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Cohort study of preoperative chronic beta blocker prescription in elderly patients as a risk factor for postoperative mortality stratified by preoperative blood pressure.

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| Complete List of Authors: | Venkatesan, Sudhir; University of Nottingham, Division of Epidemiology and Public Health, School of Medicine Joergensen, Mads; University of Copenhagen, The Cardiovascular Research Center Manning, Helen; University of Wisconsin Madison, Obstetrics & Gynecology Andersson, Charlotte; University of Copenhagen, Internal Medicine Mozid, Abdul; The London Chest Hospital, Department of Cardiology Coburn, Mark; University Hospital RWTH Aachen, Department of Anaesthesiology; Moonesinghe, Suneetha; UCL / UCLH Surgical Outcomes Research Centre, Anaesthesia and Intensive Care Foëx, Pierre; John Radcliffe, Anaesthetics Mythen, Michael; UCL, Centre for Anaesthesia Grocott, Michael; University of Southampton, Anaesthesia and Critical Care Research Unit Hardman, Jonathan; Nottingham University Hospitals NHS Trust, Department of Anaesthesia Myles, Puja; University of Nottingham, Division of Epidemiology and Public Health, School of Medicine Sanders, Robert; University of Wisconsin, Anesthesiology |
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SCHOLARONE™ Manuscripts Cohort study of pPreoperative chronic beta blocker prescription in elderly patients as a risk factor for postoperative mortality stratified by preoperative blood pressure: a cohort study

Sudhir Venkatesan MPH1, -Sudhir.Venkatesan@nottingham.ac.uk

Mads Emil Jørgensen MD², Mads.Emil.Joergensen@regionh.dk

Helen J. Manning MRCOG³, hmanning@doctors.org.uk

Charlotte Andersson MD4, ca@heart.dk

Abdul M. Mozid MRCP5, ammozid@hotmail.com

Mark Coburn⁶, mcoburn@ukaachen.de

S. Ramani Moonesinghe MRCP7, rmoonesinghe@googlemail.com

Pierre Foex FRCA⁸, pierre.foex@ndcn.ox.ac.uk

Monty Mythen FRCA7, m.mythen@ucl.ac.uk

Michael PW Grocott FRCP9, mike.grocott@soton.ac.uk

Jonathan G. Hardman FRCA¹⁰, J.Hardman@nottingham.ac.uk

Puja R. Myles PhD1, Puja.Myles2@nottingham.ac.uk

Robert D. Sanders FRCA¹¹, rsanders4@wisc.edu

- Division of Epidemiology and Public Health, School of Medicine, University of Nottingham, Nottingham, United Kingdom. UK
- 2. Department of Internal Medicine, Division of Cardiology, Glostrup Hospital, University of Copenhagen, Denmark-
- 3. Department of Obstetrics & Gynecology, University of Wisconsin, Madison, WI, USA-
- 4. The Cardiovascular Research Center, Gentofte Hospital, University of Copenhagen,

 Denmark
- 5. Department of Cardiology, Bristol Heart Institute, Bristol, United Kingdom. UK
- 6. Department of Department of Anaesthesia, Medical Faculty, RWTH Aachen University,
 Germany
- 7. University College London Hospitals & National Institute of Health Biomedical Research Centre, United Kingdom. UK
- 8. Nuffield Division of Anaesthetics, Oxford University Hospital, Oxford, United Kingdom. UK
- 9. Integrative Physiology and Critical Illness, Clinical and Experimental Sciences, Faculty of Medicine, University of Southampton; University Hospital Southampton NHS Foundation

 Trust; Southampton NIHR Respiratory Biomedical Research Unit, United Kingdom. UK
- 10. Department of Anaesthesia, University of Nottingham, United Kingdom. UK
- 11. Anesthesiology & Critical Care Trials & Interdisciplinary Outcomes Network (ACTION),

 Department of Anesthesiology, University of Wisconsin School of Medicine and Public

 Health, Madison, WI, USA-

Brief Title: Preoperative medications and surgical mortality

*Corresponding author:__Dr Robert D Sanders, Department of Anesthesiology, University of
Wisconsin School of Medicine and Public Health, 600 Highland Avenue, B6/319 CSC Madison,
WI 53792-3272 Telephone: 608-263-8100 Fax: 608-263-0575 Madison, USA. Email:
robert.sanders@wisc.edu

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The role of preoperative antihypertensive drug therapy as a risk factor for postoperative mortality as a function of preoperative arterial pressure was investigated in a large retrospective data set.

A propensity score-matched cohort study of primary care data from the UK Clinical Practice

Research Datalink including 84,633 elderly patients 65 years or over analysed mortality

following elective noncardiac surgery.

Beta blockers were associated with increased risk of 30-day mortality in patients with elevated preoperative blood pressure.

Renin angiotensin system inhibitors, calcium channel blockers, and loop diuretics were not associated with mortality, while thiazides and statins were associated with reduced risk.

<u>Prospective randomised trials are needed to confirm these findings with important implications for perioperative management.</u>

Abstract

Background: Recent data suggest that beta blockers are associated with increased perioperative risk in hypertensive patients. We investigated whether beta blockers were associated with increased risk in elderly patients with raised preoperative <u>arterial</u> blood pressure (BP).

Methods: We conducted a propensity score-matched cohort study of primary care data from the United Kingdom Clinical Practice Research Datalink (2004-2013) including 84,633 patients aged 65 years or over. Conditional logistic regression models, including factors that were significantly associated with the outcome, were constructed for 30-day mortality following elective, non-cardiac surgery. The effects of beta blockers (primary outcome), renin angiotensin system (RAS) inhibitors, calcium channel blockers, thiazides, loop diuretics and statins were investigated at systolic, and diastolic BP thresholds.

Results: Beta blockers were associated with an-increased odds of postoperative 30-day mortality in patients with systolic hypertension (defined as systolic BP >140 mmHg, adjusted odds ratio [aOR]: 1.92, 95% CI: 1.05-3.51). After excluding patients for whom prior data suggest benefit from perioperative beta blockade (patients with prior myocardial infarction or heart failure), rather than adjusting for them, the point estimate shifted slightly (aOR: 2.06, 95% CI: 1.09-3.89). Compared to no use, statins (aOR: 0.35; 95% CI: 0.17-0.75) and thiazides (aOR: 0.28; 95% CI: 0.10-0.78) were associated with lower mortality in patients with systolic hypertension.

Conclusions: These data suggest that the safety of perioperative beta blockers may be influenced by preoperative BP thresholds. A <u>randomized randomised</u> controlled trial of beta blocker withdrawal, in select populations, is required to identify <u>any a causal relationship</u>.



Introduction

For many years, beta blockers were considered as protective medications in the perioperative period, however, recent trials and meta-analyses have challenged this notion¹⁻³. Most notably the POISE study suggested that de novo institution of a high dose of slow-slow-release metoprolol may reduces the risk of myocardial infarction but increases the risk of stroke, sepsis and mortality⁴. Accumulating data from observational studies suggests that beta blockers exert a class effect⁵ and may be harmful in low_low_risk patients (defined as low revised cardiac risk index scores^{6,7}) but beneficial in those with a recent myocardial event or heart failure⁸ (reviewed in ref 3for review see³). When excluding high high risk cardiac patients, Jorgensen et al. recentlyand colleagues⁹ found that among hypertensive patients, use of beta blockers was associated with increased risk of postoperative major adverse complications⁹. Based on this work, our primary hypothesis herein was to verify whether that beta blockers were are associated with increased 30-day mortality in patients with numerically raised high preoperative blood pressure. The working hypothesis is that in patients with hypertension, the effects on cardiac output, renin-angiotensin system (RAS) and vascular tone by beta blockers in the perioperative period, outweigh the benefit of any direct cardioprotective effectbenefit. As secondary hypotheses aims we tested the association of other cardiovascular medications and postoperative 30-day mortality at different blood pressure thresholds defined as hypotension (<80 mm Hg diastolic or <120 mm Hg systolic), normotension (80-89 mm Hg diastolic or 120-139 mm Hg systolic), and or hypertension (>90 mm Hg diastolic or >140 mm Hg systolic). There are is limited guidance on the impact of other cardiovascular medications in the perioperative period, though RAS inhibitors are suggested to be withheld in some situations¹⁰. It is important to note that in our recent analysis of these data, hypertensive <u>level</u> BP values in the elderly were not associated with <u>increased</u> 30-day mortality¹¹. This *a priori* planned secondary analysis focuses on whether cardiovascular medications <u>may be are</u> an important determinant of perioperative outcomes at different BP values in <u>the</u> elderly <u>patients undergoing elective noncardiac surgery</u>.



Methods

Data source and study design

The data source for this study was the <u>UK</u> Clinical Practice Research Datalink (CPRD) which is a <u>UK</u>-primary care database representing about 6% of the country's population. We used medical codes (as listed in the Appendix of our original paper¹²) to identify patients who underwent specific non-cardiac surgeries between 1st January 2004 and 31st December 2013. We <u>then</u>-retained longitudinal data for patients aged 65 years and over, and who were registered at the GP practice for <u>at least></u> 1 year prior to their elective non-cardiac surgery. This study was approved by the Independent Scientific Advisory Committee (ISAC) for the Medicines and Healthcare Products Regulatory Agency, UK (ISAC protocol number: 11_138A).

Exposure variables

Preoperative cardiovascular medications were our the main exposure variables. They included beta blockers (primary outcome) and as secondary outcomes of statins, calcium-channel blockers, angiotensin converting enzyme inhibitors and angiotensin 2 receptor blockers referred to as RAS inhibitors, thiazide diuretics, and loop diuretics prescribed within 30 days of surgery (compared to non-users).

Outcome variable

Our outcome variable was perioperative mortality, defined as death occurring within 30 days following non-cardiac surgery. In the UK, death certificates issued by the GP are entered directly into the primary care database. Potential biases that may arise from misclassification

of death are likely to be non-differential, i.e similar between both groups:— death and survival.

We have discussed this in more detail in our previous publication ¹².

Covariates

We included the following as covariates in our multivariable models: age, gendersex, alpha2 agonists, aspirin, other antiplatelet agents, atrial fibrillation, unstable angina, valvular heart disease, myocardial infarction, cerebrovascular disease, peripheral vascular disease, chronic pulmonary disease (including asthma), heart failure, diabetes mellitus, renal disease, liver disease, cancer, body mass index (BMI) as a categorical variable (<18.5, 18.5–24.99, 25–29.99 and >30 kg m⁻²), smoking status, alcohol consumption, and socioeconomic status [using the 2010 Index of Multiple Deprivation (IMD) scores in quintiles]. We used the surgical risk score to adjust for the varying levels of risk posed by the various included surgery types. Our surgical procedural risk score was based on the a validated surgical risk scale^{13,14} and was included as an ordinal categorical variable ranging from 1 to 5 with 1 a low-low-risk procedure and 5 a high-high-risk procedure.

Statistical analysis

Based on our prior study¹¹, we considered the most recent BP measurement prior to surgery and stratified our study population by BP thresholds as follows: hypotension (<80mm Hg diastolic or <120 mm Hg systolic), normotension (80-89 mm Hg diastolic or 120-139 mm Hg systolic), and hypertension (>90mm Hg diastolic or >140 mm Hg systolic)above. We first computed propensity scores¹⁵ using multivariable logistic regression models for each of our six drugs of interest (statins, beta blockers, calcium-channel blockers, RAS inhibitors, thiazide diuretics, and loop diuretics) using the method described by of Hirano and Imbens¹⁶.

Covariates that informed the propensity score derivation models were the following comorbidities: myocardial infarction, unstable angina, heart failure and atrial fibrillation. We then generated propensity score quintiles for each drug of interest, and matched individuals on propensity score quintile, using an interval matching approach¹⁷, with a minimum 1:1 variable matching ratio.

For each of the BP threshold, we performed conditional logistic regression to investigate the association between preoperative cardiovascular medication and postoperative mortality. We ran unadjusted and adjusted models for each of the five-six cardiovascular drugs separately comparing the effect of exposure to a given medication to non-exposure. In our adjusted models we included those covariates that were statistically significantly (p-value<0.05) associated with postoperative mortality. Conditional logistic regression models for each of the five-six drugs had the following covariates in common as well as their individual propensity score: age, aspirin, other antiplatelet agents, atrial fibrillation, unstable angina, MI, cerebrovascular disease, peripheral vascular disease, chronic pulmonary disease, diabetes mellitus, renal disease, cancer, BMI, smoking status and surgical score. Additionally, (Ssignificantly associated) covariates included for specific models are presented below:

Statins: gendersex, beta blockers, RAS inhibitors, Cacalcium-channel blockers, loop diuretics, heart failure, alcohol consumption, number of BP measurements, and IMD 2010 scores.

Beta blockers: gendersex, statins, RAS inhibitors, Cacalcium-channel blockers, loop diuretics, heart failure, liver disease, alcohol consumption, number of BP measurements, and IMD 2010 scores.

Ca<u>lcium</u>-channel blockers: <u>gendersex</u>, beta blockers, RAS inhibitors, statins, loop diuretics, heart failure, and alcohol consumption.

RAS inhibitors: gendersex, statins, beta blockers, Ca-channel blockers, loop diuretics, heart failure, liver disease, alcohol consumption, number of BP measurements, and IMD 2010 scores.

Thiazides: statins, beta blockers, Cacalcium-channel blockers, loop diuretics, RAS inhibitors, liver disease, number of BP measurements, and IMD 2010 scores.

Loop diuretics: gendersex, statins, beta blockers, Cacalcium-channel blockers, thiazides, RAS inhibitors, heart failure, liver disease, alcohol consumption, number of BP measurements, and IMD 2010 scores.

Drugs were omitted from the adjusted analysis where if the sample size was insufficient. Interaction terms were specified between all drug covariates included in the multivariable model to account for the impact of drug combinations. All analyses were performed using Stata 15.0 (StataCorp, College Station, TX, USA).

Results

Our The study population included 84,633 patients aged 65 years or over who underwent elective, non-cardiac surgery (Figure 1). For the systolic BP thresholds, there were: 7,924 hypotensive patients (89.9% were also diastolic hypotensive), 34,531 normotensive patients (36.4% were also diastolic normotensive) and 41,527 hypertensive patients (17.7% were also diastolic hypertensive; Supplementary Figure 1 and Supplementary Table 1). For the diastolic BP thresholds, there were: 42,821 hypotensive patients, 32,721 normotensive patients, and 8,422 hypertensive patients. There were 495 Four hundred nighty five (1.16%) events of postoperative mortality were events recorded in our study population (Table 1). Systolic BP was missing for 651 patients and diastolic BP was missing for 669 patients.

Of those-patients who had received cardiovascular drugs, 15,578 (18.4%) only had only one prescribed cardiovascular drug, and 21,870 (25.8%) patients did not have any prescriptions for cardiovascular drugs (**Table 1**). There were 12,148 patients in our study population had received a current prescription for beta blockers. Of these, atenolol 50 mg was the most commonly prescribed beta blocker (29.6% of all beta blocker prescriptions), followed by atenolol 25 mg (17.3% of all beta blocker prescriptions) and bisoprolol 2.5 mg (10.5% of all beta blocker prescriptions).

Primary Outcome

Our primary pPropensity-matched analysis showed that, after adjustment for statistically significant confounders, in patients with systolic hypertension, beta blockers were associated with a statistically significant increase in odds of postoperative mortality (adjusted odds ratio (aOR): 1.92, 95% confidence intervals (95% CI): 1.05-3.51; Figure 2).

Secondary Outcomes

Our In a secondary analysis, that should be considered at most hypothesis hypothesis-generating, showed that in the systolic hypotensive group, statins were associated with a statistically significant decrease in the adjusted odds of postoperative mortality (aOR: 0.35, 95% CI: 0.17-0.75) in the systolic hypotensive group. For patients with systolic hypertension, statins (aOR: 0.35; 95% CI: 0.17-0.75; Figure 2) and thiazides (aOR: 0.28; 95% CI: 0.10-0.78; Figure 2) were associated with a protective effect on postoperative mortality. No significant results were observed based on diastolic BP thresholds.

Sensitivity Analyses

In 2011, beta blockers were demoted_lowered_from first-line treatment to third-line treatment for hypertension¹⁸. In order to address any resulting confounding, we performed a *post_hoc* sensitivity analysis in those patients who underwent non-cardiac surgery prior to 31 December 2011. There_wereOf 69,686 such patients, of whom 9,952 (14.3%) had been prescribed beta blockers. In this population, we found that beta blockers were associated with a two-fold increase in postoperative mortality in the systolic hypertension group (aOR: 2.07; 95% CI: 1.09 to 3.95).

Since confounding by indication could influence the observed associations above and beta blockers are associated with perioperative protection in patients with heart failure and prior myocardial infarction⁸, we conducted a *post hoc* secondary analysis by excluding patients with heart failure (n=3,063) and running our models again. In the remaining 81,570 patients, beta blockers remained statistically significant in systolic hypertension (aOR: 2.13; 95% CI: 1.18 to

3.85; **Table 2**). We conducted further analyses excluding patients with prior acute myocardial infarction (aOR: 1.95; 95% CI: 1.04 to 3.67) or both heart failure and acute myocardial infarction (aOR: 2.06; 95% CI: 1.09 to 3.89). In these groups, significant associations were observed between beta blocker <u>uses</u> and mortality in patients with systolic hypertension (**Table 2**).



Discussion

Our main finding confirmed our the hypothesis that beta blockers were are associated with increased perioperative mortality in patients with raised blood pressure. This hypothesis was based on the recent finding that beta blockers may be harmful in hypertensive patients on based determined with on a different dataset⁹. The effects of beta blockers on raised systolic BP remained throughout sensitivity analyses adjusted for year of administration and confounding from by indication of for heart failure or prior MI.

Before further inference, it is essential to note that due to the observational nature of this study, our the data do not address whether beta blocker withdrawal (or non-compliance) may be responsible for these findings as we do not have data on within hospital administration of beta blockers. This is an important limitation of our data though beta blocker continuation has been advocated by guidelines throughout the study period. Furthermore, exclusion of patients for whom continuation of beta blockers are thought to be critical (, high high risk patients with prior MI or heart failure), did not affect our the results. Our data support the notion that continuing beta blocker exposure in the perioperative period may not be advantageous in all patients as we have discussed in our a recent review³. Further epidemiological evidence is required to understand which patients may benefit from beta blocker withdrawal and who which may come to harm. However, the only way to address these issues definitively will be to conduct an appropriately powered randomized randomized controlled trial, likely targeting lower risk patients.

Beta blocker withdrawal has been associated with adverse outcomes in important epidemiological studies^{6,19}, though a more recent study provided less clear evidence with both increased mortality and lower morbidity²⁰. Our data highlight that it is unclear whether the risk/benefit ratio is the same for all levels of patient risk, with patient factors such as BP influencing the potential risk, supporting our prior work^{8,9}. One hypothesis may be, that patients with hypertension may be are vulnerable to swings in perioperative blood pressure BP due to increased vascular stiffness, or suppression of the renin-angiotensin system, combined with higher BP thresholds for organ autoregulation⁹. Simultaneously, wWe also hypothesize that they are of low enough cardiac risk so to not benefit at a population level from the cardioprotective effects of beta blockers. Hence the hypothesis is that the risk/benefit ratio of the medications is unclear in this population, though we acknowledge that we are merely at the beginning ofearly in evaluating this hypothesis and that much work is required. However, this study validates the prior finding that beta blockers may can be associated with harm in patients with hypertension⁸. Our approach leveraged the availability of preoperative BP measures that are often used by anaesthesiologists to gauge perioperative risk. We acknowledge that this is slightly discordant with the study by Jorgensen et al. and colleagues 6, in which BP measures were not available. As such oour results suggest that it is not the diagnosis of hypertension that is most critical but the actual numerical value of blood pressure (i.e. anysuch that increased risk may not pertain to well-well-controlled hypertensive patientss).

In our secondary analyses, we observed that chronic statins and thiazide diuretics were associated with protective effects against postoperative mortality in elderly patients with

systolic hypertension. If true (as these are hypothesis generating analyses) findings are confirmed, we suspect that patients on statins may have benefited from improved autoregulation, anti-inflammatory action and/or organ protection that statins are thought to afford²¹. While many observational studies have suggested that statins may improve perioperative outcomes^{22,23}, the recent Lowering the Risk of Operative Complications Using Atorvastatin Loading Dose (LOAD) randomized randomised controlled trial did not suggest benefit of *de novo* institution of statins in the perioperative period (and showed no trend to of benefit in elderly patients)²⁴. Hence the actual benefit of *de novo* perioperative statin therapy is unproven. The finding that thiazides may be protective is harder to explain. Nonetheless the finding with thiazidesbut should not necessarily be dismissed as we recently observed that thiazides are relatively protective (compared to other anti-hypertensive medications) in a cohort of Danish patients⁹. If real, any their protective effect may relate to their small stimulatory effect on the renin angiotensin system that these drugs afford²⁵.

Limitations

Our data suffers from the limitations that affect all observational analyses,—particularly that causation cannot be proven. Causality can only be concluded in the setting of experimental studies such as randomized randomised controlled trials. Furthermore, sSelection bias and confounding by indication cannot be excluded from influencing these results, hence the importance of specific hypotheses and the conduct of sensitivity analyses (such as excluding patients with heart failure). For example, it is possible that beta blockade in patients without prior MI or heart failure constitutes resistant hypertension. In this context, the risk may be conferred by the underlying pathophysiology, not the drug itself. Nonetheless gGiven the

accumulating data suggesting poor adverse outcomes may beare associated with beta blocker exposure in lower risk populations, it appears a randomized randomised controlled trial of the safety of beta blockers is warranted. Further, it is important to note that similar to all observational studies, the analysis is vulnerable to unmeasured confounding such as from variables on which we lacked data such as including comorbidities such as stable angina or specific subtypes of heart failure. Similar to many other perioperative epidemiology studies, we also lack detailed data on perioperative events that will also influence postoperative mortality, such as. These perioperative events include non-compliance and withdrawal of medication. In particular, feurther information is required needed from epidemiological datasets about the withdrawal of perioperative medications. Hence, we regard our data as hypothesis generating and require requiring confirmation in future epidemiological studies and randomized randomised controlled trials.

Conclusions

Our data suggest that beat blockers may be associated with increased risk of mortality at raised preoperative blood pressure thresholds in the elderly patients undergoing elective noncardiac surgery. Future epidemiological studies and randomized randomised trials should consider analysing results outcomes based on preoperative BP-blood pressure thresholds.

Authours' Contributors contributions

RDS, SV and PM designed the research question and study analysis plan with input from the co-authors. SV performed the analysis with input from PM and RDS. SV had full access to all

of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. All co-authors advised on the analyses. RDS and SV wrote the manuscript with significant input from PM. All authors advised on the manuscript content and contributed to editing and scientific direction. All authors approved the final manuscript.

Declaration of Interests

All authors have completed the Unified Competing Interest form at www.icmje.org/coi_disclosure.pdf (available on request from the corresponding author) and declare no competing interests that may be relevant to the submitted work.

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Figure Legends

Figure 1: STROBE diagram

Figure 2: Forest plots for the impact of various preoperative cardiovascular medications on postoperative mortality for different preoperative systolic blood pressure thresholds. Top row shows unadjusted data and bottom row shows the results after adjusting for confounders. Columns refer to different blood pressure thresholds.

Figure 3: Forest plots for the impact of various preoperative cardiovascular medications on postoperative mortality for different preoperative diastolic blood pressure thresholds. Top row shows unadjusted data and bottom row shows the results after adjusting for confounders. Columns refer to different blood pressure thresholds.

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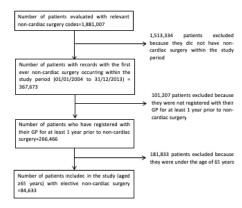


Figure 1: STROBE diagram

338x190mm (54 x 54 DPI)

Table 1: Demographic and clinical Subject characteristics. *'0' refers to those patients who never received a prescription for any of the cardiovascular drugs of interestexamined.

| | All | Survived | 30-day | p-value |
|-------------------------------------|---------------|---------------|------------|---------|
| | (n=84,633) | (n=84,138) | mortality | |
| | | | (n=495) | |
| Mean age, in yea rs (SD) | 74.7 (6.9) | 74.7 (6.9) | 81.5 (8.4) | <0.001 |
| Sex | | | | |
| Male (%) | 44,349 (52.4) | 44,088 (52.4) | 261 (52.7) | |
| Female (%) | 40,284 (47.9) | 40,050 (47.6) | 234 (47.3) | 0.884 |
| Body Mass Index (m kg-2) | | | | |
| Underweight (<18.5) | 1,528 (1.8) | 1,495 (1.8) | 33 (6.7) | |
| Normal range (18.5 to 24.99) | 25,107 (29.7) | 24,914 (29.6) | 193 (39.0) | |
| Overweight (25 to 29.99) | 31,194 (36.9) | 31,060 (36.9) | 134 (27.1) | |
| Obese (≥30) | 18,992 (22.4) | 18,937 (22.5) | 55 (11.1) | |
| Missing | 7,812 (9.2) | 7,732 (9.2) | 80 (16.2) | <0.001 |
| Smoking status | | | | |
| Non-smoker | 43,780 (51.7) | 43,557 (51.8) | 223 (45.1) | |
| Current Smoker | 8,371 (9.9) | 8,297 (9.9) | 74 (15.0) | |
| Ex-smoker | 31,193 (36.9) | 31,007 (36.9) | 186 (37.6) | |
| Missing | 1,289 (1.5) | 1,277 (1.5) | 12 (2.4) | <0.001 |
| Alcohol consumption status | | | | |
| Below limit | 38,937 (46.0) | 38,753 (46.1) | 184 (37.2) | |
| Above limit | 5,858 (6.9) | 5,836 (6.9) | 22 (4.4) | |
| Missing | 39,838 (47.1) | 39,549 (47.0) | 289 (58.4) | <0.001 |
| Comorbidities | | | | 4 |
| Atrial fibrillation (%) | 6,934 (8.2) | 6,845 (8.1) | 89 (18.0) | <0.001 |
| Other cardiac arrhythmia (%) | 48 (0.1) | 48 (0.1) | 0 (0) | 0.595 |
| Unstable angina (%) | 1,097 (1.3) | 1,086 (1.3) | 11 (2.2) | 0.068 |
| Valvular heart disease (%) | 28 (0.03) | 28 (0.03) | 0 (0) | 0.685 |
| Myocardial infarction (%) | 7,671 (9.1) | 7,586 (9.0) | 85 (17.2) | <0.001 |
| Congestive heart disease (%) | 3,063 (3.6) | 2,998 (3.6) | 65 (13.1) | <0.001 |
| Peripheral vascular disease (%) | 5,704 (6.7) | 5,618 (6.7) | 86 (17.4) | <0.001 |
| Cerebrovascular disease (%) | 5,864 (6.9) | 5,795 (6.9) | 69 (13.9) | <0.001 |
| Chronic pulmonary disease (%) | 17,023 (20.1) | 16,902 (20.1) | 121 (24.4) | 0.016 |
| Liver disease (%) | 390 (0.5) | 385 (0.5) | 5 (1.0) | 0.070 |

Commented [HCH1]: Add range

| D: 1 | 10.500 (12.5) | 10 (17 (12 () | 72 (4.4.0) | 0.455 |
|-----------------------------------|---------------|---------------|------------|------------------|
| Diabetes mellitus (%) | 10,690 (12.6) | 10,617 (12.6) | 73 (14.8) | 0.155 |
| Renal disease (%) | 11,421 (13.5) | 11,305 (13.4) | 116 (23.4) | <0.001 |
| Cancer (%) | 20,036 (23.7) | 19,853 (23.6) | 183 (37.0) | <0.001 |
| Statins | 21,617 (25.5) | 21,532 (25.6) | 85 (17.2) | 0.001 |
| Beta blockers | 12,148 (14.4) | 12,078 (14.4) | 70 (14.1) | 0.401 |
| ACE-Angiotensin converting enzyme | 20,888 (24.7) | 20,799 (24.7) | 89 (18) | 0.192 |
| inhibitors | | | | |
| Calcium channel blockers | 11,786 (13.9) | 11,743 (13.9) | 43 (8.7) | 0.005 |
| Alpha-2 agonists | 1,145 (1.4) | 1,137 (1.4) | 8 (1.6) | 0.611 |
| Thiazide diuretics | 11,657 (13.8) | 11,627 (13.8) | 30 (6.1) | <0.001 |
| Loop diuretics | 6,047 (7.1) | 5,977 (7.1) | 70 (14.1) | <0.001 |
| Aspirin | 35,336 (41.8) | 35,052 (41.7) | 284 (57.4) | <0.001 |
| Other antiplatelet drugs | 6,941 (8.2) | 6,877 (8.2) | 64 (12.9) | <0.001 |
| Number of cardiovascular drugs* | | | | |
| 0 | 21,870 (25.8) | 21,771 (25.9) | 99 (20) | |
| 1 | 15,578 (18.4) | 15,489 (18.4) | 89 (18) | |
| 2 | 13,385 (15.8) | 13,327 (15.8) | 58 (11.7) | |
| 3 | 8,773 (10.4) | 8,735 (10.4) | 38 (7.7) | |
| 4 | 3,200 (3.8) | 3,183 (3.8) | 17 (3.4) | |
| 5 | 538 (0.6) | 538 (0.6) | 0 (0) | |
| 6 | 11 (0.01) | 11 (0.01) | 0 (0) | 0.327 |
| Surgical risk score | | | | |
| Score-1 | 0 (0) | 0 (0) | 0 (0) | |
| Score-2 | 8,142 (9.6) | 8,114 (9.6) | 28 (5.7) | |
| Score-3 | 15,033 (17.8) | 14,993 (17.8) | 40 (8.1) | \boldsymbol{O} |
| Score-4 | 8,574 (10.1) | 8,533 (10.1) | 41 (8.3) | |
| Score-5 | 52,884 (62.5) | 52,498 (62.4) | 386 (78.0) | <0.001 |
| Socio-economic status | | | | |
| (IMD 2010 quintiles) | | | | |
| 1 | 13,153 (15.5) | 13,101 (15.6) | 52 (10.5) | |
| 2 | 13,491 (15.9) | 13,409 (15.9) | 82 (16.6) | |
| 3 | 10,104 (11.9) | 10,026 (11.9) | 78 (15.8) | |
| 4 | 7,745 (9.15) | 7,688 (9.1) | 57 (11.5) | |
| 5 | 5,302 (6.3) | 5,261 (6.3) | 41 (8.3) | |
| Missing | 34,838 (41.2) | 34,653 (41.2) | 185 (37.4) | <0.001 |

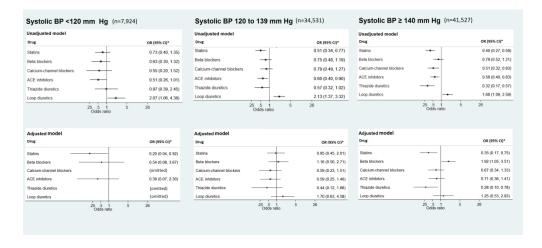


Figure 2: Forest plots for the impact of various preoperative cardiovascular medications on postoperative mortality for different preoperative systolic blood pressure thresholds. Top row shows unadjusted data and bottom row shows the results after adjusting for confounders. Columns refer to different blood pressure thresholds.

Table 2: Sensitivity analysis: Impact of beta blockers on postoperative mortality at various BP arterial pressure thresholds in patient groups excluding specific disease groups.

| | | lial myocardial minfarction | No Heart <u>heart</u> <u>Failure</u> <u>failure</u> | | No Myocardial myocardial Infarction infarction or Heart heart Failure | |
|-----------------|----------------------|--------------------------------|--|-------------------|---|-------------------|
| Systolic | Unadjusted | Adjusted | Unadjusted | Adjusted | Unadjusted | Adjusted |
| threshold | OR (95% CI) | OR (95% CI) | OR (95% CI) | OR (95% CI) | OR (95% CI) | OR (95% CI) |
| <120_mmHg | 0.79 (0.33- 0.89) | 0.28 (0.04-1.84) | 0.60 (0.26-1.43) | 0.28 (0.04-1.84) | 0.61 (0.22-1.73) | 0.28 (0.04-1.84) |
| 120-139 mmHg | 1.07 (0.64- 1.77) | 1.31 (0.50-3.43) | 1.18 (0.75-1.87) | 1.29 (0.54-3.11) | 1.11 (0.65-1.90) | 1.44 (0.50-4.13) |
| ≥140_mmHg | 0.91 (0.57- 1.47) | 1.95 (1.04-3.67)* | 0.89 (0.57-1.41) | 2.13 (1.18-3.85)* | 0.83 (0.50-1.38) | 2.06 (1.09-3.89)* |
| *p-value <0.05 | | | | | | |

^{*}p-value < 0.05

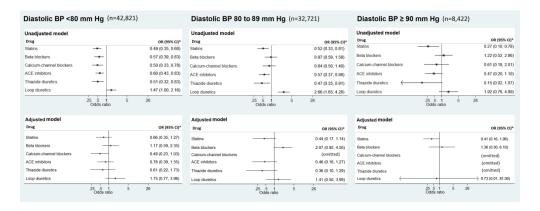


Figure 3: Forest plots for the impact of various preoperative cardiovascular medications on postoperative mortality for different preoperative diastolic blood pressure thresholds. Top row shows unadjusted data and bottom row shows the results after adjusting for confounders. Columns refer to different blood pressure thresholds.