, 2014; Leithwood et al., 2008). Indeed, when teachers have a greater sense of trust in their organisation, this is considered to be a significant predictor of their organisation commitment, and in-turn, their turnover intention (Deci et al., 1989; Somech & Bogler, 2002).

It is crucial that the role of PAS is considered as teachers’ personal strengths (e.g., competence, motivation) may still not lead to effective teaching if their work environment is dysfunctional (e.g., over-controlling leadership; Leithwood & Jantzi, 2006). Indeed, when teachers experience lower levels of PAS this has been associated with lower levels of wellbeing and may leave them vulnerable to burnout (Fernet et al., 2012; Nie et al., 2015). In the current study, we examined university lecturers’ PAS. In accordance with the JD-R model, PAS was considered a job resource because it helps university lecturers successfully achieve work targets, navigate challenging work demands, and enables personal and professional development (e.g., Schaufeli & Taris, 2014).

**Adaptability**
According to the tripartite model, adaptability refers to an individual’s capacity to use strategies to adjust and modify (manage) cognitive (thoughts), behavioural (actions) and emotional (affective) functioning in order to respond to changing, novel and uncertain circumstances, situations or events (Martin et al., 2012, 2013). For instance, imagine a university lecturer is required to teach a new class (e.g., leading a qualitative research methods module when their experience is exclusively teaching quantitative research methods), which may cause them some initial anxiety. Adaptability may involve: regulating cognitions such as thinking about how one can modify their pedagogical approach, regulating behaviour, by seeking advice from colleagues who have taught such classes, and regulating emotions such as potential anxiety to avoid procrastination, plan effectively, and deliver teaching content successfully (Collie & Martin, 2016). Although related to other constructs involving emotional regulation (e.g., resilience, buoyancy, and coping), adaptability is considered sufficiently distinct due to the focus on navigating change, novelty, and uncertainty, rather than adjusting from adversity or trauma per se (Martin et al., 2012).

We focused on adaptability as, noted earlier, the higher education (university) sector has changed significantly over recent years (Kinman & Johnson, 2019). Indeed, as job demands are increasing for lecturers (e.g., excessive workloads, learning new teaching software, adjusting to a higher volume of student numbers; Kinman & Wray, 2019), this provides a timely and pressing need to examine the extent to which university lecturers can effectively adapt to such conditions. Indeed, adaptability has been linked with a range of positive teacher and student outcomes. For example, researchers have demonstrated that more adaptable teachers tend to be motivated (Martin et al., 2019), engaged (Collie et al., 2018), and adjust their practices to meet the needs of their students (Loughland & Alonzo, 2019). Of relevance to the current study, Collie and Martin (2017) reported that when teachers were more adaptable, they also tended to report higher levels of psychological wellbeing and
organisational commitment. Moreover, the authors also found that PAS and adaptability were positively interrelated. Thus, in accordance with the JD-R model, when teachers felt autonomy from their principals (i.e., a job resource) this was linked to greater levels of adaptability (i.e., personal resource). Further, the authors reported that PAS was significantly related and indirectly related (through adaptability) to the teacher outcomes.

The current study provided the opportunity to replicate the findings observed by Collie and Martin (2017) whilst utilising the JD-R model (see Figure 1). Specifically, although adaptability was demonstrated to be an important personal resource in a school context, it is unclear as to the extent to which such a resource will extend to university lecturers. Indeed, there are some key differences between pre-tertiary and tertiary education. For example, high school teachers often have less autonomy in following a strict curriculum; but are also required to be more reactive to student needs, thus demonstrating a heightened need for adaptability. In contrast, however, university lecturers arguably have wider autonomy options and therefore, the need for adaptability may be less crucial.

The present study

To address this gap, and to examine the replicability of the findings reported in Collie & Martin (2017), we set out to investigate whether two resources—PAS (job) and adaptability (personal)—impact upon university lecturers’ workplace functioning. Specifically, we examined whether PAS was predictive of adaptability, and whether both constructs were predictive of university lecturers’ organisational commitment and psychological wellbeing. There were two principal research questions:

1. To what extent is PAS positively associated with university lecturers’ adaptability, and are PAS and adaptability positively associated with university lecturers’ organisational commitment and psychological wellbeing?
2. To what extent is PAS indirectly associated with university lecturers’ organisational commitment and psychological wellbeing via adaptability? (RQ2)

**Method**

**Sample and Procedures**

A sample of 102 university lecturers were recruited from a single ex-polytechnic higher education institution in the West Midlands, United Kingdom. All participating lecturers were based in the Faculty of Health and Life Sciences: 2% were from the School of Nursing and Midwifery; 29.4% were from the School of Psychological and Behavioural Sciences; 2% were from the School of Life Sciences; 3.9% were from the School of Health; and 62.7% of respondents did not report their School. Lecturers were 68% female and aged between 27-64 years ($M = 45.29$, $SD = 8.49$). The number of years’ lecturing experience ranged from 0-37 years ($M = 12.28$, $SD = 7.74$).

An email containing a link to the survey was distributed to all eligible lecturers within the Faculty of Health and Life Sciences by their respective Heads of School. A self-selective sample of university lectures then completed an online or paper-based survey comprising items on PAS, adaptability, organisational commitment, and psychological wellbeing. Demographic (i.e., age and gender) and other employment characteristics (i.e., years’ lecturing experience and Faculty/School in which they are based) were also recorded. Psychometric and descriptive statistics for core measures are shown in Table 1 and further elaborated on in Results. All participating lecturers were provided with a participant information sheet and gave informed consent prior to completing the survey. Participants were also given a debrief sheet after data collection.

**Measures**

**Lecturers’ perceived autonomy support (PAS).** We used Klassen et al.’s (2012) adapted version of the Work Climate Questionnaire (Baard et al., 2004) to measure lecturers’
perceived autonomy support. The scale comprised six items and was scored on a 7-point Likert scale from 1 (strongly disagree) to 7 (strongly agree). Subtle amendments were made to items to make them more relevant to our sample; for example, the word ‘principal’ (“I feel understood by my principal”) was changed to ‘management’ (“I feel understood by management”). Cronbach’s alpha in this study, was .95 indicating adequate levels of internal consistency.

**Adaptability.** We used Martin et al.’s (2013) Adaptability Scale to measure lecturers’ cognitive (e.g., “I am able to think through a number of possible options to assist me in a new situation”), behavioural (e.g., “I am able to seek out new information, helpful people, or useful resources to effectively deal with new situations”), and emotional adaptability (e.g., “I am able to reduce negative emotions [e.g., fear] to help me deal with uncertain situations”). The scale comprised nine items and was scored on a 7-point Likert scale from 1 (strongly disagree) to 7 (strongly agree). Previous psychometric work has demonstrated the reliability and validity of the scale, which has also been shown to function well when the three types of adaptability are combined into a global adaptability factor given their high interrelation (e.g., Martin et al., 2012, 2013). Cronbach’s alpha in this study, was .94 indicating adequate levels of internal consistency.

**Lecturers’ organisational commitment.** We used four items from Collie et al.’s (2016) adapted version of Vandenberghhe and Bentein’s (2009) revised Affective Commitment Scale, which focused on the school environment rather than the workplace. The scale was scored on a 7-point Likert scale from 1 (strongly disagree) to 7 (strongly agree). Subtle amendments were made to items to make them more relevant to our sample; for example, the word ‘school’ (“I am proud to belong to this school community”) was changed to ‘university’ (“I am proud to belong to this university community”). Cronbach’s alpha in this study, was .87 indicating adequate levels of internal consistency.
Lecturers’ psychological wellbeing. We used four items from Parker and Martin’s (2009) scale to measure lecturers’ psychological wellbeing (e.g., “I enjoy my work”). The scale was scored on a 7-point Likert scale from 1 (strongly disagree) to 7 (strongly agree). Previous psychometric work has confirmed the reliability and validity of the measure (e.g., Collie & Martin, 2017). Cronbach’s alpha in this study, was .91 indicating adequate levels of internal consistency.

Data Analysis

For preliminary analyses, reliability coefficients, means, standard deviations, skewness, and kurtosis were computed using SPSS Version 22. For the main analyses, confirmatory factor analysis (CFA) and structural equation modeling (SEM) were conducted using Mplus 8.4 (Muthén & Muthén, 2017) with maximum likelihood (MLR) for estimation. The CFA comprised all substantive factors (PAS, adaptability, organisational commitment, and psychological wellbeing) and covariates (gender, age, and years’ teaching experience). SEM was then employed to assess the latent correlations among factors.

Several fit indices were consulted to ascertain model fit (Hu & Bentler, 1999). We report the root mean square error of approximation (RMSEA); however, due to concerns about computing RMSEA with smaller samples (Kenny et al., 2015), we also report the comparative fit index (CFI) and the standardized root mean square residual (SRMR), which account for a smaller sample size (Byrne, 1998). We also provide the Swain (1975) corrected RMSEA and CFI indices, which account for bias in fit when estimating complex models with relatively small samples (Boomsma & Herzog, 2013; Herzog & Boomsma, 2009). Models were considered acceptable at ≤.10 for RMSEA (see MacCallum et al., 1996), ≥.90 for CFI, and ≤.08 for SRMR.

Results
Reliability coefficients, means, standard deviations, skewness, and kurtosis statistics for all variables are shown in Table 1. Reliability coefficients for all variables were acceptable, with all $\alpha$s $\geq .85$. Factor loadings were also sound (ranging from .78 to .99). Skewness and kurtosis values were occasionally high; but this was dealt with using the MLR estimator, which accounts well for nonnormality of distribution.

<<TABLE 1 ABOUT HERE>>

The CFA yielded adequate fit: $\chi^2(36) = 73.46, p < .001$, Swain-corrected RMSEA = .096 (original RMSEA = .10), Swain-corrected CFI = .952 (original CFI = .947), and SRMR = .041. Latent correlations taken from the CFA are shown in Table 2.

<<TABLE 2 ABOUT HERE>>

For the substantive factors, PAS was positively associated with adaptability ($r = .42, p < .001$), organisational commitment ($r = .60, p < .001$), and psychological wellbeing ($r = .51, p < .001$). Adaptability was found to be positively associated with organisational commitment ($r = .32, p < .01$), but not psychological wellbeing ($r = .32, p = .08$). Organisational commitment and psychological wellbeing were also significantly associated ($r = .55, p < .01$).

Turning to the covariates, the only significant association with a substantive factor, was between years’ teaching experience and PAS ($r = -.23, p < .05$) where, perhaps unsurprisingly, greater autonomy support was perceived for lecturers with less teaching experience.

Being a fully forward model (i.e., with the same number of parameters), the SEM fit was identical to the CFA: $\chi^2(36) = 73.46, p < .001$, Swain-corrected RMSEA = .096 (original RMSEA = .10), Swain-corrected CFI = .952 (original CFI = .947), and SRMR = .041. The standard beta estimates for substantive parameters are depicted in Figure 2. For completeness, the standard beta estimates for all parameters are shown in Table 3.

<<FIGURE 2 ABOUT HERE>>
Results showed that PAS predicted greater adaptability ($\beta = .43, p < .001$), organisational commitment ($\beta = .62, p < .001$), and psychological wellbeing ($\beta = .47, p < .001$). Adaptability, however, was unable to influence organisational commitment and psychological wellbeing independently of its association with perceived autonomy support. The covariates were also unable to make a significant independent contribution to the other variables in this study.

**Discussion**

In this study, we examined the influence of PAS (job resource) on adaptability (personal resource), and the influence of PAS and adaptability on organisational commitment and psychological wellbeing among university lecturers. We found that PAS was positively associated with lecturers’ adaptability and was also predictive of organisational commitment and psychological wellbeing independently of its association with adaptability. Adaptability, however, was unable to influence organisational commitment and psychological wellbeing independently of its association with PAS. These findings will be discussed in turn.

**The importance of PAS**

The first key finding was consistent with JD-R theory (Schaufeli & Bakker, 2004), which proposes that job resources—examined here via university lecturers’ PAS—support and stimulate personal and professional growth, enabling one to more effectively meet the demands of the job. University lecturers reporting greater (higher) job resources were also found to have more positive personal resources (e.g., Brackett et al., 2010; Desrumaux et al., 2015; Dicke et al., 2017; Fernet et al., 2016; Van den Broeck et al., 2013); specifically, in this study, when university lecturers perceived those in authority (management) to be autonomy supportive, this had a positive influence on their perceived ability to adjust effectively (cognitively, behaviourally, and emotionally) in situations of novelty and uncertainty (thus,
their adaptability; Martin et al., 2012, 2013). This supports prior work with pre-tertiary schoolteachers (see Collie & Martin, 2017), and may imply that higher education institutions, and management therein (i.e., those with authority), would do well to recognize the importance of the role they play (job resources) in supporting the individual competencies of their employees. Specifically, providing conditions for greater PAS, perhaps in the form of greater flexibility of working patterns and workloads, autonomy over staff appraisals and increased assertiveness, would be beneficial for the lecturers themselves and, in turn, the organisation (via organisational commitment and wellbeing), and their students (the latter association with student attainment found in prior work—see Collie & Martin, 2017—but not investigated here).

Moreover, given the direct effects observed between PAS and both organisational commitment and psychological wellbeing, it would seem that PAS may act as a buffer against health impairment and burnout and potentially supports positive psychological functioning in the workplace (Fernet et al., 2012; Nie et al., 2015; Van den Broeck et al., 2013). It could also be reasonably assumed that PAS enhances ones motivation, engagement, and feelings of trust towards their organisation (e.g., Collie et al., 2016; Collie & Martin, 2017; Fernet et al., 2015; Klassen et al., 2012; Lee & Nie, 2014; Leithwood et al., 2008) which, in turn, may positively influence their organisation commitment (Deci et al., 1989; Somech & Bogler, 2002). This key finding—that PAS directly influences adaptability, organisational commitment, and psychological wellbeing—replicates that observed with pre-tertiary schoolteachers (Collie & Martin, 2017), and shows that this pattern of relations extends beyond pre-tertiary schoolteachers to university lecturers operating in the tertiary education sector.

The ‘Less Salient’ Role of Adaptability
As a great number of changes have occurred over recent years in the higher education (university) sector (see Kinman & Johnson, 2019; Kinman & Wray, 2019), it was reasoned that ones’ ability to personally adapt effectively to these ever-changing conditions may impact upon their organisational commitment and psychological wellbeing; although, this had not yet been empirically examined in the existing literature. Indeed, this was theorized in the JD-R model (Schaufeli & Bakker, 2004) and demonstrated empirically in a sample of pre-tertiary schoolteachers (Collie & Martin, 2017). However, this was not supported by the data in the present study, where adaptability was unable to influence these outcomes independently of its association with PAS. To explain this, it could be that adaptability is less salient in tertiary (university) education as, relative to pre-tertiary education, university lecturers typically have more autonomy, less strict curriculum adherence, and less need to be reactive (e.g., dealing with unexpected behaviour issues in classrooms). It could be then, that with wider autonomy options, there is less need for adaptability in the context of teaching at university. This null finding should, however, be treated with a degree of caution, as the association between adaptability and psychological wellbeing was medium in effect size ($\beta \geq .10$; Keith, 1999) and only marginally non-significant ($p = .08$). Alternatively, as the measure of adaptability in the present study was domain general (and the others domain specific), it may have failed to fully capture the nuances associated with adjustment to change in the context of one’s place of work (university). Further empirical research is warranted to examine the replicability of these novel findings.

**Limitations and Future Directions**

In the study reported here, there are some limitations that will now be acknowledged. First, there was scope to include a more comprehensive battery of assessments for the different components in the JD-R model. For instance, to measure job resources, in addition to PAS, we might have included a measure of performance feedback (Schaufeli & Bakker,
and to measure personal resources, in addition to adaptability, we might have included a measure of optimism or resilience (Grover et al., 2018). Second, all substantive constructs in this study were measured via self-report scales; therefore, there is the risk of potential inaccurate or biased responding (Podsakoff et al., 2012); although, this could conceivably be levelled at most other studies in this area. Moreover, as this research adopted quantitative methodology, there remains scope for more in-depth qualitative research, to help understand the complexities, trends, and nuances for individuals. Such research, may help uncover how and when PAS and adaptability operate, in relation to organisational commitment and psychological wellbeing among university lecturers. Finally, as the study reported here used a concurrent, correlational design, it is not possible to make causal inferences (these can only be conceptually inferred). Future research might therefore consider embracing mixed-methodology and using longitudinal designs, incorporating a more comprehensive assessment battery, to help establish the pathways of influence and elucidate the nuances that may be at play for individuals in this context.

Conclusion

In this paper, we examined whether two resources from the JD-R model (Schaufeli & Bakker, 2004)—PAS (job) and adaptability (personal)—are predictive of each other, and organisational commitment and psychological wellbeing among a sample of university lecturers. Findings showed PAS was predictive of adaptability, organisational commitment, and psychological wellbeing among university lecturers. However, while adaptability yielded a medium effect on psychological wellbeing, it was unable to predict organisational commitment or psychological wellbeing significantly, beyond its association with PAS. Together, these findings extend prior work (e.g., in the pre-tertiary sector with schoolteachers) suggesting that, while adaptability is of importance, its influence may be more salient at pre-tertiary level—where there is typically heightened regulation and lower
autonomy—and less salient when autonomy options are wider, as is the case in higher education. Moreover, there is a strong case here, for organisations, managers, and others in positions of authority, to attend more to job resources, such as PAS, as this resource is associated with important outcomes, such as organisational commitment and psychological wellbeing.

Acknowledgements

We would like to thank the participating higher education institution (university) for supporting the data collection and sharing some of its own data with us. We also gratefully acknowledge the support of the lecturing staff who took part in this research.

References


Figure 1. The conceptual model under examination.
Figure 2. Structural equation modeling involving standardized beta coefficients for substantive parameters.
Table 1
**Descriptive Statistics and Reliability**

<table>
<thead>
<tr>
<th></th>
<th>α</th>
<th>M (or %)</th>
<th>SD</th>
<th>Skewness</th>
<th>Kurtosis</th>
<th>Factor loading range</th>
</tr>
</thead>
<tbody>
<tr>
<td>Gender (M, F)</td>
<td>.75</td>
<td>31%, 69%</td>
<td>.47</td>
<td>-3.37</td>
<td>-2.82</td>
<td>-</td>
</tr>
<tr>
<td>Age</td>
<td>.75</td>
<td>45.29</td>
<td>8.49</td>
<td>-.31</td>
<td>-1.46</td>
<td>-</td>
</tr>
<tr>
<td>Years’ experience</td>
<td>.75</td>
<td>12.28</td>
<td>7.74</td>
<td>2.54</td>
<td>.13</td>
<td>.89-0.97</td>
</tr>
<tr>
<td>PAS</td>
<td>.95</td>
<td>4.21</td>
<td>8.76</td>
<td>-1.64</td>
<td>-1.46</td>
<td>.89-0.97</td>
</tr>
<tr>
<td>Adaptability</td>
<td>.94</td>
<td>5.38</td>
<td>9.18</td>
<td>-4.97</td>
<td>4.68</td>
<td>.86-0.99</td>
</tr>
<tr>
<td>Wellbeing</td>
<td>.91</td>
<td>5.25</td>
<td>4.74</td>
<td>-2.79</td>
<td>-.47</td>
<td>.78-0.97</td>
</tr>
<tr>
<td>Organisational commitment</td>
<td>.87</td>
<td>4.69</td>
<td>5.58</td>
<td>-4.49</td>
<td>2.49</td>
<td>.87-0.93</td>
</tr>
</tbody>
</table>

*Note: Age was measured in years; Years’ experience = years’ lecturing experience; Perceived Autonomy Support (PAS), Adaptability, Wellbeing, and Organisational commitment were scored from 1-7 with higher scores corresponding to more positive levels of each construct. Gender, Age, and Years’ Experience are single-item indicators and so reliability is not computed.*
Table 2
*Latent Correlations Among Variables*

<table>
<thead>
<tr>
<th>Gender (M/F)</th>
<th>Gender</th>
<th>Age</th>
<th>Years’ experience</th>
<th>PAS</th>
<th>Adaptability</th>
<th>Wellbeing</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age</td>
<td>.20*</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Years’ experience</td>
<td>.06</td>
<td>.51***</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PAS</td>
<td>-.03</td>
<td>-.18</td>
<td>-.23*</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Adaptability</td>
<td>.11</td>
<td>-.11</td>
<td>-.01</td>
<td>.42***</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Wellbeing</td>
<td>.06</td>
<td>-.08</td>
<td>-.11</td>
<td>.51***</td>
<td>.32</td>
<td></td>
</tr>
<tr>
<td>Organisational commitment</td>
<td>.15</td>
<td>-.01</td>
<td>.01</td>
<td>.60***</td>
<td>.32**</td>
<td>.55**</td>
</tr>
</tbody>
</table>

*Note. Years’ experience = years’ lecturing experience; PAS = perceived autonomy support.  
*p<.05, **p<.01, ***p<.001*
Table 3

*Standardized Beta Coefficients from SEM*

<table>
<thead>
<tr>
<th></th>
<th>PAS</th>
<th>Adaptability</th>
<th>Well-being</th>
<th>Org. commitment</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Lecturer covariates</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Gender (M/F)</td>
<td>-.01</td>
<td>.15</td>
<td>.06</td>
<td>.16</td>
</tr>
<tr>
<td>Age</td>
<td>-.08</td>
<td>-.14</td>
<td>.02</td>
<td>-.01</td>
</tr>
<tr>
<td>Years’ experience</td>
<td>-.19</td>
<td>.16</td>
<td>-.02</td>
<td>.14</td>
</tr>
<tr>
<td><strong>Lecturer variables</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PAS</td>
<td></td>
<td></td>
<td>.43***</td>
<td>.62***</td>
</tr>
<tr>
<td>Adaptability</td>
<td></td>
<td></td>
<td>.47***</td>
<td></td>
</tr>
</tbody>
</table>

*Note. Years’ experience = years’ lecturing experience; PAS = perceived autonomy support; Org. commitment = organisational commitment.  
*p<.05, ** p<.01, *** p<.001.*