

#### The Gender Wage Gap Among Those Born in 1958: A Matching Estimator Approach

(ESRC Grant No. ES/S012583/1)

Alex Bryson UCL

UCL SRI Gender Equality Workshop

21<sup>st</sup> April 2022

# **Project Overview**

- Part of an ESRC funded project examining the GWG over the life course using birth cohort data
- The UCL team:
  - Alex Bryson (PI)
  - Heather Joshi (co-investigator)
  - David Wilkinson (co-investigator)
  - Francesca Foliano (Research Fellow)
  - Bozena Wielgoszewska (Research Fellow)
- All information on the project can be found here: <u>https://www.ucl.ac.uk/ioe/departments-and-</u> <u>centres/departments/social-science/gender-wage-gap-evidence-</u> <u>cohort-studies</u>

# Motivation

- 1. Drawbacks in parametric estimation of the gender wage gap (GWG)
  - Failure to compare 'like' men and women
- 2. Common to condition on potentially endogenous variables
  - Biases 'true' estimates of the GWG
- 3. Data from the National Child Development Survey (NCDS) provide good basis for tackling these issues:
  - Match men and women on a rich set of variables liable to impact wage formation over the life cycle which might conceivably differ by gender
  - Measured pre-labour market entry and thus less liable to be endogenous with respect to wage formation
    - Birth, 7, 11 and 16 years collected prospectively

# **Preview of Results**

- 1. Large raw GWG rising until 40s then falls but remains sizeable to age 55
- 2. Raw GWG is larger when use propensity score matching (PSM)
  - Very different to usual regression-adjusted estimates which are often half the size of the raw GWG
- 3. Implication: women have pre-labour market traits which reduce their earnings later in life relative to men
  - Chief among these is occupational expectations
  - Not true for all traits

#### **Previous Literature**

- 1. Studies indicate inverted u-shape in the GWG over the life course
  - Small in early years, widening in 30s/40s, narrowing thereafter
- 2. Falls across cohorts
- 3. Raw gap tends to close by (roughly) one half when condition on other variables
  - Depends somewhat on data set and conditioning variables
- 4. Frequently treats education and fertility decisions as exogenous when, in fact, might be endogenous and partials out some of GWG
  - Same could be said of job traits
- 5. Some exceptions using structural estimation in an effort to tackle endogenous decision-making
  - Adda, Dustmann and Stevens 2017 "The Career Costs of Children", Journal of Political Economy

# Value of Matching Estimators

- 1. Linear estimation (and decompositions on which most are based) based on unnecessarily restrictive assumptions regarding functional form
- 2. By ignoring common support, compare wages of women to men who may not be reasonable comparators
- 3. Matching may make a substantive difference to the estimation of the GWG
  - Strittmatter and Wunsch (2021) explain more of GWG when estimated with PSM
    - Substantial common support issue in their data
    - Combine exact matching on key wage determinants with PSM (radius) matching

# PSM v OLS

- 1. Both assume relevant differences between treated and non-treated are captured by their observables (conditional independence assumption)
  - violated if analysis does not incorporate all factors affecting participation and outcome of interest
  - the assumption is not testable
- 2. Advantages of PSM relative to OLS
  - semi-parametric so does not require assumption of linearity in outcome equation
  - individual causal effect is completely unrestricted so heterogeneous treatment effects can be captured (no assumption of constant additive effects)
  - highlights problem of common support since women must have 'like' counterparts in male population. Thus, avoids extrapolating beyond CS but implications if many treated individuals remain 'unmatched'

### Data and Methods

- 1. National Child Development Study (NCDS)
- 2. Log hourly wages at ages 23, 33, 42, 50, 55
  - Deflated to January 2000 prices
  - Rerun matching for each wage outcome
- 3. Propensity score matching (PSM) matching women to men on single index (the propensity score) derived from probit (0,1) if woman
- 4. Using pre-labour market covariates from mother, cohort member, teacher
  - Parental background; pregnancy/birth; ages 7, 11, 16
- 5. Theory driven as opposed to data driven (Machine Learning)
- 6. Plausibility of conditional independence assumption in this case
- 7. 5 nearest neighbours (Froelich) to recover ATT
  - enforces common support with 0.005 caliper
  - Bootstrapping
- 8. Accommodate selection into employment using zero wages

# Covariates used in matching

Wave	Variables
Pre-birth/birth	Gestation (days); mother smoked during
	pregnancy; white; birthweight (ounces);
	mother smoking 4 months after birth
Age 7	Southgate reading test score; arithmetic
	problems; N Rutter symptoms; Score on
	Bristol Social Adjustment Guide; N child
	illnesses; N hospital admissions; laterality in
	hands
Age 11	Occupational expectations when aged 25;
	standardized reading score; standardized
	maths score; type of school attended;
Age 16	In trouble with police; teacher rating on
	capability relating to maths, English,
	Languages, practical issues; mother's
	assessment of over/underweight; disability;
	alcohol consumption; smoking behaviour

Others tried: mum's and dad's social class; breast fed; region; housing tenure; siblings and household size; verbale and non-verbale test scores aged 11; female teacher; teacher rating of child aged 11; child's expectations on schooling; mum's and dad's interest in education of child; financial hardship and FSMs aged 7 and 11

#### **Occupational Expectations**

	Male	Female
Professional	9	4
Other non-manual,	6	4
scientific		
Typist, clerical	2	11
Shop assistant	1	7
Junior non-managerial	3	1
Personal services	1	9
Foreman, manual	<1	<1
Skilled manual	18	1
Semi-skilled manual	3	<2
Unskilled manual	<1	<1
Self-employed	1	1
Farm worker	2	2
HM Forces	7	<1
Sports man/woman	9	<1
Student	<1	<1
Teacher/nurse	2	20
Unclassifiable	34	38

# Match Bias

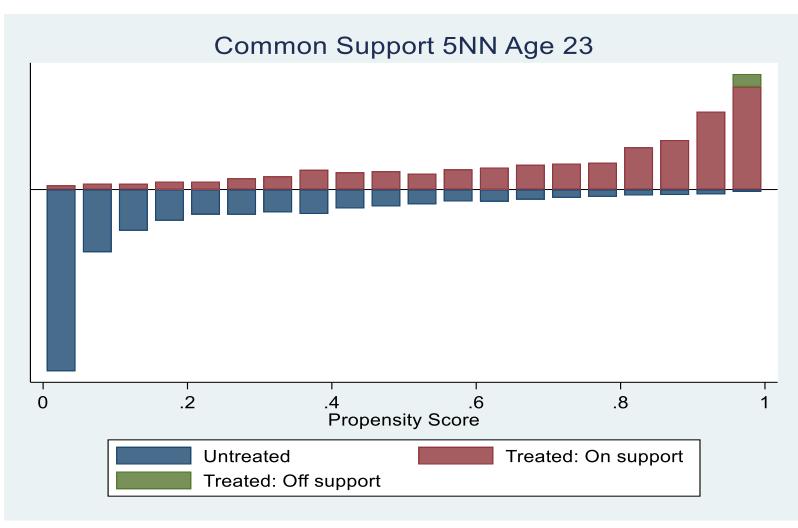
	23	33	42	50	55
Pseudo r-sq:	0.446	0.440	0.429	0.437	0.450
U M	0.032	0.046	0.041	0.042	0.060
Rubin's B	41.4*	51.4*	48.1*	48.8*	58.6*
Rubin's R	0.83	1.13	0.95	1.05	1.08

Rubin's B: absolute standardised differences of means of linear index of propensity score in treated and match non-treated groups (B<25 is ok)

Rubin's R: ratio of treated to matched non-treated variances of propensity score index (R between 0.5 and 2 is deemed balanced)

- means falls outside tolerable balance limits
- Problems in relation to reading and maths scores at age 11, Rutter, BSAG, illness

# **Common Support**



87 cases off common support. Zero at other ages

#### Results:

Log Hourly Wages, Raw Diffs and ATT from Matching

	23	33	42	50	55
Fem	1.536	1.843	1.908	2.080	2.022
Male U	1.704	2.209	2.354	2.455	2.359
Male M	1.711	2.211	2.402	2.504	2.437
Dif Raw	168	367	446	355	337
Dif Matched	176	368	495	424	415
Ν	8011	6881	7175	6031	4992

U = unmatchedM = matched

At age 23: raw = -.168 OLS = -.184 PSM = -.176 OLS with CS and match weights = -.173 ebalance = -.181

# What Next?

- 1. Regression using match weights having enforced CS
  - Compare regression-adjusted GWG with standard regression
  - Compare decomposition with decomp from standard reg-adj GWG
- 2. Specification for probit
  - Have we got the right covariates?
  - More flexible specification
- 3. Alternative matching estimators
  - NN, kernel; combine exact matching with PSM; entropy balancing
- 4. Tackline participation decision
  - Bringing in the zeros results in a much larger GWG
  - Is this the right thing to do?
- 5. Attrition
- 6. Check out wages at ages 61 and 63 (smaller Ns)
- 7. Run on BCS 1970