

1 **PSA Screening and 15-year Prostate Cancer Mortality: The CAP Randomized Clinical Trial**

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41 **Key points**

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43 **Question:** In men aged 50-69, does a single invitation for a prostate specific antigen (PSA) screening
44 test reduce prostate cancer mortality at 15-year follow-up, compared to a control group that was
45 not invited for testing?

46

47 **Findings:** In this cluster randomized trial of 415,357 men aged 50-69 randomized to a single
48 invitation for PSA screening (N=195,912) or a control group without PSA screening (N=219,445) and
49 followed-up for a median of 15-years, risks of death from prostate cancer were lower in the group
50 invited to screening (0.69% vs. 0.78%; mean difference: 0.09%), compared to the control group.

51

52 **Meaning:** Compared to a control group without routine PSA testing, a single invitation for a PSA
53 screening test reduced prostate cancer mortality at a median follow-up of 15 years, but the absolute
54 mortality benefit was small.

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58 **Abstract (Word count 383)**

59 **IMPORTANCE** The Cluster randomized trial of PSA testing for Prostate cancer (CAP) reported no
60 effect of prostate specific antigen (PSA) screening on prostate cancer mortality at median 10-year
61 follow-up (primary outcome), but the long-term effects of PSA screening on prostate cancer
62 mortality remain unclear.

63 **OBJECTIVE** To evaluate the effect of a single invitation for PSA screening on the pre-specified
64 secondary outcome of prostate cancer-specific mortality at a median of 15 years' follow-up,
65 compared to a control group not invited for screening.

66 **DESIGN, SETTING, PARTICIPANTS** Cluster randomized trial of men aged 50-69 identified from 573
67 primary-care practices in England and Wales. Primary-care practices were randomized between
68 09/25/2001 and 08/24/2007 and men were enrolled between 01/08/2002 and 01/20/2009. Follow-
69 up was completed on 03/31/2021.

70 **INTERVENTION** A single invitation for a PSA screening test with subsequent diagnostic tests if
71 PSA \geq 3.0ng/ml, compared to standard practice (control).

72 **MAIN OUTCOMES AND MEASURES** The primary outcome was reported previously. Of eight
73 prespecified secondary outcomes, results of four were reported previously. The four remaining pre-
74 specified secondary outcomes at 15-year follow-up were prostate cancer-specific mortality, all-cause
75 mortality, and prostate cancer stage and Gleason grade at diagnosis.

76 **RESULTS** Of 415,357 randomized men (mean [SD] age: 59.0 [5.6] years), 98% were analyzed in these
77 analyses. Overall, 12,013 and 12,958 men with prostate cancers were diagnosed in the intervention
78 and control groups (15-year cumulative risks 7.1% and 6.9% respectively).

79 At a median 15-year follow-up, 1,199 (0.69%) men in the intervention group and 1,451 (0.78%) men
80 in the control group died of prostate cancer (rate ratio [RR] 0.92 [95% CI 0.85, 0.99]; p=0.03).

81 Compared to the control group, the PSA screening intervention increased detection of low-grade
82 (Gleason score [GS] \leq 6; 2.2% versus 1.6%;p<0.001) and localized (T1/T2; 3.6% versus 3.1%;p<0.001)
83 disease, but not intermediate (GS=7), high-grade (GS \geq 8), locally-advanced (T3) or distally-advanced

84 (T4/N1/M1) tumors. There were 45,084 all-cause deaths (23.2%) in the intervention group and
85 50,336 deaths (23.3%) in the control group respectively (RR 0.97 [95% CI 0.94, 1.01]; p=0.11). Eight
86 deaths in the intervention and seven deaths in the control group were related to a diagnostic biopsy
87 or prostate cancer treatment.

88 **CONCLUSIONS AND RELEVANCE** A single invitation for PSA screening, compared to standard practice
89 without routine screening, reduced the secondary outcome of prostate cancer deaths at a median
90 follow-up of 15-years. However, the absolute reduction in deaths was small.

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92 **TRIAL REGISTRATION:** ISRCTN92187251

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95 **Introduction**

96 In England, the number of men diagnosed with prostate cancer increased by 68% from 28,216 in
97 2001 to 47,479 in 2019,¹ reflecting population aging and increased prostate specific antigen (PSA)
98 testing.² In the USA, approximately 3.3 million men currently live with a diagnosis of prostate
99 cancer.³ While low-risk prostate cancer progresses slowly and is associated with a low risk of
100 mortality,⁴⁻⁷ aggressive prostate cancer currently causes approximately 12,000 deaths in the UK and
101 34,700 deaths in the U.S. annually.^{3,8} The goal of PSA screening is to reduce prostate cancer
102 mortality by early detection of curable disease. However, uncertainty remains regarding the long-
103 term effect of PSA-based screening on mortality.⁹⁻¹¹

104 The CAP RCT (N=415,357) showed that, compared to a usual care (unscreened) control group, an
105 invitation to a single PSA screen increased the number of prostate cancers diagnosed during the first
106 18 months of follow-up (the time period when PSA testing and subsequent biopsies for men with an
107 elevated level of PSA took place). In this trial, rates of diagnosed prostate cancer were 2.2 per 1000
108 person-years in the control group and 10.4 per 1000 person-years in the intervention group
109 (P<0.001).¹⁰ However, at a median 10-year follow-up, the invitation for a single PSA screen did not
110 reduce prostate cancer mortality, compared to the control group (0.29% vs. 0.30%; rate ratio: 0.96;
111 95%CI;0.85-1.08;p=0.5).¹⁰ This report describes the effects of this single invitation to a PSA-screening
112 test, with subsequent diagnostic tests if PSA>3.0ng/ml, on the pre-specified secondary outcome of
113 prostate-cancer mortality at 15-year follow-up, compared to standard (unscreened) practice.¹²

114

115 **Methods**

116 The Derby National Research Ethics Service Committee East Midlands approved the study. The trial
117 **Protocol** and the **statistical analysis plan** are available as Supplementary material to the primary
118 outcome paper.¹⁰ Participants were enrolled between 01/08/2002 and 1/20/2009. Final follow-up
119 occurred 03/31/2021.

120 Men who attended PSA testing in the intervention group gave individual written informed consent
121 via the ProtecT study.¹³ Individual consent was not sought from men in the control group or from
122 non-responders in the intervention group. Instead, approval for their identification and linkage to
123 routine electronic records was obtained under Section-251 of the NHS Act 2006 from the UK Patient
124 Information Advisory Group (now Confidentiality Advisory Group).¹⁰ All clinical centers had local
125 research governance approval.

126 **Randomization**

127 The study was a primary-care based cluster RCT that tested the effects of a single invitation for a PSA
128 screening test (**eFigure 1**), compared to usual care (no screening), on the primary outcome of
129 prostate-cancer mortality at a median follow-up of 10 years. The primary outcome has been
130 reported.¹⁰ Between 2001 and 2007, 785 eligible general practices in the catchment area of 8
131 hospitals across England and Wales (located in Birmingham, Bristol, Cambridge, Cardiff, Leeds,
132 Leicester, Newcastle and Sheffield) were randomized before recruitment ('Zelen' design) to
133 intervention or control groups and practices were invited to consent to participate. Randomization
134 was blocked and stratified within groups of 10-12 neighboring practices, using a computerized
135 random number generator. Because allocation preceded the invitation for practices to participate, it
136 was not possible to conceal allocation. 573 (73%) practices, including 68% randomized to the
137 intervention group and 78% randomized to the control group, agreed to participate (**Figure 1**).

138 **Participants**

139 Men aged 50-69 years in each participating randomized general practice were included. Men with
140 prostate cancer on or before the randomization date and those registered as a patient with
141 participating practices on a temporary or emergency basis were excluded.

142 **Intervention**

143 Men in practices randomized to the intervention received a single invitation for a PSA test after
144 counselling. If the resulting PSA was 3.0-19.9ng/ml, they were offered 10-core transrectal
145 ultrasound-guided biopsies. All laboratories participated in the UK National External Quality

146 Assessment Service (UK NEQAS) for PSA testing. Test results that did not meet laboratory quality
147 assurance requirements, were lost, or if consent was ambiguous or if insufficient blood was
148 obtained, were considered non-valid. Men in the intervention group diagnosed with localized
149 prostate cancer were invited to participate in a second RCT, the ProtecT treatment trial
150 (ISRCTN20141297) which randomized participants to active monitoring (consisting of regular PSA
151 testing and clinical review), radical prostatectomy, or radical conformal radiotherapy with neo-
152 adjuvant-androgen-deprivation (**eFigure 1**).¹⁴ Men with a PSA ≥ 20 ng/ml were referred to a urologist
153 and received standard care.

154 Men in practices randomized to the control group received standard NHS management but did not
155 receive a formal invitation for PSA testing as part of this study.¹⁵ We assessed cumulative PSA testing
156 for prostate cancer detection in the control group of CAP by longitudinal analysis of a national
157 primary care database (N=434,236 men from 558 UK GP practices)².

158 **Outcomes**

159 The primary outcome of this clinical trial, 10-year prostate cancer mortality, was reported
160 previously.¹⁰ Pre-specified secondary outcomes were: definite or probable prostate cancer mortality
161 at 15-year follow-up; all-cause mortality at 10-year follow-up; all-cause mortality at 15-year follow-
162 up; all-cause mortality at 5-year follow-up; prostate cancer mortality at 5-year follow-up; disease
163 grade and staging; cost-effectiveness; and health related quality of life. The protocol did not indicate
164 the time point for assessing prostate cancer grade and staging; these were measured at median
165 follow-up time points of 10-years and 15-year follow-up. Previously reported outcomes were all-
166 cause mortality at 10-year follow-up,¹⁰ disease grade and stage at 10-year follow-up,¹⁰ cost-
167 effectiveness¹⁶ and health related quality of life.¹⁷ The current report provides results for the
168 remaining secondary outcomes of definite or probable prostate cancer mortality at 15-year follow-
169 up, all-cause mortality at 15-year follow-up, and disease grade and stage at 15-year follow-up. All-
170 cause and prostate cancer mortality at 5-year follow-up were not published separately, but five-year

171 follow-up data are shown in Kaplan Meier curves, both in the current paper and the publication of
172 the 10-year primary outcome.¹⁰

173 **Outcome ascertainment**

174 Prostate cancer mortality at 15-year follow-up was ascertained with death certificates from the
175 Office for National Statistics (ONS) at NHS England and adjudicated by an independent Cause of
176 Death Evaluation (CoDE) committee using clinical information from hospital medical records and
177 following a standardized protocol.^{18,19} Prostate cancer stage and Gleason grade were obtained from
178 the National Disease Registration Service²⁰ (NDRS, formerly Public Health England) at NHS England
179 and Public Health Wales,²¹ up to December 31st 2020.

180 **Exploratory outcomes**

181 Additional outcomes reported here that were described in the published original statistical analysis
182 plan¹⁰ were: i) mean age at diagnosis between allocated groups; and ii) a sensitivity analysis re-
183 defining the primary outcome to include: (a) definite, probable, possible and treatment-related
184 prostate cancer mortality; and (b) definite and treatment-related prostate cancer mortality.

185 **Post hoc outcomes**

186 We estimated differences in the risks of prostate cancer diagnosis between the intervention and
187 control groups at 18-months, 10-years and 15-years, to quantify changes in diagnosis rates over
188 long-term follow-up. We calculated mean sojourn time (the period in which a tumor is asymptomatic
189 but detectable by screening) from microsimulation using estimated transition parameters for single
190 episodes of screening between ages 50 to 69 and over-diagnosis rates as the difference in the
191 cumulative prostate cancer incidence between screened and unscreened groups over a lifetime
192 (further methodological details in **Supplement 1**).^{22,23}

193 **Statistical Analysis**

194 The intervention effect at a median 15-years follow-up (at March 31st 2021) was analysed comparing
195 groups as randomized using random-effects Poisson regression to estimate prostate cancer-specific
196 and all-cause mortality rate ratios (RRs) in intervention versus control practices, allowing for

197 clustering within GP practices and randomization strata. To allow for variation in the incidence of
198 prostate cancer with age, follow-up for each participant was divided into periods within five-year
199 age-groups. We present rates (per 1000 person-years) and Kaplan-Meier estimates of the cumulative
200 risk (per 100 men) of prostate cancer diagnosis, and prostate cancer and all-cause mortality.
201 In pre-specified analyses described in the original statistical analyses plan, and available as
202 Supplementary material to the primary outcome paper,¹⁰ we: i) used instrumental methods
203 (generalized method of moments estimator) to estimate the effect of attending the PSA screening
204 clinic at a median 15-years, compared with men in the control group who would have attended the
205 clinic if invited, adjusting for age-group and using robust standard errors to allow for variation
206 between practices; ii) compared mean age, and prostate cancer clinical stage (T1/T2, T3 and
207 T4/N1/M1 disease) and Gleason score (=6 [low-grade]; =7 [intermediate grade]; 8+ [high grade]) at
208 diagnosis between intervention and control groups using ordered logistic regression.
209 Prespecified subgroup analyses investigated variation in the effect of screening on prostate cancer
210 mortality by baseline age-group and quintiles of geographical area-based index of multiple
211 deprivation, a measure of socioeconomic status. An interaction test p-value was used to evaluate the
212 evidence against the null hypothesis of equal intervention effect across sub-groups.
213 In accordance with our original analysis plan,¹⁰ we did not conduct multiple imputation analyses. The
214 statistical analysis plan did not specify an intention to adjust p-values for multiple comparisons:
215 conventional adjustments assumed statistical independence between estimates, which was not the
216 case for analyses of the same outcome at 10 and 15 years. All statistical testing was for superiority
217 and p-values were 2-sided. In interpreting the results, we focused on estimated effects and
218 associated 95% CIs. Results were considered statistically significant if the P value was <.05 or not
219 statistically significant if the P value was ≥.05. All trial analyses were conducted using Stata version
220 16.1 (StataCorp).
221
222

223 **Results**

224 **Study Population**

225 911 GP practices were randomized in 99 geographical areas. Of these, 126 were subsequently
226 excluded as ineligible (**Figure 1**).¹² Consent rates were 68% (271/398) among eligible GP practices in
227 the intervention group and 78% (302/387) among eligible GP practices in the control group. Overall,
228 415,357 men registered with these practices were eligible for the intervention (N=195,912) and
229 control (N=219,445) groups. Follow-up data for cancer diagnosis and mortality at a median of 15
230 years after randomization were available for 408,721 of the eligible men (98%), including 189,326
231 (97%) randomized to the intervention and 219,395 (>99%) randomized to control (**Figure 1**).

232 Baseline characteristics were similar between intervention and control groups at practice and
233 individual level (**Table 1**). Among people randomized to the intervention who developed prostate
234 cancer (N=12,013), 9.4% were missing data for cancer stage and 10.4% were missing data for
235 Gleason grade. Among people randomized to the control group who developed prostate cancer
236 (N=12,958), 7.8% were missing data for cancer stage and 11.2% were missing data for cancer
237 Gleason grade.

238 **Rates of PSA testing**

239 Overall, 75,694 (40%) of men randomized to the intervention group underwent PSA-testing and
240 64,425 (34%) had a valid (as defined in the methods) test result. Of these, 6,855 (11%) had a PSA
241 value between 3-19.9ng/ml and were eligible for the ProtecT trial. Of these, 5,848 (85%) had a
242 prostate biopsy. Cumulative PSA testing for prostate cancer detection in the control-group was
243 indirectly estimated at 10% to 15% over 10-years median follow-up.^{2,10}

244 **Prostate cancer deaths**

245 After a median follow-up of 15.4 years (interquartile range, IQR: 14.2-16.4; range: 12.2, 19.2), there
246 were 1,199 deaths due to prostate cancer (rate: 0.47 per 1000-person years) in the intervention
247 group and 1,451 deaths (rate: 0.50 per 1000-person years) in the control-group: RR 0.92 (95% CI,
248 0.85 to 0.99; p=0.03) (**Table 2, Figure 2A**). At a median of 15-years' follow-up, the cumulative risks of

249 prostate cancer mortality were 0.69% in the intervention group and 0.78% in the control group [risk
250 difference -0.09% (95% CI, -0.15 to -0.03, P=0.02)] (**Table 2, eTable 1**). Using instrumental variable
251 analysis, the prostate cancer mortality rate ratio for the effect of screening amongst men attending
252 PSA-testing clinics was 0.83 (95% CI 0.68, 1.00; p=0.053) (**Table 2**).

253 **Overall survival**

254 There were 45,084 total deaths in the intervention group and 50,336 total deaths in the control
255 group (RR 0.97: 95% CI 0.94 to 1.01; p=0.11) (**Table 2, Figure 2B**). Other causes of death were similar
256 between the two groups (**eTable 2**).

257 **Prostate cancer grade and stage**

258 Compared to control, men in the intervention group were at higher risk of diagnosis with low-grade
259 (2.2% of men versus 1.6%; risk difference = 0.58%, 95% CI 0.50%, 0.67%), and at lower risk of high-
260 grade (1.2% versus 1.3%; risk difference = -0.15%; 95% CI: -0.22% to -0.08%), prostate cancers over
261 the 15-years follow-up (p for trend <0.001). There was a higher risk of localized (3.6% versus 3.1%;
262 risk difference = 0.56%, 95% CI 0.44%, 0.67%) prostate cancers and a lower risk of advanced-stage
263 tumors (0.9% versus 1.1%; risk difference = -0.16%; 95% CI: -0.22% to -0.10%) over the 15-years
264 follow-up in the intervention versus control group (p for trend <0.001) (**eTable 3; eFigures 2 and 3**).

265 **Exploratory results**

266 The mortality results were similar when including in the outcome definition those prostate cancer-
267 specific deaths judged as 'possible' by the Cause of Death Evaluation committee, and when
268 restricting to those judged as 'definite' prostate cancer-specific deaths (**eTable 4**). There was little
269 evidence that the intervention effect differed by age-group or socioeconomic status (p values for
270 interaction ≥ 0.46) (**Table 3**). Compared to the control group, intervention group men were a mean
271 1.22 years younger at prostate cancer diagnosis (95% CI 1.02, 1.42; p<0.001) (**eTable 3**).

272 **Post hoc results**

273 After a median 15-years follow-up, there were 12,013 (4.88 per 1000 person-years [cumulative risk:
274 7.1%]) prostate cancer diagnoses in the intervention group and 12,958 (4.60 per 1000 person-years

275 [cumulative risk: 6.9%]) in the control group (**Table 2, Figure 2C**). Differences in the risks of prostate
276 cancer diagnosis between the intervention and control groups varied markedly during follow-up:
277 cumulative risk differences per 1000 men for the intervention versus control groups were 12.23
278 (95% CI: 11.63, 12.84) at 18-months, 4.80 (95% CI: 3.53, 6.07) at 10-years, 1.38 (95% CI: -0.38, 3.14)
279 at 15-years and 0.86 (95% CI: -1.80, 3.53) at 18-years (**eTable 1**).

280 For age-groups 50-54 compared to 65-69 years, the mean sojourn time increased from 12.1 years to
281 15.3 years, and over-diagnosis from 9.2% to 20.8%, respectively (**eTable 5, eFigures 4-6**).

282 **Adverse Events**

283 Among the deaths due to prostate cancer, 8 (0.7%) in the intervention group and 7 (0.5%) in the
284 control group were related to a diagnostic biopsy or prostate cancer treatment.¹⁰ Other adverse
285 events were reported previously.^{9,11}

286

287 **Discussion**

288 In secondary analysis from this cluster RCT of 415,357 men aged 50-69, compared to usual care
289 control, a single invitation to undergo a PSA test led to an absolute reduction in prostate cancer
290 mortality of 0.09% after a median follow-up of 15 years. However, the magnitude of the effect was
291 small. There was no effect on overall survival. Policy-makers considering screening for prostate
292 cancer should consider this small reduction in deaths against the potential adverse effects
293 associated with over-diagnosis and over-treatment of prostate cancer.^{6,24}

294 This clinical trial previously reported no benefit of a single invitation to PSA screening on the primary
295 outcome of prostate cancer mortality at a median follow-up of ten years.¹⁰ PSA testing is increasingly
296 common,² particularly among men over age 60,^{2,25} and definitive evidence on the benefits and harms
297 of PSA screening remain unclear.²⁴ Analyses reported here are important because of the need for a
298 longer follow up period to evaluate the effect of PSA-detection of prostate cancers,⁵ particularly
299 because findings from the ProtecT trial showed no difference in mortality irrespective of treatment
300 over 15 years.⁶

301 The magnitude of reduction in prostate cancer mortality was smaller than the *a priori* defined effect-
302 size considered important for clinical and public health benefit.¹² The harms of PSA testing include
303 over-diagnosis, biopsy complications,⁹ adverse treatment-effects on urinary, sexual and bowel
304 function,¹¹ and the potential to miss an aggressive prostate cancer.¹⁰ This clinical trial's single
305 invitation to a PSA screen aimed to minimize over-diagnosis and over-treatment compared with
306 other screening trials, but overdiagnosis was still observed after 15-years median follow-up. The
307 European Randomized Study of Prostate Cancer Screening (ERSPC) randomized clinical trial
308 (N=162,243), which combined data from 7 centers with different protocols and screening strategies,
309 reported that PSA screening conducted every 2-4 years (mean of 1.4 tests per participant) reduced
310 prostate cancer mortality after 16 years (rate ratio: 0.80; 95% CI:0.72-0.89).²⁶ The Prostate, Lung,
311 Colorectal and Ovarian (PLCO) randomized clinical trial (N=76,683) reported little evidence of
312 prostate cancer mortality benefit after 17 years with annual PSA testing compared to usual care
313 (rate ratio: 0.93; 95% CI 0.81-1.08),²⁷ but was limited by high rates of PSA testing in the control group
314 (a mean of 2.7 routine PSA tests over the trial's 6 year intervention period²⁸) and only 35%
315 adherence to recommendations for diagnostic biopsy.²⁹ The Stockholm clinical trial compared one-
316 time PSA screening, and diagnostic investigations if PSA>10ng/ml, with an unscreened control group.
317 It demonstrated over-diagnosis of prostate cancer (persistent excess in cumulative prostate cancer
318 incidence in the screening intervention group throughout follow-up), without reduced prostate
319 cancer mortality after 20 years follow-up.³⁰ Multiple screens implemented in ERSPC and PLCO
320 increased over-diagnosis,³¹ with evidence of a strong positive correlation between the extent of the
321 absolute prostate cancer mortality reduction achieved by the screening intervention and the extent
322 of over-diagnosis (quantified as the risk difference in cumulative incidence of prostate cancer
323 between the trial arms).³²

324 **Strengths**

325 This study had several strengths. First, compared to randomizing individual patients, recruitment in
326 general practice clusters is expected to minimize volunteer bias and reduce contamination in the

327 control group, in which the intervention effects also cause greater screening in the control group.
328 Cumulative PSA testing in the control-arm of this clinical trial was indirectly estimated at 10% to 15%
329 over 10-years median follow-up, consistent with current UK policy not to recommend screening. A
330 *priori* estimates suggested that the effect on statistical power of ever undergoing PSA testing during
331 follow-up in the control group (contamination) would be minimal unless the PSA testing rate
332 reached 20%.¹² Second, all practices followed the same screening and diagnosis protocol, providing
333 consistent results. Third, among those with an elevated PSA level, adherence with recommendations
334 for biopsy was high at 85%, similar to ERSPC (81%) and higher than PLCO (35%). This feature of the
335 clinical trial would likely improve screening's potential effectiveness, which depends on patients'
336 willingness to undergo subsequent diagnostic tests. Fourth, the large sample size of this trial
337 contributed to excellent statistical power to detect a clinically meaningful effect size (a prostate
338 cancer mortality RR of 0.87), assuming a that PSA testing in the intervention-arm was between 35%
339 and 50% and that less than 20% of the control group had PSA testing.¹² Fifth, the comprehensive
340 national electronic health record linkage of all the men in this clinical trial helped attain a follow-up
341 rate of 98% over the median 15 year follow-up period.

342 **Limitations**

343 This study had several limitations. First, the screening intervention involved a single invitation for a
344 PSA screening test, which is not typical of organized screening programs. Some advanced prostate
345 cancers that might have been identified in subsequent screening rounds were likely missed. Second,
346 NHS electronic records were used to identify prostate cancer, resulting in missing data for clinical
347 characteristics and possible delay in recording diagnoses. Third, prostate cancer mortality at 15 years
348 was a secondary outcome. Fourth, after this clinical trial began, newer diagnostic methods³³ and
349 more effective treatments for advanced and metastatic prostate cancer³⁴ have been identified. Fifth,
350 few Black men, who are at higher risk of prostate cancer, were included.³⁵

351

352

353 **Conclusions**

354 A single invitation for PSA screening, compared to standard practice without routine screening,
355 reduced the secondary outcome of prostate cancer deaths at a median follow-up of 15-years.
356 However, the absolute reduction in deaths was small.

357

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388 **Role of the Funder/Sponsor**

389 The funders and sponsor had no role in the design and conduct of the study; collection,
390 management, analysis, and interpretation of the data; preparation, review, or approval of the
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392 [§]**CAP trial group.** Group members are listed in **Supplement 2.**

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410

412 **References**

- 413 1. NHS Digital. Cancer Registration Statistics, England 2019
 414 https://www.cancerdata.nhs.uk/incidence_and_mortality (accessed 30th May 2022).
- 415 2. Young GJ, Harrison S, Turner EL, Walsh EI, Oliver SE, Ben-Shlomo Y, Evans S, Lane JA, Neal
 416 DE, Hamdy FC, Donovan JL, Martin RM, Metcalfe C. Prostate-specific antigen (PSA) testing of men in
 417 UK general practice: a 10-year longitudinal cohort study. *BMJ Open* 2017; **7**(10): e017729.
- 418 3. National Cancer Institute. Surveillance E, and End Results Programme. Prostate Cancer.
 419 Cancer Stat Facts <https://seercancer.gov/statfacts/html/prosthtml> 2023 (accessed 15th May 2023).
- 420 4. Johansson J-E, Andrén O, Andersson S-O, Dickman PW, Holmberg L, Magnuson A, Adami H-
 421 O. Natural History of Early, Localized Prostate Cancer. *JAMA* 2004; **291**(22): 2713-9.
- 422 5. Bill-Axelson A, Holmberg L, Garmo H, Taari K, Busch C, Nordling S, Häggman M, Andersson S-
 423 O, Andrén O, Steineck G, Adami H-O, Johansson J-E. Radical Prostatectomy or Watchful Waiting in
 424 Prostate Cancer — 29-Year Follow-up. *New England Journal of Medicine* 2018; **379**(24): 2319-29.
- 425 6. Hamdy FC, Donovan JL, Lane JA, Metcalfe C, Davis M, Turner EL, Martin RM, Young GJ, Walsh
 426 EI, Bryant RJ, Bollina P, Doble A, Doherty A, Gillatt D, Gnanapragasam V, Hughes O, Kockelbergh R,
 427 Kynaston H, Paul A, Paez E, Powell P, Rosario DJ, Rowe E, Mason M, Catto JWF, Peters TJ, Oxley J,
 428 Williams NJ, Staffurth J, Neal DE. Fifteen-Year Outcomes after Monitoring, Surgery, or Radiotherapy
 429 for Prostate Cancer. *New England Journal of Medicine* 2023; **10.1056/NEJMoa2214122**.
- 430 7. Wilt TJ, Brawer MK, Jones KM, Barry MJ, Aronson WJ, Fox S, Gingrich JR, Wei JT, Gilhooly P,
 431 Grob BM, Nsouli I, Iyer P, Cartagena R, Snider G, Roehrborn C, Sharifi R, Blank W, Pandya P, Andriole
 432 GL, Culkin D, Wheeler T. Radical Prostatectomy versus Observation for Localized Prostate Cancer.
 433 *New England Journal of Medicine* 2012; **367**(3): 203-13.
- 434 8. Cancer Research UK. Prostate cancer statistics. [https://www.cancerresearchuk.org/health-
 435 professional/cancer-statistics/statistics-by-cancer-type/prostate-cancer/](https://www.cancerresearchuk.org/health-professional/cancer-statistics/statistics-by-cancer-type/prostate-cancer/). 2021 (accessed 24th May
 436 2021).
- 437 9. Rosario DJ, Lane JA, Metcalfe C, Donovan JL, Doble A, Goodwin L, Davis M, Catto JWF, Avery
 438 K, Neal DE, Hamdy FC. Short term outcomes of prostate biopsy in men tested for cancer by prostate
 439 specific antigen: prospective evaluation within ProtecT study. *BMJ* 2012; **344** d7894.
- 440 10. Martin RM, Donovan JL, Turner EL, Metcalfe C, Young GJ, Walsh EI, Lane JA, Noble S, Oliver
 441 SE, Evans S, Sterne JAC, Holding P, Ben-Shlomo Y, Brindle P, Williams NJ, Hill EM, Ng SY, Toole J,
 442 Tazewell MK, Hughes LJ, Davies CF, Thorn JC, Down E, Davey Smith G, Neal DE, Hamdy FC, for the
 443 CAP Trial Group. Effect of a Low-Intensity PSA-Based Screening Intervention on Prostate Cancer
 444 Mortality: The CAP Randomized Clinical Trial. *JAMA* 2018; **319**(9): 883-95.
- 445 11. Donovan JL, Hamdy FC, Lane JA, Young GJ, Metcalfe C, Walsh EI, Davis M, Steuart-Feilding T,
 446 Blazeby JM, Avery KNL, Martin RM, Bollina P, Doble A, Doherty A, Gillatt D, Gnanapragasam V,
 447 Hughes O, Kockelbergh R, Kynaston H, Paul A, Paez E, Powell P, Rosario DJ, Rowe E, Mason M, Catto
 448 JWF, Peters TJ, Wade J, Turner EL, Williams NJ, Oxley J, Staffurth J, Bryant RJ, Neal DE. Patient-
 449 Reported Outcomes 12 Years after Localized Prostate Cancer Treatment. *NEJM Evidence* 2023; **2**(4):
 450 doi:10.1056/EVIDoa2300018.
- 451 12. Turner EL, Metcalfe C, Donovan JL, Noble S, Sterne JAC, Lane JA, Avery KN, Down L, Walsh E,
 452 Davis M, Ben-Shlomo Y, Oliver SE, Evans S, Brindle P, Williams NJ, Hughes LJ, Hill EM, Davies C, Ng SY,
 453 Neal DE, Hamdy FC, Martin RM, the CAP trial group. Design and preliminary recruitment results of
 454 the Cluster randomised triAl of PSA testing for Prostate cancer (CAP). *British Journal of Cancer* 2014;
 455 **110**(12): 2829-36.
- 456 13. Donovan JL, Opmeer B, Young GJ, Mills N, Martin RM, Lane JA, Metcalfe C, Peters TJ, Davis
 457 M, Turner EL, Walsh E, Neal DE, Hamdy FC, Holding P, Mason M, Catto JWF, Rosario DJ, Staffurth J,
 458 Kynaston H, Hughes O, Bollina P, Doherty A, Gnanapragasam V, Kockelbergh R, Paul A, Paez E, Gillatt
 459 D, Rowe E, Oxley J. Factors associated with trial recruitment, preferences, and treatments received

460 were elucidated in a comprehensive cohort study. *Journal of Clinical Epidemiology* 2019; **113**: 200-
461 13.

462 14. Hamdy FC, Donovan JL, Lane JA, Mason M, Metcalfe C, Holding P, Davis M, Peters TJ, Turner
463 EL, Martin RM, Oxley J, Robinson M, Staffurth J, Walsh E, Bollina P, Catto J, Doble A, Doherty A,
464 Gillatt D, Kockelbergh R, Kynaston H, Paul A, Powell P, Prescott S, Rosario DJ, Rowe E, Neal DE. 10-
465 Year Outcomes after Monitoring, Surgery, or Radiotherapy for Localized Prostate Cancer. *New*
466 *England Journal of Medicine* 2016; **375**(15): 1415-24.

467 15. Burford D, Kirby M, Austoker J. Prostate Cancer Risk Management Programme. Information
468 for primary care; PSA testing in asymptomatic men. Evidence Document. NHS Cancer Screening
469 Programmes: Available at <http://www.cancerscreening.nhs.uk/prostate/pcrmp02.pdf>; 2010
470 (accessed 24th May 2021).

471 16. Keeney E, Sanghera S, Martin RM, Gulati R, Wiklund F, Walsh EI, Donovan JL, Hamdy F, Neal
472 DE, Lane JA, Turner EL, Thom H, Clements MS. Cost-Effectiveness Analysis of Prostate Cancer
473 Screening in the UK: A Decision Model Analysis Based on the CAP Trial. *PharmacoEconomics* 2022;
474 doi.org/10.1007/s40273-022-01191-1.

475 17. Donovan JL, Hamdy FC, Lane JA, Mason M, Metcalfe C, Walsh E, Blazeby JM, Peters TJ,
476 Holding P, Bonnington S, Lennon T, Bradshaw L, Cooper D, Herbert P, Howson J, Jones A, Lyons N,
477 Salter E, Thompson P, Tidball S, Blaikie J, Gray C, Bollina P, Catto J, Doble A, Doherty A, Gillatt D,
478 Kockelbergh R, Kynaston H, Paul A, Powell P, Prescott S, Rosario DJ, Rowe E, Davis M, Turner EL,
479 Martin RM, Neal DE. Patient-Reported Outcomes after Monitoring, Surgery, or Radiotherapy for
480 Prostate Cancer. *New England Journal of Medicine* 2016; **375**(15): 1425-37.

481 18. Turner EL, Metcalfe C, Donovan JL, Noble S, Sterne JAC, Lane JA, I Walsh E, Hill EM, Down L,
482 Ben-Shlomo Y, Oliver SE, Evans S, Brindle P, Williams NJ, Hughes LJ, Davies CF, Ng SY, Neal DE,
483 Hamdy FC, Albertsen P, Reid CM, Oxley J, McFarlane J, Robinson MC, Adolfsson J, Zietman A, Baum
484 M, Koupparis A, Martin RM. Contemporary accuracy of death certificates for coding prostate cancer
485 as a cause of death: Is reliance on death certification good enough? A comparison with blinded
486 review by an independent cause of death evaluation committee. *Br J Cancer* 2016; **115**(1): 90-4.

487 19. Williams N, Hill E, Ng S, Martin R, Metcalfe C, Donovan J, Evans S, Hughes L, Davies C, Hamdy
488 F, Neal D, Turner E, CAP Cause of Death Committee. Standardisation of information submitted to an
489 endpoint committee for cause of death assignment in a cancer screening trial - lessons learnt from
490 CAP (Cluster randomised trial of PSA testing for Prostate cancer). *BMC Medical Research*
491 *Methodology* 2015; **15**(1): 6.

492 20. National Disease Registration Service. [https://digital.nhs.uk/services/national-disease-
493 registration-service](https://digital.nhs.uk/services/national-disease-registration-service). 2023 (accessed 21st November 2023).

494 21. Public Health Wales. Data and analysis. [https://phwnhswales/services-and-
495 teams/observatory/data-and-analysis/](https://phwnhswales/services-and-teams/observatory/data-and-analysis/) 2023 (accessed 21st November 2023).

496 22. Pashayan N, Duffy SW, Pharoah P, Greenberg D, Donovan J, Martin RM, Hamdy F, Neal DE.
497 Mean sojourn time, overdiagnosis, and reduction in advanced stage prostate cancer due to
498 screening with PSA: implications of sojourn time on screening. *Br J Cancer* 2009; **100**(7): 1198-204.

499 23. Bhatt R, van den Hout A, Pashayan N. A multistate survival model of the natural history of
500 cancer using data from screened and unscreened population. *Stat Med* 2021; **40**(16): 3791-807.

501 24. Ilic D, Djulbegovic M, Jung JH, Hwang EC, Zhou Q, Cleves A, Agoritsas T, Dahm P. Prostate
502 cancer screening with prostate-specific antigen (PSA) test: a systematic review and meta-analysis.
503 *BMJ* 2018; **362**: k3519.

504 25. Vickers A, O'Brien F, Montorsi F, Galvin D, Bratt O, Carlsson S, Catto JW, Krilaviciute A,
505 Philbin M, Albers P. Current policies on early detection of prostate cancer create overdiagnosis and
506 inequity with minimal benefit. *BMJ* 2023; **381**: e071082.

507 26. Hugosson J, Roobol MJ, Månsson M, Tammela TLJ, Zappa M, Nelen V, Kwiatkowski M, Lujan
508 M, Carlsson SV, Talala KM, Lilja H, Denis LJ, Recker F, Paez A, Puliti D, Villers A, Rebillard X,
509 Kilpeläinen TP, Stenman UH, Godtman RA, Stinesen Kollberg K, Moss SM, Kujala P, Taari K, Huber A,
510 van der Kwast T, Heijnsdijk EA, Bangma C, De Koning HJ, Schröder FH, Auvinen A. A 16-yr Follow-up

511 of the European Randomized study of Screening for Prostate Cancer. *European Urology* 2019; **76**(1):
512 43-51.

513 27. Pinsky PF, Miller E, Prorok P, Grubb R, Crawford ED, Andriole G. Extended follow-up for
514 prostate cancer incidence and mortality among participants in the Prostate, Lung, Colorectal and
515 Ovarian randomized cancer screening trial. *BJU international* 2019; **123**(5): 854-60.

516 28. Gulati R, Tsodikov A, Wever EM, Mariotto AB, Heijnsdijk EA, Katcher J, de Koning HJ, Etzioni
517 R. The impact of PLCO control arm contamination on perceived PSA screening efficacy. *Cancer*
518 *causes & control : CCC* 2012; **23**(6): 827-35.

519 29. Grubb III RL, Pinsky PF, Greenlee RT, Izmirlian G, Miller AB, Hickey TP, Riley TL, Mabie JE,
520 Levin DL, Chia D, Kramer BS, Reding DJ, Church TR, Yokochi LA, Kvale PA, Weissfeld JL, Urban DA,
521 Buys SS, Gelmann EP, Ragard LR, Crawford ED, Prorok PC, Gohagan JK, Berg CD, Andriole GL.
522 Prostate cancer screening in the Prostate, Lung, Colorectal and Ovarian cancer screening trial:
523 update on findings from the initial four rounds of screening in a randomized trial. *BJU international*
524 2008; **102**(11): 1524-30.

525 30. Lundgren PO, Kjellman A, Norming U, Gustafsson O. Long-Term Outcome of a Single
526 Intervention Population Based Prostate Cancer Screening Study. *J Urol* 2018; **200**(1): 82-8.

527 31. Heijnsdijk EAM, de Carvalho TM, Auvinen A, Zappa M, Nelen V, Kwiatkowski M, Villers A,
528 Páez A, Moss SM, Tammela TLJ, Recker F, Denis L, Carlsson SV, Wever EM, Bangma CH, Schröder FH,
529 Roobol MJ, Hugosson J, de Koning HJ. Cost-effectiveness of Prostate Cancer Screening: A Simulation
530 Study Based on ERSPC Data. *Journal of the National Cancer Institute* 2015; **107**(1).

531 32. Auvinen A, Moss SM, Tammela TLJ, Taari K, Roobol MJ, Schröder FH, Bangma CH, Carlsson S,
532 Aus G, Zappa M, Puliti D, Denis LJ, Nelen V, Kwiatkowski M, Randazzo M, Paez A, Lujan M, Hugosson
533 J. Absolute Effect of Prostate Cancer Screening: Balance of Benefits and Harms by Center within the
534 European Randomized Study of Prostate Cancer Screening. *Clinical Cancer Research* 2016; **22**(1):
535 243-9.

536 33. Ahmed HU, El-Shater Bosaily A, Brown LC, Gabe R, Kaplan R, Parmar MK, Collaco-Moraes Y,
537 Ward K, Hindley RG, Freeman A, Kirkham AP, Oldroyd R, Parker C, Emberton M. Diagnostic accuracy
538 of multi-parametric MRI and TRUS biopsy in prostate cancer (PROMIS): a paired validating
539 confirmatory study. *The Lancet* 2017; **389**(10071): 815-22.

540 34. de Bono J, Mateo J, Fizazi K, Saad F, Shore N, Sandhu S, Chi KN, Sartor O, Agarwal N, Olmos
541 D, Thiery-Vuillemin A, Twardowski P, Mehra N, Goessl C, Kang J, Burgents J, Wu W, Kohlmann A,
542 Adelman CA, Hussain M. Olaparib for Metastatic Castration-Resistant Prostate Cancer. *New England*
543 *Journal of Medicine* 2020; **382**(22): 2091-102.

544 35. Ben-Shlomo Y, Evans S, Ibrahim F, Patel B, Anson K, Chingwundoh F, Corbishley C, Dorling
545 D, Thomas B, Gillatt D, Kirby R, Muir G, Nargund V, Popert R, Metcalfe C, Persad R. The Risk of
546 Prostate Cancer amongst Black Men in the United Kingdom: The PROCESS Cohort Study. *European*
547 *Urology* 2008; **53**(1): 99-105.

548

549 **Figure Titles and Footnotes**

550

551 **Figure 1: Recruitment, randomization, and flow of practices and patients in a trial of PSA testing**
552 **for prostate cancer**

553

554

555 Footnotes

556 *Shaded boxes: Flow of GP practices through trial recruitment; unshaded boxes: flow of men through trial*
557 *recruitment; ^a Pseudo-anonymised follow-up; ^b NHS digital national data opt-outs (previously type-2 opt-outs)*
558 *preventing NHS data being used for research. [https://www.nhs.uk/using-the-nhs/about-the-nhs/opt-out-of-](https://www.nhs.uk/using-the-nhs/about-the-nhs/opt-out-of-sharing-your-health-records/)*
559 *[sharing-your-health-records/](https://www.nhs.uk/using-the-nhs/about-the-nhs/opt-out-of-sharing-your-health-records/)*

560 **Practices were randomized prior to invitation to take part in the trial. Randomization was blocked and*
561 *stratified by geographical area based on groups of 10-12 neighboring primary care practices and using a*
562 *computerized random number generator to allocate near-equal number of practices in each stratum to*
563 *intervention and comparison groups. The intervention was a single invitation to prostate specific antigen (PSA)*
564 *screening.*

565 ***Numbers of men are as of November 2021 and are subject to small changes over time because of continued*
566 *updates from NHSD e.g. changes to the trace status of the men (e.g. men newly successfully traced). Note that*
567 *not all men traced at 15 years were traced at 10 years.*

568 *Follow-up was through routine NHS electronic vital status and cancer registry databases for diagnoses and*
569 *deaths notified by Nov 2021 but that occurred up to 31st March 2021.*

570 **Figure 2: The Effect of the Trial Intervention on the Cumulative Incidence of Prostate Cancer**
571 **Mortality and Diagnosis, and All-Cause Mortality After a Median 15-Years Follow-Up. The**
572 **intervention was a single invitation to PSA screening.**

573

574 **Figure 2A: Prostate cancer mortality, by group**

575 **Figure 2B: All-cause mortality, by group**

576 **Figure 2C: Prostate cancer detection, by group**

577

578 Footnote

579 P-values from random-effects Poisson model (see Statistical Analysis section).

580

581

582 **Table 1: Individual and practice level characteristics at baseline amongst consented GP practices**
583 **and men included in the analysis (adapted from Turner et al¹² and Martin et al.¹⁰)**
584

	Intervention group	Control group
Individual Characteristics	n= 189,326 men	n= 219,395 men
Median age (IQR)	58.5 (54.3, 63.5)	58.6 (54.3, 63.5)
Median Index of Multiple Deprivation score, England (IQR)	17.5 (10.1, 33.2)	16.9 (9.8, 32.4)
Median Index of Multiple Deprivation score, Wales (IQR)	17.6 (9.2, 29.5)	13.7 (7.1, 29)
Urban area (%) ^a	163,701 (86%)	189,667 (86%)
Race (%White) ^b	98% ^b	Not available
Practice Characteristics	n= 271 practices	n= 302 practices
Median practice list size (IQR) ^c	6,300 (4,150, 9,107)	6,300 (3,793, 9,000)
Number of urban practices (%)	244 (90%)	267 (88%)
Number of multiple partner GP practices (%)	242 (89%)	267 (88%)
Single partner practices ^d	21 (8%)	29 (10%)
Small practices (2-3)	60 (22%)	61 (20%)
Medium/large practices (4+)	128 (47%)	146 (48%)
Missing	62 (23%)	66 (22%)
Median QOF points achieved (%) ^e (IQR); n	98.9 (97.4, 99.6); 224	99 (97.4, 99.7); 266
Median Index of Multiple Deprivation score, England (IQR); n	21.8 (12.7, 44.1); 231	23.6 (13.3, 46.7); 271
Median Index of Multiple Deprivation score, Wales (IQR); n	18.8 (11.9, 22.9); 40	20.1 (7.6, 34.5); 31
<i>Mean prevalence^f, %</i>		
All cancers (s.d)	0.6 (0.3)	0.5 (0.2)
Diabetes (s.d)	3.6 (1.0)	3.7 (1.0)
Obesity (s.d)	8.0 (2.8)	7.8 (2.8)
Coronary heart disease (s.d)	4.1 (1.4)	3.9 (1.3)

585 *Index of Multiple Deprivation, a measure of relative deprivation for small areas: a higher score indicates more*
586 *deprivation, range 0-100. English and Welsh IMD scores are not directly comparable and are reported*
587 *separately. The Index of Multiple Deprivation for the practice refers to the area of the practice not where*
588 *patients live; QOF = Quality and Outcomes Framework, a system for performance management and payment*
589 *of GPs based on the quality of their care: data are % of total QOF points achieved; IQR = interquartile range*
590 *(25th percentile, 75th percentile); s.e. = standard error; ^aRural/urban classification 2004, a measure of*
591 *population density and sparseness, urban defined as areas >10,000 people; ^bRace/ethnicity for men attending*
592 *the intervention group PSA test clinic were ascertained by a nurse using a standardized questionnaire as one of*
593 *a range of baseline characteristics to assess generalisability.¹³ Race/ethnicity were defined using UK Office for*
594 *National Statistics Census categories and recoded as White and Other (all other categories collapsed due to low*
595 *numbers of non-White participants). Race/ethnicity data were not available from NHS routine data we had*
596 *access to at the time, so we could not compute these data for the control group. ^cThe total number of*
597 *individuals registered at GP practices (primary care practices). ^dSingle partner GP practices are primary care*
598 *practices with a single General Practitioner registered and practicing from there. ^eBased on 2007/2008 data,*

599 *England only. Quality and Outcomes Framework (QoF) scores are measured from 135 indicators and one*
600 *measure of depth of care (holistic care) and are split across clinical, organisational, patient experience and*
601 *additional services domains (maximum score 1,000 points).^fThe prevalence of medical conditions across*
602 *practices obtained from the clinical domain indicators of QoF: practices reported counts of patients with each*
603 *condition and practice list size, enabling calculation of mean prevalence.*

604

605 **Table 2: Effect of the trial intervention on prostate cancer specific and all-cause mortality and prostate cancer diagnosis by random allocation and by instrumental**
 606 **variable analysis, after a median 15-years of follow-up (median 10-year estimate can be obtained from Martin et al¹⁰).**

Intervention group n=189,326; 2,543,298 person years)				Control group (n=219,395; 2,885,418 person-years)			Estimated effect of intervention versus control		
Events	Rate/1000 person years (95% CI)	Risk [%] at 15 years (95% CI) ^a	Events	Rate/1000 person years (95% CI)	Risk [%] at 15 years (95% CI) ^a	Risk difference [%] at 15 years (95% CI)	Rate ratio (95% CI) ^b	P value ^{b,c}	
15-year prostate cancer mortality^d									
As randomized	1,199	0.47 (0.45, 0.50)	0.69 (0.65, 0.73)	1,451	0.50 (0.48, 0.53)	0.78 (0.73, 0.82)	-0.09 (-0.15, -0.03)	0.92 (0.85, 0.99)	0.033
IV analysis ^e	-	-	-	-	-	-	-	0.83 (0.68, 1.00)	0.053
15-year all-cause mortality									
As randomized	45,084	17.7 (17.6, 17.9)	23.2 (23.0, 23.4)	50,336	17.4 (17.3, 17.6)	23.3 (23.1, 23.5)	-0.07 (-0.35, 0.21)	0.97 (0.94, 1.01)	0.11
IV analysis ^e	-	-	-	-	-	-	-	1.01 (0.91, 1.12)	0.85
15-year prostate cancer diagnoses									
As randomized	12,013	4.88 (4.80, 4.97)	7.08 (6.95, 7.21)	12,958	4.60 (4.52, 4.68)	6.94 (6.82, 7.06)	0.14 (-0.04, 0.31)	1.06 (1.02, 1.09)	0.001

607 CI = confidence interval. IV: Instrumental variable. Median follow-up time was 15.43 years (interquartile range: 14.23-16.43; range: 12.19, 19.23). The intervention was a
 608 single invitation to PSA screening. ^aThe numbers of deaths for the cumulative 15-year risk by intervention versus control group are 1,018 and 1,288, respectively. ^bAdjusted
 609 for current age using a lexis diagram approach; variation between randomisation cluster and GP practice accommodated by random effects in a three-level model.
 610 ^cLikelihood ratio test of the null hypothesis “no difference between the groups”. ^dDefined as definite or probable prostate cancer death or intervention related death by an
 611 independent cause of death committee. ^eInstrumental variable analysis to estimate the effect of screening amongst those attending the PSA testing clinic, using a
 612 generalized method of moments (gmm) estimator with random allocation as the instrumental variable.

613
614

615 **Table 3: Exploratory analysis of prostate cancer mortality rate ratios comparing intervention versus control groups, by age and deprivation scores, after a**
616 **median 15-years follow-up**

	Intervention group (n=189,326) Person years = 2,543,298			Control group (n=219,395) Person-years = 2,885,418			Estimated effect of intervention versus control		
	Deaths	Rate/1000 person years (95% CI)	Risk [%] at 15 years (95% CI)	Deaths	Rate/1000 person years (95% CI)	Risk [%] at 15 years (95% CI)	Risk difference [%] at 15 years (95% CI)	Rate ratio (95% CI) ^a	P value for interaction ^a
Age at baseline									
50-54	132	0.17 (0.14, 0.20)	0.22 (0.18, 0.27)	154	0.18 (0.15, 0.21)	0.25 (0.21, 0.30)	-0.03 (-0.09, 0.03)	0.96 (0.76, 1.22)	0.75
55-59	251	0.33 (0.29, 0.38)	0.47 (0.41, 0.54)	300	0.35 (0.31, 0.39)	0.54 (0.47, 0.61)	-0.07 (-0.16, 0.02)	0.92 (0.78, 1.10)	
60-64	368	0.64 (0.58, 0.71)	0.97 (0.87, 1.09)	465	0.70 (0.64, 0.77)	1.10 (1.00, 1.22)	-0.13 (-0.28, 0.02)	0.90 (0.77, 1.04)	
65-69+	448	1.05 (0.96, 1.15)	1.61 (1.45, 1.78)	532	1.07 (0.99, 1.17)	1.76 (1.60, 1.93)	-0.15 (-0.38, 0.08)	0.98 (0.86, 1.12)	
IMD area deprivation tertile England^b									
Most affluent	326	0.44 (0.40, 0.50)	0.61 (0.54, 0.69)	425	0.47 (0.43, 0.52)	0.71 (0.64, 0.79)	-0.11 (-0.21, 0.00)	0.92 (0.79, 1.07)	0.46
Mid-level	373	0.51 (0.46, 0.56)	0.76 (0.68, 0.85)	463	0.53 (0.49, 0.58)	0.84 (0.76, 0.93)	-0.08 (-0.20, 0.04)	0.94 (0.82, 1.07)	
Most deprived	351	0.48 (0.44, 0.54)	0.74 (0.66, 0.83)	444	0.55 (0.50, 0.61)	0.86 (0.77, 0.95)	-0.11 (-0.23, 0.01)	0.85 (0.74, 0.99)	
IMD area deprivation tertile Wales^c									
Most affluent	45	0.41 (0.31, 0.55)	0.52 (0.37, 0.73)	43	0.34 (0.25, 0.46)	0.47 (0.34, 0.65)	+0.05 (-0.19, 0.28)	1.16 (0.76, 1.77)	0.84
Mid-level	48	0.37 (0.28, 0.49)	0.62 (0.46, 0.84)	36	0.40 (0.29, 0.56)	0.60 (0.43, 0.84)	+0.02 (-0.25, 0.30)	0.89 (0.55, 1.43)	
Most deprived	56	0.49 (0.37, 0.63)	0.66 (0.49, 0.89)	39	0.41 (0.30, 0.56)	0.72 (0.52, 1.02)	-0.07 (-0.38, 0.25)	1.23 (0.82, 1.85)	

617 ^aAdjustment for age stratum and practice cluster effects apart from age which was not adjusted for age stratum. ^bIndex of Multiple Deprivation. Scores range from 0 to 100 with
618 higher scores indicating higher levels of deprivation. Tertile 1 has scores ranging from 1.08 to 12.17, tertile 2 has scores ranging 12.18 to 25.95 and tertile 3 has scores ranging from
619 25.97 to 79.98. ^cScores range from 0 to 100 (England and Wales do not share the same scale) with higher scores indicating higher levels of deprivation. Tertile 1 has scores ranging
620 from 1.40 to 10.30, tertile 2 has scores ranging 10.40 to 23.30 and tertile 3 has scores ranging from 23.40 to 78.90.

621

622

n=911 Primary care practices randomised (within 99 geographical areas England and Wales)*

26 Practices excluded in 6 geographical areas:
12 No control group practices provided consent in 3 areas
9 No intervention group practices provided consent in 2 areas
5 Not approached because recruiting center had already closed in 1 area

26 Practices excluded in 6 geographical areas:
12 No control group practices provided consent in 3 areas
9 No intervention group practices provided consent in 2 areas
5 Not approached because recruiting center had already closed in 1 area

Practices randomised to **intervention** group n= 440 (93 areas)

Practices randomised to **control** group n= 419 (93 areas)

Practices excluded: 42
Consented but out of time: 13
Involved in other prostate cancer study: 10
Atypical population/unable to produce list: 8
Ceased to exist: 6
Randomised in error: 5

Practices excluded: 32
Ceased to exist: 19
Involved in other prostate cancer study involving screening: 13

Practices eligible for the **intervention** group n= 398 (93 areas)

Practices eligible for the **control** group n= 387 (93 areas)

Refused: 127 (32%)
Did not respond to invitation: 85
Refused to participate: 42

Refused: 85 (22%)
Did not respond to invitation: 45
Refused to participate: 40

Practices participating in **intervention** group n= 271 (68%)
Median list size: 6,883 (IQR: 9,107 – 4,150 = 4,957)
Men 50-69 years in intervention arm practices n= 197,938

Practices participating in **control** group n= 302 (78%)
Median list size: 6,777 (IQR: 9,000 – 3,793 = 5,207)
Men 50-69 years in control arm practices n= 221,644

2026 Men excluded **

Prostate cancer pre-randomisation	n=1,433
No record of registration with NHS digital	n=257
Death pre-randomisation	n=176
Failed to trace at NHS digital	n=160

2199 men excluded **

Prostate cancer pre-randomisation	n=1,688
Death pre-randomisation	n=286
No record of registration with NHS digital	n=127
Failed to trace at NHS digital	n=95
Refused	n=3

Men eligible in **intervention** group n= 195,912

Men eligible in **control** group n= 219,445

6,586 Men excluded from primary analysis**

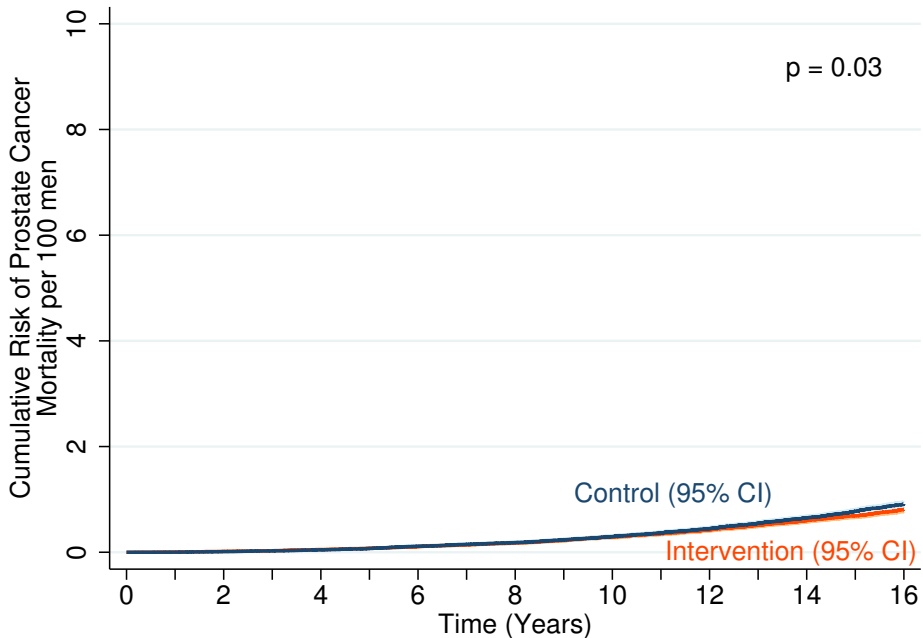
Did not wish to participate	n=6311 ^a
No consent for flagging	n=198
Lost to follow-up including NHS national data opt outs ^b and embarkation	n=62
Event date on list date	n=8
Date of birth missing	n=7

50 Men excluded from primary analysis**

Lost to follow-up including NHS national data opt outs ^b and embarkation	n=45
Event on list date	n=5

Men analysed in **intervention** group n= 189,326

Men analysed in **control** group n= 219,395



Number at risk

Intervention : 189326

Control : 219395

177962

206205

164154

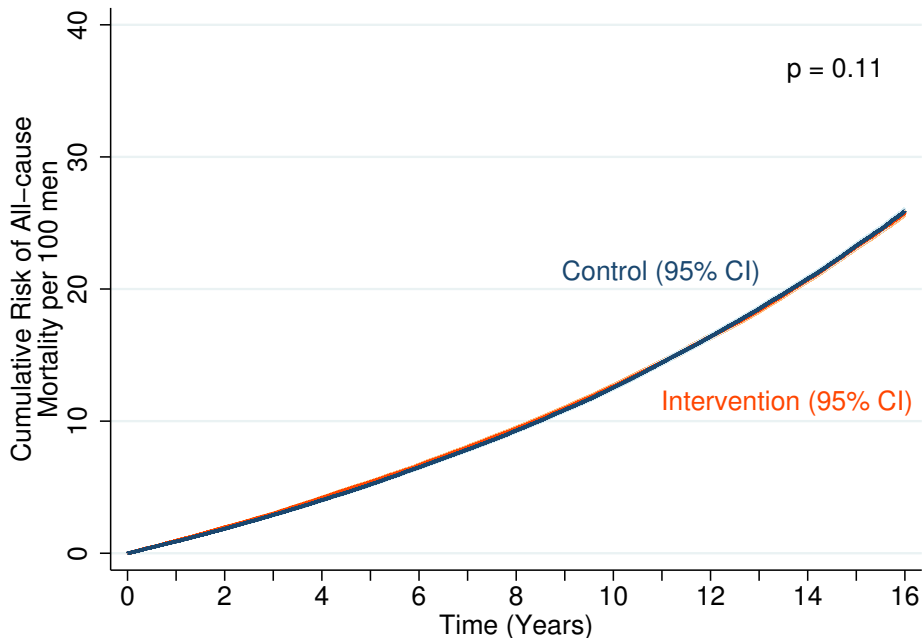
189599

146469

166375

51975

40988



Number at risk

Intervention : 189326

Control : 219395

177962

206205

164154

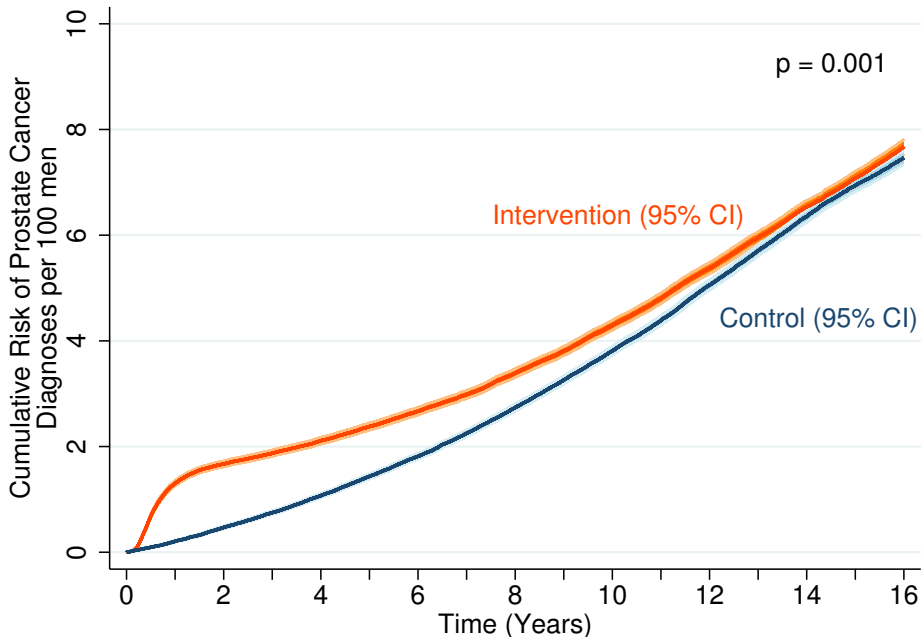
189599

146469

166375

51975

40988



Number at risk

Intervention : 189326

Control : 219395

174289

204203

158876

184887

139138

158863

48427

38396

Supplemental Online Content

Martin RM, Turner EL, Young GJ, et al; CAP Trial Group. Prostate-specific antigen screening and 15-year prostate cancer mortality. *JAMA*. Published online April 10, 2024.

doi:10.1001/jama.2024.4011

eMethods.

eTable 1. Prostate cancer-specific diagnoses and mortality and all-cause mortality at 10-years, 15-years and 18-years post-randomisation (and at 18 months for prostate cancer diagnoses) by random allocation and an as randomized estimate of the difference between groups

eTable 2. Underlying causes of deaths in intervention versus control groups at 15-year median follow-up (not including prostate cancer)

eTable 3. Effect of the CAP trial intervention on characteristics of prostate cancer cases at diagnosis

eTable 4. Sensitivity analyses employing alternative definitions of prostate cancer deaths

eTable 5. Estimated mean and median sojourn time and probability of overdiagnosis

eFigure 1. CAP trial design

eFigure 2. Cumulative incidence of prostate cancer by TNM stage at diagnosis

eFigure 3. Cumulative incidence of prostate cancer by Gleason score at diagnosis

eFigure 4. Comparing simulated data to empirical data for the cumulative prostate cancer incidence and cancer-specific and all-other cause mortality risk among the screened men and the unscreened group

eFigure 5. Comparison number of subjects per 100, 000 cohorts at death from all causes by ages between simulated data and CAP data

eFigure 6. Transition diagram for multi-state survival models

This supplemental material has been provided by the authors to give readers additional information about their work.

Methods used to estimate overdiagnosis and mean sojourn time.

We simulated a cohort of three million men aged 50-69 years and followed to death, calibrated against CAP data – prostate cancer incidence rate, and cancer-specific and all-other cause mortality rates (**eFigure 4**) and age at death (**eFigure 5**). We applied multistate survival model with parametric hazards and the following states: healthy, screen-detectable, screen-detected, clinically diagnosed, cancer-specific death, and all-other cause deaths, to estimate the natural history parameters and time to death after a cancer diagnosis (**eFigure 6**). The transition between healthy and screen-detectable states was assumed to follow the Weibull distribution, while other transitions were assumed to follow the Gompertz distribution. We estimated the transition hazards between the states and the misclassification of states (i.e., 1-episode sensitivity¹) by maximising the likelihood functions.² We derived the mean sojourn time and overdiagnosis from microsimulation using the estimated transition parameters and one-off screening between ages 50 to 69 and assuming 85% of men with elevated PSA level undertake biopsy. We calculated the sojourn time as the length of time in the screen-detectable state given a transition to a clinically diagnosed state (i.e. the time by which diagnosis is advanced by screening [lead time]). We estimated overdiagnosed cases as the difference in cumulative prostate cancer incidence between screened and unscreened groups over lifetime. The probability of overdiagnosis was the fraction overdiagnosed among screen-detected cases.

Supplementary Tables

eTable 1: Prostate cancer-specific diagnoses and mortality and all-cause mortality at 10-years, 15-years and 18-years post-randomisation (and at 18 months for prostate cancer diagnoses) by random allocation and an as randomized estimate of the difference between groups.

	Intervention group (<i>n</i> =189,326)		Control group (<i>n</i> =219,395)		Cumulative risk difference per 1000 men (95% CI)
	N	Cumulative risk per 1000 men (95% CI)	N	Cumulative risk per 1000 men (95% CI)	
Prostate cancer mortality					
At 10-years	488	2.89 (2.65, 3.16)	575	2.95 (2.72, 3.21)	-0.06 (-0.41, 0.29)
At 15-years	1,018	6.90 (6.48, 7.34)	1,288	7.76 (7.34, 8.21)	-0.86 (-1.48, -0.25)
At 18-years	1,185	10.92 (10.14, 11.76)	1,440	12.09 (11.19, 13.07)	-1.17 (-2.41, 0.07)
All-cause mortality					
At 10-years	23,212	126.30 (124.79, 127.83)	26,581	125.37 (123.97, 126.79)	0.92 (-1.15, 3.00)
At 15-years	40,001	232.08 (230.06, 234.12)	46,073	232.75 (230.86, 234.65)	-0.68 (-3.46, 2.10)
At 18-years	44,747	316.15 (313.03, 319.29)	50,045	320.46 (316.91, 324.03)	-4.27 (-9.01, 0.47)
Prostate cancer diagnoses					
At 18-months	2,912	15.51 (14.96, 16.08)	711	3.28 (3.05, 3.53)	12.23 (11.63, 12.84)
At 10-years	7,558	42.92 (41.98, 43.88)	7,554	38.12 (37.28, 38.97)	4.80 (3.53, 6.07)
At 15-years	11,291	70.78 (69.51, 72.08)	12,368	69.40 (68.21, 70.62)	1.38 (-0.38, 3.14)
At 18-years	12,001	86.30 (84.53, 88.12)	12,938	85.44 (83.48, 87.44)	0.86 (-1.80, 3.53)

N is numbers of deaths and diagnoses as shown in the row headers. CI: Confidence interval. This table differs from Table 2, in that it reports cumulative risks at specific time points (10, 15 and 18 years), while Table 2 reports the data after a median 15 years of follow-up (range: 12.2 to 19.2 years).

eTable 2: Underlying causes of death^a in intervention versus control groups at 15-year median follow-up (not including prostate cancer).

Cause of death	Intervention n (%)	Control n (%)
Any (not incl. prostate cancer)	43,885 (100%)	48,885 (100%)
Other cancers	16,553(38%)	18,440 (38%)
Circulatory diseases	12,419 (28%)	13,662 (28%)
Respiratory disease	5,287 (12%)	5,796 (12%)
Digestive disease	2,316 (5%)	2,612 (5%)
Infectious disease	385 (1%)	402 (1%)
Genitourinary diseases	445 (1%)	503 (1%)
Blood, immune, endocrine	644 (1%)	736 (2%)
Nervous system disease	1,862 (4%)	2,217 (5%)
Accident	1,126 (3%)	1,278 (3%)
Other	2,705(6%)	3,074 (6%)
No ICD10 code	143 (<1%)	165 (<1%)

^aUnderlying cause of death for non-prostate cancer deaths was determined by death certificate.

There were 95,420 all-cause deaths in total, including 308 deaths without an ICD10 code and 2,650 prostate cancer deaths (N=92,462 non prostate cancer deaths with an ICD-10 code).

eTable 3. Effect of the CAP trial intervention on characteristics of prostate cancer cases at diagnosis^a

		Intervention group			Controls
		Attended PSA clinic 75,694	Did not attend PSA clinic 113,632	All invited 189,326	219,395
Number of prostate cancers (%)^b		6,554 (8.7%)	5,459 (4.8%)	12,013 (6.3%)	12,958 (5.9%)
Clinical characteristics at diagnosis					
Person-years of follow up		1,043,530	1,416,377	2,459,907	2,815,181
Rate per 1000-person years		6.28 (6.13, 6.43)	3.85 (3.75, 3.96)	4.88 (4.80, 4.97)	4.60 (4.52, 4.68)
Mean age (SD)		67.28 (6.54)	69.21 (5.91)	68.16 (6.33)	69.38 (5.90)
Median years between randomization and diagnosis (IQR)		5.90 (0.67, 11.26)	9.15 (5.25, 12.15)	7.84 (1.69, 11.76)	8.93 (5.29, 12.02)
Grade (%)					
	Grade recorded ^c	5,991 (91.4%)	4,769 (87.4%)	10,760 (89.6%)	11,501 (88.8%)
	≤6 ^b	2,704 (3.6%)	1,407 (1.2%)	4,111 (2.2%)	3,482 (1.6%)
	7 ^b	2,305 (3.0%)	2,097 (1.8%)	4,402 (2.3%)	5,082 (2.3%)
	3+4	1,011 (1.3%)	1,074 (0.9%)	2,085 (1.1%)	2,708 (1.2%)
	4+3	468 (0.6%)	570 (0.5%)	1,038 (0.5%)	1,443 (0.7%)
	Unknown ^d	826 (1.1%)	453 (0.4%)	1,279 (0.7%)	931 (0.4%)
	≥8 ^b	982 (1.3%)	1,265 (1.1%)	2,247 (1.2%)	2,937 (1.3%)
Stage (%)					
	Stage recorded ^c	5,952 (90.8%)	4,933 (90.4%)	10,885 (90.6%)	11,945 (92.2%)
	T1/T2 ^b	4,227 (5.6%)	2,647 (2.3%)	6,874 (3.6%)	6,746 (3.1%)
	T3 ^b	1,160 (1.5%)	1,146 (1.0%)	2,306 (1.2%)	2,871 (1.3%)
	T4/N1/M1 ^b	565 (0.7%)	1,140 (1.0%)	1,705 (0.9%)	2,328 (1.1%)

The intervention was a single invitation to PSA screening. The PSA clinic was the clinic men were invited to have the PSA test explained, consider having a PSA test and give written informed consent with a 24-hour period cooling off period.

IQR = interquartile range (25th percentile, 75th percentile). CI = confidence interval. ^aDiagnoses were collected from routine data sources, NHS England data were used in the first instance (n=23,415 cancers) and additional cases were included if present in data provided by Public Health Wales (n=930) or the National Disease Registration Service (NDRS, formerly Public Health England) (n=626). ^bDenominators are column header totals. ^cDenominators are N of prostate cancers in each column. ^dMissing primary and secondary Gleason grade to enable 3+4 and 4+3 subdivision.

eTable 4: Sensitivity analyses employing alternative definitions of prostate cancer deaths.

	Intervention group (n=189,326) Person years=2,543,298		Control group (n=219,395) Person years=2,885,418		As randomized estimate	
	Events	Rate/1000 person years (95% CI)	Events	Rate/1000 person years (95% CI)	Rate ratio (95% CI)	P value ^a
Including 'possible' prostate cancer death ^d	1230	0.48 (0.46, 0.51)	1498	0.52 (0.49, 0.55)	0.91 (0.85, 0.99)	P=0.020
Definite prostate cancer death only ^e	1028	0.40 (0.38, 0.43)	1254	0.43 (0.41, 0.46)	0.91 (0.84, 0.99)	P=0.030

^aLikelihood ratio test of the null hypothesis “no difference in prostate cancer mortality between the groups”, adjusted for randomisation cluster and age using a lexis diagram approach. ^bDefined as definite, probable or possible prostate cancer death or intervention related death by an independent cause of death committee. ^cDefined as definite prostate cancer death or intervention related death by an independent cause of death committee.

eTable 5: Estimated mean and median sojourn time and probability of overdiagnosis.

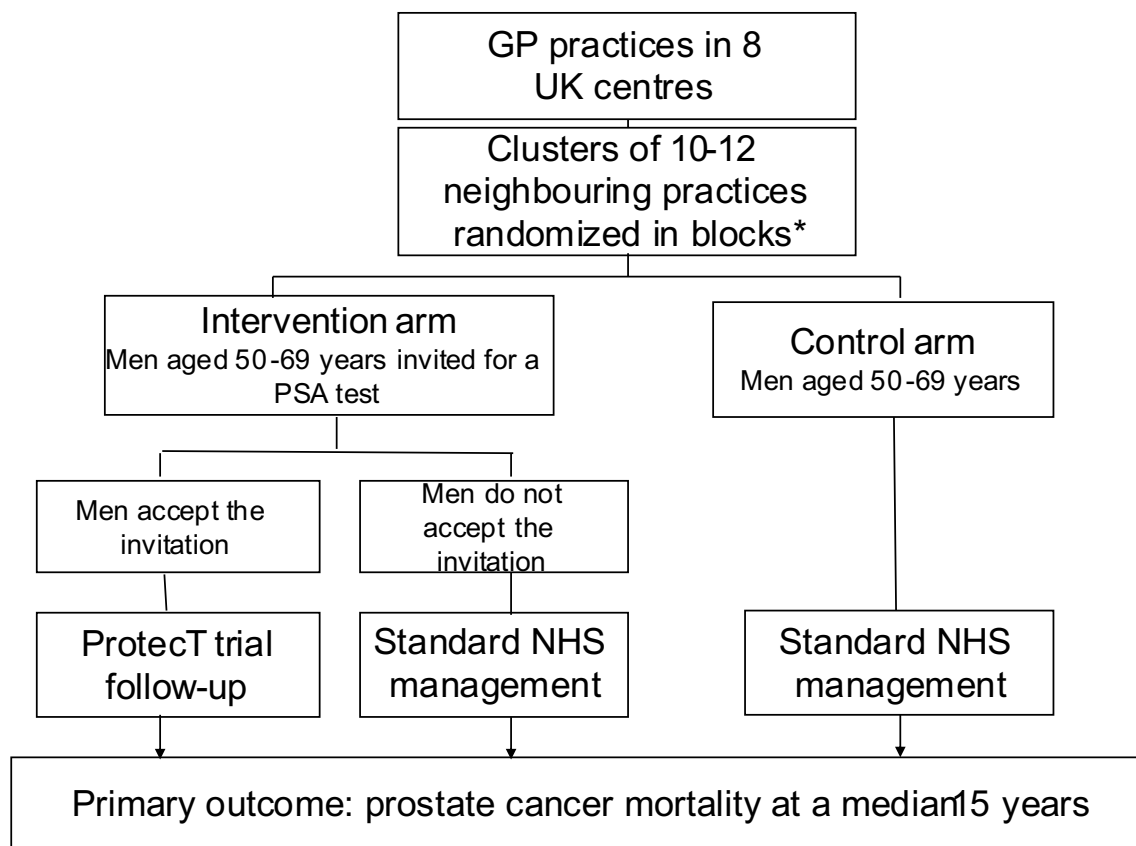
Age group	Mean sojourn time (years)	95% confidence interval (years)	Median sojourn time (years)	Interquartile range (years)
50-54	12.1	12.1 – 12.2	10.6	5.0 – 17.5
55-59	13.2	13.1 – 13.2	11.9	5.5 – 19.3
60-64	14.2	14.2 – 14.3	13.0	5.9 – 21.4
65-69	15.3	15.2 – 15.3	13.8	6.2 – 23.4
50-69	13.4	13.4 -13.4	12.0	5.5 -19.8
	Mean overdiagnosis %	95% confidence interval (%)	Median overdiagnosis %	Interquartile range %
50-54	9.2	8.9 – 9.4	9.3	8.0 – 10.4
55-59	13.3	13.1 – 13.5	13.4	12.4 – 14.3
60-64	17.1	17.0 – 17.3	17.2	16.4 – 17.9
65-69	20.8	20.6 – 21.0	20.8	20.0 – 21.3
50-69	15.0	14.4 – 15.5	14.8	13.6 – 15.8

The sojourn time represents the duration of the preclinical screen-detectable period for each of the 3 million men who transition from screen-detectable to clinically diagnosed state. Sojourn time varies between individuals.

*Overdiagnosis estimates are based on simulation of 200 cohorts of 3 million men aged 50 to 69 followed to death. The episode sensitivity¹ (the ability of the full diagnostic process – testing and biopsy – to find cancer in the detectable preclinical phase) increased from 50.0% to 85.3% for ages 50 to 69.

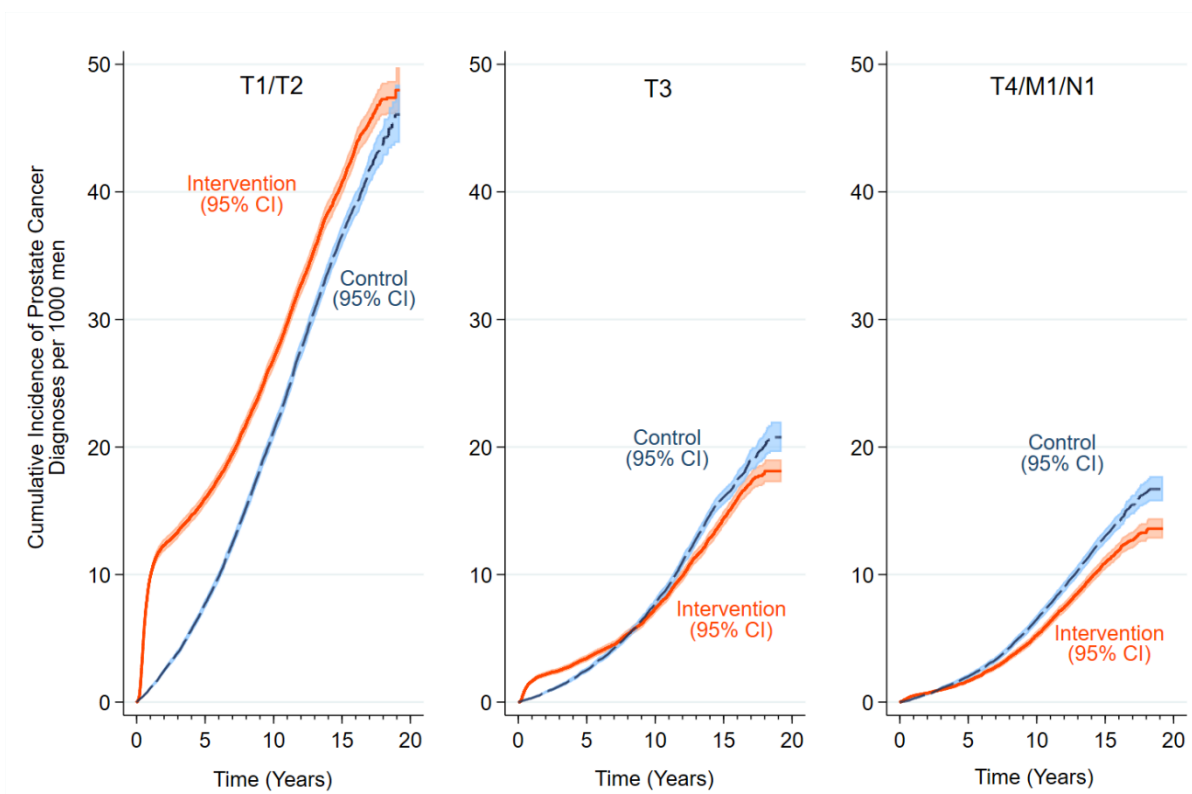
Supplementary Figures

eFigure 1: CAP trial design.



CAP is a UK-wide cluster RCT in which 573 GP practices in 8 UK centres (Sheffield, Newcastle, Bristol, Birmingham, Cardiff, Leeds, Cambridge, Leicester) were randomised and consented to either PSA testing and prostate cancer diagnosis (ProtecT trial) or the routine-practice comparison arm. Pre-specified Prostate cancer mortality outcomes were collected at a median 10-years (reached 31st March 2016) and 15-years (reached 31st March 2021) follow-up.

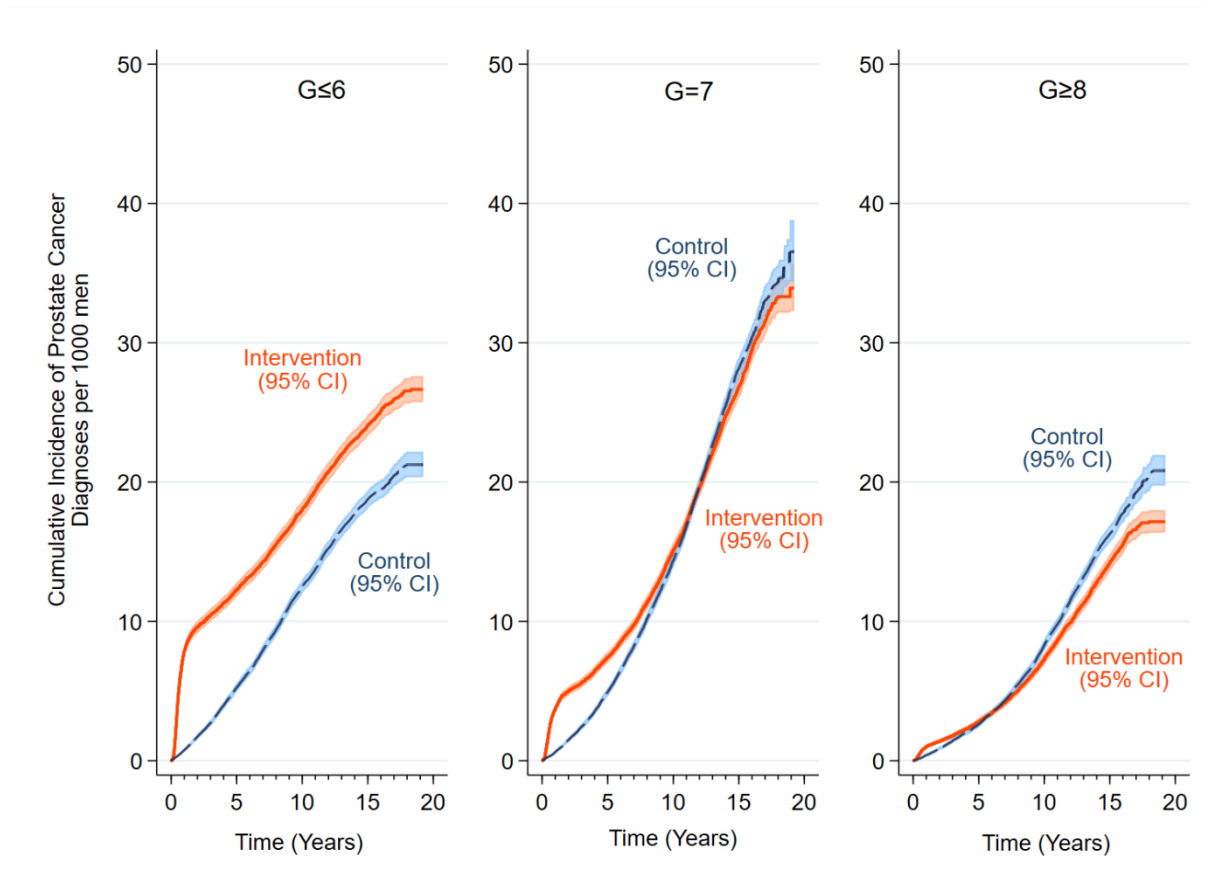
eFigure 2: Cumulative incidence of prostate cancer by TNM stage at diagnosis.



Number at risk at the start of each two-year period (number of prostate cancer diagnoses in that period)												
Time (year)	Median (IQR) follow up	0	2	4	6	8	10	12	14	16	18	20
A: Clinical stage T1/T2												
Intervention	14.20 (11.42, 16.04)	189,326 (2302)	180,957 (418)	174,289 (514)	167,024 (697)	158,876 (805)	149,145 (868)	139,138 (763)	103,163 (386)	48,427 (119)	12,794 (2)	0 (0)
Control	14.35 (11.09, 15.67)	219,395 (531)	212,352 (665)	204,203 (844)	194,558 (1,044)	184,887 (1,083)	172,125 (1,079)	158,863 (936)	119,810 (459)	38,396 (94)	9,687 (11)	0 (0)
B: Clinical stage T3												
Intervention	14.20 (11.42, 16.04)	189,326 (404)	180,957 (136)	174,289 (183)	167,024 (220)	158,876 (302)	149,145 (374)	139,138 (388)	103,163 (230)	48,427 (68)	12,794 (1)	0 (0)
Control	14.35 (11.09, 15.67)	219,395 (168)	212,352 (223)	204,203 (300)	194,558 (357)	184,887 (447)	172,125 (547)	158,863 (541)	119,810 (230)	38,396 (52)	9,687 (6)	0 (0)
C: Clinical stage T4/M1/N1^a												
Intervention	14.20 (11.42, 16.04)	189,326 (131)	180,957 (97)	174,289 (152)	167,024 (223)	158,876 (266)	149,145 (316)	139,138 (309)	103,163 (161)	48,427 (46)	12,794 (4)	0 (0)
Control	14.35 (11.09, 15.67)	219,395 (133)	212,352 (181)	204,203 (227)	194,558 (314)	184,887 (397)	172,125 (417)	158,863 (402)	119,810 (207)	38,396 (47)	9,787 (3)	0 (0)

CI: confidence interval, IQR: interquartile range, ^aIf any of these conditions were satisfied patients were categorized as T4, e.g. a patient with T3, N0 and M1 would be categorized as T4/N1/M1.

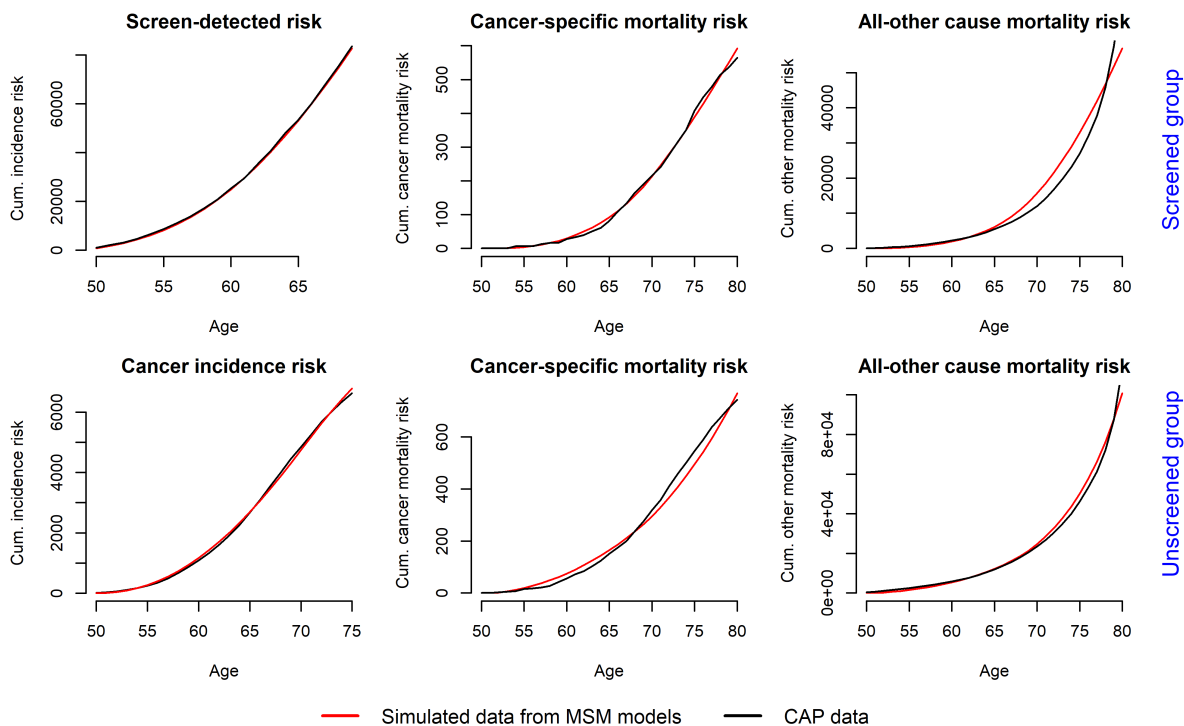
eFigure 3: Cumulative incidence of prostate cancer by Gleason score at diagnosis.



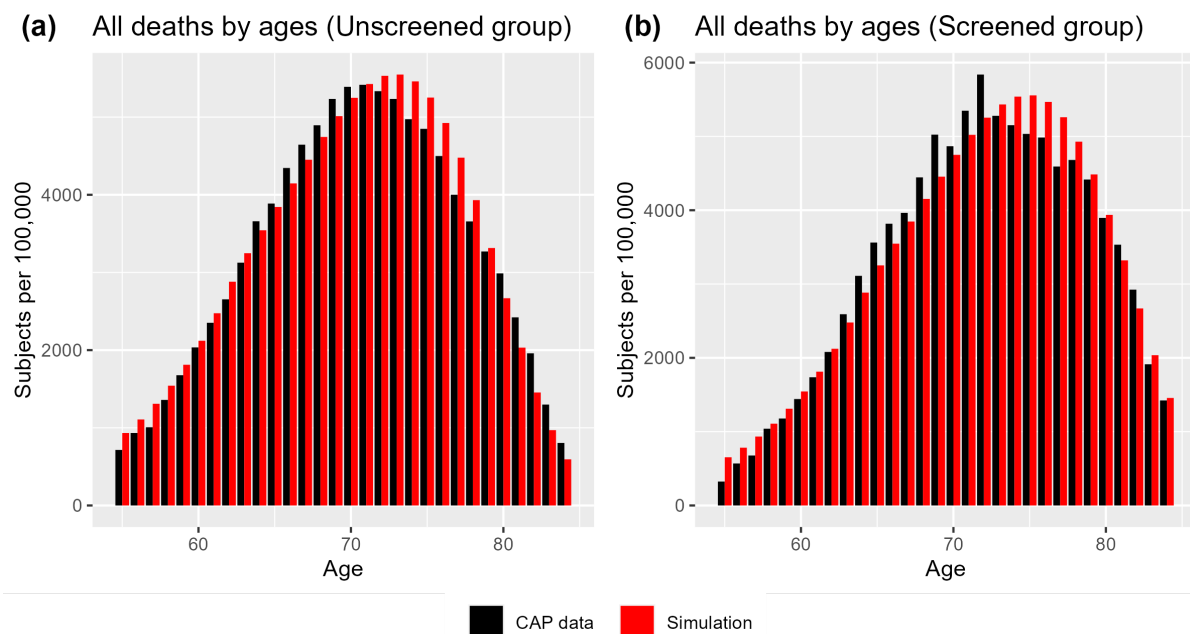
Number at risk at the start of each two-year period (number of prostate cancer diagnoses in that period)												
Time (year)	Median (IQR) follow up	0 ^b	2	4	6	8	10	12	14	16	18	20
<i>A: Gleason ≤6</i>												
Intervention	14.20 (11.42, 16.04)	189,326 (1790)	180,957 (313)	174,289 (330)	167,024 (378)	158,876 (390)	149,145 (402)	139,138 (307)	103,163 (153)	48,427 (47)	12,794 (1)	0 (0)
Control	14.35 (11.09, 15.67)	219,395 (374)	212,352 (458)	204,203 (503)	194,558 (565)	184,887 (556)	172,125 (463)	158,863 (373)	119,810 (157)	38,396 (33)	9,687 (0)	0 (0)
<i>B: Gleason 7</i>												
Intervention	14.20 (11.42, 16.04)	189,326 (930)	180,957 (256)	174,289 (366)	167,024 (460)	158,876 (578)	149,145 (664)	139,138 (676)	103,163 (352)	48,427 (118)	12,794 (2)	0 (0)
Control	14.35 (11.09, 15.67)	219,395 (317)	212,352 (432)	204,203 (593)	194,558 (705)	184,887 (754)	172,125 (899)	158,863 (863)	119,810 (426)	38,396 (84)	9,687 (9)	0 (0)
<i>C: Gleason ≥8</i>												
Intervention	14.20 (11.42, 16.04)	189,326 (257)	180,957 (155)	174,289 (198)	167,024 (266)	158,876 (342)	149,145 (389)	139,138 (375)	103,163 (208)	48,427 (57)	12,794 (0)	0 (0)
Control	14.35 (11.09, 15.67)	219,395 (193)	212,352 (288)	204,203 (288)	194,558 (391)	184,887 (503)	172,125 (539)	158,863 (502)	119,810 (236)	38,396 (52)	9,687 (5)	0 (0)

CI: confidence interval, IQR: interquartile range

eFigure 4: Comparing simulated data to empirical data for the cumulative prostate cancer incidence and cancer-specific and all-other cause mortality risk among the screened men and the unscreened group. Average of 200 simulations of three million men aged 50-69 years with one-off screening in the screened group.

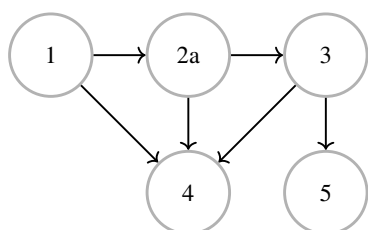


eFigure 5: Comparison number of subjects per 100,000 cohorts at death from all causes by ages between simulated data and CAP data. Average of 200 simulations of three million men aged 50-69 years with one-off screening in the screened group.

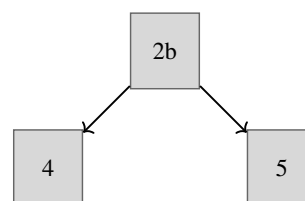


eFigure 6: Transition diagram for multi-state survival models a. Natural history model with states 1-Healthy, 2a – Screen-detectable, 3-clinically diagnosed, 4-all-other cause death, 5-cancer-specific death; b. survival model for screen-detected cancers with states 2b-screen-detected, 4-all-other cause death, 5-cancer-specific death.

Model a.



Model b.



Supplementary Material References

1. Hakama M, Auvinen A, Day NE, Miller AB. Sensitivity in cancer screening. *Journal of Medical Screening* 2007; **14**(4): 174-7.
2. Bhatt R, van den Hout A, Pashayan N. A multistate survival model of the natural history of cancer using data from screened and unscreened population. *Stat Med* 2021; **40**(16): 3791-807.