

# If Exchange Rates Are Random Walks, Then Almost Everything We Say About Monetary Policy Is Wrong

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The key question asked of standard monetary models used for policy analysis is: how do changes in short-term interest rates affect the economy? The standard answer is that such policy changes affect the economy by changing the means of macroeconomic aggregates while having no effect on their conditional variances. Unfortunately, the data on exchange rates imply nearly the opposite. Fluctuations in interest rates are associated with nearly one-for-one changes in conditional variances and almost no changes in conditional means. With regard to monetary policy analysis, this means that standard monetary models capture nothing of what is going on in the data. Therefore, almost everything we say about monetary policy, based on these models, is wrong.

Standard log-linear models of monetary policy of both the New Keynesian and neoclassical variety link nominal interest rates, through an Euler equation, to the conditional means of the log of two variables: the representative agent's marginal utility growth and inflation. (Changes in these two variables are loosely thought of as reflecting the real and nominal effects of monetary policy.) The main debate among standard modelers has been about how much interest rate changes affect each of the two variables. They do not debate a common assumption of their models—that interest rate changes have no

effect on the conditional variances of marginal utilities and inflation.

That common assumption, however, is grossly inconsistent with a well-established feature of the data—nominal rates of exchange between major currencies are well approximated by random walks.<sup>1</sup> Mechanically, that fact implies that when a central bank changes its interest rate relative to the rates on other major currencies, the change is reflected almost entirely as a change in the excess returns on its bonds over the returns on foreign bonds. Interpreted in a standard model, the exchange rate fact implies that changes in a domestic interest rate relative to a foreign interest rate lead to one-for-one changes in conditional variances and nearly no changes in conditional means. The fact implies that, at least when they are analyzing changes in domestic interest rates relative to those of foreign interest rates, standard monetary models are of little use.

Clearly, to analyze monetary policy, we need a new approach that captures the effects of interest rate changes on conditional variances. We have tried one in which such effects are interpreted as time-varying risk. (For elaboration, see Alvarez, Atkeson, and Kehoe 2006.)

## I. Standard Models of Monetary Policy

Standard models of monetary policy start with a presumption that a monetary authority controls the short-term nominal interest rate on bonds or other assets, denominated in its own currency. Most of these models assume a representative consumer who participates in all asset markets. We begin by describing these representative consumer models and their assumption that interest

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<sup>1</sup> This finding dates back at least to the work of Richard A. Meese and Kenneth Rogoff (1983) and has been confirmed by Yin-Wong Cheung, Menzie D. Chinn, and Antonio Garcia Pascual (2005). As discussed in Cheung, Chinn, and Pascual (2005), some evidence suggests exchange rates are not random walks, but predictable, at least at long horizons.

rate changes affect only the conditional means of variables. Then, we show how that description generalizes beyond those models.

### A. Representative Consumer Models

The short-term nominal interest rate enters standard representative consumer models through an Euler equation of the form

$$(1) \quad \frac{1}{1+i_t} \equiv \exp(-i_t) = \beta E_t \left[ \frac{U_{c,t+1}}{U_{c,t}} \frac{1}{\pi_{t+1}} \right],$$

where  $i_t$  is the logarithm of the short-term nominal interest rate  $1+i_t$ ,  $\beta$ ,  $U_{c,t}$  is the discount factor and the marginal utility of the representative consumer, and  $\pi_{t+1}$  is the inflation rate. Analysts commonly assume the data are well approximated by a conditionally log-normal model, so that this Euler equation can be written as

$$(2) \quad i_t = E_t \left[ -\log \frac{U_{c,t+1}}{U_{c,t}} \frac{1}{\pi_{t+1}} \right] - \frac{1}{2} \text{var}_t \left[ \log \frac{U_{c,t+1}}{U_{c,t}} \frac{1}{\pi_{t+1}} \right].$$

The critical question in monetary policy analysis is what terms on the right side of (2) change when the monetary authority changes the interest rate  $i_t$ . The standard assumption is that the conditional variances are constant, so that the second term in (2) is constant. This leaves the familiar version of the Euler equation:

$$(3) \quad i_t = -E_t \log \frac{U_{c,t+1}}{U_{c,t}} + E_t \log \pi_{t+1} + \text{constant}.$$

Changes in the nominal interest rate can be broken down into the change in the expected growth in the marginal utility of consumption of the representative agent and the change in expected inflation. Loosely speaking, we think of the first component as reflecting the *real effect* of monetary policy on the economy and the second as reflecting the *nominal effect*. The debate in monetary policy analysis is over how

changes in the nominal interest rate are divided into these two types of effects. For example, in the simplest flexible price models, monetary policy is neutral, its real effects are zero, and changes in nominal interest rates change only expected inflation. In more complicated models, frictions of various sorts, such as sticky prices, imply that changes in interest rates have real and nominal effects, and the details of the model determine their decomposition.

Regardless of which side of the debate a particular standard model of monetary policy represents, it assumes that changes in interest rates affect only the conditional means of endogenous variables, not conditional variances or other higher moments. This is a serious problem for representative consumer models.

### B. More General Models

More general models, which do not assume a representative consumer, have this problem, for they, too, limit the effects of monetary policy changes to the conditional means of variables.

To see this, note that equations (1)–(3) can be written more abstractly in terms of a nominal *pricing kernel* (or stochastic discount factor)  $m_{t+1}$  as

$$(4) \quad \exp(-i_t) = E_t m_{t+1}.$$

In a model with a representative agent, this pricing kernel is  $m_{t+1} = \beta U_{c,t+1}/(U_{c,t}\pi_{t+1})$ , and (4) is the representative agent's first-order condition for optimal bond holdings. In some segmented market models, (4) is the first-order condition for agents who participate in the bond market, while in others, (4) is no single agent's first-order condition. In general, equation (4) is implied by lack of arbitrage possibilities in the financial market. With log-normality, (4) implies that

$$(5) \quad i_t = E_t[-\log m_{t+1}] - \frac{1}{2} \text{var}_t[\log m_{t+1}],$$

and, with constant conditional variances, that  $i_t = -E_t \log m_{t+1} + \text{constant}$ . Thus, the more general assumption made in the literature is that monetary policy affects only the conditional mean of the log of the pricing kernel, not its conditional variance.

C. Adding a Foreign Country

Below, we use data on interest rate differentials and exchange rates to flesh out the major problem with the standard approach. Here, to set up that analysis, we consider the implications of adding a foreign country to our model with its own currency and its own monetary policy.

Assuming conditional log-normality gives the foreign analog of (5), namely, that

$$(6) \quad i_t^* = E_t[-\log m_{t+1}^*] - \frac{1}{2} \text{var}_t[\log m_{t+1}^*],$$

where asterisks denote foreign variables. When the foreign pricing kernel comes from a representative consumer in the foreign country,  $m_{t+1}^* = \beta U_{c,t+1}^*/(U_{c,t}^* \pi_{t+1}^*)$ , and (6) is the foreign representative consumer's Euler equation for foreign bonds.

Combining (5) and (6) allows the interest differential to be written as

$$(7) \quad i_t - i_t^* = E_t[\log m_{t+1}^* - \log m_{t+1}] - p_t,$$

where

$$(8) \quad p_t = \frac{1}{2} [\text{var}_t \log m_{t+1}^* - \text{var}_t \log m_{t+1}].$$

Note that under the standard assumption of constant conditional variances, the term  $p_t$  is constant. (For a similar derivation, see David K. Backus, Silverio Foresi, and Chris I. Telmer 2001.)

The standard approach to analyzing monetary policy assumes that when the monetary authorities in two countries change the interest differential,  $i_t - i_t^*$ , what changes are the conditional means in (7), not the conditional variances in (8).

II. The Problem

The problem is the data contradict that assumption. One of the most robust features of the data on nominal exchange rates between major currencies is that they are well approximated by random walks. This fact means the standard models have the analysis backward—when the interest differential changes, what changes are the conditional variances, not the conditional means.

A. A Contradiction

We demonstrate how the data contradict the standard model by linking exchange rates to nominal pricing kernels. Lack of arbitrage and complete financial markets imply that

$$(9) \quad m_{t+1}^* = m_{t+1} \frac{e_{t+1}}{e_t},$$

where  $e_t$  is the nominal exchange rate. To derive this equation in a standard model, add into that model the opportunity for a home investor to purchase a foreign currency-denominated asset with stochastic return  $R_{t+1}^*$ . The home currency return on this asset is given by  $R_{t+1}^* e_{t+1}/e_t$ . Hence, lack of arbitrage for the home investor implies that  $1 = E_t m_{t+1} (e_{t+1}/e_t) R_{t+1}^*$ . The pricing kernel  $m_{t+1}^*$  defined by (9) also prices foreign currency returns, so that  $1 = E_t m_{t+1}^* R_{t+1}^*$ . Under the assumption of complete markets, the pricing kernel is unique, and this gives the result (9). The assumption of complete markets is sufficient to obtain this result, but it is not necessary, as we will discuss.

Taking logs and then conditional expectations of (9) gives

$$(10) \quad E_t \log e_{t+1} - \log e_t = E_t \log m_{t+1}^* - E_t \log m_{t+1}.$$

Using (10) in (7) gives

$$(11) \quad i_t - i_t^* = E_t[\log e_{t+1} - \log e_t] - p_t,$$

where, recall,  $p_t$  represents an expression involving the conditional variances.

Now, compare equation (11) to the data. In the data, interest differentials show large and persistent movements over time. But since exchange rates are well approximated by random walks, the expected change in the exchange rate,  $E_t[\log e_{t+1} - \log e_t]$ , must be approximately a constant.<sup>2</sup> Hence, (10) and (11) imply that

<sup>2</sup> Indeed, at least since Eugene F. Fama's (1984) seminal work, this conditional expectation has been found to comove negatively with interest differentials. In particular, in a regression of the form  $\log e_{t+1} - \log e_t = a + b(i_t - i_t^*) + \varepsilon_t$ , the estimated value of  $b$  is almost always smaller than one and is often negative. A negative value of  $b$  strengthens

when the interest differential  $i_t - i_t^*$  moves, what moves are the conditional variances in  $p_t$ , not the conditional means in (10).

Why should this discrepancy trouble users of standard monetary models? Because it reveals that their standard debates about how to divvy up the effects of interest rate changes into real and nominal effects are debates about terms that are essentially constant. The standard monetary models have nothing to say about the terms that are affected by interest rate changes—the conditional variances.

### B. An Interpretation

Changes in conditional variances are abstract model expressions, but they can be interpreted as critical economic variables—changes in risk premia. Under this interpretation, standard models are missing a link between monetary policy changes and risk.

To understand this interpretation, consider a simple example. Let the foreign currency be the British pound and the home currency be the US dollar. Define the (log) excess return for a pound-denominated bond as the expected log dollar return on a pound bond minus the log dollar return on a dollar bond. Let  $\exp(i_t)$  and  $\exp(i_t^*)$  be the nominal interest rates on the dollar and pound bonds, and  $e_t$  be the price of pounds in units of dollars, or the exchange rate between the currencies, in a time period  $t$ . The dollar return on a pound bond,  $\exp(i_t^*)e_{t+1}/e_t$ , is obtained by converting a dollar in period  $t$  to  $1/e_t$  pounds, buying a pound bond paying interest  $\exp(i_t^*)$ , and then converting the resulting pounds back to dollars in  $t + 1$  at the exchange rate  $e_{t+1}$ . The (log) excess expected return  $p_t$  is then defined as the difference between the expected log dollar return on a pound bond and the log return on a dollar bond:

$$(12) \quad p_t = i_t^* + E_t \log e_{t+1} - \log e_t - i_t.$$

Clearly, the dollar return on the pound bond is risky, because the future exchange rate  $e_{t+1}$  is not known in  $t$ . The excess return compensates the holder of the pound bond for this exchange rate risk.

In the model that we have laid out, the excess expected return  $p_t$  in (12) can be expressed in terms of conditional variances of nominal pricing kernels, as in (8). Hence, we interpret changes in these conditional variances as changes in risk. (Other possible interpretations of  $p_t$  are that it represents compensation to the holder of the foreign bond for differences in liquidity services, transaction costs, or tax rates, none of which are measured across these bonds.)

With our interpretation, we can restate our point. The fact that exchange rates are approximately random walks implies that most of the fluctuations in interest rate differentials are changes in risk—a feature standard models do not link to monetary policy changes.

### III. Extensions

So far, in order to derive equation (9), the link between exchange rates and nominal pricing kernels, we have assumed complete asset markets. Here, we show how our argument extends to models with incomplete markets and to models with other financial frictions.

Consider simple incomplete market models which allow the trading of only a limited set of financial assets. In such models, pricing kernels are not unique. As discussed by Michael W. Brandt, John H. Cochrane, and Pedro Santa-Clara (2006), however, even with incomplete markets, equation (9) holds for the minimum variance pricing kernels. Hence, with such kernels, our argument goes unchanged.

Consider, next, a version of our argument that applies even if asset markets are extremely incomplete, for example, if a home consumer has access to only three assets: a home currency bond, a foreign currency bond, and foreign currency. We show that in such a situation, if the exchange rate is a random walk, then fluctuations in interest differentials correspond to fluctuations in conditional variances and covariances, not to fluctuations in conditional means.

To see that, let  $m_{t+1}$  be any kernel that prices home currency returns. This kernel must satisfy  $1 = E_t m_{t+1} R_{t+1}$  for any asset with the home currency return  $R_{t+1}$  at  $t + 1$ . In particular, the kernel must satisfy (4) for home currency bonds, and  $1 = E_t m_{t+1}(e_{t+1}/e_t) \exp(i_t^*)$  for the home currency return on an investment in foreign currency

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our argument, but for simplicity, we focus on what happens with  $b = 0$ , or when the exchange rate is a random walk.

bonds. With some simple manipulations, conditional log-normality of all variables implies that

$$(13) \quad i_t - i_t^* = [E_t \log e_{t+1} - \log e_t] \\ + \left[ \frac{1}{2} \text{var}_t \left( \log \frac{e_{t+1}}{e_t} \right) \right. \\ \left. + \text{cov}_t \left( \log m_{t+1}, \log \frac{e_{t+1}}{e_t} \right) \right].$$

If exchange rates are random walks, then the first term on the right side of (13) is constant. So fluctuations in interest differentials must lead to one-for-one changes in the second term. Hence, changes in monetary policy are shown by the data to be changes in conditional variances, whereas the standard models assume they are changes in conditional means. Thus, our argument applies even for extensions of standard models that include extreme forms of market incompleteness.

Our argument applies more generally to the large class of models with financial frictions in which the pricing kernels satisfy equation (9), and which assume that the conditional variances of these pricing kernels are constant.

#### IV. Implications

Our analysis of the standard approach to modeling monetary policy tells us that economists need new models, and we have some suggestions on how to get them. Our analysis also has something to say about how US monetary policy has worked in recent years. And it implies that the old standard models need not be discarded completely. They can help us understand cross-section patterns of average short-term interest differentials.

##### A. Arguing Causality

In making our point, we have not needed to argue the direction of causality between changes in interest rates and changes in risk. Does risk in financial markets change for some reason unrelated to monetary policy, and does the monetary authority react, changing the nominal interest rate in order to accommodate the risk change? Or does the monetary authority's interest rate

change result in a change in financial market risk? With our exchange rate analysis in mind, a brief review of recent US and UK monetary policy suggests that, at least lately, the causality has been from changes in interest rates to changes in risk premia.<sup>3</sup>

A graphical view of the recent monetary policies of the two countries suggests this. Figure 1A, which is available at [www.e-aer.org/data/may07/P07027\\_app.zip](http://www.e-aer.org/data/may07/P07027_app.zip), plots monthly data on the US federal funds rate and the Bank of England's official bank rate from January 2000 through November 2006. This figure shows the Federal Reserve's decision to dramatically reduce the federal funds rate over the first half of this time period and then to raise it over the second half. The corresponding policy moves by the Bank of England were much less dramatic. The figure shows that these differences in monetary policy between the United States and the United Kingdom led to large and persistent movements in the interest differential between the dollar and the pound. Market observers have attributed these policy decisions to a variety of factors, none of which includes accommodating changes in the conditional volatility of consumption growth or inflation or, more generally, in pricing kernels.

The interest differential movements do not correlate well with changes in exchange rates over this period. Figure 1B, which is available at [www.e-aer.org/data/may07/P07027\\_app.zip](http://www.e-aer.org/data/may07/P07027_app.zip), is a scatterplot of the dollar-pound interest differential,  $i_t - i_t^*$ , against the corresponding change in exchange rates,  $\log e_{t+1} - \log e_t$ , with both series expressed in annualized units.<sup>4</sup> The widely dispersed plots are consistent with the idea that the expected change in the dollar-pound exchange rate was essentially unrelated to the dollar-pound interest differential over this time period.

If we accept that monthly exchange rate changes are unrelated to interest rate differentials, then Figure 1A and Figure 1B indicate that

<sup>3</sup> There have been other episodes in which observers have argued that the Federal Reserve has changed policy in response to changes in financial market risk. These include the stock market crash of October 1987, the Russian debt crisis in 1998, and the period after September 11, 2001.

<sup>4</sup> The difference between the US federal funds rate and the UK official bank rate is nearly identical to the interest differential relevant for exchange rate arbitrage, namely, the one-month dollar-pound forward premium.

at the beginning of 2004, investors required an expected excess return of almost three percentage points to hold British pounds, while at the beginning of 2006, that requirement was zero. Figure 1 seems to imply that recent US monetary policy actions have had their main impact on risk and not on the factors that standard analyses focus on.

### B. *Using Old Models*

We have argued that the standard models for monetary policy analysis are not useful for understanding how fluctuations in interest differentials affect the economy. Are these models useful at all? The data suggest they are. Standard models do a reasonable job of accounting for cross-section data on long-run averages of differences in interest rates across countries.

To investigate this issue, we use monthly data for the period from January 1976 to March 1998 to construct average one-month interest rate differentials with the US rate for 14 countries, as well as corresponding average rates of exchange rate change over this period. Figure 2, available at [www.e-aer.org/data/may07/P07027\\_app.zip](http://www.e-aer.org/data/may07/P07027_app.zip), displays a scatterplot of these data. It shows a clear positive relationship between the averages, with slope close to 1. This relationship supports the idea that regardless of its problem with monetary policy analysis, the standard model with constant conditional variances is a reasonable approximation for cross-section data on long averages of differences in short-term interest rates across countries.

### C. *Designing New Models*

The data on exchange rates push us to the view that analysts of monetary policy must look in new directions for tools to help us understand how policy changes affect the economy. One possibly fruitful direction is to develop models in which the excess return on foreign bonds fluctuates at the monthly level due to fluctuations in differential liquidity services, differential transaction costs, or differential tax rates across bonds. A more promising direction is simpler—to develop models in which changes in monetary policy affect the economy primarily by changing risk. In ongoing research (Alvarez, Atkeson, and Kehoe 2006), we built such a model based on the

idea that asset markets are segmented and that monetary policy affects risk by endogenously changing the degree of market segmentation. We have shown that this model can generate, qualitatively, the type of systematic variation in risk premia called for by the data on interest rates and exchange rates. Our work, of course, represents only a first, simple step toward building models in which changes in monetary policy affect the economy primarily by changing risk.

### V. **Concluding Remarks**

Must monetary models be able to account for fluctuations in excess returns? Indeed, hasn't modern business cycle theory been quite successful at accounting for fluctuations in aggregate quantities even though it has done a fairly miserable job at accounting for asset prices, particularly the large movements in excess returns that are part of asset prices? This sort of skepticism is implicit in much of the business cycle literature. Accounting for asset prices seems to be thought of as of second-order importance when thinking about the determination of economic aggregates such as consumption, investment, and employment, which are at the heart of business cycle theory.

Regardless of the merits of that view, it is inappropriate for analyzing monetary policy. Determining how changes in an asset price and the short-term interest rate affect the economy is clearly at the heart of monetary policy analysis. As we have argued, the data on exchange rates imply that movements in interest rate differentials are reflected almost entirely in fluctuations in excess returns. Thus, for monetary policy, accounting for fluctuations in these excess returns is essential, and monetary models, which cannot account for them, cannot help us understand the effects of interest rate changes on the economy.

We have used data on exchange rates to rethink the analyses of interest rate changes in standard monetary models. We could have used data on the excess returns on long-term domestic bonds over short-term domestic bonds, since another well-established fact is that these excess returns vary systematically with variables plausibly controlled by the Federal Reserve, such as the term spread. In standard models, however, these excess returns are all constant. These models cannot account for term spread movements either.

We have focused on exchange rates rather than the term structure of interest rates because the implications of exchange rates are so striking. Specifically, if exchange rates are random walks, then all of the fluctuations in interest differentials are accounted for by fluctuations in conditional variances and none by fluctuations in conditional means. The data are so opposite of what standard models assume that even the most die-hard defenders of them should take note. If these data are accurate, then almost everything we say about monetary policy is wrong.

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